

# Probabilistic Forecasts for Hierarchical Time Series

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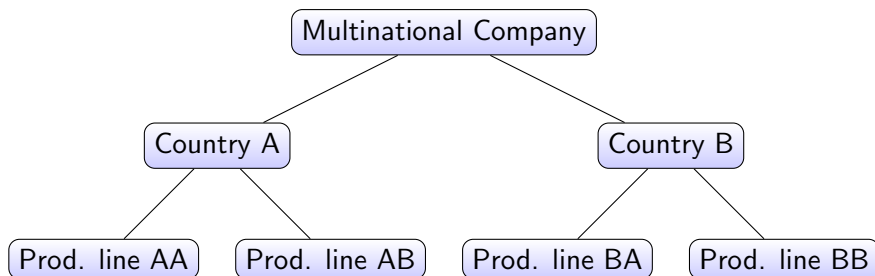
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- 1 Project 1: Hierarchical Forecast Reconciliation: <sup>a</sup> in Geometric View
- 2 Project 2: Probabilistic Forecasts for Hierarchical Time Series
  - Part 1: Definitions and <sup>or</sup> Parametric Approach
  - Part 2: A non-parametric bootstrap approach for ~~probabilistic forecasts reconciliation~~
- 3 Hierarchical forecasts <sup>for</sup> in macroeconomic variables - An application to Australian GDP ~~forecasts~~
- 4 Summary and time plan for completion

# Project 1: Hierarchical Forecast Reconciliation in Geometric View

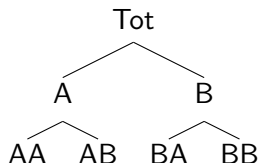
# Project 1: Motivation

## ■ Example:



- **Hierarchical time series:** A collection of multiple time series that has an inherent aggregation structure.
- Forecasts should add up. We call it *coherent*.
- **Objective:** Defining the coherency and reconciliation of point forecasts in terms of geometrical concepts.

# Project 1: Notations and Preliminaries



$$\mathbf{y}_t = [y_{Tot,t}, y_{A,t}, y_{B,t}, y_{AA,t}, y_{AB,t}, y_{BA,t}, y_{BB,t}]^T$$

$$\mathbf{b}_t = [y_{AA,t}, y_{AB,t}, y_{BA,t}, y_{BB,t}]^T$$

$$m = 4$$

$$n = 7$$

$$\mathbf{S} = \begin{pmatrix} 1 & 1 & 1 & 1 \\ 1 & 1 & 0 & 0 \\ 0 & 0 & 1 & 1 \end{pmatrix}_{4 \times 4}$$

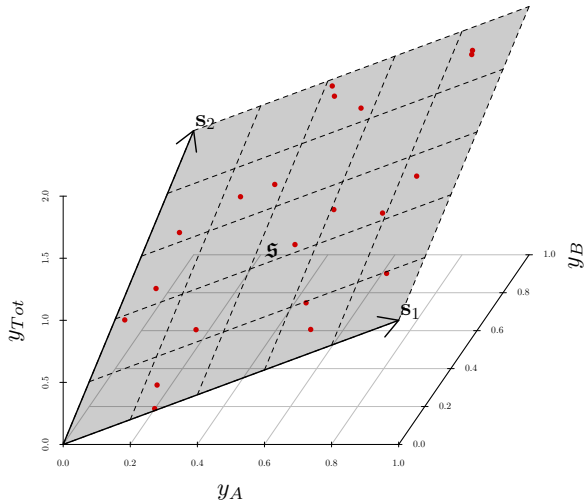
- Due to the aggregation nature of the hierarchy we have,

$$\mathbf{y}_t = \mathbf{S}\mathbf{b}_t$$

## Coherent subspace

The  $m$ -dimensional linear subspace  $\mathfrak{s} \subset \mathbb{R}^n$  that is spanned by the columns of  $\mathbf{S}$ , i.e.  $\mathfrak{s} = \text{span}(\mathbf{S})$ , is defined as the *coherent space*.

# Project 1: *Coherent forecasts*



# Project 1: Point forecast reconciliation

- Let  $\hat{\mathbf{y}} \in \mathbb{R}^n$  be an incoherent forecast and  $g(\cdot)$  be a function  $g : \mathbb{R}^n \rightarrow \mathbb{R}^m$ .

## Definition

A **point forecast**  $\tilde{\mathbf{y}}$  <sup>avg</sup> is reconciled with respect to  $g(\cdot)$  iff

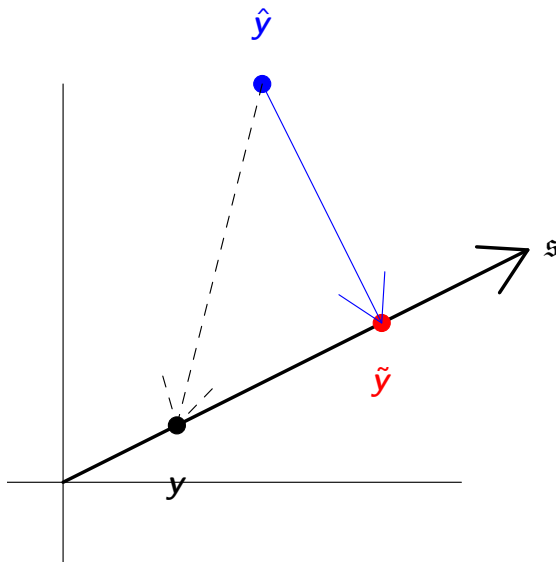
$$\tilde{\mathbf{y}} = \mathbf{S}g(\hat{\mathbf{y}})$$

- If  $g(\cdot)$  is linear,

$$\tilde{\mathbf{y}}_{T+h} = \mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{T+h}$$

# Project 1: *Point forecast reconciliation - OLS*

$$\mathbf{S}\mathbf{G} = \mathbf{S}(\mathbf{S}'\mathbf{S})^{-1}\mathbf{S}' \text{ (Hyndman et al., 2011)}$$





# Project 1: *Point forecast reconciliation - MinT*

- Minimises the trace of mean squared reconciled forecast errors (Wickramasuriya, Athanasopoulos, and Hyndman, 2018).

