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问题 1.

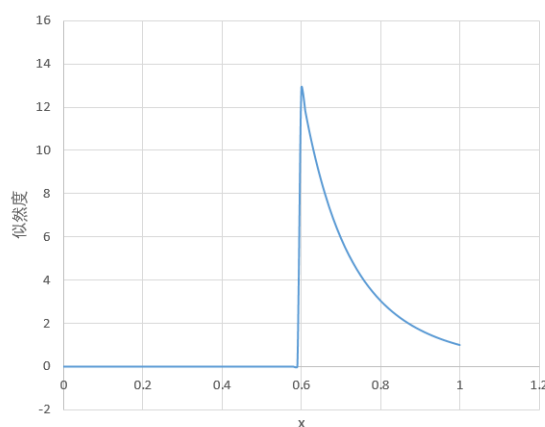
(1)

似然度函数： $L(\theta) = \prod_{i=1}^n p(x_i | \theta) = \frac{1}{\theta^n} \quad (\theta \geq \max\{D\})$

对似然度函数求偏导： $\frac{\partial L}{\partial \theta} = \frac{-n}{\theta^{n+1}} < 0$

由于似然度函数在定义域内单调递减，所以 $\hat{\theta} = \max\{D\}$

(2)



右图为似然度函数

由 (1) 式的似然度函数可知，当 $\theta < \max\{D\}$ 时，不可能出现样本 $\max\{D\}$ 。当且仅当 $\theta \geq \max\{D\}$ 时，样本集 D 的所有样本才会存在，所以不需要知道其他四个点的值。

问题 2.

(1)

后验估计：

$$\ln(l(\mu)p(\mu)) = -\frac{n}{2} \ln((2\pi)^d |\Sigma|) - \frac{1}{2} \sum_{k=1}^n (x_k - \mu)^\top \Sigma^{-1} (x_k - \mu)$$

$$- \frac{1}{2} \ln((2\pi)^d |\Sigma_0|) - \frac{1}{2} (\mu - m_0)^\top \Sigma_0^{-1} (\mu - m_0)$$

$$\text{求偏导得: } \frac{\partial \ln(l(\mu)p(\mu))}{\partial \mu} = \sum_{k=1}^n \Sigma^{-1} (x_k - \mu) + \Sigma_0^{-1} (m_0 - \mu)$$

$$\text{令偏导等于 0 得: } \hat{\mu} = (n\Sigma^{-1} + \Sigma_0^{-1})^{-1} (\sum_{k=1}^n \Sigma^{-1} x_k + \Sigma_0^{-1} m_0)$$

$$\text{即均值 } \mu \text{ 的 MAP 估计为: } \hat{\mu} = (n\Sigma^{-1} + \Sigma_0^{-1})^{-1} (\sum_{k=1}^n \Sigma^{-1} x_k + \Sigma_0^{-1} m_0)$$

(2)

$$\text{经过线性变换后, } \mu \sim N(Am_0, A\Sigma_0A^\top), x \sim N(A\mu, A\Sigma A^\top)$$

似然度函数：

$$\begin{aligned}\ln(l(\mu)p(\mu)) &= -\frac{n}{2} \ln((2\pi)^d |A\Sigma A^\top|) - \frac{1}{2} \sum_{k=1}^n (A(x_k - \mu))^\top (A\Sigma A^\top)^{-1} (A(x_k - \mu)) \\ &\quad - \frac{1}{2} \ln((2\pi)^d |A\Sigma_0 A^\top|) - \frac{1}{2} (A(\mu - m_0))^\top (A\Sigma_0 A^\top)^{-1} (A(\mu - m_0)) \\ &= K - \frac{1}{2} \sum_{k=1}^n (x_k - \mu)^\top A^\top (A^\top)^{-1} \Sigma^{-1} A^{-1} A(x_k - \mu) \\ &\quad - \frac{1}{2} (\mu - m_0)^\top A^\top (A^\top)^{-1} \Sigma_0^{-1} A^{-1} A(\mu - m_0) \\ &= K - \frac{1}{2} \sum_{k=1}^n (x_k - \mu)^\top \Sigma^{-1} (x_k - \mu) - \frac{1}{2} (\mu - m_0)^\top \Sigma_0^{-1} (\mu - m_0)\end{aligned}$$

