

2.10.1

(1)

$$\begin{aligned}\bar{y} &= \frac{\sum_{i=1}^m y_0 i + \sum_{i=m+1}^{2m} y_1 i}{2m} \\ \bar{y}_0 + \bar{y}_1 &= \frac{1}{m} \sum_{i=1}^m y_0 i + \frac{1}{m} \sum_{i=m+1}^{2m} y_1 i = \frac{1}{m} (\sum_{i=1}^m y_0 i + \sum_{i=m+1}^{2m} y_1 i) \\ \therefore \bar{y} &= \frac{\bar{y}_0 + \bar{y}_1}{2}\end{aligned}$$

(2)

$$\bar{x} = \frac{\sum_{i=1}^m x_i + \sum_{i=m+1}^{2m} x_i}{2m} = \frac{0 + 1 \cdot m}{2m} = \frac{1}{2}$$

(3)

$$SXX = \sum_{i=1}^{2m} (x_i - \bar{x})x_i = \sum_{i=1}^m (x_i - \frac{1}{2})x_i + \sum_{i=m+1}^{2m} (x_i - \frac{1}{2})x_i = 0 + \frac{1}{2}m = \frac{m}{2}$$

(4)

$$\begin{aligned}SXY &= \sum_{i=1}^{2m} (x_i - \bar{x})y_i = \sum_{i=1}^m (x_i - \frac{1}{2})y_0 i + \sum_{i=m+1}^{2m} (x_i - \frac{1}{2})y_1 i = \sum_{i=1}^m -\frac{1}{2}y_0 i + \\ &\sum_{i=m+1}^{2m} \frac{1}{2}y_1 i = -\frac{1}{2}\bar{y}_0 m + \frac{1}{2}\bar{y}_1 m = \frac{m(\bar{y}_1 - \bar{y}_0)}{2}\end{aligned}$$

2.10.2

$$\hat{\beta}_1 = \frac{SXY}{SXX} = \frac{\frac{m(\bar{y}_1 - \bar{y}_0)}{\frac{m}{2}}}{\frac{m}{2}} = \bar{y}_1 - \bar{y}_0$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} = \frac{\bar{y}_0 + \bar{y}_1}{2} - (\bar{y}_1 - \bar{y}_0)\bar{x} = \bar{y}_0$$

OLS estimators minimize the residual sum of squares function, which least the inherent asymmetry in the roles of the response and the predictor in this problem.

2.10.3

For $x_i = 0$, where $i = 1, \dots, m$, all fitted values $= \hat{\beta}_0 + \hat{\beta}_1 x_i = \bar{y}_0 + (\bar{y}_1 - \bar{y}_0)x_i = \bar{y}_0$. Residuals $= y_i - \bar{y}_0$, where $i = 1, \dots, m$.

For $x_i = 1$, where $i = m+1, \dots, 2m$, all fitted values $= \hat{\beta}_0 + \hat{\beta}_1 x_i =$

$\bar{y}_0 + (\bar{y}_1 - \bar{y}_0)x_i = \bar{y}_1$. Residuals = $y_i - \bar{y}_1$, where $i = m + 1, \dots, 2m$.

$$RSS = \sum_{i=1}^{2m} (y_i - \bar{y}_0)^2 + \sum_{i=m+1}^{2m} (y_i - \bar{y}_1)^2 = SYY_0 + SYY_1$$

$$\hat{\sigma}^2 = \frac{RSS}{m_0 + m_1 - 2} = \frac{SYY_0 + SYY_1}{m_0 + m_1 - 2}$$

2.10.4

$$\begin{aligned} \text{t-statistic for test } \beta_1 = 0: \quad & \frac{\hat{\beta}_1}{\sigma \sqrt{\frac{1}{SXX}}} = \frac{\bar{y}_1 - \bar{y}_0}{\sqrt{\frac{RSS}{m_0 + m_1 - 2}} \sqrt{\frac{1}{SXX}}} = \frac{\bar{y}_1 - \bar{y}_0}{\sqrt{\frac{RSS}{m_0 + m_1 - 2}} \sqrt{\frac{2}{m}}} = \\ & \frac{\bar{y}_1 - \bar{y}_0}{\sqrt{\frac{RSS}{m_0 + m_1 - 2}} \sqrt{\frac{m_0 + m_1}{m_0 m_1}}} = \frac{\bar{y}_1 - \bar{y}_0}{\hat{\sigma} \sqrt{\frac{1}{m_0} + \frac{1}{m_1}}} \end{aligned}$$

The usual two-sample t-test for comparing two groups with an assumption of equal within-group variance:

$$\begin{aligned} & \frac{\bar{y}_1 - \bar{y}_0}{\sqrt{\frac{(m_0 - 1)SD_0^2 + (m_1 - 1)SD_1^2}{m_0 + m_1 - 2}} \sqrt{\frac{m_0 + m_1}{m_0 m_1}}} = \frac{\bar{y}_1 - \bar{y}_0}{\sqrt{\frac{(m_0 - 1)SD_0^2 + (m_1 - 1)SD_1^2}{m_0 + m_1 - 2}} \sqrt{\frac{2}{m}}} \\ & \because (m_0 - 1) \frac{SYY_0}{m_0 - 1} + (m_1 - 1) \frac{SYY_1}{m_1 - 1} = SYY_0 + SYY_1 = RSS \end{aligned}$$

\therefore they are the same.