- **Geometry**

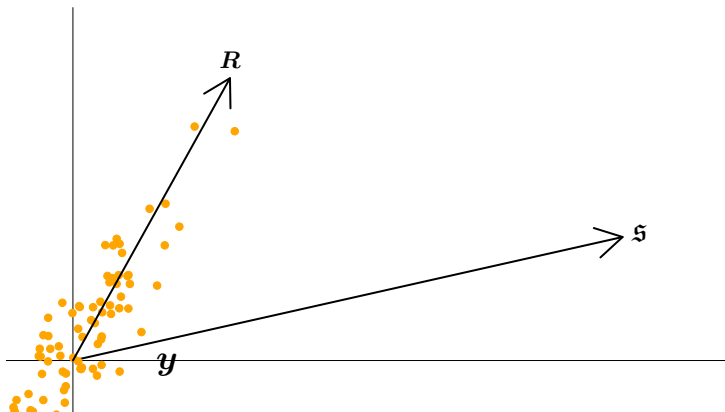
- Consider the covariance matrix of  $\mathbf{y}_{T+h} - \hat{\mathbf{y}}_{T+h}$ .
- This can be estimated using in-sample forecast errors.

$$\mathbf{W} = \sum_{t=1}^T (\mathbf{y}_t - \hat{\mathbf{y}}_t)(\mathbf{y}_t - \hat{\mathbf{y}}_t)'$$

- This provides information about the likely direction of an error.
- Projecting along this direction is more likely to result in a reconciled forecast that is closer to the target.

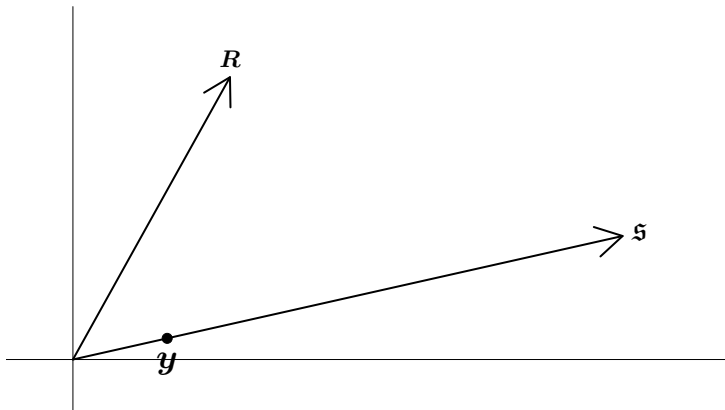
# Project 1: *Point forecast reconciliation - MinT*

$$SG = S (S' W_{T+h}^{-1} S)^{-1} S' W_{T+h}^{-1}$$



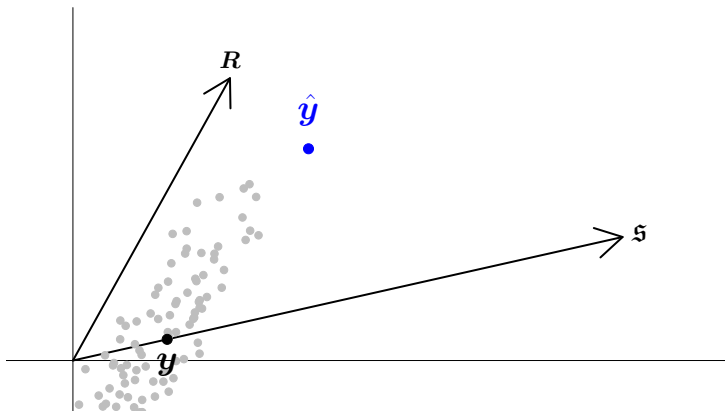
## Project 1: *Point forecast reconciliation - MinT*

$$SG = S (S' W_{T+h}^{-1} S)^{-1} S' W_{T+h}^{-1}$$



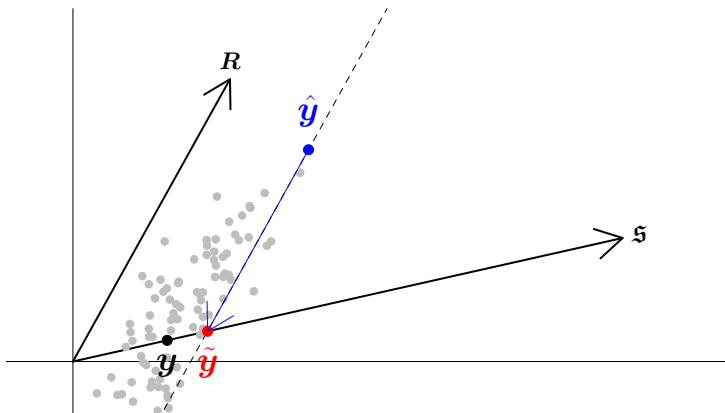
# Project 1: *Point forecast reconciliation - MinT*

$$SG = S (S' W_{T+h}^{-1} S)^{-1} S' W_{T+h}^{-1}$$



# Project 1: Point forecast reconciliation - MinT

$$SG = S (S' W_{T+h}^{-1} S)^{-1} S' W_{T+h}^{-1}$$



# Project 1: *Why projections?*

- Projections preserve ~~the~~ unbiasedness.
- What if the <sup>base?</sup> incoherent forecasts are biased?
- Can we bias correct and proceed with projections?  
We investigate this further,

# Project 2: Probabilistic Forecasts for Hierarchical Time Series

# Part 1: Definitions and Parametric Approach



## Project 2: *Motivation and Objectives*

- Extending the “reconciliation” method into probabilistic framework
- Probabilistic forecasts should reflect the inherent properties of real data. In particular,
  - ★ Aggregation structure
  - ★ Correlation structure
- Existing literature
  - (Ben Taieb et al., [2017](#))
  - (Jeon, Panagiotelis, and Petropoulos, [2018](#))
- **Objectives:**
  - 1 Defining coherency and reconciliation for probabilistic forecasts
  - 2 Probabilistic forecast reconciliation in the parametric framework
  - 3 Probabilistic forecast reconciliation in the non-parametric framework

## Project 2: *Coherent probabilistic forecasts*

Let  $(\mathbb{R}^m, \mathcal{F}_{\mathbb{R}^m}, \nu)$  and  $(\mathfrak{s}, \mathcal{F}_{\mathfrak{s}}, \mu)$  be probability triples on  $m$ -dimensional space and the coherent subspace respectively.