$$\text{求偏导得: } \frac{\partial \ln(l(\mu)p(\mu))}{\partial \mu} = \sum_{k=1}^n \Sigma^{-1} (x_k - \mu) + \Sigma_0^{-1} (m_0 - \mu)$$

$$\text{令偏导等于 } 0 \text{ 得: } \hat{\mu} = (n\Sigma^{-1} + \Sigma_0^{-1})^{-1} (\sum_{k=1}^n \Sigma^{-1} x_k + \Sigma_0^{-1} m_0)$$

可得 μ 经过线性变换后, MAP 估计值与未变换的估计值相同, 即经过线性变换后, MAP 仍能正确估计 μ

问题 3.

(1)

$$\begin{aligned}Q(\theta; \theta^0) &= \ln P(x_1 | \theta) + \ln(x_2 | \theta) + \int_{-\infty}^{+\infty} \ln p(x_3 | \theta) \frac{P(2, x_{32} | \theta^0)}{\int_{-\infty}^{+\infty} p(2, x_{32} | \theta^0) dx_{32}} dx_{32} \\ &= -2 \ln \theta_1 - 2 \ln \theta_2 - 4\theta_1 + \int_{-\infty}^{+\infty} \ln p(x_3 | \theta) \frac{P(2, x_{32} | \theta^0)}{\frac{1}{\theta_1} e^{-2\theta_1}} dx_{32} \\ &= -2 \ln \theta_1 - 2 \ln \theta_2 - 4\theta_1 + \int \ln \frac{1}{\theta_1} e^{-2\theta_1} \frac{1}{\theta_2} \cdot \frac{P(2, x_{32} | \theta^0)}{\frac{1}{\theta_1} e^{-2\theta_1}} dx_{32}\end{aligned}$$

$$\text{当 } \theta_2 \geq 4 \text{ 时, } \int \ln \frac{1}{\theta_1} e^{-2\theta_1} \frac{1}{\theta_2} \frac{P(2, x_{32} | \theta^0)}{\frac{1}{\theta_1} e^{-2\theta_1}} dx_{32} = \int_0^4 \ln \frac{1}{\theta_1} e^{-2\theta_1} \frac{1}{\theta_2} \frac{1}{4} dx_{32} = -\ln \theta_1 - \ln \theta_2 - 2\theta_1$$

$$\text{当 } 3 \leq \theta_2 < 4 \text{ 时, } \int \ln \frac{1}{\theta_1} e^{-2\theta_1} \frac{1}{\theta_2} \cdot \frac{P(2, x_{32} | \theta^0)}{\frac{1}{\theta_1} e^{-2\theta_1}} dx_{32} = \int_0^{\theta_2} \ln \frac{1}{\theta_1} e^{-2\theta_1} \frac{1}{\theta_2} \cdot \frac{1}{4} dx_{32} = \frac{\theta_2}{4} (-\ln \theta_1 - \ln \theta_2 - 2\theta_1)$$

$$\text{当 } \theta_2 < 3 \text{ 时, } \int \ln \frac{1}{\theta_1} e^{-2\theta_1} \frac{1}{\theta_2} \cdot \frac{P(2, x_{32} | \theta^0)}{\frac{1}{\theta_1} e^{-2\theta_1}} dx_{32} = NULL$$

$$\text{综上所述, } Q(\theta; \theta^0) = \begin{cases} -3 \ln \theta_1 - 3 \ln \theta_2 - 6\theta_1 & 4 \geq \theta_2 \\ -2 \ln \theta_1 - 2 \ln \theta_2 - 4\theta_1 - \frac{\theta_1 \theta_2}{2} - \frac{\theta_2 \ln \theta_1}{4} - \frac{\theta_2 \ln \theta_2}{4} & 3 \leq \theta_2 < 4 \\ NULL & \text{else} \end{cases}$$

