# Concomitant Lasso with Repetitions (CLaR): beyond averaging multiple realizations of correlated noise

**Quentin Bertrand**

Joint work with:
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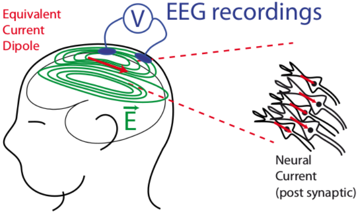
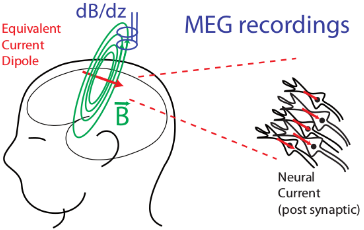
**Alexandre Gramfort** (INRIA)
**Joseph Salmon** (IMAG, Univ Montpellier, CNRS)

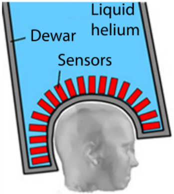
# M/EEG inverse problem for brain imaging

- sensors: electric and magnetic fields during a cognitive task
- goal: which parts of the brain are responsible for the signals?
- applications: epilepsy treatment, brain aging, anesthesia risks



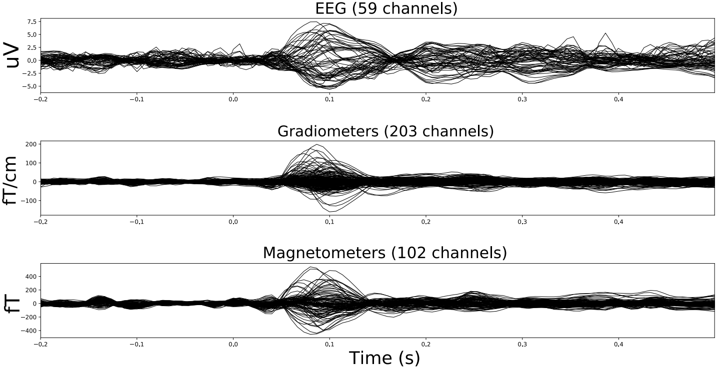




First EEG
recordings
in 1929
by H. Berger



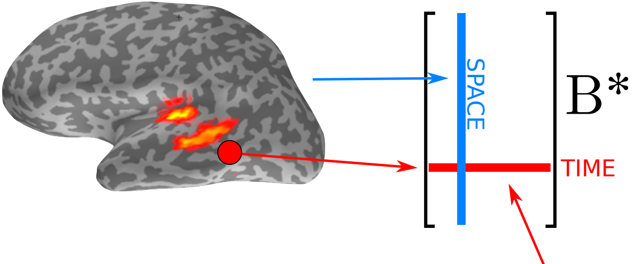

Hôpital La Timone
Marseille, France

# M/EEG data

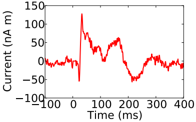



- 3 different types of sensor

# Source modeling (discretization with voxels)

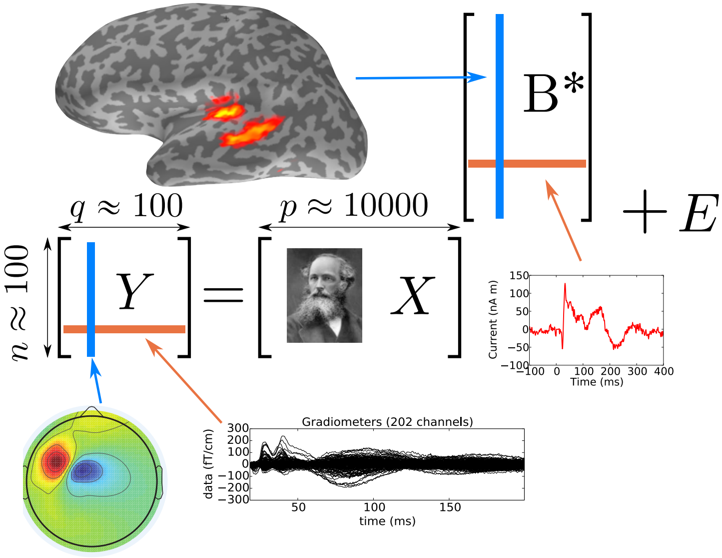



Position a few thousands candidate sources over the brain (*e.g.*, every 5mm)




