

A constrained ℓ_1 minimization approach for estimating multiple Sparse Gaussian or Nonparanormal Graphical Models

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Outline

1 Introduction

- Motivation
- Previous Studies

2 Method

- Proposed Model: SIMULE
- Solution and Variation

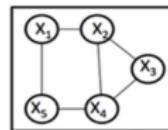
3 Theoretical and Experimental Results

- Theoretical Results
- Experimental Results

Motivation: Structure Learning from Heterogeneous Samples

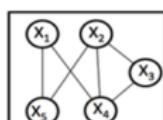
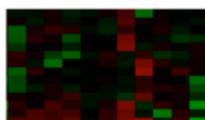
- Learning relational graph structure among features/variables from an observed sample dataset is an important task in Machine Learning.

Context/Task(1)



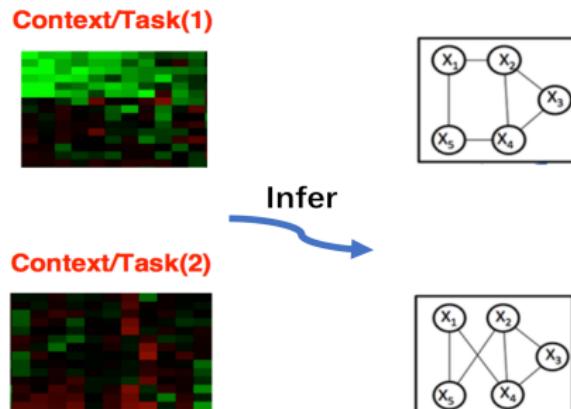
Infer

Context/Task(2)



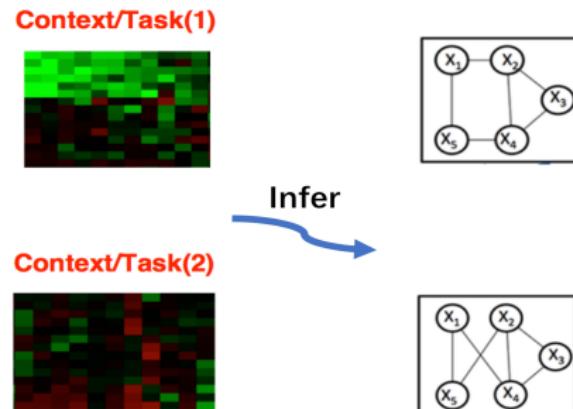
Motivation: Structure Learning from Heterogeneous Samples

- Learning relational graph structure among features/variables from an observed sample dataset is an important task in Machine Learning.
- This paper focuses on inferring graph structures from **multiple related datasets** (heterogeneous samples) about the same set of variables.



Motivation: Structure Learning from Heterogeneous Samples

- Learning relational graph structure among features/variables from an observed sample dataset is an important task in Machine Learning.
- This paper focuses on inferring graph structures from **multiple related datasets** (heterogeneous samples) about the same set of variables.
- We mainly focus on estimating **conditional dependency graphs** using the **sparse Gaussian Graphical Model (sGGM)**.



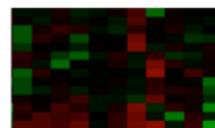
When Working on Multiple Different but Related Datasets:

- Samples of many real applications take the form of multiple **different** but **related** data matrices.
 - Blood cancer samples vs. Breast cancer samples;
 - Normal patient samples vs. Cancel patient samples;

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Context/Task(2)



Case I:



Case II:



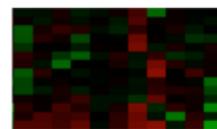
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- Samples of many real applications take the form of multiple **different** but **related** data matrices.
 - Blood cancer samples vs. Breast cancer samples;
 - Normal patient samples vs. Cancel patient samples;
- A **multi-task** learning setting: to investigate the **commonalities** and **differences** among different datasets.

