

# Hierarchical Forecasts Reconciliation

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## **Abstract**

TBC

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# 1 Introduction

## 2 Coherent forecasts

### 2.1 Notation and preliminaries

We briefly define the concept of a *hierarchical time series* in a fashion similar to Wickramasuriya et al. (2018), Hyndman & Athanasopoulos (2018) and others, before elaborating on some of the limitations of this understanding. A *hierarchical time series* is a collection of  $n$  variables indexed by time, where some variables are aggregates of other variables. We let  $\mathbf{y}_t \in \mathbb{R}^n$  be a vector comprising observations of all variables in the hierarchy at time  $t$ . The *bottom-level series* are defined as those  $m$  variables that cannot be formed as aggregates of other variables; we let  $\mathbf{b}_t \in \mathbb{R}^m$  be a vector comprised of observations of all bottom-level series at time  $t$ . The hierarchical structure of the data implies that the following holds for all  $t$

$$\mathbf{y}_t = \mathbf{S}\mathbf{b}_t, \tag{1}$$

where  $\mathbf{S}$  is an  $n \times m$  constant matrix that encodes the aggregation constraints.

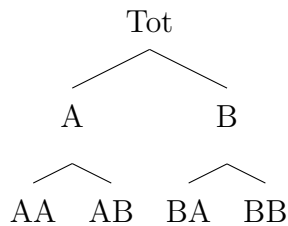


Figure 1: An example of a two level hierarchical structure.

To clarify these concepts consider the example of the hierarchy in Figure 1. For this hierarchy,  $n = 7$ ,  $\mathbf{y}_t = [y_{Tot,t}, y_{A,t}, y_{B,t}, y_{C,t}, y_{AA,t}, y_{AB,t}, y_{BA,t}, y_{BB,t}]'$ ,  $m = 4$ ,  $\mathbf{b}_t = [y_{AA,t}, y_{AB,t}, y_{BA,t}, y_{BB,t}]'$  and

$$\mathbf{S} = \begin{pmatrix} 1 & 1 & 1 & 1 \\ 1 & 1 & 0 & 0 \\ 0 & 0 & 1 & 1 \\ \mathbf{I}_4 \end{pmatrix},$$

where  $\mathbf{I}_4$  is the  $4 \times 4$  identity matrix.

While such a definition is completely serviceable, it obscures the full generality of the literature on so-called hierarchical time series. In fact, concepts such as coherence and reconciliation, defined in full below, only require the data to have two important characteristics; the first is that they are multivariate, the second is that they adhere to linear constraints.

## 2.2 Coherence

The property that data adhere to some linear constraints is referred to as *coherence*. We now provide definitions aimed at providing geometric intuition of hierarchical time series.

**Definition 2.1** (Coherent subspace). The  $m$ -dimensional linear subspace  $\mathfrak{s} \subset \mathbb{R}^n$  for which a set of linear constraints holds for all  $\mathbf{y} \in \mathfrak{s}$  is defined as the *coherent subspace*.

To further illustrate, Figure 2 depicts the most simple three variable hierarchy where  $y_{Tot,t} = y_{A,t} + y_{B,t}$ . The coherent subspace is depicted as a grey 2-dimensional plane within 3-dimensional space, i.e.  $m = 2$  and  $n = 3$ . It is worth noting that the coherent subspace is spanned by the columns of  $\mathbf{S}$ , i.e.  $\mathfrak{s} = \text{span}(\mathbf{S})$ . In Figure 2, these columns are  $\vec{s}_1 = (1, 1, 0)'$  and  $\vec{s}_2 = (1, 0, 1)'$ . However, it is equally important to recognise that the hierarchy could

also have been defined in terms of  $y_{Tot,t}$  and  $y_{A,t}$  rather than the bottom level series,  $y_{A,t}$  and  $y_{B,t}$ . In this case the corresponding ‘ $\mathbf{S}$  matrix’ would have columns  $(1, 0, 1)'$  and  $(0, 1, -1)'$ . However, while there are multiple ways to define an  $\mathbf{S}$  matrix, in all cases the columns will span the same coherent subspace, which is unique.

