#### Probabilistic Forecasts for Hierarchical Time Series

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Supervisors:

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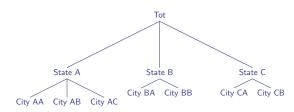
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#### Overview

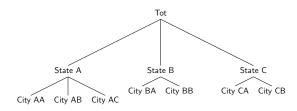
- Introduction
- Probabilistic forecast reconciliation for hierarchical time series
- Probabilistic forecasts in a non-parametric framework: A bootstrap approach
- 4 Probabilistic forecasts for Australian domestic tourism flows
- **5** Conclusions and future work

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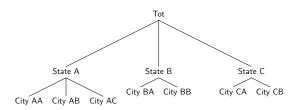
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■ Forecasts of these hierarchical time series should be *coherent* 

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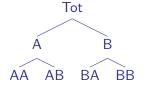
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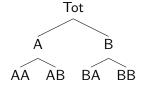
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- **Objective:** Reconciliation of probabilistic forecasts for hierarchical time series such that they preserve the inherent properties of the hierarchical nature

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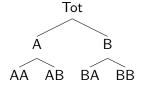
# Probabilistic forecast reconciliation for hierarchical time series

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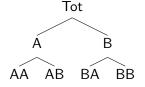


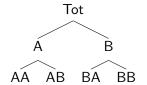


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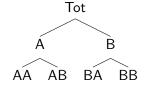




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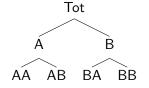


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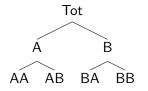
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$$\mathbf{S} = \begin{pmatrix} 1 & 1 & 1 & 1 \\ 1 & 1 & 0 & 0 \\ 0 & 0 & 1 & 1 \\ 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$



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■ Due to the aggregation nature of the hierarchy we have,

$$\mathbf{y}_t = \mathbf{S}\mathbf{b}_t$$

# Definition 1: Coherent subspace

Suppose an n-dimensional time series  $\mathbf{y}_t \in \mathbb{R}^n$  is subject to the linear aggregation constraint  $\mathbf{y}_t = \mathbf{S}\mathbf{b}_t$  where  $\mathbf{b}_t \in \mathbb{R}^m$  and  $\mathbf{S}$  is an  $n \times m$  constant matrix. Let  $\mathbb{C}^m$  be an m-dimensional subspace of  $\mathbb{R}^n$ , where  $\mathbb{C}^m \subset \mathbb{R}^n$ . Then  $\mathbb{C}^m$  is said to be a coherent space if it is spanned by the columns of  $\mathbf{S}$ .

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# Definition 2: Coherent point forecasts

Suppose  $\mathbf{\breve{y}}_{t+h|t} \in \mathbb{R}^n$  denotes point forecasts of each series in the hierarchy at time t+h. Then  $\mathbf{\breve{y}}_{t+h|t}$  is *coherent* if  $\mathbf{\breve{y}}_{t+h|t} \in \mathbb{C}^m$ .

## Definition 3: Coherent probabilistic forecasts

Let  $(\mathbb{R}^m, \mathscr{F}^m, \nu^m)$  be a probability triple where  $\mathscr{F}^m$  is a  $\sigma$ -algebra on  $\mathbb{R}^m$ . Then,  $(\mathbb{C}^m, \mathscr{F}_{\mathbf{S}}, \widecheck{\nu})$  is said to be a coherent probability triple iff

$$\breve{\nu}(\mathbf{S}(\mathbf{A})) = \nu^{m}(\mathbf{A}) \quad \forall \mathbf{A} \in \mathscr{F}^{m},$$

where S(A) denotes the image of subset A under S.

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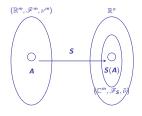
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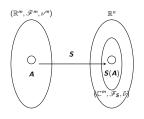
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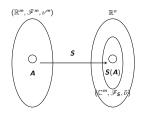


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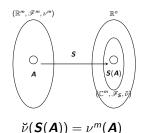
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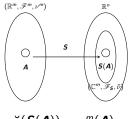
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$$\begin{split} & \boldsymbol{\check{y}}_{t+h} = [\check{y}_{Tot,t+h}, \check{y}_{A,t+h}, \check{y}_{B,t+h}] \\ & \check{y}_{Tot,t+h} = \check{y}_{A,t+h} + \check{y}_{B,t+h} \end{split}$$

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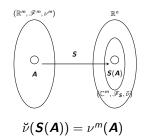
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$$f(\breve{y}_{t+h}) = 0 \quad \forall \breve{y}_{t+h} \in \mathbb{N}^{n-m}$$

#### Definition 4: Point forecast reconciliation

Let  $\hat{\mathbf{y}}_{t+h} \in \mathbb{R}^n$  be any set of incoherent forecasts at time t+h, and let

$$\tilde{\mathbf{y}}_{t+h} = \mathbf{S} \circ \mathbf{g}(\hat{\mathbf{y}}_{t+h}),$$

where  $\mathbf{g}: \mathbb{R}^n \to \mathbb{R}^m$  and  $\mathbf{S} \circ \mathbf{g}(\cdot)$  is a mapping of  $\mathbf{g}(\cdot)$  onto  $\mathbb{C}^m$ . Then  $\tilde{\mathbf{y}}_{t+h}$  is said to be "reconciled" if  $\tilde{\mathbf{y}}_{t+h} \in \mathbb{C}^m$ .

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- Definition 4 holds for linear as well as non-linear reconciliation
- Previous studies have only focused on linear case,

$$g(\hat{y}) = P\hat{y},$$

where P is a  $m \times n$  matrix. Thus giving

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$$\tilde{\mathbf{y}}_{t+h} = \mathbf{SP}\hat{\mathbf{y}}_{t+h}$$

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• Suppose  $S = \{s_1, \ldots, s_m\}$  spans  $\mathbb{C}^m$ ,  $R = \{r_1, \ldots, r_{n-m}\}$  spans  $\mathbb{R}^{n-m}$ . Then  $(S : R) = B = \{s_1, \ldots, s_m, r_1, \ldots, r_{n-m}\}$  spans  $\mathbb{R}^n$ .

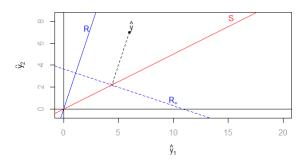
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■ Step 1: Change of coordinates (linear transformation step)

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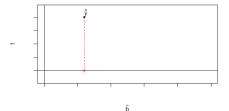
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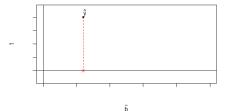
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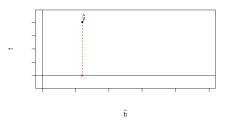


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■ Step 1: Change of coordinates (linear transformation step)

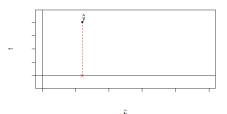
$$( ilde{m{b}}_{t+h}', m{t}_{t+h}')'$$
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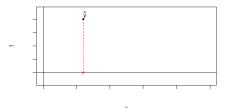


$$(\mathbf{S} \stackrel{.}{:} \mathbf{R})(\tilde{\mathbf{b}}'_{t+h}, \ \mathbf{t}'_{t+h})' = \hat{\mathbf{y}}_{t+h},$$
  $(\tilde{\mathbf{b}}'_{t+h}, \ \mathbf{t}'_{t+h})' = (\mathbf{S} \stackrel{.}{:} \mathbf{R})^{-1}\hat{\mathbf{y}}_{t+h}.$ 

$$(\tilde{\boldsymbol{b}}'_{t+h}, \ \boldsymbol{t}'_{t+h})' = (\boldsymbol{S} \ \dot{\cdot} \ \boldsymbol{R})^{-1} \hat{\boldsymbol{y}}_{t+h}$$

Step 1: Change of coordinates (linear transformation step)

