

# Concomitant Lasso with Repetitions (CLaR): beyond averaging multiple realizations of correlated noise

**Quentin Bertrand**

Joint work with:

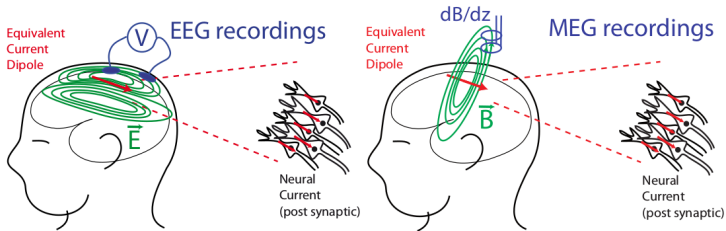
**Mathurin Massias** (INRIA)

**Alexandre Gramfort** (INRIA)

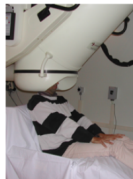
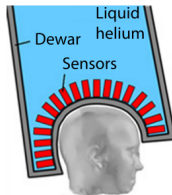
**Joseph Salmon** (IMAG, Univ Montpellier, CNRS)

# M/EEG inverse problem for brain imaging

- ▶ sensors: electric and magnetic fields during a cognitive task
- ▶ goal: which parts of the brain are responsible for the signals?
- ▶ applications: epilepsy treatment, brain aging, anesthesia risks