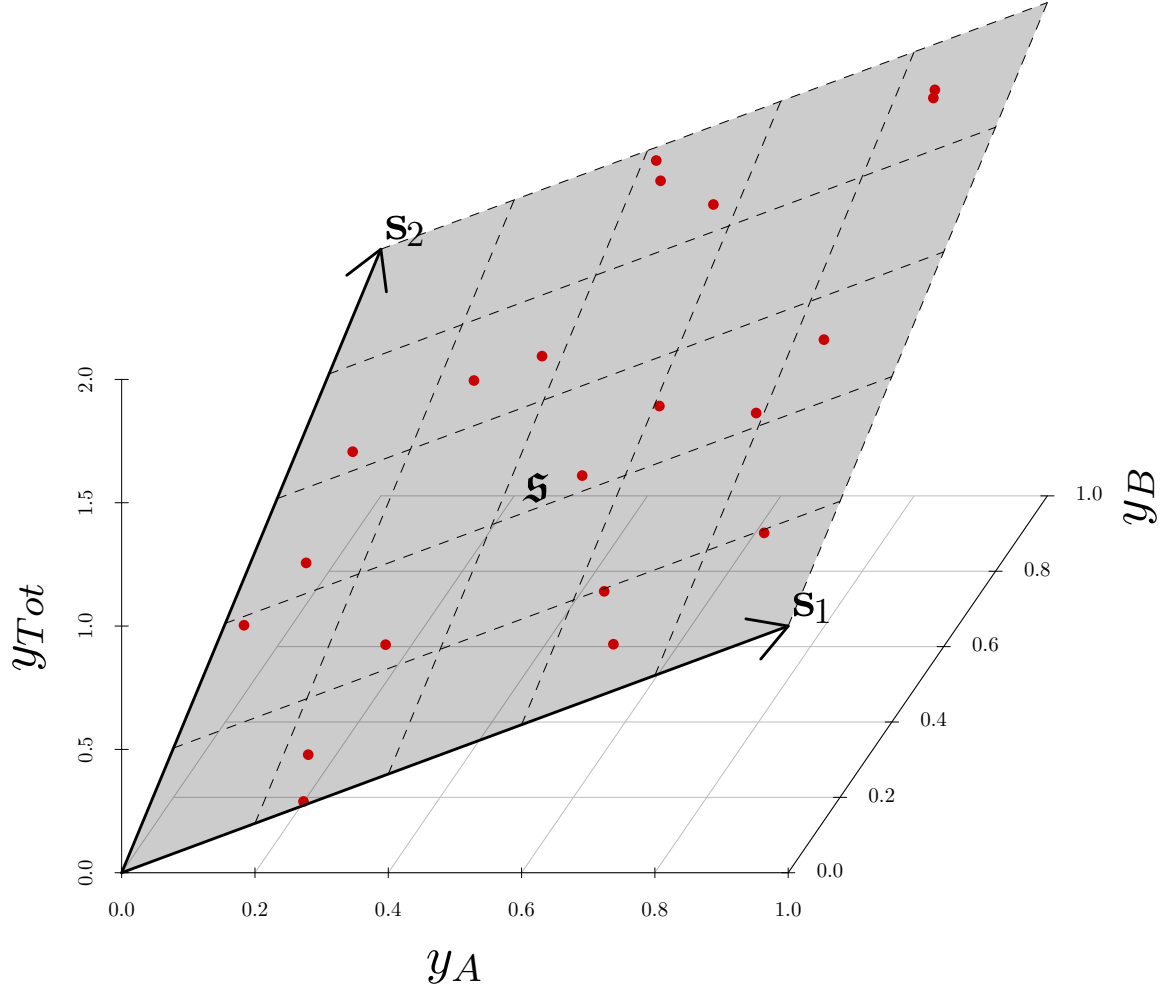


Figure 2: Depiction of a three dimensional hierarchy with  $y_{Tot} = y_A + y_B$ . The gray colour two dimensional plane reflects the coherent subspace  $\mathfrak{s}$  where  $\vec{s}_1 = (1, 1, 0)'$  and  $\vec{s}_2 = (1, 0, 1)'$  are basis vectors that spans  $\mathfrak{s}$ . The points in  $\mathfrak{s}$  represents realisations or coherent forecasts

**Definition 2.2** (Hierarchical Time Series). A hierarchical time series is an  $n$ -dimensional multivariate time series such that all observed values  $\mathbf{y}_1, \dots, \mathbf{y}_T$  and all future values  $\mathbf{y}_{T+1}, \mathbf{y}_{T+2}, \dots$  lie in the coherent subspace, i.e.  $\mathbf{y}_t \in \mathfrak{s} \quad \forall t$ .

Despite the common use of the term *hierarchical time series*, it should be clear from the definition that the data need not necessarily follow a hierarchy. Also notable by its absence in the above definition is any reference to *aggregation*. In some ways, terms such as *hierarchical* and *aggregation* can be misleading since the literature has covered instances that cannot easily be depicted in a similar fashion to Figure 1 and or do not involve aggregation. **Include brief summary of all non-traditional hierarchies - e.g. grouped hierarchies, temporal hierarchies with wierd overlapping, problems where we look at differences between variables etc.** Finally, although Definition 2.2 makes clear reference to time series, this definition can be easily generalised to any vector-valued data for which some linear constraints are known to hold for all realisations.

**Definition 2.3** (Coherent Point Forecasts). Let  $\check{\mathbf{y}}_{t+h|t} \in \mathbb{R}^n$  be a vector of point forecasts of all series in the hierarchy at time  $t+h$ , made using information up to and including time  $t$ . Then  $\check{\mathbf{y}}_{t+h|t}$  is *coherent* if  $\check{\mathbf{y}}_{t+h|t} \in \mathfrak{s}$ .

Without any loss of generality, that above definition could also be applied to prediction for multivariate data in general, rather than just forecasting of time series. While the observed data will be coherent by definition, it is important to note that there are a number of reasons why forecasts or predictions may be incoherent.