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$$(oldsymbol{\mathcal{S}}\stackrel{.}{:}oldsymbol{\mathcal{R}})( ilde{oldsymbol{b}}'_{t+h},\ oldsymbol{t}'_{t+h},\ oldsymbol{t}'_{t+h},\ oldsymbol{t}'_{t+h})'=(oldsymbol{\mathcal{S}}\stackrel{.}{:}oldsymbol{\mathcal{R}})^{-1}\hat{oldsymbol{\mathcal{Y}}}_{t+h}.$$

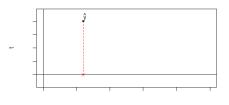
$$(\tilde{oldsymbol{b}}_{t+h}^\prime,\ oldsymbol{t}_{t+h}^\prime)^\prime = (oldsymbol{S}\ \dot{oldsymbol{R}})^{-1}\hat{oldsymbol{y}}_{t+h}$$

Since 
$$(\mathbf{\textit{S}} \stackrel{:}{:} \mathbf{\textit{R}})^{-1} = \begin{pmatrix} (\mathbf{\textit{R}}'_{\perp}\mathbf{\textit{S}})^{-1}\mathbf{\textit{R}}'_{\perp} \\ & \ddots \\ (\mathbf{\textit{S}}'_{\perp}\mathbf{\textit{R}})^{-1}\mathbf{\textit{S}}'_{\perp} \end{pmatrix}$$
 we have,

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Step 1: Change of coordinates (linear transformation step)

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$$(\mathbf{S}\stackrel{.}{:}\mathbf{R})(\tilde{\mathbf{b}}'_{t+h},\ \mathbf{t}'_{t+h})'=\hat{\mathbf{y}}_{t+h},$$
  $(\tilde{\mathbf{b}}'_{t+h},\ \mathbf{t}'_{t+h})'=(\mathbf{S}\stackrel{.}{:}\mathbf{R})^{-1}\hat{\mathbf{y}}_{t+h}.$ 

$$(\tilde{\boldsymbol{b}}'_{t+h},\ \boldsymbol{t}'_{t+h})' = (\boldsymbol{S} \ \dot{\boldsymbol{c}} \ \boldsymbol{R})^{-1} \hat{\boldsymbol{y}}_{t+h}$$

Since 
$$(\mathbf{S} \stackrel{.}{:} \mathbf{R})^{-1} = \begin{pmatrix} (\mathbf{R}_{\perp}' \mathbf{S})^{-1} \mathbf{R}_{\perp}' \\ \cdots \\ (\mathbf{S}_{\perp}' \mathbf{R})^{-1} \mathbf{S}_{\perp}' \end{pmatrix}$$
 we have,

$$egin{pmatrix} \hat{oldsymbol{b}}_{t+h} \ \cdots \ oldsymbol{t}_{t+h} \end{pmatrix} = egin{pmatrix} (oldsymbol{R}'_oldsymbol{S})^{-1}oldsymbol{R}'_oldsymbol{L} \ \cdots \ (oldsymbol{S}'_oldsymbol{L}oldsymbol{R})^{-1}oldsymbol{S}'_oldsymbol{L} \end{pmatrix} \hat{oldsymbol{y}}_{t+h}.$$

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■ Step 2: Ignoring the null coordinates

$$\tilde{\boldsymbol{b}}_{t+h} = (\boldsymbol{R}_{\perp}'\boldsymbol{S})^{-1}\boldsymbol{R}_{\perp}'\hat{\boldsymbol{y}}_{t+h}. \tag{1}$$

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■ Step 3: Reconciled point forecasts for the whole hierarchy

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Step 2: Ignoring the null coordinates

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- Step 3: Reconciled point forecasts for the whole hierarchy
  - Mapping  $\tilde{\boldsymbol{b}}_{t+h} \in \mathbb{R}^m$  to  $\mathbb{C}^m$  through  $\boldsymbol{S}$  we get

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Step 2: Ignoring the null coordinates

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$$\tilde{\mathbf{y}}_{t+h} = \mathbf{S}(\mathbf{R}_{\perp}'\mathbf{S})^{-1}\mathbf{R}_{\perp}'\hat{\mathbf{y}}_{t+h}, \qquad \tilde{\mathbf{y}}_{t+h} \in \mathbb{C}^m \subset \mathbb{R}^n.$$
 (2)

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Step 2: Ignoring the null coordinates

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 (2)

 $m{P}$  Previous definitions on point forecast reconciliation coincides if  $m{P} = (m{R}'_\perp m{S})^{-1} m{R}'_\perp$ 

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Step 2: Ignoring the null coordinates

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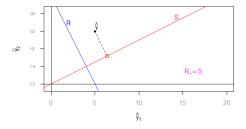
- Previous definitions on point forecast reconciliation coincides if  $P = (R'_{\perp} S)^{-1} R'_{\perp}$
- Further we find  $R_{\perp}$  s.t  $(R'_{\perp}S)^{-1}R'_{\perp}S = I$ . This coincides with the unbiased condition SPS = S (Hyndman et al., 2011)

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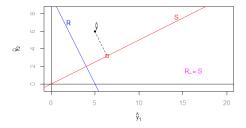
OLS Reconciliation

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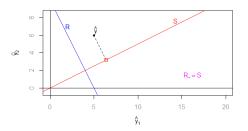
#### OLS Reconciliation



#### OLS Reconciliation



#### OLS Reconciliation



Method	Reconciled point forecasts	$R'_{\perp}$
OLS	$ ilde{oldsymbol{y}}_{t+h}^{OLS} = oldsymbol{\mathcal{S}}(oldsymbol{\mathcal{S}}'oldsymbol{\mathcal{S}})^{-1}oldsymbol{\mathcal{S}}'\hat{oldsymbol{y}}_{t+h}$	S'
MinT	$ ilde{oldsymbol{y}}_{t+h}^{ extit{Min} extsf{T}} = oldsymbol{S}(oldsymbol{S}'oldsymbol{W}_h^{-1}oldsymbol{S})^{-1}oldsymbol{S}'oldsymbol{W}_h^{-1}\hat{oldsymbol{y}}_{t+h}$	$S'W_h^{-1}$

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#### Definition 5: Probabilistic forecast reconciliation

Suppose  $(\mathbb{R}^n, \mathscr{F}^n, \hat{\nu})$  is an incoherent probability triple and  $(\mathbb{R}^m, \mathscr{F}^m, \nu^m)$  is a probability triple defined on  $\mathbb{R}^m$ . Let  $\mathbf{g}: \mathbb{R}^n \to \mathbb{R}^m$ . Then the probability measure on the reconciled bottom levels is such that

$$u^m(\mathbf{A}) = \hat{\nu}(\mathbf{g}^{-1}(\mathbf{A})), \quad \forall \quad \mathbf{A} \in \mathscr{F}^m.$$

Further the probability measure of the whole reconciled hierarchy is given by

$$ilde{
u}(oldsymbol{S}(oldsymbol{A})) = \hat{
u}(oldsymbol{g}^{-1}(oldsymbol{A})) \qquad orall oldsymbol{A} \in \mathscr{F}^m,$$

where  $\boldsymbol{S}:\mathbb{R}^m \to \mathbb{C}^m$ ,  $\tilde{\nu}(\cdot)$  is the probability measure on the measure space  $(\mathbb{C}^m, \mathscr{F}_S)$  and  $\boldsymbol{g}^{-1}(\boldsymbol{A})$  is the pre-image of  $\boldsymbol{A}$  in  $\mathbb{R}^n$ 

#### Definition 5: Probabilistic forecast reconciliation

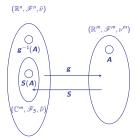
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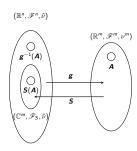
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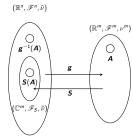