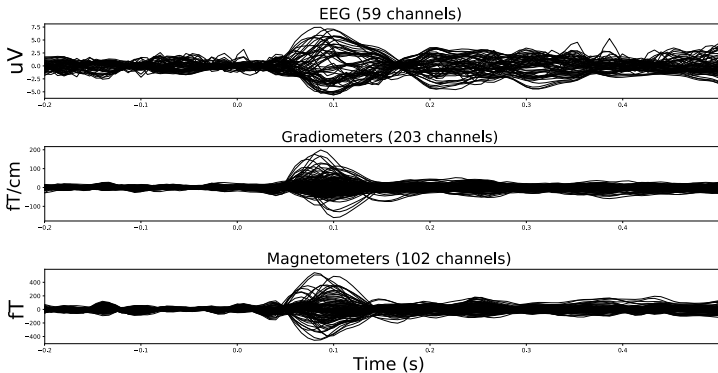


First EEG  
recordings  
in 1929  
by H. Berger



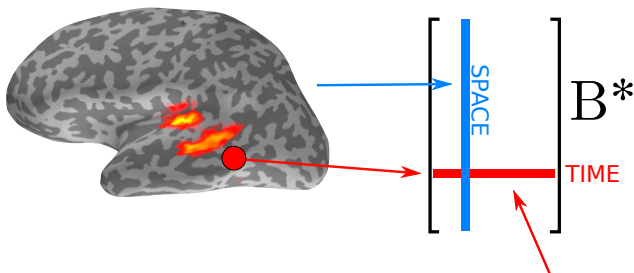
Hôpital La Timone  
Marseille, France

# M/EEG data

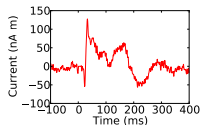


- 3 different types of sensor

# Source modeling (discretization with voxels)

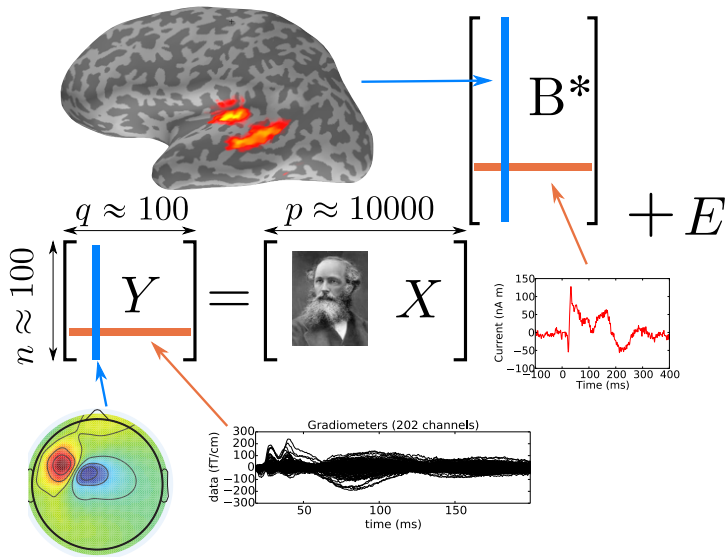


Position a few thousands candidate sources over the brain (e.g., every 5mm)



$$B^* \in \mathbb{R}^{p \times q}$$

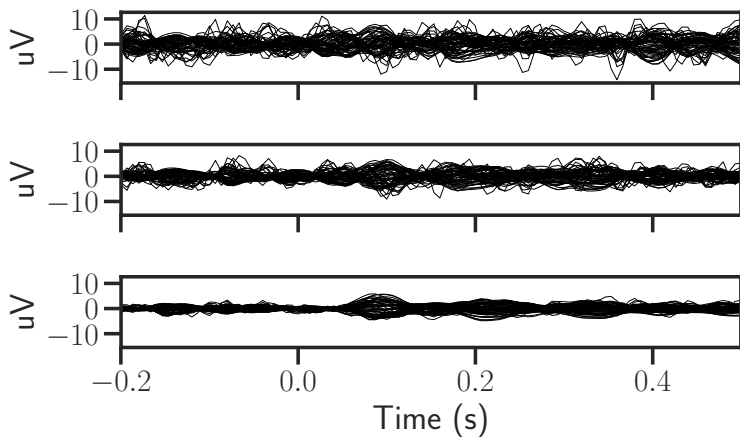
# The M/EEG inverse problem: modeling



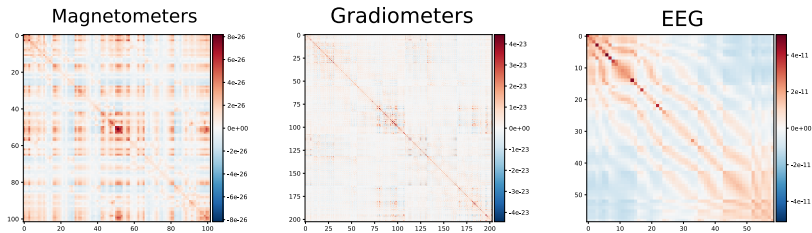
$$n \ll p$$

## Very noisy data: must repeat recordings

- average of 5 (top) / 10 (middle) / 50 (bottom) repetitions



# Noise covariance for each type of sensor



► 3 different sensors  $\implies$  3 different noise structures

# A Multi-Task framework

## Multi-Task regression notation:

- ▶  $n$  observations (e.g., number of sensors)
- ▶  $q$  tasks (e.g., temporal information)
- ▶  $p$  features
- ▶  $r$  number of repetitions
- ▶  $Y^{(1)}, \dots, Y^{(r)} \in \mathbb{R}^{n \times q}$  observation matrices;  $\bar{Y} = \frac{1}{r} \sum_l Y^{(l)}$
- ▶  $X \in \mathbb{R}^{n \times p}$  design matrix (known)

$$Y^{(l)} = XB^* + SE^{(l)}$$

where

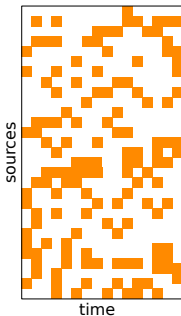
- ▶  $B^* \in \mathbb{R}^{p \times q}$  : true source activity matrix (unknown)
- ▶  $S \in \mathbb{S}_{++}^n$  co-standard deviation matrix (unknown)
- ▶  $E^{(1)}, \dots, E^{(r)} \in \mathbb{R}^{n \times q}$  : white Gaussian noise