First, since applications of hierarchical forecasting tend to be very high dimensional a common strategy in practice is to produce forecasts for each time series independently using univariate models. Second, even where a multivariate model is used for the full vector of observations, it may be difficult to capture the linear constraints inherent in the

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data particularly for complicated non-linear models. Third, in some cases judgemental adjustments may be made inducing incoherent forecasts.

### 3 Forecast reconciliation

As discussed in the previous section, coherence is not guaranteed for a number of reasons. To ensure aligned decision making, it is desirable to adjust forecasts ex post to ensure coherence. This process is referred to as *reconciliation*. In the most general terms, reconciliation can be defined as follows

**Definition 3.1** (Reconciled forecasts). Let  $\psi$  be a mapping,  $\psi : \mathbb{R}^n \rightarrow \mathfrak{s}$ . The point forecast  $\tilde{\mathbf{y}}_{t+h|t} = \psi(\hat{\mathbf{y}}_{t+h|t})$  is said to “reconcile”  $\hat{\mathbf{y}}_{t+h|t}$  with respect to the mapping  $\psi(\cdot)$

All reconciliation methods that we are aware of consider a linear mapping for  $\psi$ , which involves pre-multiplying base forecasts by an  $n \times n$  matrix that has  $\mathfrak{s}$  as its image. One way to achieve this is with a matrix  $\mathbf{SG}$ , where  $\mathbf{G}$  is an  $(n - m) \times n$  matrix (some authors use  $\mathbf{P}$  used in place of  $\mathbf{G}$ ). This facilitates an interpretation of reconciliation as a two-step process, in the first step, base forecasts  $\hat{\mathbf{y}}_{t+h|t}$  are combined to form a new set of bottom level forecasts, in the second step, these mapped to a full vector of coherent forecasts via pre-multiplication by  $\mathbf{S}$ .

Although pre-multiplying base forecasts by  $\mathbf{SG}$  will result in coherent forecasts, a number of desirable properties arise when  $\mathbf{SG}$  has the specific structure of a *projection* matrix onto  $\mathfrak{s}$ . In general a projection matrix has the idempotence property, i.e.  $\mathbf{SG}^2 = \mathbf{SG}$ . However a much more important property of projection matrices, used in multiple instances below, is that any vector lying in the image of the projection will be mapped onto itself by that projection. In our context this means that for any  $\mathbf{v} \in \mathfrak{s}$ ,  $\mathbf{SG}\mathbf{v} = \mathbf{v}$ .



We begin by considering the special case of an orthogonal projection whereby  $\mathbf{G} = (\mathbf{S}'\mathbf{S})^{-1}\mathbf{S}'$ . This is equivalent to so called OLS reconciliation as introduced by Hyndman et al. (2011). We refrain from any discussion of regression models focusing instead on geometric interpretations. However the connection between OLS and orthogonal projection should be clear, in the context of regression modelling predicted values from OLS are obtained via an orthogonal projection of the data onto the span of the regressors.

### 3.1 Orthogonal projection

In this section we discuss two sensible properties that can be achieved by reconciliation via orthogonal projection. The first is that reconciliation should adjust the base forecasts as little as possible, i.e. the base and reconciled forecast should be ‘close’. The second is that reconciliation in some sense should improve forecast accuracy, or more loosely, that the reconciled forecast should be ‘closer’ to the realisation that is the target of the forecast.

To address the first of these properties we make the concept of closeness more concrete, by considering the Euclidean distance between the base forecast  $\hat{\mathbf{y}}$  and the reconciled forecast  $\tilde{\mathbf{y}}$ . A property of an orthogonal projection is that the distance between  $\hat{\mathbf{y}}$  and  $\tilde{\mathbf{y}}$  will be as small as possible while still ensuring that  $\tilde{\mathbf{y}} \in \mathfrak{s}$ . In this sense reconciliation via orthogonal projection leads to the smallest possible adjustments of the base forecasts.

The second property introduced above has been the focus of theoretical results in the forecast reconciliation literature. Here we provide a more streamlined version of proofs by van Erven & Cugliari (2014) and Wickramasuriya et al. (2018) before providing geometrical intuition aimed at simplifying the reader’s understanding of these proofs.

Consider the Euclidean distance between a forecast and the target. This is equivalent to the root of the sum of squared errors over the entire hierarchy. Let  $\mathbf{y}_{t+h}$  be the realisation of the data generating process at time  $t + h$ , and let  $\|\mathbf{v}\|_2$  be the  $L_2$  norm of vector  $\mathbf{v}$ . The

following theorem shows that reconciliation never increases, and in most cases reduces, the sum of squared errors of point forecasts.