$$B^* \in \mathbb{R}^{p \times q}$$

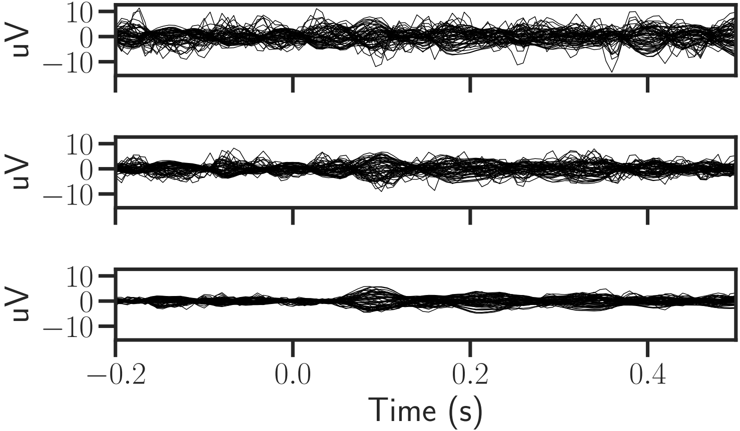
# The M/EEG inverse problem: modeling



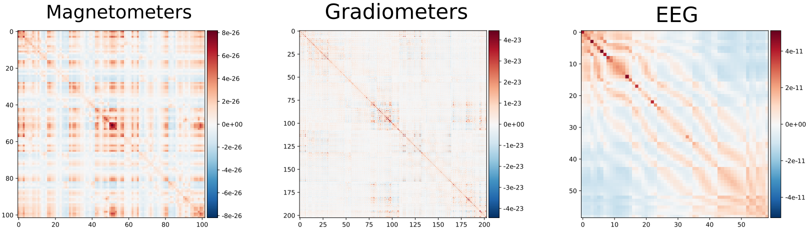

$$n << p$$

# Very noisy data: must repeat recordings

- average of 5 (top) / 10 (middle) / 50 (bottom) repetitions



# Noise covariance for each type of sensor




- 3 different sensors $\Longrightarrow$ 3 different noise structures

# A Multi-Task framework

Multi-Task regression notation:

- $n$ observations (*e.g.*, number of sensors)
- $q$ tasks (*e.g.*, temporal information)
- $p$ features
- $r$ number of repetitions
- $Y^{(1)}, \dots, Y^{(r)} \in \mathbb{R}^{n \times q}$ observation matrices; $\bar{Y} = \frac{1}{r} \sum_l Y^{(l)}$
- $X \in \mathbb{R}^{n \times p}$ design matrix (known)

$$\boxed{Y^{(l)} = X\mathrm{B}^* + S\mathrm{E}^{(l)}}$$

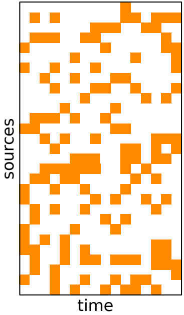
where

- $\mathrm{B}^* \in \mathbb{R}^{p \times q}$ : true source activity matrix (unknown)
- $S \in \mathbb{S}^n_{++}$ co-standard deviation matrix (unknown)
- $\mathrm{E}^{(1)}, \dots, \mathrm{E}^{(r)} \in \mathbb{R}^{n \times q}$ : white Gaussian noise

# Multi-Task penalties[1]

Popular convex penalties considered:

$$\hat{\mathrm{B}} \in \mathop{\arg\min}_{\mathrm{B} \in \mathbb{R}^{p \times q}} \left( \frac{1}{2nq} \|\bar{Y} - X\mathrm{B}\|_F^2 + \lambda \Omega(\mathrm{B}) \right)$$




Parameter $\hat{\mathrm{B}} \in \mathbb{R}^{p \times q}$

Sparse support: no structure

Penalty: **Lasso type**

$$\Omega(\mathrm{B}) = \|\mathrm{B}\|_1 = \sum_{j=1}^{p} \sum_{k=1}^{q} |\mathrm{B}_{j,k}|$$

[1] G. Obozinski, B. Taskar, and M. I. Jordan. "Joint covariate selection and joint subspace selection for multiple classification problems". In: *Statistics and Computing* 20.2 (2010), pp. 231–252.