# Multi-Task penalties<sup>(1)</sup>

Popular convex penalties considered:

$$\hat{\mathbf{B}} \in \arg \min_{\mathbf{B} \in \mathbb{R}^{p \times q}} \left( \frac{1}{2nq} \|\bar{\mathbf{Y}} - \mathbf{X}\mathbf{B}\|_F^2 + \lambda \Omega(\mathbf{B}) \right)$$



Sparse support: no structure

Penalty: **Lasso type**

$$\Omega(\mathbf{B}) = \|\mathbf{B}\|_1 = \sum_{j=1}^p \sum_{k=1}^q |\mathbf{B}_{j,k}|$$

Parameter  $\hat{\mathbf{B}} \in \mathbb{R}^{p \times q}$

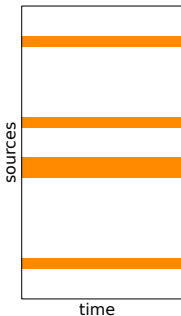
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<sup>(1)</sup>G. Obozinski, B. Taskar, and M. I. Jordan. "Joint covariate selection and joint subspace selection for multiple classification problems". In: *Statistics and Computing* 20.2 (2010), pp. 231–252.

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$$\hat{\mathbf{B}} \in \arg \min_{\mathbf{B} \in \mathbb{R}^{p \times q}} \left( \frac{1}{2nq} \|\bar{\mathbf{Y}} - \mathbf{X}\mathbf{B}\|_F^2 + \lambda \Omega(\mathbf{B}) \right)$$



Sparse support: group structure

Penalty: **Group-Lasso type**

$$\Omega(\mathbf{B}) = \|\mathbf{B}\|_{2,1} = \sum_{j=1}^p \|\mathbf{B}_{j,:}\|_2$$

Parameter  $\hat{\mathbf{B}} \in \mathbb{R}^{p \times q}$

where  $\mathbf{B}_{j,:}$ : the  $j$ -th row of  $\mathbf{B}$

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# Multi-Task data-fitting term

- Classical Multi-Task estimator: use averaged signal

$$\hat{\mathbf{B}} \in \arg \min_{\mathbf{B} \in \mathbb{R}^{p \times q}} \left( \frac{1}{2nq} \left\| \bar{\mathbf{Y}} - \mathbf{X}\mathbf{B} \right\|_F^2 + \lambda \Omega(\mathbf{B}) \right)$$

- How to take advantage of the number of repetitions?

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- **How to take advantage of the number of repetitions?**

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# Reminder on the Lasso theory<sup>(2)(3)</sup>

## (i.i.d. case, Single-Task)

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### Theorem

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- ▶ i.i.d. Gaussian noise
- ▶ +  $X$  satisfying the “Restricted Eigenvalue” property
- ▶ +  $\lambda = 2\sigma_* \sqrt{\frac{2\log(p/\delta)}{n}}$
- ▶  $\implies$  with probability  $1 - \delta$ :

$$\frac{1}{n} \|X\beta^* - X\hat{\beta}^{(\lambda)}\|^2 \leq \frac{18}{\kappa_{s^*}^2} \frac{\sigma_*^2 s^*}{n} \log\left(\frac{p}{\delta}\right)$$

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BUT  $\sigma_*$  is unknown in practice !

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# Reminder on the Square root Lasso<sup>(4)</sup>(5)(6)

## (i.i.d. case, Single-Task)

$$\hat{\beta} \in \arg \min_{\beta \in \mathbb{R}^p} \frac{1}{\sqrt{n}} \|y - X\beta\|_2 + \lambda \|\beta\|_1$$

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$\lambda$  does not depend on  $\sigma_*$  anymore!