**Theorem 3.1** (Distance reducing property). *If  $\tilde{\mathbf{y}}_{t+h|t} = \mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{t+h|t}$ , where  $\mathbf{G}$  is such that  $\mathbf{S}\mathbf{G}$  is an orthogonal projection onto  $\mathfrak{s}$ , then the following inequality holds:*

$$\|(\tilde{\mathbf{y}}_{t+h|t} - \mathbf{y}_{t+h})\|_2^2 \leq \|(\hat{\mathbf{y}}_{t+h|t} - \mathbf{y}_{t+h})\|_2^2. \quad (2)$$

*Proof.* Since the aggregation constraints must hold for all realisations,  $\mathbf{y}_{t+h} \in \mathfrak{s}$  and  $\mathbf{y}_{t+h} = \mathbf{S}\mathbf{G}\mathbf{y}_{t+h}$  since  $\mathbf{S}\mathbf{G}$  is a projection onto  $\mathfrak{s}$ . Therefore,

$$\|(\tilde{\mathbf{y}}_{t+h|t} - \mathbf{y}_{t+h})\|_2 = \|(\mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{t+h|t} - \mathbf{S}\mathbf{G}\mathbf{y}_{t+h})\|_2 \quad (3)$$

$$= \|\mathbf{S}\mathbf{G}(\hat{\mathbf{y}}_{t+h|t} - \mathbf{y}_{t+h})\|_2. \quad (4)$$

The Cauchy-Schwarz inequality can be used to show that orthogonal projections are bounded operators (Hunter & Nachtergaele 2001), therefore

$$\|\mathbf{S}\mathbf{G}(\hat{\mathbf{y}}_{t+h|t} - \mathbf{y}_{t+h})\|_2 \leq \|(\hat{\mathbf{y}}_{t+h|t} - \mathbf{y}_{t+h})\|_2.$$

□

The inequality is strict whenever  $\hat{\mathbf{y}}_{t+h|t} \notin \mathfrak{s}$ .

The simple geometric intuition behind the proof is demonstrated in Figure ?? . In this schematic, the coherent subspace is depicted as a black arrow. The base forecast  $\hat{\mathbf{y}}$  is shown as a blue dot. Since it is incoherent it does not lie in  $\mathfrak{s}$ . Reconciliation is an orthogonal projection from  $\hat{\mathbf{y}}$  to the coherent subspace yielding the reconciled forecast  $\tilde{\mathbf{y}}$  shown in red. Finally, the target of the forecast  $\mathbf{y}$  is displayed as a black point, and although its exact location is unknown to the forecaster, it is known that it will lie somewhere along the coherent subspace.

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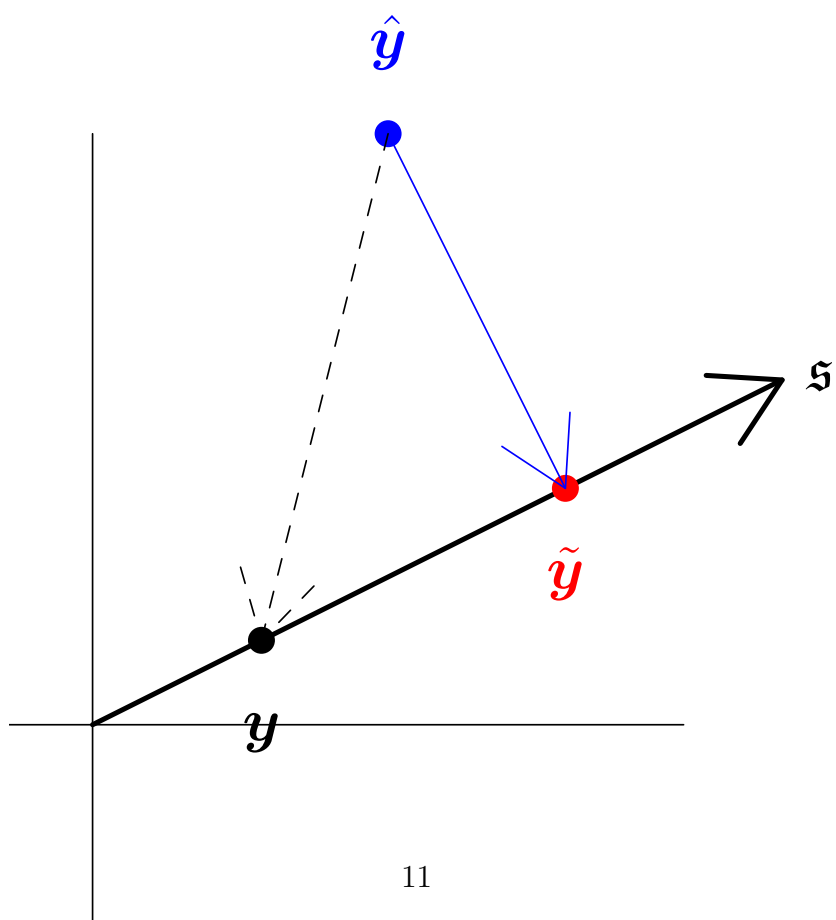


Figure ?? clearly shows that  $\hat{\mathbf{y}}$ ,  $\tilde{\mathbf{y}}$  and  $\mathbf{y}$  form a right angled triangle. In this triangle the distance between  $\mathbf{y}$  and  $\hat{\mathbf{y}}$  is the hypotenuse and therefore must be longer than the distance between  $\mathbf{y}$  and  $\tilde{\mathbf{y}}$ . As such reconciliation is guaranteed to reduce the squared error of the forecast.