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$$u^m(\mathbf{A}) = \hat{\nu}(\mathbf{g}^{-1}(\mathbf{A})), \quad \forall \mathbf{A} \in \mathscr{F}^m.$$

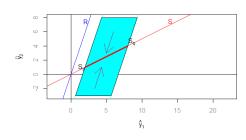
$$\tilde{\nu}(\mathbf{S}(\mathbf{A})) = \hat{\nu}(\mathbf{g}^{-1}(\mathbf{A})) \qquad \forall \mathbf{A} \in \mathscr{F}^m.$$



$$u^m(\mathbf{A}) = \hat{\nu}(\mathbf{g}^{-1}(\mathbf{A})), \quad \forall \mathbf{A} \in \mathscr{F}^m.$$

$$\tilde{\nu}(\boldsymbol{S}(\boldsymbol{A})) = \hat{\nu}(\boldsymbol{g}^{-1}(\boldsymbol{A})) \qquad \forall \boldsymbol{A} \in \mathscr{F}^m.$$

■ Back to 2-dimensional schematic



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■ **Step 1: Change of coordinates** (linear transformation step)

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Recall  $(\tilde{b}'_{t+h}, t'_{t+h})'$ : coordinates of  $\hat{y}_{t+h}$  with respect to (S : R).

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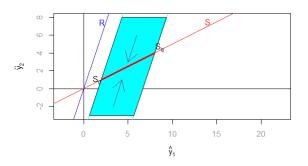
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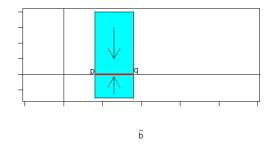
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■ Let  $\hat{f}(\cdot)$  be the probability density of  $\hat{y}_{t+h}$ 

- Let  $\hat{f}(\cdot)$  be the probability density of  $\hat{y}_{t+h}$
- Let  $f_B(\cdot)$  be the density of the coordinates of  $\hat{y}_{t+h}$  with respect to basis B

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- lacktriangle Recall that  $(m{S}\ \dot{m{E}}\ R)( ilde{m{b}}'_{t+h},\ m{t}'_{t+h})'=\hat{m{y}}_{t+h},$  and  $\hat{m{y}}_{t+h}=m{S} ilde{m{b}}_{t+h}+m{R}m{t}_{t+h}$

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■ Step 2: Marginalizing over the null space

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$$f_{\mathcal{B}}(\tilde{m{b}}_{t+h},m{t}_{t+h}) = \hat{m{f}}(m{S}\tilde{m{b}}_{t+h} + m{R}m{t}_{t+h}) \, \Big| m{S} \, \vdots \, m{R} \Big|$$

■ Step 2: Marginalizing over the null space i.e.,

$$\hat{\mathbf{f}}_b(\tilde{\mathbf{b}}_{t+h}) = \int \hat{\mathbf{f}}(\mathbf{S}\tilde{\mathbf{b}}_{t+h} + \mathbf{R}\mathbf{t}_{t+h}) \left| \mathbf{S} \right| \mathbf{R} d\mathbf{t}_{t+h}$$

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$$\hat{f}_b(\tilde{b}_{t+h}) = \int \hat{f}(S\tilde{b}_{t+h} + Rt_{t+h}) \left| S : R \right| dt_{t+h}$$

■ Step 3: Reconciled density of the whole hierarchy

$$\tilde{\mathbf{f}}(\tilde{\mathbf{y}}_{t+h}) = \tilde{\mathbf{f}}_b(\mathbf{S}\tilde{\mathbf{b}}_{t+h}), \quad [\tilde{\mathbf{y}}_{t+h} = \mathbf{S}\tilde{\mathbf{b}}_{t+h}]$$

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Assuming Gaussianity, let an incoherent forecast distribution at time T + h be given by

$$\mathcal{N}(\hat{\mathbf{y}}_{T+h}, \mathbf{W}_{T+h}) \stackrel{d}{\leftrightarrow} \hat{\mathbf{f}}(.)$$

where  $\hat{\mathbf{y}}_{T+h}$  is the incoherent mean and

$$W_{T+h} = E_{\mathbf{y}_{T+h}}[(\mathbf{y}_{T+h} - \hat{\mathbf{y}}_{T+h})(\mathbf{y}_{T+h} - \hat{\mathbf{y}}_{T+h})^T | \mathcal{I}_T]$$
 is incoherent variance

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 Gaussian densities are closed under affine transformation and marginalization

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- Gaussian densities are closed under affine transformation and marginalization
- From the linear transformation

$$f_{B}(\cdot) = \frac{\exp\left\{-\frac{1}{2}\left((\boldsymbol{s} \stackrel{:}{\cdot} \boldsymbol{R})\left(\tilde{\boldsymbol{b}}_{T+h}^{T+h}\right) - \hat{\boldsymbol{y}}_{T+h}\right)'\boldsymbol{w}_{T+h}^{-1}\left((\boldsymbol{s} \stackrel{:}{\cdot} \boldsymbol{R})\left(\tilde{\boldsymbol{b}}_{T+h}^{T+h}\right) - \hat{\boldsymbol{y}}_{T+h}\right)\right\}}{(2\pi)^{\frac{n}{2}}\left|\boldsymbol{w}_{T+h}\right|^{\frac{1}{2}}\left|(\boldsymbol{s} \stackrel{:}{\cdot} \boldsymbol{R})^{-1}\right|}$$

■ Marginalizing over  $t_{T+h}$  🗚

$$\tilde{\mathbf{f}}_b(\tilde{\mathbf{b}}_{T+h}) = \frac{\exp\left\{-\frac{1}{2}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})'(\mathbf{PW}_{T+h}\mathbf{P}')^{-1}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})\right\}}{(2\pi)^{\frac{n}{2}} \left|\mathbf{PW}_{T+h}\mathbf{P}'\right|^{\frac{1}{2}}}$$

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■ This implies,  $\tilde{\pmb{b}}_{T+h} \sim \mathscr{N}(\pmb{P}\hat{\pmb{y}}_{T+h}, \pmb{P}\pmb{W}_{T+h}\pmb{P}')$  where  $\pmb{P} = (\pmb{R}'_{\perp}\pmb{S})^{-1}\pmb{R}'_{\perp}$ 

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■ Marginalizing over  $t_{T+h}$  🔼

$$\tilde{\mathbf{f}}_b(\tilde{\mathbf{b}}_{T+h}) = \frac{\exp\left\{-\frac{1}{2}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})'(\mathbf{PW}_{T+h}\mathbf{P}')^{-1}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})\right\}}{(2\pi)^{\frac{n}{2}} \left|\mathbf{PW}_{T+h}\mathbf{P}'\right|^{\frac{1}{2}}}$$

- This implies,  $\tilde{\boldsymbol{b}}_{T+h} \sim \mathcal{N}(\boldsymbol{P}\hat{\boldsymbol{y}}_{T+h}, \boldsymbol{PW}_{T+h}\boldsymbol{P}')$  where  $\boldsymbol{P} = (\boldsymbol{R}'_{\perp}\boldsymbol{S})^{-1}\boldsymbol{R}'_{\perp}$
- Therefore the reconciled Gaussian distribution of the whole hierarchy is given by

$$ilde{m{f}}( ilde{m{y}}_{t+h}) = ilde{m{f}}_b(m{S} ilde{m{b}}_{T+h}) \overset{d}{\leftrightarrow} \mathscr{N}(m{S}m{P}\hat{m{y}}_{T+h}, \ m{S}m{P}m{W}_{T+h}m{P}'m{S}'),$$
 where  $m{P} = (m{R}_+'m{S})^{-1}m{R}_+'$ 

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■ Marginalizing over  $t_{T+h}$  🔼

$$\tilde{\mathbf{f}}_b(\tilde{\mathbf{b}}_{T+h}) = \frac{\exp\left\{-\frac{1}{2}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})'(\mathbf{PW}_{T+h}\mathbf{P}')^{-1}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})\right\}}{(2\pi)^{\frac{n}{2}} \left|\mathbf{PW}_{T+h}\mathbf{P}'\right|^{\frac{1}{2}}}$$