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<sup>(4)</sup> A. Belloni, V. Chernozhukov, and L. Wang. “Square-root Lasso: pivotal recovery of sparse signals via conic programming”. In: *Biometrika* 98.4 (2011), pp. 791–806.

<sup>(5)</sup> T. Sun and C.-H. Zhang. “Scaled sparse linear regression”. In: *Biometrika* 99.4 (2012), pp. 879–898.

<sup>(6)</sup> C. Giraud. *Introduction to high-dimensional statistics*. Vol. 138. CRC Press, 2014.

# The Smoothed Concomitant Lasso<sup>(7)</sup>

## (i.i.d. case, Single-Task)

$$\hat{\beta}^{(\lambda)} \in \arg \min_{\beta \in \mathbb{R}^p} \underbrace{\frac{1}{\sqrt{n}} \|y - X\beta\|_2}_{\text{non-smooth}} + \lambda \underbrace{\|\beta\|_1}_{\text{non-smooth}}$$

Idea: replacing  $\|\cdot\|_2$  by  $\underbrace{\|\cdot\|_2 \square \underline{\sigma} \omega\left(\frac{\cdot}{\underline{\sigma}}\right)}_{\text{smooth}}(z) = \min_{\sigma \geq \underline{\sigma}} \left( \frac{\|z\|_2^2}{2\sigma} + \frac{\sigma}{2} \right)$

$$(\hat{\beta}^{(\lambda)}, \hat{\sigma}^{(\lambda)}) \in \arg \min_{\beta \in \mathbb{R}^p, \sigma \geq \underline{\sigma}} \frac{\|y - X\beta\|_2^2}{2n\sigma} + \frac{\sigma}{2} + \lambda \|\beta\|_1$$

► jointly convex: alternate minimization

Question: can this estimator (with unknown  $\sigma^*$ ) generalize for correlated Gaussian noise?

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<sup>(7)</sup>E. Ndiaye et al. "Efficient Smoothed Concomitant Lasso Estimation for High Dimensional Regression". In: *Journal of Physics: Conference Series* 904.1 (2017), p. 012006.

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# Generalization ? Yes !

## (correlated Gaussian noise, Multi-Task)

$$\text{SGCL}^{(8)}: (\hat{\mathbf{B}}^{\text{SGCL}}, \hat{\mathbf{S}}^{\text{SGCL}}) \in \arg \min_{\substack{\mathbf{B} \in \mathbb{R}^{p \times q} \\ \mathbf{S} \in \mathbb{S}_{++}^n, \mathbf{S} \succeq \underline{\sigma}}} \underbrace{\frac{\|\bar{\mathbf{Y}} - \mathbf{X}\mathbf{B}\|_{\mathbf{S}^{-1}}^2}{2nq}}_{\text{smooth}} + \frac{\text{Tr}(\mathbf{S})}{2n} + \underbrace{\lambda \|\mathbf{B}\|_{2,1}}_{\text{separable}}$$

### Benefits

- ▶ jointly convex formulation

### Drawbacks:

- ▶ Statistically:  $\mathcal{O}(n^2)$  parameters to estimate for  $\mathbf{S}$  only  $nq$  observations

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<sup>(8)</sup> M. Massias et al. "Generalized Concomitant Multi-Task Lasso for Sparse Multimodal Regression". In: *AISTATS*. vol. 84. 2018, pp. 998–1007.

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# Can take advantage of repetitions? Yes!

**CLaR**<sup>(9)</sup>:

$$(\hat{\mathbf{B}}^{\text{CLaR}}, \hat{\mathbf{S}}^{\text{CLaR}}) \in \arg \min_{\substack{\mathbf{B} \in \mathbb{R}^{p \times q} \\ \mathbf{S} \in \mathbb{S}_{++}^n, \mathbf{S} \succeq \underline{\sigma}}} \frac{\sum_{l=1}^r \left\| \mathbf{Y}^{(l)} - \mathbf{X} \mathbf{B} \right\|_{\mathbf{S}^{-1}}^2}{2nqr} + \frac{\text{Tr}(\mathbf{S})}{2n} + \lambda \left\| \mathbf{B} \right\|_{2,1}$$

- Statistically:  $\mathcal{O}(n^2)$  parameters to estimate for  $\mathbf{S}$  with  $nqr$  observations ( $r$  = number of repetitions)

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<sup>(9)</sup> Bertrand\_Massias\_Gramfort\_Salmon19.