### Definition

The probability measure  $\mu$  is coherent if

$$\nu(\mathcal{B}) = \mu(s(\mathcal{B})) \quad \forall \mathcal{B} \in \mathcal{F}_{\mathbb{R}^m}$$

where  $s(\mathcal{B})$  is the image of  $\mathcal{B}$  under premultiplication by  $\mathbf{S}$

## Project 2: *Reconciled Probabilistic Forecast*

Let  $g : \mathbb{R}^n \rightarrow \mathbb{R}^m$  be a function. Then

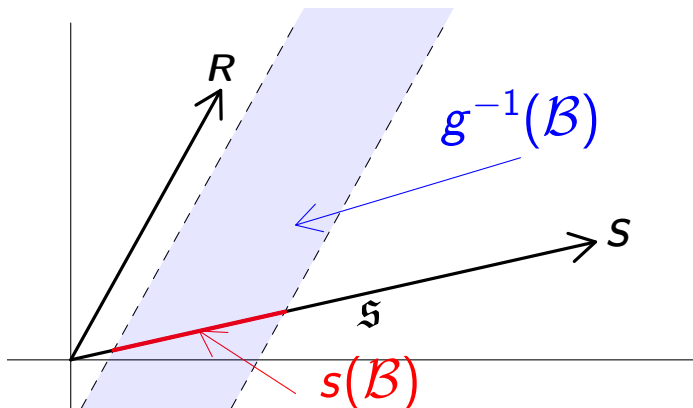
### Definition

The probability triple  $(\mathfrak{s}, \mathcal{F}_{\mathfrak{s}}, \tilde{\nu})$  reconciles the probability triple  $(\mathbb{R}^n, \mathcal{F}_{\mathbb{R}^n}, \hat{\nu})$  with respect to  $g$  iff

$$\tilde{\nu}(s(\mathcal{B})) = \nu(\mathcal{B}) = \hat{\nu}(g^{-1}(\mathcal{B})) \quad \forall \mathcal{B} \in \mathcal{F}_{\mathbb{R}^m}$$

where  $g^{-1}$  is the pre-image of  $g$ .

## Project 2: Geometry



## Project 2: *Analytically*

If we have an unreconciled density the reconciled density can be obtained by linear transformations and marginalisation.

$$\begin{aligned}\Pr(\tilde{\mathbf{b}} \in \mathcal{B}) &= \Pr(\hat{\mathbf{y}} \in g^{-1}(\mathcal{B})) \\ &= \int_{g^{-1}(\mathcal{B})} f(\hat{\mathbf{y}}) d\hat{\mathbf{y}} \\ &= \int_{\mathcal{B}} \int f(\mathbf{S}\tilde{\mathbf{b}} + \mathbf{R}\tilde{\mathbf{a}}) |(\mathbf{S} \ \mathbf{R})| d\tilde{\mathbf{a}} d\tilde{\mathbf{b}}\end{aligned}$$

## Project 2: *For elliptical distributions*

Consider <sup>the</sup> case where the base and true predictive distributions are elliptical.

### Theorem

*There exists a matrix  $\mathbf{G}$  such that the true predictive distribution can be recovered by linear reconciliation.*

This follows from the closure property of elliptical distributions under affine transformations and marginalisation.

## Project 2: Assuming Gaussian distribution

- Let  $\mathcal{N}(\hat{\mathbf{y}}_{T+h}, \mathbf{W}_{T+h})$  be an incoherent forecast distribution at time  $T+h$  where  $\hat{\mathbf{y}}_{T+h}$  is the incoherent mean and  $\mathbf{W}_{T+h} = E_{\mathbf{y}_{T+h}}[(\mathbf{y}_{T+h} - \hat{\mathbf{y}}_{T+h})(\mathbf{y}_{T+h} - \hat{\mathbf{y}}_{T+h})^T | \mathcal{I}_T]$  is <sup>the</sup> incoherent variance
- The reconciled Gaussian distribution is given by,

$$\mathcal{N}(\mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{T+h}, \mathbf{S}\mathbf{G}\mathbf{W}_{T+h}\mathbf{G}'\mathbf{S}')$$

- $\mathbf{G} = (\mathbf{S}'\mathbf{W}_{T+h}^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}_{T+h}^{-1}$  minimizes the energy score in the limiting case A1
- Simulation study <sup>evidence?</sup> evident for improved predictive performance in reconciled Gaussian forecast distributions

# Today's talk

- A non-parametric bootstrap approach for probabilistic forecasts/  
reconciliation
- Hierarchical forecasts <sup>for</sup> in macroeconomic variables - An application to  
Australian GDP ~~forecasts~~



# Part 2: A non-parametric bootstrap approach for probabilistic forecasts~~s~~ reconciliation

# Probabilistic forecast reconciliation: Non-parametric approach

? unsuitable?

- Often parametric densities are unavailable but we can simulate a sample from the predictive distribution
- Suppose  $\hat{\mathbf{y}}_{T+h}^{[1]}, \dots, \hat{\mathbf{y}}_{T+h}^{[J]}$  is a sample from the incoherent predictive distribution
- Then setting  $\tilde{\mathbf{y}}_{T+h}^{[j]} = \mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{T+h}^{[j]}$  produces a sample from the reconciled predictive distribution with respect to  $\mathbf{G}$

# Probabilistic forecast reconciliation: Non-parametric approach

- 1 Fit univariate models at each node using data up to time  $T$
- 2 Let  $\mathbf{\Gamma}_{(T \times n)} = (\mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_T)'$  be a matrix of in-sample residuals where  $\mathbf{e}_t = \mathbf{y}_t - \hat{\mathbf{y}}_t$
- 3 Let  $\mathbf{\Gamma}_{(H \times n)}^b = (\mathbf{e}_1^b, \dots, \mathbf{e}_H^b)'$  be a block bootstrap sample of size  $h$  from  $\mathbf{\Gamma}$  and repeat this for  $b = 1, \dots, B$ .
- 4 Generate  $h$ -step ahead sample paths from the fitted models incorporating  $\mathbf{\Gamma}^b$ . Denote these by  $\hat{\mathbf{y}}_{T+h}^b$ , for  $h = 1, \dots, H$ .
- 5 Repeat step 4 for  $b = 1, \dots, B$  times
- 6 Setting  $\tilde{\mathbf{y}}_{T+h,j}^b = \mathbf{SG} \hat{\mathbf{y}}_{T+h,j}^b$  produces a sample from the reconciled distribution

# Optimal reconciliation of future paths

- We propose to find an optimal  $\mathbf{G}$  matrix by minimizing Energy score

$$\operatorname{argmin}_{\mathbf{G}_h} \mathbb{E}_Q[\mathbf{eS}(\mathbf{S}\mathbf{G}_h \hat{\mathbf{Y}}'_{T+h}, \mathbf{y}_{T+h})],$$

where,

$$\mathbf{eS}(\tilde{\mathbf{Y}}_{T+h}, \mathbf{y}_{T+h}) = \mathbb{E}_{\tilde{\mathbf{F}}} \|\tilde{\mathbf{Y}}_{T+h} - \mathbf{y}_{T+h}\|^\alpha - \frac{1}{2} \mathbb{E}_{\tilde{\mathbf{F}}} \|\tilde{\mathbf{Y}}_{T+h} - \tilde{\mathbf{Y}}_{T+h}^*\|^\alpha, \\ \alpha \in (0, 2]$$