- This implies,  $\tilde{\pmb{b}}_{T+h} \sim \mathscr{N}(\pmb{P}\hat{\pmb{y}}_{T+h}, \pmb{P}\pmb{W}_{T+h}\pmb{P}')$  where  $\pmb{P} = (\pmb{R}'_{\perp}\pmb{S})^{-1}\pmb{R}'_{\perp}$
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Puwasala Gamakumara Probabilistic Hierarchical

■ Marginalizing over  $t_{T+h}$  All

$$\tilde{\mathbf{f}}_b(\tilde{\mathbf{b}}_{T+h}) = \frac{\exp\left\{-\frac{1}{2}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})'(\mathbf{PW}_{T+h}\mathbf{P}')^{-1}(\tilde{\mathbf{b}}_{T+h} - \mathbf{P}\hat{\mathbf{y}}_{T+h})\right\}}{(2\pi)^{\frac{n}{2}} \left|\mathbf{PW}_{T+h}\mathbf{P}'\right|^{\frac{1}{2}}}$$

- This implies,  $\hat{\boldsymbol{b}}_{T+h} \sim \mathcal{N}(\boldsymbol{P}\hat{\boldsymbol{y}}_{T+h}, \boldsymbol{P}\boldsymbol{W}_{T+h}\boldsymbol{P}')$  where  $P = (R'_{\perp} S)^{-1} R'_{\perp}$
- Therefore the reconciled Gaussian distribution of the whole hierarchy is given by

$$\tilde{\mathbf{f}}(\tilde{\mathbf{y}}_{t+h}) = \tilde{\mathbf{f}}_b(\mathbf{S}\tilde{\mathbf{b}}_{T+h}) \stackrel{d}{\leftrightarrow} \mathscr{N}(\mathbf{S}\mathbf{P}\hat{\mathbf{y}}_{T+h}, \ \mathbf{S}\mathbf{P}\mathbf{W}_{T+h}\mathbf{P}'\mathbf{S}'),$$
 where  $\mathbf{P} = (\mathbf{R}'_{\perp}\mathbf{S})^{-1}\mathbf{R}'_{\perp}$ 

■ Goal: Solve for P by minimizing a proper scoring rule

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■ We use the Energy Score, which is a proper scoring rule

$$ES(\tilde{\mathbf{Y}}_{T+h}, \mathbf{y}_{T+h}) = E_{\mathbf{Y}_{T+h}} ||\tilde{\mathbf{Y}}_{T+h} - \mathbf{y}_{T+h}||^{\alpha} - \frac{1}{2} E_{\mathbf{Y}_{T+h}} ||\tilde{\mathbf{Y}}_{T+h} - \tilde{\mathbf{Y}}_{T+h}^*||^{\alpha},$$
for  $\alpha \in (0, 2]$ ,

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■ We use the Energy Score, which is a proper scoring rule

 $\mathcal{N}(SP\hat{\mathbf{v}}_{T+h}, SPW_{T+h}P'S')$ .

$$ES(\tilde{\mathbf{Y}}_{T+h}, \mathbf{y}_{T+h}) = E_{\mathbf{Y}_{T+h}} ||\tilde{\mathbf{Y}}_{T+h} - \mathbf{y}_{T+h}||^{\alpha} - \frac{1}{2} E_{\mathbf{Y}_{T+h}} ||\tilde{\mathbf{Y}}_{T+h} - \tilde{\mathbf{Y}}_{T+h}^*||^{\alpha},$$
 for  $\alpha \in (0, 2]$ , where  $\tilde{\mathbf{Y}}_{T+h}$  and  $\tilde{\mathbf{Y}}_{T+h}^*$  are independent random variables from

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 for  $\alpha \in (0,2]$ , where  $\tilde{\boldsymbol{Y}}_{T+h}$  and  $\tilde{\boldsymbol{Y}}_{T+h}^*$  are independent random variables from  $\mathcal{N}(\boldsymbol{SP}\hat{\boldsymbol{y}}_{T+h}, \boldsymbol{SPW}_{T+h} \boldsymbol{P'S'}).$ 

- There is no closed form expression for  $ES(\tilde{Y}_{T+h}, y_{T+h})$  for  $\alpha \in (0, 2)$  under the Gaussian predictive distribution
- However in the upper limit of  $\alpha$ , i.e. when  $\alpha = 2$ ,

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$$ES(\tilde{\mathbf{Y}}_{T+h}, \mathbf{y}_{T+h}) = ||\mathbf{SP}\hat{\mathbf{y}}_{T+h} - \mathbf{y}_{T+h}||^2$$

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■ Then our objective function is

$$\operatorname*{arg\,min}_{\boldsymbol{P}} \mathcal{T}r\{E_{\boldsymbol{y}_{T+h}}[(\boldsymbol{y}_{T+h}-\boldsymbol{SP}\hat{\boldsymbol{y}}_{T+h})(\boldsymbol{y}_{T+h}-\boldsymbol{SP}\hat{\boldsymbol{y}}_{T+h})^T|\mathcal{I}_T]\},$$
 where  $\mathcal{I}_T=\{\boldsymbol{y}_1,....,\boldsymbol{y}_T\}$ 

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$$\underset{\boldsymbol{P}}{\text{arg min }} Tr\{E_{\boldsymbol{y}_{T+h}}[(\boldsymbol{y}_{T+h}-\boldsymbol{SP}\hat{\boldsymbol{y}}_{T+h})(\boldsymbol{y}_{T+h}-\boldsymbol{SP}\hat{\boldsymbol{y}}_{T+h})^T|\mathcal{I}_T]\},$$

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Assuming unbiasedness for coherent forecasts

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where  $\mathcal{I}_{\mathcal{T}} = \{ \mathbf{\textit{y}}_1, ...., \mathbf{\textit{y}}_{\mathcal{T}} \}$ 

Assuming unbiasedness for coherent forecasts

$$\underset{\boldsymbol{P}}{\operatorname{arg\,min}} Tr\{\boldsymbol{SPW}_{T+h}\boldsymbol{P}^T\boldsymbol{S}^T\}$$
Subject to  $\boldsymbol{SPS} = \boldsymbol{S}$ 

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■ Wickramasuriya, Athanasopoulos, and Hyndman (2018) have shown that a unique solution to this optimization problem attain at

$$\boldsymbol{P} = (\boldsymbol{S}^T \boldsymbol{W}_{T+h}^{-1} \boldsymbol{S})^{-1} \boldsymbol{S}^T \boldsymbol{W}_{T+h}^{-1}.$$

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Then our objective function is

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where  $\mathcal{I}_{T} = \{ y_{1}, ...., y_{T} \}$ 

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■ Thus  $R'_{\perp} = S'W_{T+h}^{-1}$ . This is referred to as MinT solution

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$$R'_{\perp} = S'W_{T+h}^{-1}$$

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$$R'_{\perp} = S'W_{T+h}^{-1}$$

■ The optimally reconciled Gaussian forecast density is given by  $\mathcal{N}[S(R'_{\perp}S)^{-1}R'_{\perp}\hat{y}_{T+h}, S(R'_{\perp}S)^{-1}R'_{\perp}W_{T+h}R_{\perp}(R'_{\perp}S)^{-1}S'],$ 

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- lacktriangle Different estimators of  $oldsymbol{W}_{T+h}$  yield different  $oldsymbol{R}'_{\perp}$  matrices

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- Different estimators of  $W_{T+h}$  yield different  $R'_{\perp}$  matrices