Context/Task(1)



Context/Task(2)



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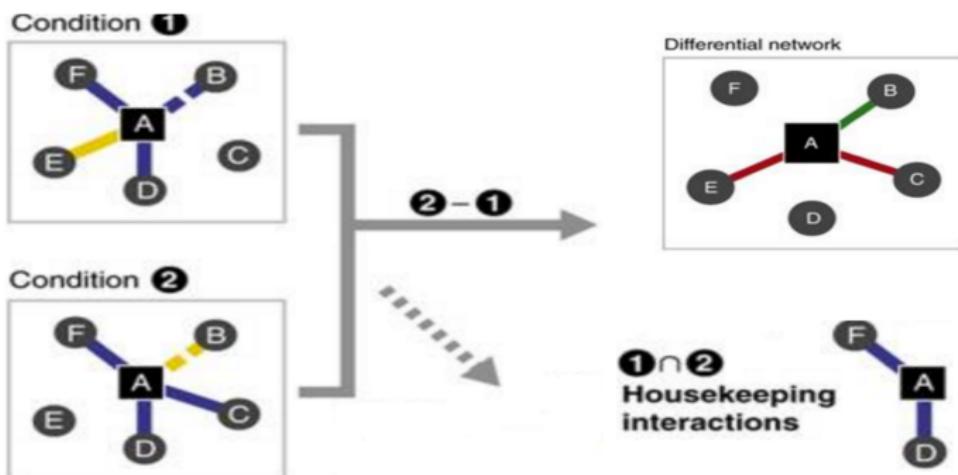


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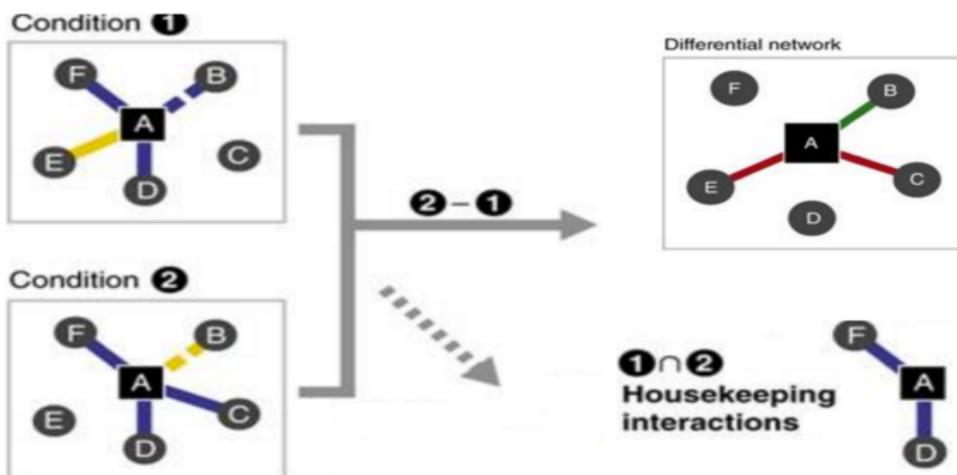
Our Aim: Shared and Task-specific Graph Structures

- We aim to obtain **shared** and **task-specific** graph structures from heterogeneous samples.



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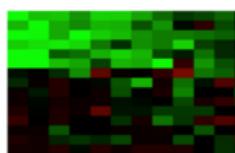
- We aim to obtain **shared** and **task-specific** graph structures from heterogeneous samples.
- For example, in computational biology [Ideker and Krogan(2012)] urges to estimate **housekeeping interactions** and **differential network** among genes or proteins.



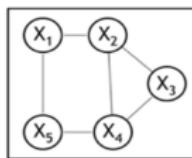
Our Aim: To Learn Shared and Task-specific Graph Structures from Multiple Related Datasets

- Main Task: How to estimate / learn **shared** (Ω_S) and **task-specific** ($\Omega_I^{(i)}$) graph structures among feature variables from multiple **different** but **related** datasets about the same set of features.