Theorem 3.1 is in some ways more powerful than perhaps previously understood. It is often stated in terms of expectations. However, the distance reducing property result is even stronger since it will hold for any realisation and any forecast. Nothing needs to be assumed about the statistical properties of the data generating process or the process by which forecasts are made.

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However, in other ways, Theorem 3.1 is weaker than perhaps often understood. First, when improvements in forecast accuracy are discussed in the context of the theorem, this refers to a very specific measure of forecast accuracy. In particular, this measure is the root of the sum of squared errors of *all* variables in the hierarchy. As such, while forecast improvement is guaranteed for the hierarchy overall, reconciliation can lead to worse forecasts for individual series. Second, although orthogonal projections are guaranteed to improve on base forecasts both for all realisations and in expectation, they are not necessarily the projection that leads to the greatest improvement in forecast accuracy. As such referring to reconciliation via orthogonal projections as ‘optimal’ is somewhat misleading since it does not have the optimality properties of some oblique projections, in particular MinT. It is to oblique projections that we now turn our attention.

## 3.2 Oblique Projections

One justification for using an orthogonal projection is that it leads to improved forecast accuracy in terms of the root of the sum of squared errors of *all* variables in the hierarchy. A clear shortcoming of this measure of forecast accuracy is that forecasts errors in all series

should not necessarily be treated equally. For example, in hierarchies, top-level series tend to have a much larger scale than bottom level series. Even when two series are on a similar scale, series that are more predictable or less variable will tend to be downweighted by simply aggregating square errors. An even more sophisticated understanding may take the correlation between series into account. All of these considerations lead towards reconciliation of the form  $\tilde{\mathbf{y}} = \mathbf{S}(\mathbf{S}'\mathbf{W}^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}^{-1}\hat{\mathbf{y}}$ , where  $\mathbf{W}$  is a symmetric matrix. Generally, it is assumed that  $\mathbf{W}$  is invertible, otherwise a pseudo inverse can be used.

It should be noted that  $\mathbf{S}(\mathbf{S}'\mathbf{W}^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}^{-1}$  is an oblique, rather than an orthogonal projection matrix. However this matrix can be considered to be an orthogonal projection for a different geometry, namely the generalised Euclidean geometry with respect to  $\mathbf{W}^{-1}$ . One way to understand this geometry is that it is the same as Euclidean geometry when all vectors are first transformed by pre-multiplying by  $\mathbf{W}^{-1/2}$ . This leads to a transformed  $\mathbf{S}$  matrix  $\mathbf{S}^* = \mathbf{W}^{-1/2}\mathbf{S}$  and transformed  $\hat{\mathbf{y}}$  and  $\tilde{\mathbf{y}}$  vectors  $\hat{\mathbf{y}}^* = \mathbf{W}^{-1/2}\hat{\mathbf{y}}$  and  $\tilde{\mathbf{y}}^* = \mathbf{W}^{-1/2}\tilde{\mathbf{y}}$ . The transformed reconciled forecast results from an orthogonal projection in the transformed space since

$$\tilde{\mathbf{y}}^* = \mathbf{W}^{-1/2}\tilde{\mathbf{y}} \tag{5}$$

$$= \mathbf{W}^{-1/2}\mathbf{S}(\mathbf{S}'\mathbf{W}^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}^{-1}\hat{\mathbf{y}} \tag{6}$$

$$= \mathbf{S}^* \left( \mathbf{S}^{*\prime} \mathbf{S}^* \right)^{-1} \mathbf{S}^{*\prime} \hat{\mathbf{y}}^* \tag{7}$$

Another way of understanding the generalised Euclidean geometry is that it is defined by the norm  $\mathbf{v}'\mathbf{W}^{-1}\mathbf{v}$ . This interpretation is quite instructive when it comes to thinking about the connection between distances and loss functions. In the generalised Euclidean geometry, the distance between the reconciled forecast and the realisation is given by  $(\hat{\mathbf{y}} - \mathbf{y})'\mathbf{W}^{-1}(\hat{\mathbf{y}} - \mathbf{y})$ . For diagonal  $\mathbf{W}$  this is equivalent to a weighted sum of squared

error loss function and when  $\mathbf{W}$  is a covariance matrix this is equivalent to a Mahalanobis distance. As such Theorem 3.1 can easily be generalised. If the objective function is some weighted sum of squared errors, or a Mahalanobis distance, then the projection matrix  $\mathbf{S}(\mathbf{S}'\mathbf{W}^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}^{-1}$  is guaranteed to improve forecast accuracy over base forecasts, for an appropriately selected  $\mathbf{W}$ .