- Monte-Carlo approximation to the above objective function is,

$$\operatorname{argmin}_{\mathbf{G}} \sum_{i=1}^N \left\{ \frac{1}{B} \sum_{j=1}^B \|\mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{T+h,i,j}^b - \mathbf{y}_{T+h}\| - \right. \\ \left. \frac{1}{2(B-1)} \sum_{j=1}^{B-1} \|\mathbf{S}\mathbf{G}(\hat{\mathbf{y}}_{T+h,i,j}^b - \hat{\mathbf{y}}_{T+h,i,j+1}^b)\| \right\}$$

# Optimal reconciliation of future paths *Cont.*

- We impose the following structure to the  $\mathbf{G}$  matrix

$$\mathbf{G} = (\mathbf{S}'\mathbf{W}\mathbf{S})^{-1} \mathbf{S}'\mathbf{W} \quad (1)$$

- We propose four methods to optimise  $\mathbf{G}$

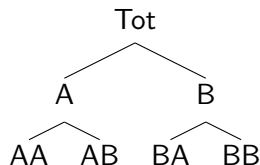
**Method 1:** Optimising  $\mathbf{W}$

**Method 2:** Optimising cholesky decomposition of  $\mathbf{W}$   
 $\mathbf{W} = \mathbf{R}'\mathbf{R}$  where  $\mathbf{R}$  is an upper triangular matrix

**Method 3:** Optimising cholesky of  $\mathbf{W}$  - restricted for scaling  
 $\mathbf{W} = \mathbf{R}'\mathbf{R}$  s.t.  $\mathbf{i}'\mathbf{W}\mathbf{i} = 1$  where  $\mathbf{i} = (1, 0, \dots, 0)'$

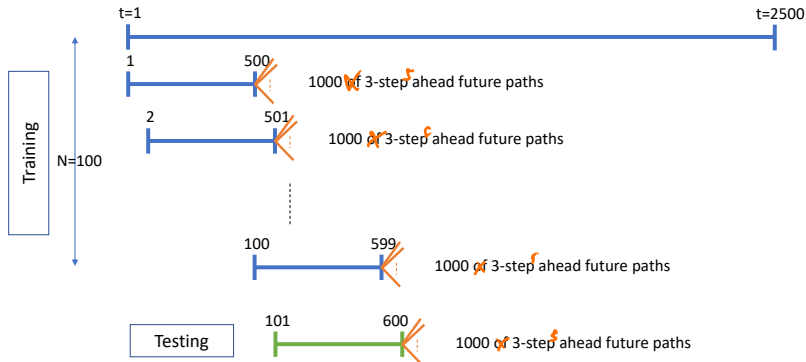
**Method 4:** Optimising  $\mathbf{G}$  such that  $\mathbf{G}\mathbf{S} = \mathbf{I}$

- Data generating process



- DGP was designed such that we have much noisier series in the bottom level

# Monte-Carlo Simulation



## ■ Simulation setup:

- First 2500 observations <sup>were</sup> generated
- Fit Univariate ARIMA models for a rolling window of 500 observations
- $B = 1000$  of  $h = 1, 2, 3$  steps-ahead bootstrap future paths generated
- Training window is rolled one observation ahead and process was repeated until we ~~get~~ <sup>were generated</sup>  $N = 100$  of incoherent,  $h = 1, 2, 3$  steps-ahead future paths
- We find the optimal  $\mathbf{G}_h$  for  $h = 1, 2, 3$  that reconcile <sup>s</sup>  $h$ -step-ahead future paths giving minimal average Energy score
- This optimal  $\mathbf{G}_h$  is then used to reconcile the incoherent future paths ~~obtained~~ for the test set
- <sup>The</sup> Process was repeated ~~for~~ 1000 times and average scores were calculated for the test set



# Monte-Carlo Simulation *Cont.*

Optimisation method	Hierarchy 1				Hierarchy 2			
	$h = 1$		$h = 3$		$h = 1$		$h = 3$	
	ES	VS	ES	VS	ES	VS	ES	VS
Method 1 - Optimising <b><i>W</i></b>	2.48	0.11	2.75	0.11	5.36	1.21	5.83	1.38
Method 2 - Optimising <b><i>R</i></b>	2.48	0.11	2.75	0.11	5.37	1.21	5.83	1.37
Method 3 - Optimising <b><i>R</i></b> (Restricted)	2.48	0.11	2.75	0.11	5.37	1.21	5.83	1.37
Method 4 - Optimising <b><i>G</i></b>	2.48	0.11	2.75	0.11	5.38	1.21	5.83	1.38

- Parameterisation does not matter

## ■ Comparison with point forecast reconciliation methods

Reconciliation method	Hierarchy 1				Hierarchy 2			
	$h = 1$		$h = 3$		$h = 1$		$h = 3$	
	ES	VS	ES	VS	ES	VS	ES	VS
Optimal <b>G</b>	2.48*	0.106	2.75*	0.106	5.36*	1.21*	5.83*	1.38*
MinT(Shrink)	2.47*	0.105	2.74*	0.105	5.33*	1.19*	5.77*	1.34*
WLS	2.46*	0.105	2.74*	0.105	5.43*	1.23	5.98*	1.40*
OLS	2.54*	0.105	2.80*	0.105	5.51*	1.23	5.98*	1.40*
Base	2.67	0.105	2.94	0.105	5.71	1.28	6.27	1.49

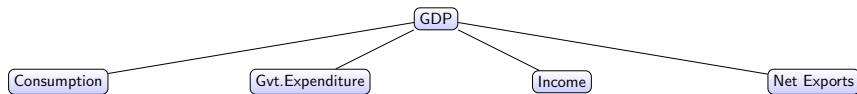
\* Add definition

## ■ Reconciliation methods perform better than Base forecasts

## ■ MinT(Shrink) is at least as good as Optimal method. Thus going forward with MinT projection



# Hierarchical forecasts <sup>for</sup> ~~in~~ macroeconomic variables - An application to Australian GDP ~~forecasts~~

# Macroeconomic forecasting



- Common forecasting approaches involves univariate methods of multivariate methods such as VAR. ?
- The era of big data led to the use of regularization and shrinkage methods - dynamic factor models, Lasso, Bayesian VARs. ^
- The predictors in these methods commonly include the components of the variables of interest.
- This might fail to reflect the deterministic relationship between macroeconomic variables in the forecasts.