Method	Estimate of $W_h$	Estimate of $R'_{\perp}$
OLS MinT(Sample) MinT(Shrink) MinT(WLS)	$oldsymbol{\hat{N}}_{T+1}^{sam} \ \hat{oldsymbol{W}}_{T+1}^{shr} =  au  ext{Diag}(\hat{oldsymbol{W}}_{T+1}^{sam}) + (1- au)\hat{oldsymbol{W}}_{T+1}^{sam}$ A2 $\hat{oldsymbol{W}}_{T+1}^{wls} =  ext{Diag}(\hat{oldsymbol{W}}_{T+1}^{shr})$	$egin{array}{c} m{S}' \ m{S}'(\hat{m{W}}_{T+1}^{sam})^{-1} \ m{S}'(\hat{m{W}}_{T+1}^{shr})^{-1} \ m{S}'(\hat{m{W}}_{T+1}^{wls})^{-1} \end{array}$

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$$R'_{\perp} = S'W_{T+h}^{-1}$$

- The optimally reconciled Gaussian forecast density is given by  $\mathcal{N}[S(R'_{\perp}S)^{-1}R'_{\perp}\hat{y}_{T+h},S(R'_{\perp}S)^{-1}R'_{\perp}W_{T+h}R_{\perp}(R'_{\perp}S)^{-1}S'],$
- Different estimators of  $W_{T+h}$  yield different  $R'_{\perp}$  matrices

Method	Estimate of $W_h$	Estimate of $R'_{\perp}$
OLS MinT(Sample) MinT(Shrink) MinT(WLS)	$oldsymbol{\hat{N}}_{T+1}^{sam} \ \hat{oldsymbol{W}}_{T+1}^{shr} =  au  ext{Diag}(\hat{oldsymbol{W}}_{T+1}^{sam}) + (1- au)\hat{oldsymbol{W}}_{T+1}^{sam}$ A2 $\hat{oldsymbol{W}}_{T+1}^{wls} =  ext{Diag}(\hat{oldsymbol{W}}_{T+1}^{shr})$	$egin{array}{c} m{S}' \ m{S}'(\hat{m{W}}_{T+1}^{sam})^{-1} \ m{S}'(\hat{m{W}}_{T+1}^{shr})^{-1} \ m{S}'(\hat{m{W}}_{T+1}^{wls})^{-1} \end{array}$

March 21, 2018

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Method	Estimate of $W_h$	Estimate of $R'_{\perp}$
OLS	1	<b>S</b> '
MinT(Sample)	$\hat{W}_{T+1}^{sam}$	$oldsymbol{\mathcal{S}'}(\hat{oldsymbol{\mathcal{W}}}_{T+1}^{sam})^{-1}$
MinT(Shrink)	$\hat{m{W}}_{T+1}^{shr} =  au$ Diag $(\hat{m{W}}_{T+1}^{sam}) + (1- au)\hat{m{W}}_{T+1}^{sam}$ A2	$S'(\hat{W}_{T+1}^{shr})^{-1}$
MinT(WLS)	$\hat{m{W}}_{T+1}^{wls} = Diag(\hat{m{W}}_{T+1}^{shr})$	$S'(\hat{W}_{T+1}^{wls})^{-1}$

■ Predictive ability of the reconciled Gaussian densities from these methods will be evaluated in a simulation setting

### Evaluating probabilistic forecasts

Energy score (Gneiting et al., 2008) 
$$= \mathsf{E}_{\breve{F}} \| \breve{\mathbf{Y}}_{T+h} - \mathbf{y}_{T+h} \|^{\alpha} - \frac{1}{2} \mathsf{E}_{\breve{F}} \| \breve{\mathbf{Y}}_{T+h} - \breve{\mathbf{Y}}_{T+h}^{*} \|^{\alpha}, \quad \alpha \in (0,2]$$
 Log score (Gneiting and Raftery, 2007) 
$$\mathsf{LS}(\breve{F}, \mathbf{y}_{T+h}) = -\log \breve{\mathbf{f}}(\mathbf{y}_{T+h})$$

$$VS(\breve{F}, y_{T+h}) = \sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij} (|y_{T+h,i} - y_{T+h,j}|^{p} - E_{\breve{F}} |\breve{Y}_{T+h,i} - \breve{Y}_{T+h,j}|^{p})^{2}$$

CRPS (Gneiting and Raftery, 2007)  
CRPS(
$$\check{F}_i, y_{T+h,i}$$
) =  $\mathsf{E}_{\check{F}_i} |\check{Y}_{T+h,i} - y_{T+h,i}| - \frac{1}{2} \mathsf{E}_{\check{F}_i} |\check{Y}_{T+h,i} - \check{Y}_{T+h,i}^*|$ 

 $reve{Y}_{T+h}$  and  $reve{Y}_{T+h}^*$ Independent random vectors from the coherent forecast distribution **F**.

Vector of realizations.  $\mathbf{y}_{T+h}$ 

 $oldsymbol{y}_{T+h}$  :  $oldsymbol{Y}_{T+h,i}$  and  $oldsymbol{Y}_{T+h,j}$  : ith and ith components of the vector  $\mathbf{Y}_{T+h}$ Non-negative weights Wii

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■ To compare two competitive forecasting methods, we use Skill score (Gneiting and Raftery, 2007)

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■ To compare two competitive forecasting methods, we use Skill score (Gneiting and Raftery, 2007)

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- To compare two competitive forecasting methods, we use Skill score (Gneiting and Raftery, 2007)
- For a given forecasting method, evaluated by a particular scoring rule  $S(\cdot)$ , the skill score is given by

$$SS[S_B(\cdot)] = \frac{S_B(\boldsymbol{Y}, \boldsymbol{y})^{\text{ref}} - S_B(\boldsymbol{\check{Y}}, \boldsymbol{y})}{S_B(\boldsymbol{Y}, \boldsymbol{y})^{\text{ref}}} \times 100\%,$$

#### where

 $S_B(\cdot)$  : Average score over B samples  $S_B(\mathbf{Y},\mathbf{y})^{\mathrm{ref}}$  : Average score of the reference forecasting methods

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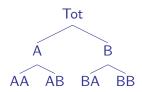
where

 $S_B(\cdot)$  : Average score over B samples  $S_B(\mathbf{Y},\mathbf{y})^{\mathrm{ref}}$  : Average score of the reference forecasting methods

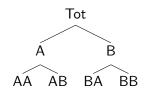
■  $SS[S_B(\cdot)]$  gives the percentage improvement of the preferred forecasting method relative to the reference method

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■ Data generating process A3



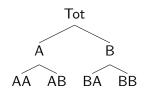
Data generating process A3



■ DGP was designed such that we have noisier series in the bottom levels than in the aggregate levels

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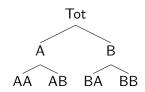
Data generating process



- DGP was designed such that we have noisier series in the bottom levels than in the aggregate levels
- Simulation setup

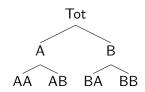
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Data generating process



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- Simulation setup
  - 501 observations were generated

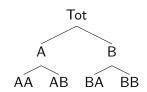
Data generating process A3



- DGP was designed such that we have noisier series in the bottom levels than in the aggregate levels
- Simulation setup
  - 501 observations were generated
  - Univariate ARIMA models were fitted to the first 500 observations and
     1- step ahead incoherent forecasts were generated

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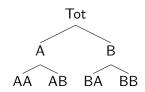
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  - $\bullet$  Mean and variance forecasts of incoherent Gaussian densities were reconciled through different estimates of  ${\pmb R}'_\perp$

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     1- step ahead incoherent forecasts were generated
  - Mean and variance forecasts of incoherent Gaussian densities were reconciled through different estimates of  $R'_{\perp}$
  - This process was replicated using 1000 different series from the same (DGP)

#### Monte-Carlo simulation Cont..

■ Comparison of reconciled vs bottom-up Gaussian forecasts

Forecasting	Energ	y score	Log	score	Variogram score		
method	Average score	Skill score (%)	Average score	Skill score (%)	Average score	Skill score (%)	
MinT(Shrink)	7.47	10.11	11.34	6.44	3.05	4.69	
MinT(Sample)	7.47	10.11	11.33	6.52	3.05	4.69	
MinT(WLS)	7.91	4.81	12.64	-4.29	3.23	-0.94	
OLS	10.14	-22.02	135.13	-1014.93	4.60	-43.75	
Bottom up	8.31		12.12		3.20		
Note:		tive) entry in "Sk asting method ov			ne percentage incr	rease(/decrease) of	

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#### Monte-Carlo simulation Cont...