First, the linear subspace onto which all points are projected, or the image of the projection, must be defined. In our context this can be defined by the  $m$  columns of the matrix  $\mathbf{S}$ . Second, the direction along which points are projected must be defined. This will be achieved by defining a matrix  $\mathbf{R}$  with  $n - m$  columns then span the direction of projection. A schematic of this is presented . A projection matrix can then be constructed as  $\mathbf{S}(\mathbf{R}'_{\perp}\mathbf{S})^{-1}\mathbf{R}'_{\perp}$  where,  $\mathbf{R}_{\perp}$  is an  $n \times m$  orthogonal complement to  $\mathbf{R}$  such that  $\mathbf{R}'_{\perp}\mathbf{R} = \mathbf{0}$ . It is simple to verify that this construction satisfies the properties of a projection matrix, namely symmetry and idempotence.

A straightforward choice of  $\mathbf{R}$  for the most simple three variable hierarchy where  $y_{1,t} = y_{2,t} + y_{3,t}$ , is the vector  $(1, -1, -1)$  which is orthogonal (in the Euclidean sense) to the columns of  $\mathbf{S}$ . In this case, the matrix  $\mathbf{R}$  can be interpreted as a ‘restrictions’ matrix since it has the property that  $\mathbf{R}'\mathbf{y} = \mathbf{0}$  for coherent  $\mathbf{y}$ . In OLS reconciliation,  $\mathbf{R}'_{\perp} = \mathbf{S}'$  whereas in MinT or WLS reconciliation  $\mathbf{R}'_{\perp}$  takes the form  $\mathbf{S}'\mathbf{W}^{-1}$ . We will be discussing these projections distinctly in the next subsection.

In MinT reconciliation,  $\mathbf{R}'_{\perp}$  is taking the form  $\mathbf{S}'\mathbf{W}^{-1}$ , where it can be thought of as orthogonal projections after pre-multiplying by  $\mathbf{W}^{-1/2}$ . That is, the coordinates of incoherent space will be scaled by  $\mathbf{W}^{-1/2}$  which is then followed by the orthogonal projection. Alternatively this can be interpreted as an oblique projections in Euclidean space where the columns of  $\mathbf{R}$  is the ‘direction’ along which incoherent point forecasts are projected onto the coherent space  $\mathfrak{s}$  as depicted in Figure ???. In terms of distances, MinT minimises

the Euclidean distance between  $\hat{\mathbf{y}}_{t+h|t}$  and  $\tilde{\mathbf{y}}_{t+h|t}$  in the transformed space which is same as the scaled Euclidean distance in the original space. Latter is also referred to as the Mahalanobis distance. We also note that the WLS is a special case of MinT where  $\mathbf{W}^{-1}$  is a diagonal matrix.

Wickramasuriya et al. (2018) showed that the MinT is optimal with respect to the mean squared forecast errors. We can provide a more general geometrical explanation to this optimality using the schematic in Figure 4. Consider the h-step ahead reconciled forecast errors. These can be always approximated by the insample h-step ahead forecast errors. Since these errors are coherent, they lie in a direction that is closer to the coherent subspace  $\mathfrak{s}$ . Therefore if you project  $\hat{\mathbf{y}}$  along the direction of these in-sample forecast errors, then you can get closer to the true value  $\mathbf{y}$  as depicted in the schematic. Further, unlike OLS, the squared error for MinT reconciled forecasts is not always less than that of base forecasts in every single replication although it outperforms on average.

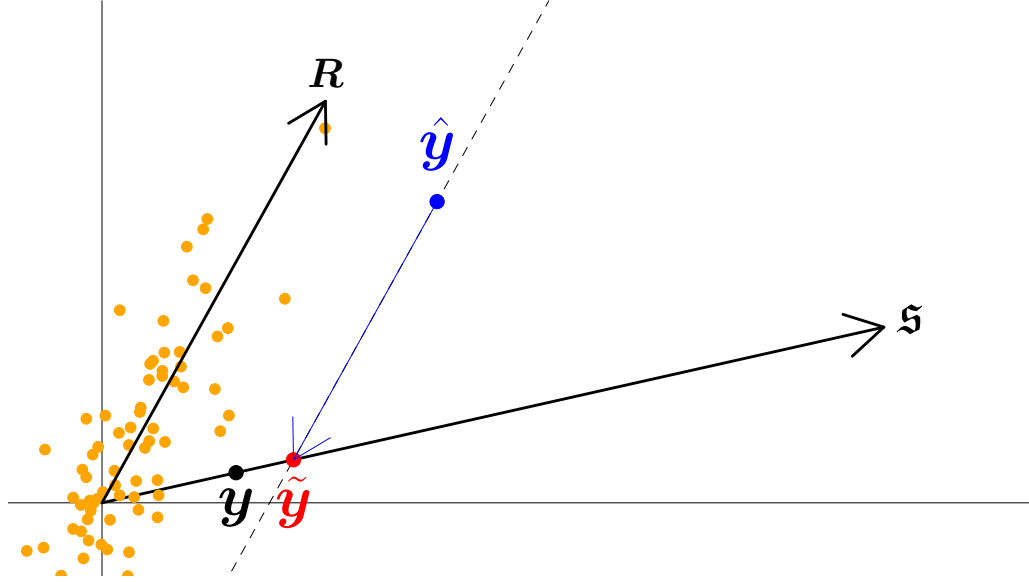


Figure 4: A schematic to represent MinT reconciliation. Points in orange colour represent the insample errors.  $\mathbf{R}$  shows the direction of the insample errors.  $\hat{\mathbf{y}}$  is projected onto  $\mathbf{s}$  along the the direction of  $\mathbf{R}$ .