# Macroeconomic forecasting

- Both aligned decision making and forecast accuracy are key concerns for economic agents and policy makers 
- Thus we propose to use hierarchical forecasting methods in macroeconomic forecasts.
- Related literature: Only one application on point forecasting for inflation (Capistrán, Constandse, and Ramos-Francia, [2010](#); Weiss, [2018](#))
- To the best of our knowledge we use hierarchical forecasting methods for point as well as probabilistic forecasts for the first time in macroeconomic literature 

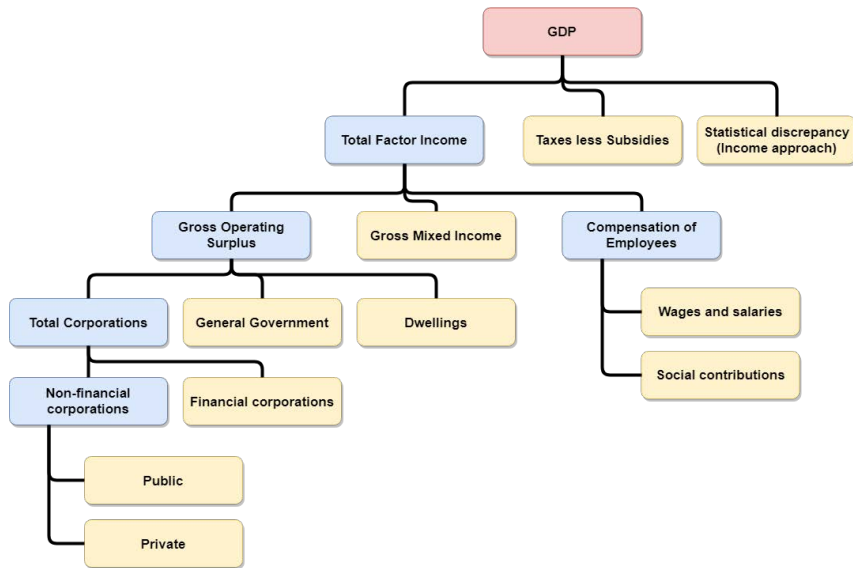
- We consider Gross Domestic Product (GDP) of Australia with quarterly data spanning the period 1984:Q4–2018:Q3
- The Australian Bureau of Statistics (ABS) measures GDP using three main approaches - Production, Income and Expenditure  
The final GDP figure is obtained as an average of these three figures.
- We restrict our attention to nominal, seasonally unadjusted data
- Thus we concentrate on the Income and Expenditure approaches

## Income approach

$$\begin{aligned} \text{GDP} = & \text{Gross operating surplus} + \text{Gross mixed income} \\ & + \text{Compensation of employees} \\ & + \text{Taxes less subsidies on production and imports} \\ & + \text{Statistical discrepancy (I)}. \end{aligned}$$

- The hierarchy has two levels of aggregation below the top-level, with a total of  $n = 16$  series and  $m = 10$  bottom level series

# Australian GDP : *Data structures - Income approach*



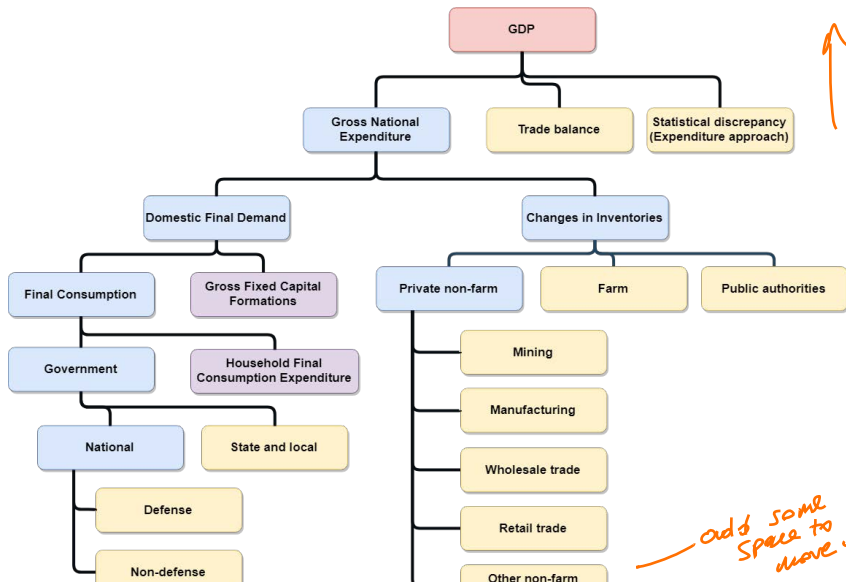


## Expenditure approach

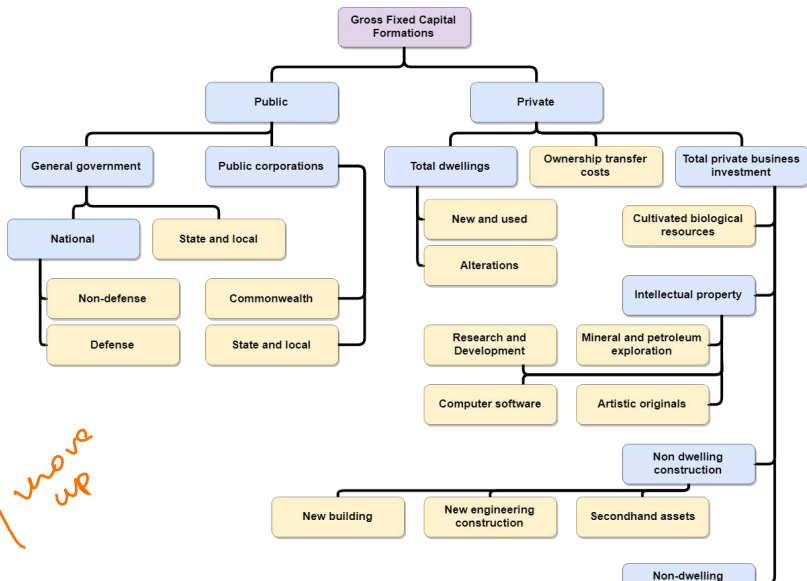
$$\begin{aligned} \text{GDP} = & \text{Final consumption expenditure} + \text{Gross fixed capital formation} \\ & + \text{Changes in inventories} + \text{Trade balance} \\ & + \text{Statistical discrepancy (E)}. \end{aligned}$$