Context/Task(1)

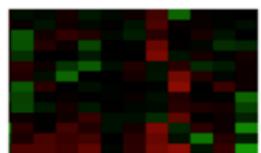


$$(x_1^{(1)}, x_2^{(1)}, \dots, x_p^{(1)}) \in \mathbb{R}^p$$

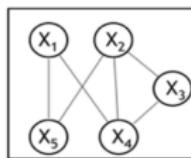


$$\Omega^1$$

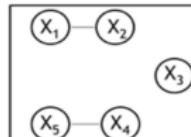
Context/Task(2)



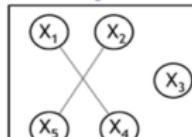
$$(x_1^{(2)}, x_2^{(2)}, \dots, x_p^{(2)}) \in \mathbb{R}^p$$



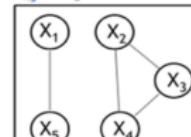
$$\Omega^2$$



$$\Omega_I^1$$



$$\Omega_I^2$$



$$\Omega_S$$

Individual(1) Individual(2)

Shared

Notations

- \mathbf{X} Data matrix.
- Σ Covariance matrix.
- Ω Inverse of covariance matrix (precision matrix).
- $\mathbf{X}^{(i)}$ The i -th data matrix.
- $\Sigma^{(i)}$ The i -th covariance matrix.
- $\Omega^{(i)}$ The i -th precision matrix.
- p The total number of feature variables.
- n_i The number of samples in the i -th data matrix.
- K The total number of tasks.

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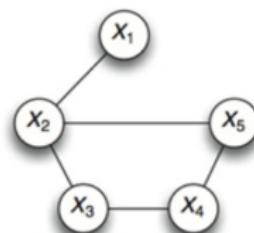
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Background: Sparse Gaussian Graphical Model (sGGM)

- $X \sim N(\mu, \Sigma)$.

Inverse Covariance Matrix

$$\begin{pmatrix} 1 & 0.2 & 0 & 0 & 0 \\ 0.2 & 1 & 0.2 & 0 & 0.2 \\ 0 & 0.2 & 1 & 0.2 & 0 \\ 0 & 0 & 0.2 & 1 & 0.2 \\ 0 & 0.2 & 0 & 0.2 & 1 \end{pmatrix}$$

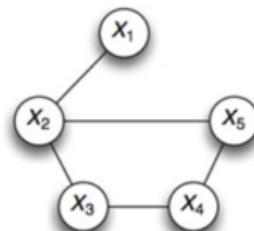


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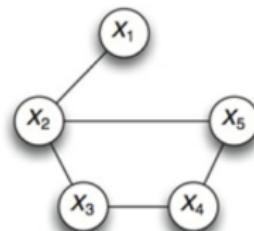


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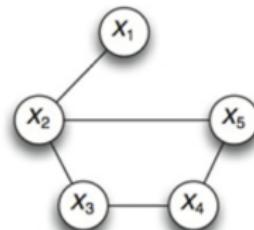


Background: Sparse Gaussian Graphical Model (sGGM)

- $X \sim N(\mu, \Sigma)$.
- Covariance matrix Σ can be calculated from X
- Precision matrix Ω is the inverse of covariance matrix Σ
- The sparsity pattern of Ω captures the conditional dependency pattern among variables.
- For example,

Inverse Covariance Matrix

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Background: Graphical Lasso for sGGM Structure Learning

- Traditionally, we estimate sGGM from samples (of a single task) using an ℓ_1 penalized MLE formulation.

Graphical Lasso

[Friedman et al.(2008) Friedman, Hastie, and Tibshirani]

$$\operatorname{argmin}_{\Omega} -\ln \det(\Omega) + \text{tr} \left(\Omega \widehat{\Sigma} \right) + \lambda_n \|\Omega\|_1 \quad (1.1)$$

Previous Methods: Joint Graphical Lasso (JGL) for Jointly Estimating Multiple sGGMs

- Most previous studies add **a second penalty function $P()$ into** the penalized likelihood formulation.