同时由概率密度函数积分为 1, 且 $\theta_1 > 0$:

$$\int_0^{+\infty} \frac{1}{\theta_1} e^{-x\theta_1} dx = -\frac{1}{\theta_1^2} e^{-x\theta_1} \Big|_0^{+\infty} = \frac{1}{\theta_1^2} = 1 \Rightarrow \theta_1 = 1$$

$$\text{于是: } Q(\theta; \theta^0) = \begin{cases} -3 \ln \theta_2 - 6 & 4 \geq \theta_2 \\ -2 \ln \theta_2 - 4 - \frac{\theta_2}{2} - \frac{\theta_2 \ln \theta_2}{4} & 3 \leq \theta_2 < 4 \\ NULL & \text{else} \end{cases}$$

(2)

当 $\theta_2 \geq 4$ 时, $\frac{\partial Q}{\partial \theta_2} < 0$, Q 随着 θ_2 的增大而减小, 于是此时 $\hat{\theta}_2 = 4, Q = -10.158883$

当 $3 \leq \theta_2 < 4$ 时, $\frac{\partial Q}{\partial \theta_2} = -\frac{2}{\theta_2} - \frac{1}{2} - \frac{\ln \theta_2 + 1}{4} < 0$, 此时 $\hat{\theta}_2 = 3, Q = -8.521183$

综上所述, $\hat{\theta}_2 = 3$

问题 4.

前向过程:

$$\alpha_i(1) = \pi_i b_i(O_1)$$

$$\alpha_i(t+1) = b_i(O_{t+1}) \sum_{j=1}^C \alpha_j(t) a_{ji}$$

式中 i 表示时刻为 t 时的状态, O_t 表示 t 时刻的观测值。 $\alpha_i(t)$ 表示 t 时刻状态为 i , 观测序列为 O_1, O_2, \dots, O_t 的概率。通过上式可得, 得到所有的 $\alpha_i(t)$ 复杂度为 $O(TC^2)$

反向过程:

$$\beta_i(T) = 1$$

$$\beta_i(t) = \sum_{j=1}^C \beta_j(t+1) a_{ij} b_j(O_{t+1})$$

$\beta_i(t)$ 表示 t 时刻下状态为 i , 剩余观测序列为 $O_{t+1}, O_{t+2}, \dots, O_T$ 的概率。通过上式可得, 得到所有的 $\beta_i(t)$ 复杂度为 $O(TC^2)$

$\gamma_i(t)$ 表示给定观测序列 O , 在 t 时刻状态为 i 的概率, 如下所示

$$\gamma_i(t) = P(X_t = i | O, \theta) = \frac{P(X_t = i, O | \theta)}{P(O | \theta)} = \frac{\alpha_i(t) \beta_i(t)}{\sum_{j=1}^C \alpha_j(t) \beta_j(t)}$$

$\xi_{ij}(t)$ 表示给定观测序列 O , 在 t 时刻状态为 i , $t+1$ 时刻状态为 j 的概率, 如下所示

$$\xi_{ij}(t) = P(X_t = i, X_{t+1} = j | O, \theta) = \frac{P(X_t = i, X_{t+1} = j, O | \theta)}{P(O | \theta)} = \frac{\alpha_i(t) a_{ij} \beta_j(t+1) b_j(O_{t+1})}{\sum_{j=1}^C \alpha_j(t) \beta_j(t)}$$

$$\text{定义公式 } 1_{O_t=v_k} = \begin{cases} 1 & \text{if } O_t = v_k \\ 0 & \text{otherwise} \end{cases}$$

$$\text{则更新公式为: } a_{ij}^* = \frac{\sum_{t=1}^{T-1} \xi_{ij}(t)}{\sum_{t=1}^{T-1} \gamma_i(t)}, \quad b_i^*(v_k) = \frac{\sum_{t=1}^T 1_{O_t=v_k} \gamma_i(t)}{\sum_{t=1}^T \gamma_i(t)}$$