- The hierarchy has three levels of aggregation below the top-level, with a total of  $n = 80$  series and  $m = 53$  series at the bottom level. 📌

# Australian GDP : *Data structures - Expenditure approach*

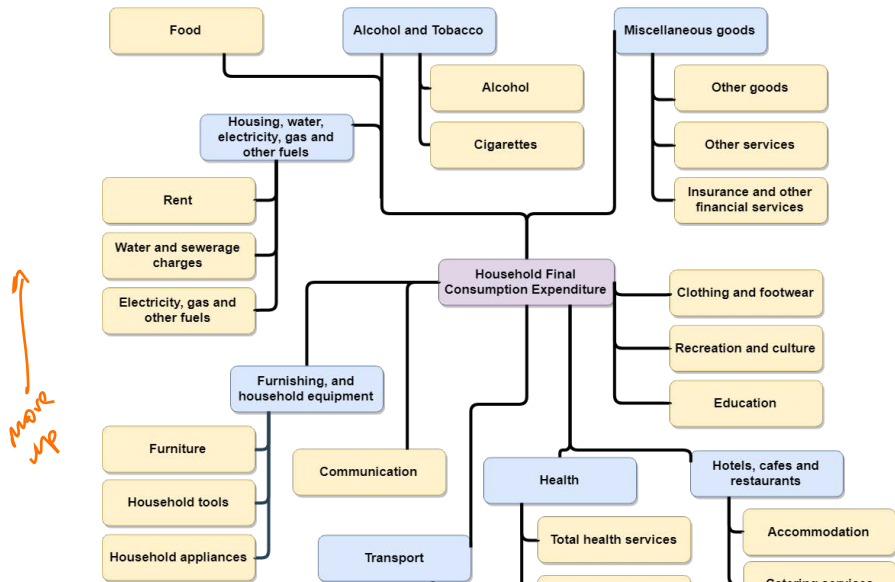


# Australian GDP : *Data structures - Expenditure approach*

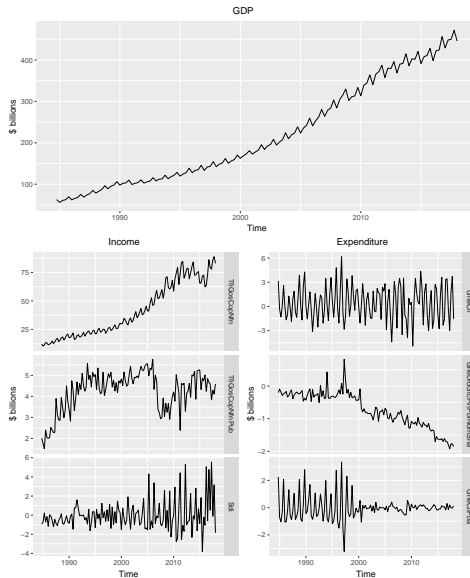


↑  
move  
up

# Australian GDP : *Data structures - Expenditure approach*



# Australian GDP : *Time plots for different levels*



## ■ Analysis set up:

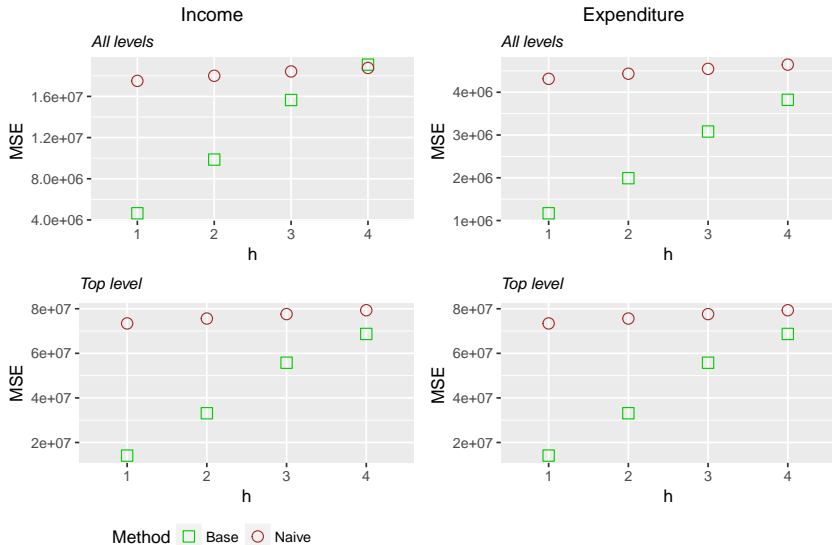
- First training sample is set from 1984:Q4 to 1994:Q3 and forecasts produces for four quarters ahead (1994:Q4 to 1995:Q3).
- Then the training window is expanded by one quarter at a time.
- This leads to 94 1-step-ahead, 93 2-steps-ahead, 92 3-steps-ahead and 91 4-steps-ahead forecasts available for evaluation.

## ■ Base forecasting models:

- Univariate ARIMA and ETS models were fitted for each training set.
- Four-steps ahead forecasts were generated using the fitted models.

One to four

# Point forecasts: *Base* vs *Naive*



# Point forecasts: *Reconciliation*

Reconciled forecasts are given by,

$$\tilde{\mathbf{y}}_{T+h} = \mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{T+h}$$

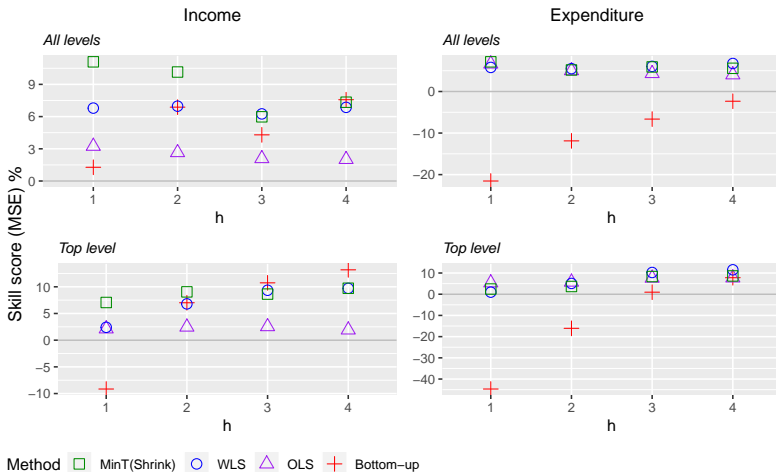
Method	$\mathbf{G}$
BU	$(\mathbf{0}_{m \times n-m} \quad \mathbf{I}_{m \times m})$
OLS	$(\mathbf{S}'\mathbf{S})^{-1}\mathbf{S}'$
WLS	$(\mathbf{S}'\hat{\mathbf{W}}_{T+1}^{wls}\mathbf{S})^{-1}\mathbf{S}'\hat{\mathbf{W}}_{T+1}^{wls}$
MinT(Shrink)	$(\mathbf{S}'\hat{\mathbf{W}}_{T+1}^{shr}\mathbf{S})^{-1}\mathbf{S}'\hat{\mathbf{W}}_{T+1}^{shr}$