Joint Graphical Lasso (JGL)
[Danaher et al.(2013) Danaher, Wang, and Witten]

$$\begin{aligned} \operatorname{argmin}_{\Omega^{(i)}} & - \sum_i n_i (\ln \det(\Omega^{(i)}) + \text{tr}(\Omega^{(i)} \widehat{\Sigma}^{(i)})) \\ & + \lambda_1 \sum_i \|\Omega^{(i)}\|_1 + \lambda_2 P(\Omega^{(1)}, \Omega^{(2)}, \dots, \Omega^{(K)}) \end{aligned} \quad (1.2)$$

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- $P(\Omega^{(1)}, \Omega^{(2)}, \dots, \Omega^{(K)})$ captures a certain assumption about relationships between multiple graphs.
- For example, **fused norm** to push graphs similar:
$$P(\Omega^{(1)}, \Omega^{(2)}, \dots, \Omega^{(K)}) = \sum_{i>j} ||\Omega^{(i)} - \Omega^{(j)}||_1.$$

Joint Graphical Lasso (JGL)

[Danaher et al.(2013) Danaher, Wang, and Witten]

$$\begin{aligned} & \underset{\Omega^{(i)}}{\operatorname{argmin}} - \sum_i n_i (\ln \det(\Omega^{(i)}) + \operatorname{tr} (\Omega^{(i)} \widehat{\Sigma}^{(i)})) \\ & + \lambda_1 \sum_i ||\Omega^{(i)}||_1 + \lambda_2 P(\Omega^{(1)}, \Omega^{(2)}, \dots, \Omega^{(K)}) \end{aligned} \tag{1.2}$$

Previous Studies: Drawbacks

- Two possible ways to infer multiple sGGMs from heterogeneous samples:
 - (1) Estimating one by one using graphical lasso by assuming the graphs are not related.
 - (2) Using JGL: joint graphical lasso by designing the appropriate second penalty function $P()$.

Previous Studies: Drawbacks

- Two possible ways to infer multiple sGGMs from heterogeneous samples:
 - (1) Estimating one by one using graphical lasso by assuming the graphs are not related.
 - (2) Using JGL: joint graphical lasso by designing the appropriate second penalty function $P()$.
- **Drawbacks:**
 - **I:** Both of them **can not directly output the shared structure among multiple graphs.**
 - **II:** Need extra steps to decode and can not control estimating the shared and task-specific pattern among graphs.
 - **III: No theoretical analysis** in the previous JGL studies to prove why jointly learning graphs is helpful?

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Goals

Our model aims to have the following properties:

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- It provides a strong **theoretical guarantee**.

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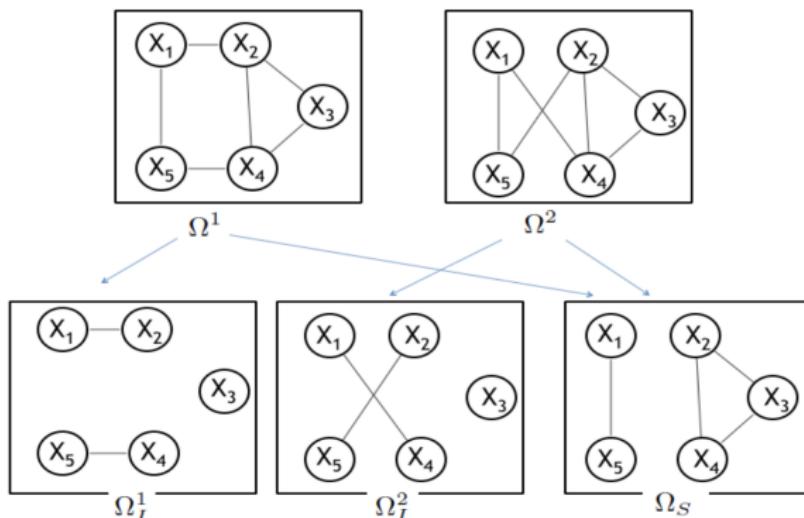
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- It estimates the shared and task-specific graph patterns **explicitly** and simultaneously.
- It can **control** the estimation of shared versus the task-specific patterns.
- It provides a strong **theoretical guarantee**.
- It achieves **good empirical** performance.