计算 $\alpha_i(t)$ 、 $\beta_i(t)$ 复杂度均为 $O(TC^2)$ ，更新时 a_{ij}^* 与 $b_i^*(v_k)$ 复杂度均为 $O(TC^2)$ 。综上所述，更新一次全部的参数复杂度为 $O(4TC^2)$

问题 5.

(1)

$$\begin{aligned}
 \bar{p}_n(x) &= E(P_n(x)) \\
 &= \frac{1}{n} \sum_{i=1}^n E\left(\frac{1}{h_n} \varphi\left(\frac{x-x_i}{h_n}\right)\right) \\
 &= \int \frac{1}{h_n} \varphi\left(\frac{x-v}{h_n}\right) p(v) dv \\
 &= \int \frac{1}{h_n} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-v}{h_n}\right)^2} \cdot \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{v-u}{\sigma}\right)^2} dv \\
 &= \int \frac{1}{2\pi\sigma h_n} e^{-\frac{1}{2}\left(\frac{v^2-2uv+u^2}{\sigma^2} + \frac{x^2-2vx+v^2}{h_n^2}\right)} dv \\
 &= \int \frac{1}{2\pi\sigma h_n} e^{-\frac{1}{2}\left(\frac{(\sigma^2+h_n^2)v^2}{\sigma^2+h_n^2} - \frac{2uh_n^2+2x\sigma^2}{\sigma^2 h_n^2} v + \frac{u^2 h_n^2 + \sigma^2 x^2}{\sigma^2 h_n^2}\right)} dv \\
 &= e^{-\frac{1}{2}\left(\frac{u^2 h_n^2 + \sigma^2 x^2}{\sigma^2 h_n^2}\right)} \cdot \frac{1}{2\pi\sigma h_n} \int e^{-\frac{1}{2}\left(\frac{\sigma^2+h_n^2}{\sigma^2 h_n^2} v^2 - \frac{2uh_n^2+2x\sigma^2}{\sigma^2 h_n^2} v\right)} dv \\
 \text{记 } a^2 &= \frac{\sigma^2 h_n^2}{\sigma^2 + h_n^2} \quad b = \frac{uh_n^2 + x\sigma^2}{\sigma^2 h_n^2} \cdot \frac{\sigma^2 h_n^2}{\sigma^2 + h_n^2} \\
 &\int e^{-\frac{1}{2}\left(\frac{1}{a^2} v^2 - 2\frac{1}{a^2} bv\right)} dv \\
 &= \int e^{-\frac{1}{2}\frac{1}{a^2}(v^2 - 2bv)} dv \\
 &= \int e^{-\frac{1}{2}\left(\frac{v-b}{a}\right)^2} \cdot e^{\frac{b^2}{2a^2}} dv \\
 &= e^{\frac{b^2}{2a^2}} \sqrt{2\pi} a \\
 \bar{P}_n(x) &= \frac{\sqrt{2\pi}\sigma h_n}{\sqrt{\sigma^2 + h_n^2}} \cdot \frac{1}{2\pi\sigma h_n} e^{-\frac{1}{2}\left(\frac{x^2}{h_n^2} - \frac{\left(\frac{x\sigma^2 + uh_n^2}{\sigma^4 h_n^4}\right) \cdot \frac{\sigma^4 h_n^4}{(\sigma^2 + h_n^2)^2} + \frac{u^2}{\sigma^2}\right)} \\
 &= \frac{1}{\sqrt{2\pi}\sqrt{\sigma^2 + h_n^2}} \cdot e^{\frac{1}{2}\frac{(x-u)^2}{\sigma^2 + h_n^2}} \\
 \text{即: } \bar{P}_n(x) &\sim N(\mu, \sigma^2 + h_n^2)
 \end{aligned}$$