#### ■ Comparison of coherent vs incoherent Gaussian forecasts

Forecasting	-	Total	S	eries - A	Se	Series - B		
method	CRPS	LogS	CRPS	LogS	CRPS	LogS		
MinT(Shrink)	1.12	0.34	10.07	2.93	5.41	1.52		
MinT(Sample)	1.12	0.34	10.07	2.93	5.41	1.52		
MinT(WLS)	-2.61	-2.02	5.28	-4.40	2.70	-4.24		
OLS	-38.06	-698.99	-24.70	-1368.33	-24.86	-1159.09		
Bottom up	-89.55	-21.83	-8.87	-2.35	-9.46	-2.73		
Incoherent	2.68	2.97	4.17	3.41	3.70	3.30		

Forecasting	Seri	Series - AA		Series - AB		ies - BA	Series - BB	
method	CRPS	LogS	CRPS	LogS	CRPS	LogS	CRPS	LogS
MinT(Shrink)	8.71	2.71	10.57	3.04	5.95	1.86	7.91	2.46
MinT(Sample)	8.71	2.71	10.57	3.04	5.95	1.86	8.19	2.46
MinT(WLS)	5.54	0.30	5.96	0.30	2.43	-0.62	5.08	0.62
OLS	-22.43	-931.63	-22.49	-886.32	-26.01	-834.67	-23.45	-812.92
Incoherent	3.79	3.32	3.69	3.29	3.46	3.23	3.54	3.25

Note: "Incoherent" row represents the average score for incoherent forecasts.

Each entry above this row represent the percentage skill score with reference to the Incoherent forecasts.

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- 4 Generate h-step ahead sample paths from the fitted models incorporating  $\mathbf{E}^b$ . Denote these by  $\mathbf{y}_{T+h}^b$

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- **5** Repeat step 3-5 for J times
- $\tilde{y}_{T+h}^b = S(R'_{\perp}S)^{-1}R'_{\perp}y_{T+h}^b,$  where  $\tilde{y}_{T+h}^b$  denotes h-step ahead coherent sample paths

■ Simulation setup

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#### Simulation setup

• 1510 observations were generated

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- 1510 observations were generated
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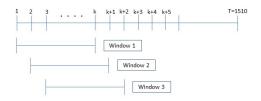
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#### Comparison of reconciled vs bottom-up forecasts

Forecasting	h=	h=1		h=2		h=3		h=4		h=5	
method	ES	VS	ES	VS	ES	VS	ES	VS	ES	VS	
MinT(Shrink)	6.39	5.83	6.48	5.54	5.62	5.43	3.94	3.98	3.45	3.64	
MinT(Sample)	6.39	5.83	6.35	5.19	5.29	4.89	3.47	3.09	3.04	2.73	
MinT(WLS)	3.83	0.89	3.89	1.04	3.35	1.09	2.44	1.33	2.06	1.28	
OLS	3.35	0	3.24	0.35	2.69	0.27	2.06	0.66	1.73	0.73	
Bottom Up	6.26	2.23	7.72	2.89	9.26	3.68	10.67	4.52	12.16	5.49	
"Bottom-up" row represent the average score for incoherent forecasts.											
Note:	Each e	ntry abo	ve this r	Each entry above this row represent the percentage skill score with							

reference to the Bottom-up forecasts.

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#### Comparison of coherent vs incoherent forecasts

Note:

Forecasting method	Total	Series-A	Series-B	Series-AA	Series-AB	Series-BA	Series-BB
				h=1			
MinT(Shrink)	0.28	7.17	6.57	10.04	2.93	7.35	0.41
MinT(Sample)	0.29	7.10	6.45	10.04	3.02	7.22	0.32
MinT(WLS)	-1.78	6.46	5.80	5.17	-0.03	6.37	-0.68
OLS	-6.30	5.62	5.49	4.50	0.09	5.42	-1.59
Bottom Up	-69.68	-5.25	-4.51	0.00	0.00	0.00	0.00
Incoherent	1.86	2.73	2.68	2.93	2.78	2.98	2.35
				h=5			
MinT(Shrink)	0.63	6.30	0.67	7.56	0.35	1.63	-2.01
MinT(Sample)	0.56	6.39	0.58	7.15	0.02	1.29	-3.47
MinT(WLS)	1.24	5.45	0.62	5.39	-3.10	2.41	-3.36
OLS `	1.16	5.07	1.20	4.64	-2.78	2.43	-5.33
Bottom Up	-9.44	-0.52	-1.42	0.00	0.00	0.00	0.00
Incoherent	8.72	7.46	5.96	6.78	4.66	5.95	2.68

"Incoherent" row represent the average score for incoherent forecasts.

Each entry above this row represent the percentage skill score with
reference to the Incoherent forecasts.

 Domestic tourism in Australia can be disaggregated based on a geographical hierarchy

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 Domestic tourism in Australia can be disaggregated based on a geographical hierarchy

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 Domestic tourism in Australia can be disaggregated based on a geographical hierarchy

#### Geographical hierarchical structure for Australia

Level	No.Series per level
Total (Australia)	1
Level-1 (States)	7
Level-2 (Zones)	27
Level-3 (Regions)	75

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 Domestic tourism in Australia can be disaggregated based on a geographical hierarchy

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- Reconciliation of point forecasts (Athanasopoulos, Ahmed, and Hyndman, 2009, Hyndman et al., 2011, Wickramasuriya, Athanasopoulos, and Hyndman, 2018)
- We use a non-parametric bootstrap approach to obtain reconciled probabilistic forecasts

■ We consider monthly "overnight trips" over the period January 1998 - December 2016

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- Univariate ARIMA models were fitted using a rolling window of 100 observations
- 5000 reconciled future sample paths were generated for forecast horizons h=1,...,5

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- We consider monthly "overnight trips" over the period January 1998 -December 2016
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#### Analysis:

- Univariate ARIMA models were fitted using a rolling window of 100 observations
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- The training window was rolled one observation ahead and the process repeated until we get 123 reconciled future sample paths

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#### ■ Comparison of reconciled vs. bottom-up probabilistic forecasts

Forecasting	h	=1	h	h=2		=3	h	h=4		=5	
method	ES	VS	ES	VS	ES	VS	ES	VS	ES	VS	
					States						
MinT(Shrink)	19.58	27.95	14.65	23.12	12.05	21.25	7.76	14.09	9.50	15.22	
MinT(WLS)	4.50	6.58	4.37	6.31	3.32	6.01	1.89	3.52	3.58	4.67	
OLS	7.00	3.15	5.58	5.01	2.65	3.31	1.83	-0.66	3.58	2.01	
Bottom Up	345.63	271.05	348.68	281.63	355.98	290.47	367.41	300.25	364.81	299.75	
		Zones									
MinT(Shrink)	13.11	19.30	10.94	19.48	9.40	18.24	6.99	15.10	7.62	15.26	
MinT(WLS)	2.32	1.99	2.55	2.99	1.95	2.61	1.05	1.28	1.27	1.97	
OLS	1.91	-5.82	1.58	-2.36	0.47	-1.78	0.25	-2.07	1.29	-1.57	
Bottom Up	228.70	2366.28	229.54	2440.32	234.09	2510.51	238.98	2556.95	237.98	2552.26	
					Regions						
MinT(Shrink)	10.43	15.96	9.03	16.11	7.87	15.05	6.28	12.81	6.96	13.21	
MinT(WLS)	1.54	1.43	1.75	2.16	1.30	1.83	0.68	0.95	0.97	1.46	
OLS	1.37	-3.01	1.28	-0.77	0.49	-0.90	0.30	-1.08	1.14	-0.18	
Bottom Up	194.19	10956.41	193.50	11145.87	196.26	11355.30	198.25	11461.30	197.76	11438.31	
				ent the ave	_						
Note:	Each ei	ntry above	this ro	v represent	the per	centage ski	ill score	with			

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reference to the Bottom-up forecasts.