## 4 Bias in forecast reconciliation

Before turning our attention to the issue of bias itself it is important to state a sensible property that any reconciliation method should have. That is if base forecasts are already coherent then reconciliation should not change the forecast. As stated in Section 3, this



property holds when  $\mathbf{SG}$  is a projection matrix. This implies for arbitrary  $\mathbf{G}$ , reconciliation may in fact change an already coherent forecast.

The property that projections map all vectors in the coherent subspace onto themselves is useful in proving the unbiasedness preserving property of reconciliation . Before restating this proof using a clear geometric interpretation we discuss in a precise fashion what is meant by unbiasedness.

Suppose that the target of a point forecast is  $\boldsymbol{\mu}_{t+h|t} := \mathbb{E}(\mathbf{y}_{t+h} \mid \mathbf{y}_1, \dots, \mathbf{y}_t)$  where the expectation is taken over the predictive density. Our point forecast can be thought of as an estimate of this quantity. The forecast is random due to uncertainty in the training sample and it is with respect to this uncertainty that unbiasedness refers. More concretely, the point forecast will be unbiased if  $\mathbb{E}_{1:t}(\hat{\mathbf{y}}_{t+h|t}) = \boldsymbol{\mu}_{t+h|t}$ , where the subscript  $1:t$  denotes an expectation taken over the training sample.

**Theorem 4.1** (Unbiasedness preserving property). *For unbiased  $\hat{\mathbf{y}}_{t+h|t}$ , the reconciled point forecast is also an unbiased prediction as long as  $\mathbf{SG}$  is a projection onto  $\mathfrak{s}$ .*

*Proof.* The expected value of the reconciled forecast is given by

$$\mathbb{E}_{1:t}(\tilde{\mathbf{y}}_{t+h|t}) = \mathbb{E}_{1:t}(\mathbf{SG}\hat{\mathbf{y}}_{t+h|t}) = \mathbf{SG}\mathbb{E}_{1:t}(\hat{\mathbf{y}}_{t+h|t}) = \mathbf{SG}\boldsymbol{\mu}_{t+h|t}.$$

Since  $\boldsymbol{\mu}_{t+h|t}$  is an expectation taken with respect to the degenerate predictive density it must lie in  $\mathfrak{s}$ . We have already established that when  $\mathbf{SG}$  is a projection onto  $\mathfrak{s}$  then it maps all vectors in  $\mathfrak{s}$  onto themselves. As such  $\mathbf{SG}\boldsymbol{\mu}_{t+h|t} = \boldsymbol{\mu}_{t+h|t}$  when  $\mathbf{SG}$  is a projection matrix.  $\square$

We note that the above result holds when the projection  $\mathbf{SG}$  is only onto the coherent subspace  $\mathfrak{s}$  and not for all projection matrices in general. To describe this more explicitly suppose  $\mathbf{SG}$  has as its image  $\mathfrak{L}$  which is itself a lower dimensional linear subspace of  $\mathfrak{s}$ , i.e.

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$\mathfrak{L} \subset \mathfrak{s}$ . Then for  $\{\boldsymbol{\mu}_{t+h|t} : \boldsymbol{\mu}_{t+h|t} \in \mathfrak{s}, \boldsymbol{\mu}_{t+h|t} \notin \mathfrak{L}\}$ ,  $\mathbf{SG}\boldsymbol{\mu}_{t+h|t} \neq \boldsymbol{\mu}_{t+h|t}$ . This is depicted in Figure 5 where  $\boldsymbol{\mu}_{t+h|t}$  is projected to a point  $\bar{\boldsymbol{\mu}}$  in  $\mathfrak{L}$ . Therefore in this case, the reconciled forecast will have as its expectation  $\bar{\boldsymbol{\mu}}$  rather than  $\boldsymbol{\mu}_{t+h|t}$  and be biased. This result has implications in practice, in particular, the top-down method (Gross & Sohl 1990) has

$$\mathbf{G} = \begin{pmatrix} \mathbf{p} & \mathbf{0}_{(m \times n-1)} \end{pmatrix} \quad (8)$$

where  $\mathbf{p} = (p_1, \dots, p_m)'$  is an  $m$ -dimensional vector consisting a set of proportions which is use to disaggregate the top-level forecasts along the hierarchy. In this case it can be verified that  $\mathbf{SG}$  is idempotent, i.e.  $\mathbf{SGSG} = \mathbf{SG}$  and therefore  $\mathbf{SG}$  is a projection matrix. However the image of this projection is not an  $m$ -dimensional subspace but a 1-dimensional subspace. As such, top-down reconciliation will bias base forecasts when those base forecasts are unbiased.