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### Proposition

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Link with the Trace norm<sup>(10)</sup>

$$\hat{\mathbf{B}}^{\text{CLaR}} = \arg \min_{\mathbf{B} \in \mathbb{R}^{p \times q}} (\|\cdot\|_{\text{Tr}} \square \omega_{\underline{\sigma}})(Z) + \lambda n \|\mathbf{B}\|_{2,1} .$$

where  $Z = \frac{1}{\sqrt{q}}[Y^{(1)} - \mathbf{X}\mathbf{B} | \dots | Y^{(r)} - \mathbf{X}\mathbf{B}]$ .

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- ▶ justification for the estimator introduced heuristically
- ▶ generalization of van de Geer<sup>(11)</sup>

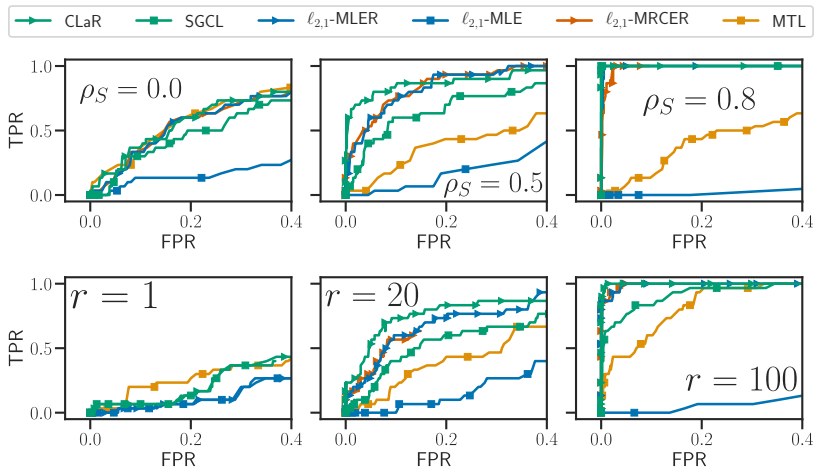
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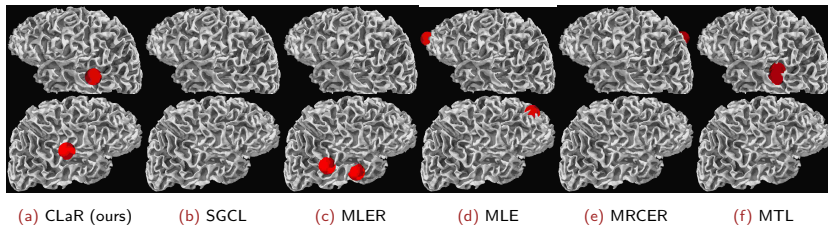
<sup>(11)</sup> S. van de Geer. *Estimation and testing under sparsity*. Vol. 2159. Lecture Notes in Mathematics. Lecture notes from the 45th Probability Summer School held in Saint-Flour, 2015, École d'Été de Probabilités de Saint-Flour. Springer, 2016, pp. xiii+274.

# Simulated scenarios

- ▶  $n = 150, p = 500, q = 100$
- ▶  $X$  Toeplitz-correlated
- ▶  $S^*$  Toeplitz matrix:  $S^*_{i,j} = \rho_{S^*}^{|i-j|}$ ,  $\rho_{S^*} \in ]0, 1[$



# Real data



**Figure:** *Real data, left auditory stimulations* ( $n = 102$ ,  $p = 7498$ ,  $q = 76$ ,  $r = 63$ ) Sources found in the left hemisphere (top) and the right hemisphere (bottom) after left auditory stimulations.