### Probabilistic forecasts for Australian domestic tourism flow

■ Comparison of coherent vs. incoherent probabilistic forecasts in total Australia and each states

Forecasting method	Australia	NSW	Victoria	Queensland		Western Australia	Tasmania	Northern Territory
				h=1				
MinT(Shrink)	8.07	9.61	16.84	22.38	23.38	11.25	13.38	20.41
MinT(WLS)	-12.31	-7.54	2.64	4.95	1.12	1.31	6.93	2.55
OLS	1.39	2.96	5.14	7.68	1.58	-0.48	-13.04	-22.4
Bottom Up	-20.00	-16.22	-1.12	3.07	-4.64	-6.52	4.43	-0.96
Incoherent	478.81	188.48	159.49	136.27	51.29	52.97	35.23	24.27
				h=5				
MinT(Shrink)	1.72	6.65	11.96	6.30	14.61	3.66	16.54	24.58
MinT(WLS)	-6.83	-0.30	2.01	2.21	2.26	-1.77	1.40	6.45
OLS `	-0.29	2.20	3.41	4.23	9.50	2.58	9.99	-4.52
Bottom Up	-9.07	-3.33	0.24	1.48	0.44	-7.35	-2.59	4.70
Incoherent	552.45	217.09	170.55	145.01	54.99	59.59	38.76	33.87
Note:		y above t	his row re	he average s present the porecasts.		skill scor		

### Probabilistic forecasts for Australian domestic tourism flow

■ Comparison of coherent vs. incoherent probabilistic forecasts in Zones and Regions

Forecasting			Zones					Regions		
method	h=1	h=2	h=3	h=4	h=5	h=1	h=2	h=3	h=4	h=5
MinT(Shrink)	14.38	11.25	10.20	8.67	9.39	9.96	8.58	7.71	6.53	6.98
MinT(WLS)	2.88	2.15	1.85	1.76	1.75	1.54	1.73	1.42	0.82	0.94
OLS	1.79	0.55	-0.12	1.00	2.22	-0.11	-0.07	-0.51	-0.20	0.74
Bottom Up	0.21	-0.66	-0.43	0.35	0.26					
Incoherent	31.75	31.87	32.71	33.73	33.49	14.36	14.45	14.70	14.87	14.81
	"Incol	nerent"	row rep	resent	the ave	rage sc	ore for ii	ncoheren	t forecas	sts.
Note:	Each entry above this row represent the percentage skill score with reference to the Incoherent forecasts.									

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### Probabilistic forecasts for Australian domestic tourism flow

■ Comparison of coherent vs. incoherent mean forecasts

Forecasting			Australia					States		
method	h=1	h=2	h=3	h=4	h=5	h=1	h=2	h=3	h=4	h=5
MinT(Shrink)	11.09	5.56	4.61	2.17	1.51	16.64	11.28	12.19	6.39	8.10
MinT(WLS)	-9.90	-5.33	-2.17	-0.71	-5.48	0.08	0.26	1.31	-0.16	0.41
OLS	1.62	1.53	1.05	1.16	0.12	2.26	0.47	0.43	-0.22	3.06
Bottom Up	-18.88	-12.33	-7.45	-3.66	-8.23	-5.86	-4.64	-2.60	-2.64	-2.16
Incoherent	669.21	697.51	744.54	812.17	764.24	128.71	132.00	138.92	142.92	142.07
			Zones					Regions		
	h=1	h=2	h=3	h=4	h=5	h=1	h=2	h=3	h=4	h=5
MinT(Shrink)	14.64	11.15	10.71	8.58	9.16	10.87	8.81	8.05	6.71	6.88
MinT(WLS)	2.77	2.13	1.92	1.59	1.68	1.75	1.80	1.54	0.88	0.95
OLS `	1.59	0.55	0.04	0.55	1.80	-0.07	-0.17	-0.53	-0.41	0.31
Bottom Up	-0.09	-0.90	-0.66	0.26	0.28					
Incoherent	43.95	44.38	45.58	47.01	46.53	20.11	20.23	20.58	20.85	20.72
	"Incohere	ent" row re	epresent tl	ne averag	ge MAE	for inco	herent f	orecasts.		
Note:	: Each entry above this row represent the percentage improvement with reference to the Incoherent forecasts.									

# Conclusions and future work

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■ We define probabilistic forecast reconciliation

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- We define probabilistic forecast reconciliation
- Our definition holds for linear and non-linear reconciliation

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Puwasala Gamakumara Probabilistic Hierarchii

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- We define probabilistic forecast reconciliation
- Our definition holds for linear and non-linear reconciliation
- Linear reconciliation involves two steps
  - Step 1: Transform coordinates of incoherent probabilistic forecasts
  - Step 2: Marginalize over the null space of the coherent subspace
- MinT not only produces optimal point forecasts, but also optimally reconcile Gaussian forecast densities

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■ We introduce a novel non-parametric bootstrap approach for producing reconciled probabilistic forecasts

Puwasala Gamakumara Probabilistic

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 We introduce a novel non-parametric bootstrap approach for producing reconciled probabilistic forecasts

- We introduce a novel non-parametric bootstrap approach for producing reconciled probabilistic forecasts
- Results from a simulation study favored MinT for reconciling bootstrapped future paths

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- We introduce a novel non-parametric bootstrap approach for producing reconciled probabilistic forecasts
- Results from a simulation study favored MinT for reconciling bootstrapped future paths
- Finally we apply the non-parametric approach to obtain coherent probabilistic forecasts for domestic tourism flow in Australia

## Time plan

	Thesis Chapter	Task description	Time duration	Progress
2.	Literature Review	Writing the chapter.	May/2018 - July/2018	Incomplete
3.	Probabilistic forecast reconciliation for hierarchical time series	Completing the paper.	Mar/2018 - Apr/2018	90% Completed
4.	Probabilistic forecast reconciliation in non-parametric framework	Methodology. Simulation study. Providing theoretical foundation.	Nov/2016 - Feb/2017 Mar/2017 - June/2017 Apr/2018 - July/2018	Completed Completed Incomplete
5.	Application	Forecasting Australian domestic tourism flow. Forecasting Walmart sales.	Dec/2017 - Mar/2017 Oct/2017 - Mar/2019	Completed

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 Propositionic Horizonic (1700cc) structure
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# Thank You!!

## **Appendix**

**▲** A1

### Appendix

#### **▲** A1

■ By substituting the Gaussian distribution function for  $\hat{f}(\cdot)$  in  $f_B(\cdot)$  we get,

$$f_{B}(\cdot) = \frac{1}{(2\pi)^{\frac{n}{2}} \left| \left( \stackrel{P}{Q} \right) W_{t+h} \left( \stackrel{P}{Q} \right)' \right|^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} \begin{pmatrix} \tilde{b}_{t+h} - P \hat{y}_{t+h} \\ t_{t+h} - Q \hat{y}_{t+h} \end{pmatrix}' \right.$$
$$\left[ \left( \stackrel{P}{Q} \right) W_{t+h} \left( \stackrel{P}{Q} \right)' \right]^{-1} \begin{pmatrix} \tilde{b}_{t+h} - P \hat{y}_{t+h} \\ t_{t+h} - Q \hat{y}_{t+h} \end{pmatrix} \right\}.$$

where

$$(\mathbf{S} \stackrel{:}{:} \mathbf{R})^{-1} = \begin{pmatrix} (\mathbf{R}'_{\perp} \mathbf{S})^{-1} \mathbf{R}'_{\perp} \\ \cdots \\ (\mathbf{S}'_{\perp} \mathbf{R})^{-1} \mathbf{S}'_{\perp} \end{pmatrix} = \begin{pmatrix} \mathbf{P} \\ \mathbf{Q} \end{pmatrix},$$

and  $\mathbf{\textit{P}} = (\mathbf{\textit{R}}'_{\perp} \mathbf{\textit{S}})^{-1} \mathbf{\textit{R}}'_{\perp}$  and  $\mathbf{\textit{Q}} = (\mathbf{\textit{S}}'_{\perp} \mathbf{\textit{R}})^{-1} \mathbf{\textit{S}}'_{\perp}$ .