Proposed Method: Our "SIMULE" Formulation

We model each task's precision matrix $\Omega^{(i)}$ as a sum of task-specific $\Omega_I^{(i)}$ and task-shared Ω_S :

$$\Omega^{(i)} = \Omega_I^{(i)} + \Omega_S \quad (2.1)$$



Proposed method: Overview Figure

X^1_{p*n}

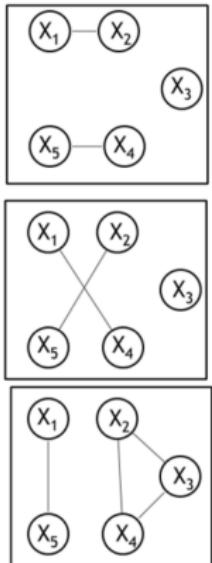
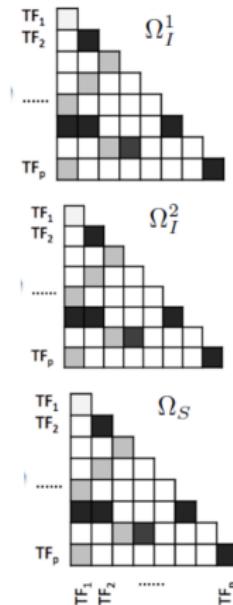
$\Sigma = \text{Cov}(X) =$

$$\begin{bmatrix} \sigma_{11} & \sigma_{12} & \cdots & \sigma_{1n} \\ \sigma_{21} & \sigma_{22} & \cdots & \sigma_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{n1} & \sigma_{n2} & \cdots & \sigma_{nn} \end{bmatrix}$$

X^2_{p*n}

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Why JGL Estimators Can't Get "SIMULE"

- JGL estimators are mostly solved by ADMM based optimization.

CLIME estimator [Cai et al.(2011)Cai, Liu, and Luo]

$$\operatorname{argmin}_{\Omega} \|\Omega\|_1 \quad (2.2)$$

Subject to: $\|\widehat{\Sigma}\Omega - I\|_{\infty} \leq \lambda_n$

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- With "SIMULE" formulation, **difficult to separate the optimization** into independent ADMM sub-procedures. Because,
 - The derivative of "SIMULE" in the JGL, i.e., gradient of $\ln \det(\Omega_I^{(i)} + \Omega_S)$ gets inverse of matrix summation.
 - Inverse of the summation of two matrices makes the optimization not separable.

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 - Inverse of the summation of two matrices makes the optimization not separable.
- Therefore, we use an **alternative formulation for sGGM: A constrained ℓ_1 minimization formulation.**

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SIMULE: to Infer Shared and Individual Parts of MULtiple sGGM Explicitly

- By using a constrained ℓ_1 minimization formulation, our estimator SIMULE can **jointly learn multiple graphs** from multiple **different** but **related** sample datasets (on the same set of feature variables).

SIMULE

$$\widehat{\Omega}_I^{(1)}, \widehat{\Omega}_I^{(2)}, \dots, \widehat{\Omega}_I^{(K)}, \widehat{\Omega}_S = \operatorname{argmin}_{\Omega_I^{(i)}, \Omega_S} \sum_i \|\Omega_I^{(i)}\|_1 + \epsilon K \|\Omega_S\|_1 \quad (2.3)$$

Subject to: $\|\widehat{\Sigma}^{(i)}(\Omega_I^{(i)} + \Omega_S) - I\|_\infty \leq \lambda_n, i = 1, \dots, K$

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Method: Optimization Solution

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- In detail, suppose $\beta^{(i)}, \beta^s$ are a column of $\Omega_I^{(i)}, \Omega_S$.

$$\operatorname{argmin}_{\beta^{(i)}, \beta^s} \sum_i \|\beta^{(i)}\|_1 + \epsilon K \|\beta^s\|_1 \quad (2.4)$$

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- We have proved the "SIMULE" formulation guarantees a unique optimal solution.
- We use ϵ to control the sparsity of shared versus task-specific graph patterns.