$$\hat{\mathbf{W}}_{T+1}^{shr} = \tau \text{Diag}(\hat{\mathbf{W}}_{T+1}^{sam}) + (1 - \tau)\hat{\mathbf{W}}_{T+1}^{sam}$$

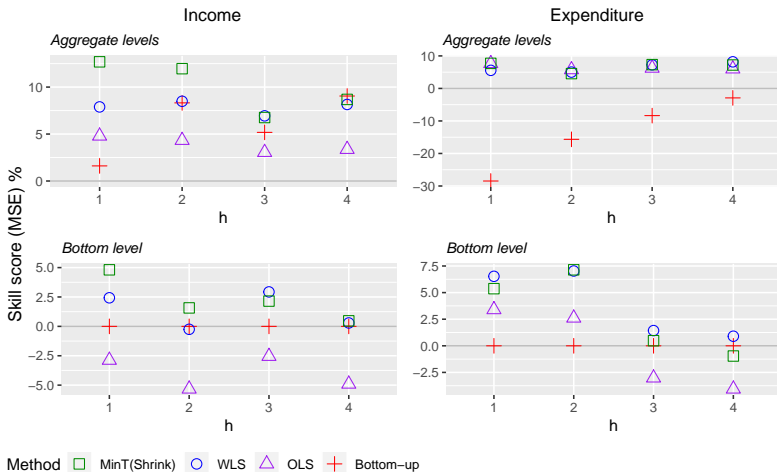
$$\hat{\mathbf{W}}_{T+1}^{wls} = \text{Diag}(\hat{\mathbf{W}}_{T+1}^{shr})$$



# Reconciled Point forecasts - *Results*



# Reconciled Point forecasts - Results



- Gaussian approach :

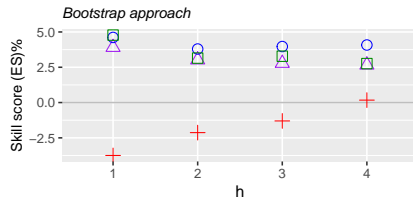
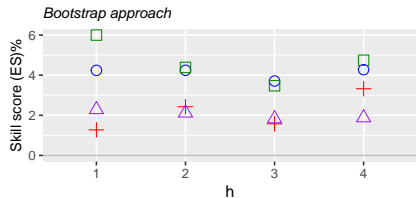
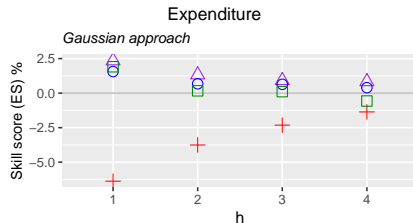
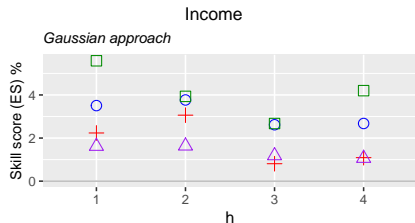
$$\mathcal{N}(\mathbf{SG}\hat{\mathbf{y}}_{T+h}, \mathbf{SGW}_{T+h}\mathbf{G}'\mathbf{S}')$$

- Non-parametric Bootstrap approach :

$$\tilde{\mathbf{Y}}_{T+h} = \mathbf{SG}\hat{\mathbf{Y}}'_{T+h}$$

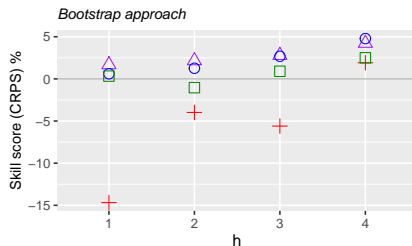
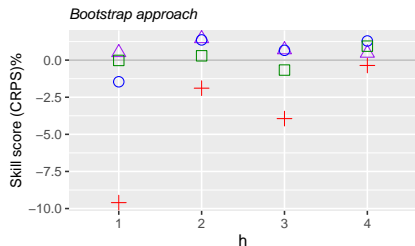
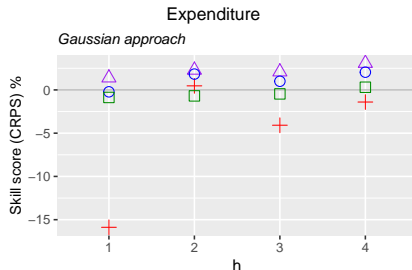
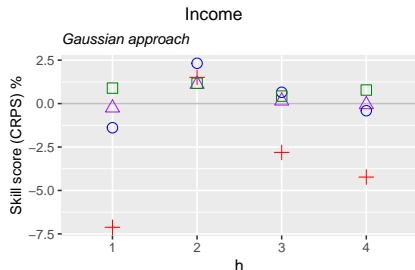
where,  $\hat{\mathbf{Y}}_{T+h} = (\hat{\mathbf{y}}_{T+h}^1, \dots, \hat{\mathbf{y}}_{T+h}^B)'$

# Reconciled Probabilistic Forecasts



Method ■ MinT(Shrink) ○ WLS △ OLS + Bottom-up

# Reconciled Probabilistic Forecasts



Method MinT(Shrink) WLS OLS Bottom-up

# Summary and time plan for completion

# Summary

- We define point and probabilistic forecast reconciliation in geometric terms.
- We propose a parametric approach for probabilistic forecast reconciliation.
- We introduce a novel non-parametric bootstrap approach for producing reconciled probabilistic forecasts.
- Simulation study *provides* evidence <sup>(2)</sup> that the optimal reconciliation with respect to energy score is equivalent to reconciling each sample path via MinT approach.
- We apply hierarchical forecast reconciliation methods to forecast Australian GDP in point as well as probabilistic framework.