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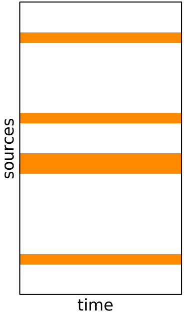


Parameter $\hat{\mathrm{B}} \in \mathbb{R}^{p \times q}$

Sparse support: group structure

Penalty: **Group-Lasso type**

$$\Omega(\mathrm{B}) = \|\mathrm{B}\|_{2,1} = \sum_{j=1}^{p} \|\mathrm{B}_{j,:}\|_2$$

where $\mathrm{B}_{j,:}$ the $j$-th row of $\mathrm{B}$

[1] G. Obozinski, B. Taskar, and M. I. Jordan. "Joint covariate selection and joint subspace selection for multiple classification problems". In: *Statistics and Computing* 20.2 (2010), pp. 231–252.

# Multi-Task data-fitting term

- Classical Multi-Task estimator: use averaged signal

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# Reminder on the Lasso theory[2][3] (i.i.d. case, Single-Task)

**Theorem**

- i.i.d. Gaussian noise
- $+\ X$ satisfying the "Restricted Eigenvalue" property
- $+\ \lambda = 2\sigma_* \sqrt{\frac{2\log(p/\delta)}{n}}$
- $\Longrightarrow$ with probability $1-\delta$:

$$\frac{1}{n}\left\| X\beta^* - X\hat{\beta}^{(\lambda)} \right\|^2 \leq \frac{18}{\kappa_{s^*}^2} \frac{\sigma_*^2 s^*}{n} \log\left(\frac{p}{\delta}\right)$$

**BUT** $\sigma_*$ is unknown in practice !

[2] P. J. Bickel, Y. Ritov, and A. B. Tsybakov. "Simultaneous analysis of Lasso and Dantzig selector". In: *Ann. Statist.* 37.4 (2009), pp. 1705–1732.

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# Reminder on the Square root Lasso[4][5][6] (i.i.d. case, Single-Task)

$$\hat{\beta} \in \arg\min_{\beta \in \mathbb{R}^p} \frac{1}{\sqrt{n}} \|y - X\beta\|_2 + \lambda \|\beta\|_1$$

**Theorem**

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$\lambda$ does not depend on $\sigma_*$ anymore!

[4] A. Belloni, V. Chernozhukov, and L. Wang. "Square-root Lasso: pivotal recovery of sparse signals via conic programming". In: *Biometrika* 98.4 (2011), pp. 791–806.

[5] T. Sun and C.-H. Zhang. "Scaled sparse linear regression". In: *Biometrika* 99.4 (2012), pp. 879–898.

[6] C. Giraud. *Introduction to high-dimensional statistics*. Vol. 138. CRC Press, 2014.

# The Smoothed Concomitant Lasso[7] (i.i.d. case, Single-Task)

$$\hat{\beta}^{(\lambda)} \in \underset{\beta \in \mathbb{R}^p}{\arg\min} \frac{1}{\sqrt{n}} \underbrace{\|y - X\beta\|_2}_{\text{non-smooth}} + \lambda \underbrace{\|\beta\|_1}_{\text{non-smooth}}$$

Idea: replacing $\|\cdot\|_2$ by $\underbrace{\|\cdot\|_2 \,\square\, \underline{\sigma}\,\omega\left(\frac{\cdot}{\underline{\sigma}}\right)}_{\text{smooth}}(z) = \min_{\sigma \geq \underline{\sigma}} \left(\frac{\|z\|_2^2}{2\sigma} + \frac{\sigma}{2}\right)$

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- jointly convex: alternate minimization

Question: can this estimator (with unknown $\sigma^*$) generalize for correlated Gaussian noise?

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# Generalization ? Yes ! (correlated Gaussian noise, Multi-Task)

**SGCL**[8]:

$$(\hat{\mathrm{B}}^{\mathrm{SGCL}}, \hat{S}^{\mathrm{SGCL}}) \in \mathop{\arg\min}_{\substack{\mathrm{B}\in\mathbb{R}^{p\times q}\\ S\in\mathbb{S}^n_{++}, S\succeq \underline{\sigma}}} \underbrace{\frac{\|\bar{Y} - X\mathrm{B}\|^2_{S^{-1}}}{2nq}}_{\text{smooth}} + \frac{\mathrm{Tr}(S)}{2n} + \underbrace{\lambda \|\mathrm{B}\|_{2,1}}_{\text{separable}}$$

Benefits

- jointly convex formulation

Drawbacks:

- Statistically: $\mathcal{O}(n^2)$ parameters to estimate for $S$ only $nq$ observations

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# Can take advantage of repetitions? Yes!