Model Variation: NSIMULE for jointly estimating multiple nonparanormal Graphical Models

- The Gaussian assumption of our model can extend easily to a **more general distribution** family: **nonparanormal**.

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- **The only necessary change:** by simply replacing the sample covariance matrices $\widehat{\Sigma}^{(i)}$ in Equation 2.3 into the kendal's tau correlation matrices $\widehat{\mathbf{S}}^{(i)}$.
- We denote this estimator as **nonparanormal SIMULE** (NSIMULE).

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Theoretical Results

- Comparing SIMULE v. CLIME w.r.t the statistical convergence rate for estimating K graphs:

Multi-task:	K Single-task:
$O\left(\frac{\log(Kp)}{n_{tot}}\right)$	$\sum_i O\left(\frac{\log p}{n_i}\right)$

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- By assuming $n_i = \frac{n_{tot}}{K}$:
- We can conclude that $\frac{\log(Kp)}{n_{tot}} < K \frac{\log p}{n_{tot}}$
- This indicates that the multi-task estimator is better!!!

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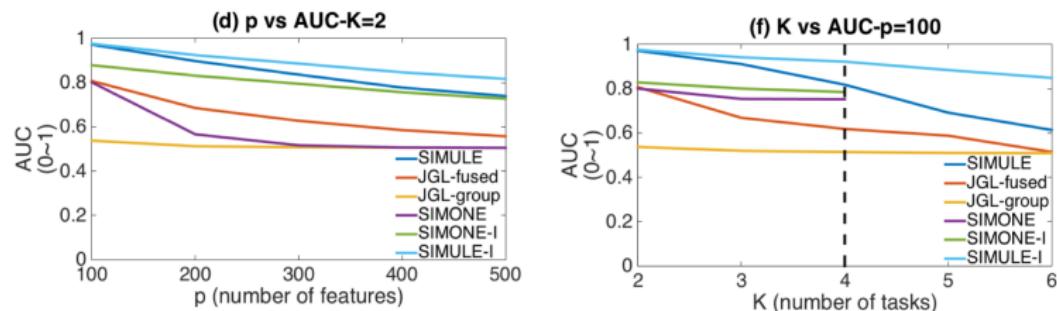
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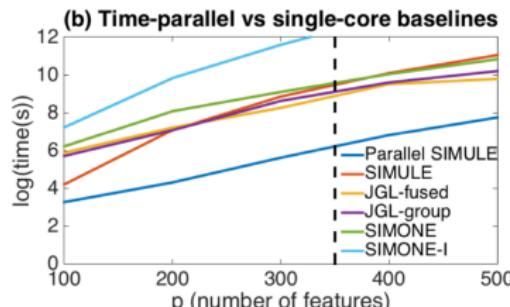
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Results on Synthetic Datasets: Accuracy and Parallelization

- Accuracy (AUC with a varying p and a varying K):