- ▶ expected: 2 sources (one in each auditory cortex)
- ▶  $\lambda$  chosen such that  $\|\hat{\mathbf{B}}\|_{2,0} = 2$
- ▶ deep sources for SGCL and  $\ell_{2,1}$ -MRCER (not visible)

# Conclusion and perspectives

- ▶ New estimator to handle correlated noise and repetitions in Multi-Task
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# Merci!

*"All models are wrong but some come with good open source implementation and good documentation to use these."*

A. Gramfort

- ▶ Python code online for CLaR <https://github.com/QB3/CLaR>
- ▶ Papers: arXiv<sup>(12)</sup>,<sup>(13)</sup>



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# Competitors

- (smoothed)  $\ell_{2,1}$ -MLE

$$(\hat{\mathbf{B}}, \hat{\Sigma}) \in \arg \min_{\substack{\mathbf{B} \in \mathbb{R}^{p \times q} \\ \Sigma \succeq \underline{\sigma}^2 / r^2}} \left\| \bar{\mathbf{Y}} - X\mathbf{B} \right\|_{\Sigma^{-1}}^2 - \log \det(\Sigma^{-1}) + \lambda \|\mathbf{B}\|_{2,1} \quad ,$$

- and its repetitions version ( $\ell_{2,1}$ -MLER):

$$(\hat{\mathbf{B}}, \hat{\Sigma}) \in \arg \min_{\substack{\mathbf{B} \in \mathbb{R}^{p \times q} \\ \Sigma \succeq \underline{\sigma}^2}} \sum_1^r \left\| \mathbf{Y}^{(l)} - X\mathbf{B} \right\|_{\Sigma^{-1}}^2 - \log \det(\Sigma^{-1}) + \lambda \|\mathbf{B}\|_{2,1} \quad .$$

- $\ell_{2,1}$ -MLE and  $\ell_{2,1}$ -MLER are bi-convex but not jointly convex

# Smoothing of matrix norm

## Huber-like formula for the Frobenius norm

$$\|\cdot\|_F \square \underline{\sigma} \omega \left( \frac{\cdot}{\underline{\sigma}} \right) (Z) = \begin{cases} \frac{\|Z\|_F^2}{2\underline{\sigma}} + \frac{\underline{\sigma}}{2}, & \text{if } \|Z\|_F \leq \underline{\sigma} \\ \|Z\|_F, & \text{if } \|Z\|_F > \underline{\sigma} \end{cases}$$
$$= \min_{\sigma \geq \underline{\sigma}} \left( \frac{\|Z\|_F^2}{2\sigma} + \frac{\sigma}{2} \right)$$

What about other norms ?

# Smoothing of matrix norm

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## Huber-like formula for the nuclear/trace norm

$$\begin{aligned}\|\cdot\|_{s,1} \square \omega_{\underline{\sigma}}(Z) &= \begin{cases} \frac{\|Z\|_F^2}{2\underline{\sigma}} + \frac{\underline{\sigma}}{2} n \wedge q, & \text{if } \|Z\|_{\infty} \leq \underline{\sigma} \\ \frac{1}{2\underline{\sigma}} \sum_i \gamma_i^2 - (\gamma_i \wedge \underline{\sigma} - \gamma_i)^2, & \text{if } \|Z\|_{\infty} > \underline{\sigma} \end{cases} \\ &= \min_{S \succeq \underline{\sigma}} \frac{1}{2} \|Z\|_{S^{-1}}^2 + \frac{1}{2} \text{Tr}(S)\end{aligned}$$

$\gamma_i$ : singular values of  $Z$

$\|Z\|_{S^{-1}}^2 := \text{Tr}(Z^{\top} S^{-1} Z)$  Mahalanobis distance

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