**CLaR**[9]:

$$(\hat{\mathrm{B}}^{\mathrm{CLaR}}, \hat{S}^{\mathrm{CLaR}}) \in \mathop{\arg\min}_{\substack{\mathrm{B} \in \mathbb{R}^{p \times q} \\ S \in \mathbb{S}^n_{++}, S \succeq \underline{\sigma}}} \frac{\sum_{l=1}^{r} \left\| Y^{(l)} - X\mathrm{B} \right\|^2_{S^{-1}}}{2nqr} + \frac{\mathrm{Tr}(S)}{2n} + \lambda \left\| \mathrm{B} \right\|_{2,1}$$

- Statistically: $\mathcal{O}(n^2)$ parameters to estimate for $S$ with $nqr$ observations ($r$ = number of repetitions)

[9] **Bertrand_Massias_Gramfort_Salmon19**.

**Proposition**

Link with the Trace norm[10]

$$\hat{\mathrm{B}}^{\mathrm{CLaR}} = \underset{\mathrm{B} \in \mathbb{R}^{p \times q}}{\arg\min} \left( \|\cdot\|_{\mathrm{Tr}} \,\square\, \omega_{\underline{\sigma}} \right)(Z) + \lambda n \|\mathrm{B}\|_{2,1} \quad .$$

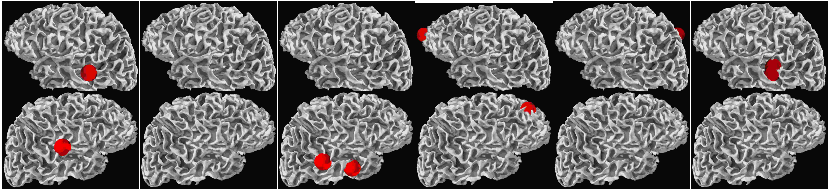
where $Z = \frac{1}{\sqrt{q}}[Y^{(1)} - X\mathrm{B}| \dots |Y^{(r)} - X\mathrm{B}]$.

- justification for the estimator introduced heuristically
- generalization of van de Geer[11]

(10) **Bertrand_Massias_Gramfort_Salmon19**.

(11) S. van de Geer. *Estimation and testing under sparsity*. Vol. 2159. Lecture Notes in Mathematics. Lecture notes from the 45th Probability Summer School held in Saint-Four, 2015, École d'Été de Probabilités de Saint-Flour. Springer, 2016, pp. xiii+274.

# Real data



(a) CLaR (ours) (b) SGCL (c) MLER (d) MLE (e) MRCER (f) MTL

Figure: *Real data, left auditory stimulations ($n = 102$, $p = 7498$, $q = 76$, $r = 63$)* Sources found in the left hemisphere (top) and the right hemisphere (bottom) after left auditory stimulations.

- expected: 2 sources (one in each auditory cortex)
- $\lambda$ chosen such that $\|\hat{\mathrm{B}}\|_{2,0} = 2$
- deep sources for SGCL and $\ell_{2,1}$-MRCER (not visible)

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# Merci!

"*All models are wrong but some come with good open source implementation and good documentation to use these.*"

A. Gramfort

- Python code online for CLaR https://github.com/QB3/CLaR
- Papers: arXiv[12],[13]

[12] M. Massias et al. "Generalized Concomitant Multi-Task Lasso for Sparse Multimodal Regression". In: *AISTATS*. vol. 84. 2018, pp. 998–1007.

[13] **Bertrand_Massias_Gramfort_Salmon19**.

# Competitors

- (smoothed) $\ell_{2,1}$-MLE

$$(\hat{\mathrm{B}}, \hat{\Sigma}) \in \underset{\substack{\mathrm{B} \in \mathbb{R}^{p \times q} \\ \Sigma \succeq \underline{\sigma}^2 / r^2}}{\arg\min} \left\| \bar{Y} - X\mathrm{B} \right\|_{\Sigma^{-1}}^2 - \log\det(\Sigma^{-1}) + \lambda \left\| \mathrm{B} \right\|_{2,1} \quad ,$$