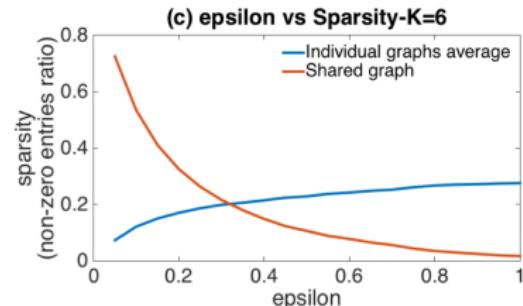
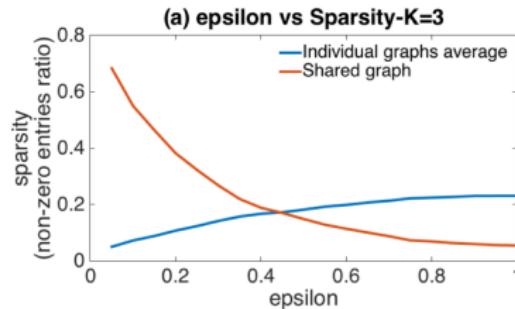


- Computation time cost with a varying p :

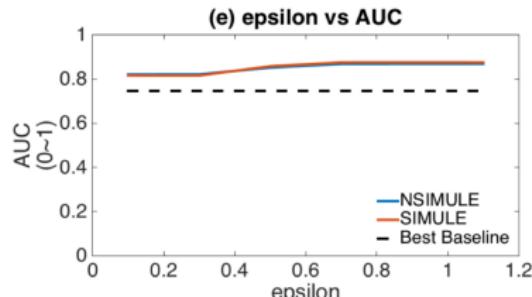


Results on Synthetic Datasets: Sensitivity of Hyperparameter ϵ

- The hyperpara ϵ controls the differences of sparsity among the shared graph and task-specific graphs.

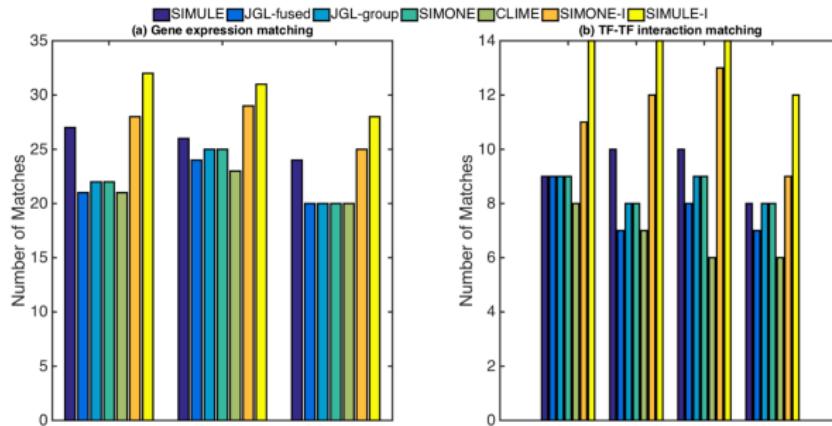


- The sensitivity of ϵ vs. accuracy.



Results on Two Real-World Datasets: Number of Matched Edges versus the Existing Domain Databases

- Two real world datasets:
 - (1) Gene expressions of samples in 2 different cell types
 - (2) Transcription Factors' ENCODE ChIP-seq measurements across 3 different cell lines
- Validation by counting the overlapped interactions according to the existing bio-databases (MInact).
- Our methods obtain the most matches compared to the state-of-the-art baselines.



R Package is Available !!!

- The project website: <http://jointggm.org/>
- R package "simule":
 - `install.packages("simule")`
 - `demo(simuleDemo) !`
 - `https://cran.r-project.org/web/packages/simule/index.html`

References

-  T. Cai, W. Liu, and X. Luo.
A constrained l1 minimization approach to sparse precision matrix estimation.
Journal of the American Statistical Association, 106(494):594–607, 2011.
-  P. Danaher, P. Wang, and D. M. Witten.
The joint graphical lasso for inverse covariance estimation across multiple classes.
Journal of the Royal Statistical Society: Series B (Statistical Methodology), 2013.
-  J. Friedman, T. Hastie, and R. Tibshirani.
Sparse inverse covariance estimation with the graphical lasso.
Biostatistics, 9(3):432–441, 2008.
-  T. Ideker and N. J. Krogan.
Differential network biology.
Molecular systems biology, 8(1):565, 2012.