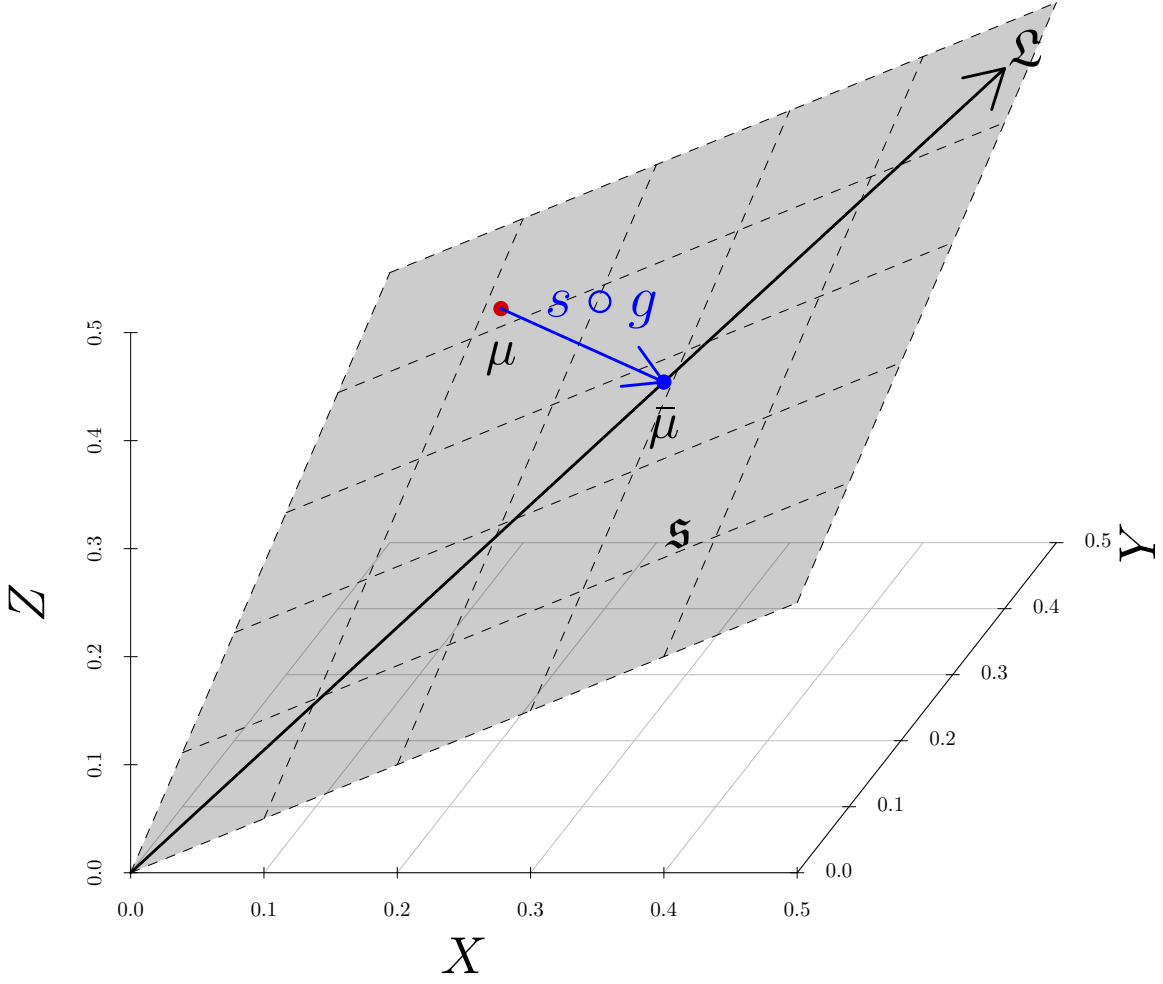


Figure 5:  $\mathfrak{L}$  is a linear subspace of the coherent subspace  $\mathfrak{s}$ . If  $s \circ g$  is a projection not onto  $\mathfrak{s}$  but onto  $\mathfrak{L}$ , then  $\mu \in \mathfrak{s}$  will be moved to  $\bar{\mu} \in \mathfrak{L}$ .

Finally, it is often stated that an assumption required to prove the unbiasedness pre-

serving property is that  $\mathbf{S}\mathbf{G}\mathbf{S} = \mathbf{S}$  or alternatively that  $\mathbf{G}\mathbf{S} = \mathbf{I}$ . Both of these conditions are equivalent to assuming that  $\mathbf{S}\mathbf{G}$  is a projection matrix. When the problem is viewed through the prism of imposing a constraint  $\mathbf{G}\mathbf{S} = \mathbf{I}$  to ensure unbiasedness is preserved, one may be tempted to deal with biased forecasts by selecting  $\mathbf{G}$  in an unconstrained manner. However, equipped with a geometric understanding of the problem, we would advise against this approach. Our own solution to dealing with biased forecasts is discussed in detail in the next section.

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## 5 Bias correction

## 6 Application

## 7 Conclusions

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