- and its repetitions version ($\ell_{2,1}$-MLER):

$$(\hat{\mathrm{B}}, \hat{\Sigma}) \in \underset{\substack{\mathrm{B} \in \mathbb{R}^{p \times q} \\ \Sigma \succeq \underline{\sigma}^2}}{\arg\min} \sum_{1}^{r} \left\| Y^{(l)} - X\mathrm{B} \right\|_{\Sigma^{-1}}^2 - \log\det(\Sigma^{-1}) + \lambda \left\| \mathrm{B} \right\|_{2,1} \quad .$$

- $\ell_{2,1}$-MLE and $\ell_{2,1}$-MLER are bi-convex but not jointly convex

# Smoothing of matrix norm

**Huber-like formula for the Frobenius norm**

$$\left\|\cdot\right\|_F \,\square\, \underline{\sigma}\,\omega\left(\frac{\cdot}{\underline{\sigma}}\right)(Z) = \begin{cases} \frac{\left\|Z\right\|_F^2}{2\underline{\sigma}} + \frac{\underline{\sigma}}{2}, & \text{if } \left\|Z\right\|_F \le \underline{\sigma} \\ \left\|Z\right\|_F, & \text{if } \left\|Z\right\|_F > \underline{\sigma} \end{cases}$$

$$= \min_{\sigma \ge \underline{\sigma}} \left(\frac{\left\|Z\right\|_F^2}{2\sigma} + \frac{\sigma}{2}\right)$$

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What about other norms ?

**Huber-like formula for the nuclear/trace norm**

$$\|\cdot\|_{s,1} \,\Box\, \omega_{\underline{\sigma}}(Z) = \begin{cases} \frac{\|Z\|_F^2}{2\underline{\sigma}} + \frac{\underline{\sigma}}{2} n \wedge q, & \text{if } \|Z\|_\infty \leq \underline{\sigma} \\ \frac{1}{2\underline{\sigma}} \sum_i \gamma_i^2 - (\gamma_i \wedge \underline{\sigma} - \gamma_i)^2, & \text{if } \|Z\|_\infty > \underline{\sigma} \end{cases}$$

$$= \min_{S \succeq \underline{\sigma}} \frac{1}{2} \|Z\|_{S^{-1}}^2 + \frac{1}{2} \operatorname{Tr}(S)$$

$\gamma_i$: singular values of $Z$

$\|Z\|_{S^{-1}}^2 := \operatorname{Tr}(Z^\top S^{-1} Z)$ **Mahalanobis distance**

# Smoothing of matrix norm

## Huber-like formula for the Frobenius norm

$$\|\cdot\|_F \,\square\, \underline{\sigma}\,\omega\left(\frac{\cdot}{\underline{\sigma}}\right)(Z) = \begin{cases} \frac{\|Z\|_F^2}{2\underline{\sigma}} + \frac{\underline{\sigma}}{2}, & \text{if } \|Z\|_F \leq \underline{\sigma} \\ \|Z\|_F\,, & \text{if } \|Z\|_F > \underline{\sigma} \end{cases}$$

$$= \min_{\sigma \geq \underline{\sigma}} \left( \frac{\|Z\|_F^2}{2\sigma} + \frac{\sigma}{2} \right)$$

What about other norms ?

## Huber-like formula for the nuclear/trace norm

$$\|\cdot\|_{s,1} \,\square\, \omega_{\underline{\sigma}}(Z) = \begin{cases} \frac{\|Z\|_F^2}{2\underline{\sigma}} + \frac{\underline{\sigma}}{2} n \wedge q, & \text{if } \|Z\|_\infty \leq \underline{\sigma} \\ \frac{1}{2\underline{\sigma}} \sum_i \gamma_i^2 - (\gamma_i \wedge \underline{\sigma} - \gamma_i)^2, & \text{if } \|Z\|_\infty > \underline{\sigma} \end{cases}$$

$$= \min_{S \succeq \underline{\sigma}} \tfrac{1}{2} \|Z\|_{S^{-1}}^2 + \tfrac{1}{2} \operatorname{Tr}(S)$$

$\gamma_i$: singular values of $Z$

$\|Z\|_{S^{-1}}^2 := \operatorname{Tr}(Z^\top S^{-1} Z)$ **Mahalanobis distance**

# Simulated scenarios

- $n = 150$, $p = 500$, $q = 100$
- $X$ Toeplitz-correlated
- $S^*$ Toeplitz matrix: ${S^*}_{i,j} = \rho_{S^*}^{|i-j|}$, $\rho_{S^*} \in ]0, 1[$