# Time plan for completion

	Thesis Chapter	Task description	Time duration	Progress
1 and 2.	Introduction and Background Review	Writing the chapter.	September/2019 - October/2019	40% complete
3.	Hierarchical forecast reconciliation in Geometric view	Bias correction and application	May/2019 - July/2019	75% Completed
4.	Probabilistic forecast reconciliation for hierarchical time series	Completing the paper.	June/2019 - August/2019	90% Completed
5.	Application	Forecasting Australian GDP		100% Completed



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# References II

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Thank You!!

# Project 2: Probabilistic forecasts evaluation

A1

**Energy score** (Gneiting et al., 2008)

$$\text{eS}(\check{\mathbf{Y}}_{T+h}, \mathbf{y}_{T+h}) = \mathbb{E}_{\check{\mathbf{F}}} \|\check{\mathbf{Y}}_{T+h} - \mathbf{y}_{T+h}\|^\alpha - \frac{1}{2} \mathbb{E}_{\check{\mathbf{F}}} \|\check{\mathbf{Y}}_{T+h} - \check{\mathbf{Y}}_{T+h}^*\|^\alpha, \quad \alpha \in (0, 2]$$

**Log score** (Gneiting and Raftery, 2007)

$$\text{LS}(\check{\mathbf{F}}, \mathbf{y}_{T+h}) = -\log \check{\mathbf{f}}(\mathbf{y}_{T+h})$$

**Variogram score** (Scheuerer and Hamill, 2015)

$$\text{VS}(\check{\mathbf{F}}, \mathbf{y}_{T+h}) = \sum_{i=1}^n \sum_{j=1}^n w_{ij} \left( |y_{T+h,i} - y_{T+h,j}|^p - \mathbb{E}_{\check{\mathbf{F}}} |\check{Y}_{T+h,i} - \check{Y}_{T+h,j}|^p \right)^2$$

**CRPS** (Gneiting and Raftery, 2007)

$$\text{CRPS}(\check{F}_i, y_{T+h,i}) = \mathbb{E}_{\check{F}_i} |\check{Y}_{T+h,i} - y_{T+h,i}| - \frac{1}{2} \mathbb{E}_{\check{F}_i} |\check{Y}_{T+h,i} - \check{Y}_{T+h,i}^*|$$

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$\check{\mathbf{Y}}_{T+h}$  and  $\check{\mathbf{Y}}_{T+h}^*$  : Independent random vectors from the coherent forecast distribution  $\check{\mathbf{F}}$ .

$\mathbf{y}_{T+h}$  : Vector of realizations.

$\check{Y}_{T+h,i}$  and  $\check{Y}_{T+h,j}$  :  $i$ th and  $j$ th components of the vector  $\check{\mathbf{Y}}_{T+h}$

- ◀ A2 Shrinkage estimator for 1-step ahead base forecast errors

$$\hat{\Sigma}_{T+1}^{shr} = \tau \hat{\Sigma}_{T+1}^D + (1 - \tau) \hat{\Sigma}_{T+1},$$

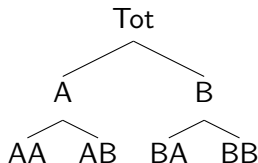
where  $\hat{\Sigma}_{T+1}^D$  is the diagonal matrix comprising diagonal entries of  $\hat{\Sigma}_{T+1}$  and

$$\tau = \frac{\sum_{i \neq j} \text{Var}(\hat{r}_{ij})}{\sum_{i \neq j} \hat{r}_{ij}^2}$$

is a shrinkage parameter.  $\hat{r}_{ij}$  is the  $ij$ -th element of sample correlation matrix. In this estimation, the off-diagonal elements of 1-step ahead sample covariance matrix will be shrunk to zero depending on the sparsity.

# Monte-Carlo simulation

## ■ Data generating process ◀ A3



- $\{w_{AA,t}, w_{AB,t}, w_{BA,t}, w_{BB,t}\} \sim ARIMA(p, d, q)$
- $p \in \{1, 2\}$  and  $d \in \{0, 1\}$
- $\{\epsilon_{AA,t}, \epsilon_{AB,t}, \epsilon_{BA,t}, \epsilon_{BB,t}\} \sim \mathcal{N}(\mathbf{0}, \Sigma)$
- Parameters for *AR* and *MA* components were randomly and uniformly generated from  $[0.3, 0.5]$  and  $[0.3, 0.7]$  respectively

## ■ $y_t$ are then generated as follows

Bottom level	Aggregate level 1	Total
$y_{AA,t} = w_{AA,t} + u_t - 0.5v_t$	$y_{A,t} = w_{AA,t} + w_{AB,t} - v_t$	$y_{Tot,t} = w_{AA,t} + w_{AB,t} + w_{BA,t} + w_{BB,t}$
$y_{AB,t} = w_{AB,t} - u_t - 0.5v_t$	$y_{B,t} = w_{BA,t} + w_{BB,t} + v_t$	
$y_{BA,t} = w_{BA,t} + u_t + 0.5v_t$		
$y_{BB,t} = w_{BB,t} - u_t + 0.5v_t$		

- To get less noisier series at aggregate levels, we choose  $\Sigma$ ,  $\sigma_u^2$  and  $\sigma_v^2$  such that,

$$\text{Var}(\epsilon_{AA,t} + \epsilon_{AB,t} + \epsilon_{BA,t} + \epsilon_{BB,t}) \leq \text{Var}(\epsilon_{AA,t} + \epsilon_{AB,t} - v_t) \leq \text{Var}(\epsilon_{AA,t} + u_t - 0.5v_t),$$

- Thus we choose,  $\Sigma = \begin{pmatrix} 5.0 & 3.1 & 0.6 & 0.4 \\ 3.1 & 4.0 & 0.9 & 1.4 \\ 0.6 & 0.9 & 2.0 & 1.8 \\ 0.4 & 1.4 & 1.8 & 3.0 \end{pmatrix}$ ,  $\sigma_u^2 = 19$  and  $\sigma_v^2 = 18$ .

# Sample version of the scoring rules

- For a possible finite sample of size  $B$  from the multivariate forecast density  $\check{\mathbf{F}}$ , the variogram score is defined as,

$$VS(\check{\mathbf{F}}, \mathbf{y}_{T+h}) = \sum_{i=1}^n \sum_{j=1}^n w_{ij} \left( |y_{T+h,i} - y_{T+h,j}|^p - \frac{1}{B} \sum_{k=1}^B |\check{Y}_{T+h,i}^k - \check{Y}_{T+h,j}^k|^p \right)^2$$