原式得证

(2)

$$\begin{aligned}\text{Var}(p(x)) &= \sum_{i=1}^n E \left[\left(\frac{1}{nh_n} \varphi \left(\frac{\mathbf{x} - \mathbf{x}_i}{h_n} \right) - \frac{1}{n} \bar{p}_n(\mathbf{x}) \right)^2 \right] \\ &= nE \left[\frac{1}{n^2 h_n^2} \varphi^2 \left(\frac{\mathbf{x} - \mathbf{x}_i}{h_n} \right) \right] - \frac{1}{n} \bar{p}_n^2(\mathbf{x}) \\ &= \frac{1}{nh_n^2} \int \varphi^2 \left(\frac{\mathbf{x} - \mathbf{v}}{h_n} \right) p(\mathbf{v}) d\mathbf{v} - \frac{1}{n} \bar{p}_n^2(\mathbf{x}) \\ &\approx \frac{1}{nh_n^2} \int \varphi^2 \left(\frac{x - v}{h_n} \right) p(v) dv\end{aligned}$$

$$\begin{aligned}&\int p^2 \left(\frac{x - v}{h_n} \right) p(v) dv \\ &= \int \frac{1}{2\pi} e^{-\left(\frac{x-v}{h_n}\right)^2} \cdot \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{v-u}{\sigma}\right)^2} dv \\ &= \frac{1}{(2\pi)^{\frac{3}{2}}\sigma} \int e^{-\left(\frac{v^2}{2\sigma^2} + \frac{v^2}{h_n^2} - \frac{uv}{\sigma^2} - \frac{2xv}{h_n^2} + \frac{x^2}{h_n^2} + \frac{u^2}{2\sigma^2}\right)} dv \\ &= \frac{\sqrt{\frac{\sigma^2 h_n^2}{h_n^2 + 2\sigma^2}}}{2\pi\sigma} e^{-\frac{1}{2}\frac{(x-u)^2}{\frac{h_n^2}{2} + \sigma^2}} \int \frac{1}{\sqrt{2\pi}\sqrt{\frac{\sigma^2 h_n^2}{h_n^2 + 2\sigma^2}}} e^{-\frac{1}{2}\left(\frac{v - \frac{uh_n^2 + 2x\sigma^2}{h_n^2 + 2\sigma^2}}{\sqrt{\frac{\sigma^2 h_n^2}{h_n^2 + 2\sigma^2}}}\right)^2} dv\end{aligned}$$

$$\begin{aligned}&= \frac{\sqrt{\frac{\sigma^2 h_n^2}{h_n^2 + 2\sigma^2}}}{2\pi\sigma} e^{-\frac{1}{2}\frac{(x-u)^2}{\frac{h_n^2}{2} + \sigma^2}} \\ &\frac{1}{nh_n^2} \int \varphi^2 \left(\frac{x - v}{h_n} \right) p(v) dv \\ &= \frac{1}{2nh_n\sqrt{\pi}} \cdot \frac{1}{\sqrt{2\pi}\sqrt{\frac{h_n^2}{2} + \sigma^2}} e^{-\frac{1}{2}\frac{(x-u)^2}{\frac{h_n^2}{2} + \sigma^2}} \\ &\approx \frac{1}{2nh_n\sqrt{\pi}} \cdot \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\frac{(x-u)^2}{\sigma^2}}\end{aligned}$$

$$= \frac{1}{2nh_n\sqrt{\pi}} p(x)$$

$$\text{即: } \text{Var}(p(x)) \approx \frac{1}{2nh_n\sqrt{\pi}} p(x)$$

原式得证