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### **Appendix**

■ Shrinkage estimator for 1-step ahead base forecast errors

$$\hat{\Sigma}_{T+1}^{shr} = \tau \hat{\Sigma}_{T+1}^D + (1-\tau)\hat{\Sigma}_{T+1},$$

where  $\hat{\Sigma}_{T+1}^D$  is the diagonal matrix comprising diagonal entries of  $\hat{\Sigma}_{T+1}$  and

$$au = rac{\sum_{i 
eq j} \hat{Var}(\hat{r}_{ij})}{\sum_{i 
eq j} \hat{r}_{ij}^2}$$

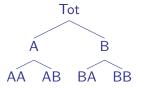
is a shrinkage parameter.  $\hat{r}_{ij}$  is the ij-th element of sample correlation matrix. In this estimation, the off-diagonal elements of 1-step ahead sample covariance matrix will be shrunk to zero depending on the sparsity.

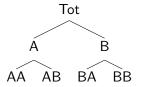
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■ Data generating process ► A3

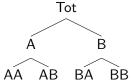
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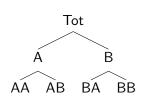
■ Data generating process ► A3



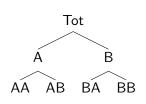


■ Data generating process ► A3

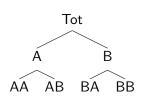




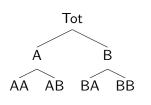
- $\qquad \{ \textit{w}_{AA,t}, \textit{w}_{AB,t}, \textit{w}_{BA,t}, \textit{w}_{BB,t} \} \sim \textit{ARIMA}(\textit{p},\textit{d},\textit{q})$
- $p \in \{1, 2\}$  and  $d \in \{0, 1\}$



- $p \in \{1, 2\}$  and  $d \in \{0, 1\}$
- $\quad \blacksquare \ \left\{ \epsilon_{AA,t}, \epsilon_{AB,t}, \epsilon_{BA,t}, \epsilon_{BB,t} \right\} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$

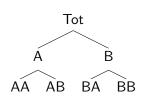


- $p \in \{1, 2\}$  and  $d \in \{0, 1\}$
- $\blacksquare \ \{\epsilon_{\mathsf{AA},t},\epsilon_{\mathsf{AB},t},\epsilon_{\mathsf{BA},t},\epsilon_{\mathsf{BB},t}\} \sim \mathcal{N}(\mathbf{0},\boldsymbol{\Sigma})$
- Parameters for *AR* and *MA* components were randomly and uniformly generated from [0.3, 0.5] and [0.3, 0.7] respectively



- $p \in \{1, 2\}$  and  $d \in \{0, 1\}$
- $\blacksquare \ \{\epsilon_{\mathsf{AA},t},\epsilon_{\mathsf{AB},t},\epsilon_{\mathsf{BA},t},\epsilon_{\mathsf{BB},t}\} \sim \mathcal{N}(\mathbf{0},\boldsymbol{\Sigma})$
- Parameters for *AR* and *MA* components were randomly and uniformly generated from [0.3, 0.5] and [0.3, 0.7] respectively

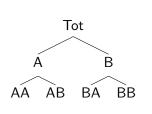
### ■ Data generating process (A3)



- $\qquad \{ w_{AA,t}, w_{AB,t}, w_{BA,t}, w_{BB,t} \} \sim ARIMA(p,d,q)$
- $p \in \{1, 2\} \text{ and } d \in \{0, 1\}$
- $\blacksquare \ \{\epsilon_{\mathsf{AA},t},\epsilon_{\mathsf{AB},t},\epsilon_{\mathsf{BA},t},\epsilon_{\mathsf{BB},t}\} \sim \mathcal{N}(\mathbf{0},\boldsymbol{\Sigma})$
- Parameters for AR and MA components were randomly and uniformly generated from [0.3, 0.5] and [0.3, 0.7] respectively

lacktriangle  $oldsymbol{y}_t$  are then generated as follows

### ■ Data generating process ► A3



- $p \in \{1, 2\} \text{ and } d \in \{0, 1\}$
- $\qquad \{\epsilon_{AA,t},\epsilon_{AB,t},\epsilon_{BA,t},\epsilon_{BB,t}\} \sim \mathcal{N}(\mathbf{0},\boldsymbol{\Sigma})$
- Parameters for *AR* and *MA* components were randomly and uniformly generated from [0.3, 0.5] and [0.3, 0.7] respectively
- $\mathbf{y}_t$  are then generated as follows

Bottom level	Aggregate level 1	Total
$y_{AA,t} = w_{AA,t} + u_t - 0.5v_t$ $y_{AB,t} = w_{AB,t} - u_t - 0.5v_t$ $y_{BA,t} = w_{BA,t} + u_t + 0.5v_t$ $y_{BB,t} = w_{BB,t} - u_t + 0.5v_t$		$y_{Tot,t} = w_{AA,t} + w_{AB,t} + w_{BA,t} + w_{BB,t}$

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### Monte-Carlo simulation Cont...

■ To get less noisier series at aggregate levels, we choose  $\Sigma$ ,  $\sigma_u^2$  and  $\sigma_v^2$  such that,

$$\mathsf{Var}(\epsilon_{AA,t} + \epsilon_{AB,t} + \epsilon_{BA,t} + \epsilon_{BB,t}) \leq \mathsf{Var}(\epsilon_{AA,t} + \epsilon_{AB,t} - v_t) \leq \mathsf{Var}(\epsilon_{AA,t} + u_t - 0.5v_t),$$

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### Monte-Carlo simulation Cont...

■ To get less noisier series at aggregate levels, we choose  $\Sigma$ ,  $\sigma_u^2$  and  $\sigma_v^2$  such that,

$$\mathsf{Var}(\epsilon_{AA,t} + \epsilon_{AB,t} + \epsilon_{BA,t} + \epsilon_{BB,t}) \leq \mathsf{Var}(\epsilon_{AA,t} + \epsilon_{AB,t} - \nu_t) \leq \mathsf{Var}(\epsilon_{AA,t} + u_t - 0.5\nu_t),$$

■ Thus we choose, 
$$\Sigma = \begin{pmatrix} 5.0 & 3.1 & 0.6 & 0.4 \\ 3.1 & 4.0 & 0.9 & 1.4 \\ 0.6 & 0.9 & 2.0 & 1.8 \\ 0.4 & 1.4 & 1.8 & 3.0 \end{pmatrix}$$
,  $\sigma_u^2 = 19$  and  $\sigma_u^2 = 18$ .

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### Sample version of the scoring rules

■ For a possible finite sample of size B from the multivariate forecast density  $\boldsymbol{\check{F}}$ , the variogram score is defined as,

$$VS(\breve{F}, y_{T+h}) = \sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij} \left( |y_{T+h,i} - y_{T+h,j}|^{p} - \frac{1}{B} \sum_{k=1}^{B} |\breve{Y}_{T+h,i}^{k} - \breve{Y}_{T+h,j}^{k}|^{p} \right)^{2}$$

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