

Martin Lindquist

Department of Biostatistics
Johns Hopkins
Bloomberg School of Public Health

Tor Wager

Department of Psychology and
Neuroscience and the
Institute for Cognitive Science
University of Colorado, Boulder

GLM Estimation

GLM

A standard GLM can be written:

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon} \quad \boldsymbol{\varepsilon} \sim N(\mathbf{0}, \mathbf{V})$$

where

$$\begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & X_{11} & \cdots & X_{1p} \\ 1 & X_{21} & \cdots & X_{2p} \\ \vdots & \vdots & & \vdots \\ 1 & X_{n1} & \cdots & X_{np} \end{bmatrix} \times \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_p \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

↑ ↑ ↑ ↑

fMRI Data Design matrix Regression coefficients Noise

\mathbf{V} is the covariance matrix whose format depends on the noise model.

Problem Formulation

- Assume the model:

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon} \quad \boldsymbol{\varepsilon} \sim N(\mathbf{0}, \mathbf{I}\sigma^2)$$

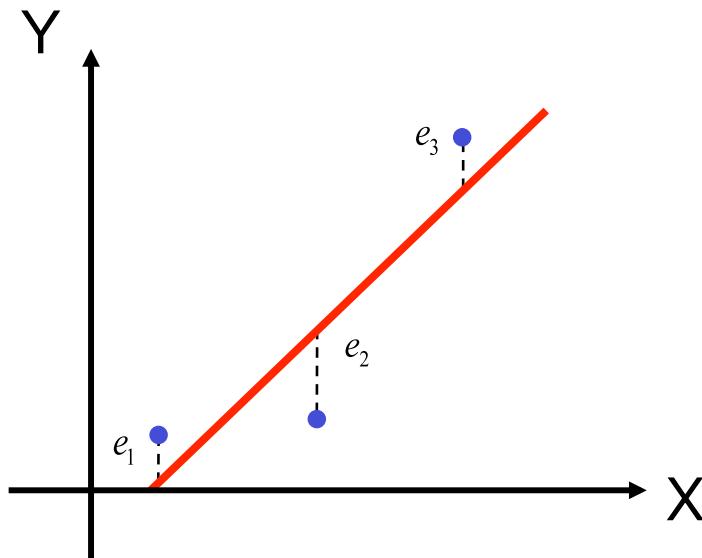
- The matrices \mathbf{X} and \mathbf{Y} are assumed to be known, and the noise is assumed to be uncorrelated.
- Our goal is to find the value of $\boldsymbol{\beta}$ that minimizes:

$$(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})^T (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})$$

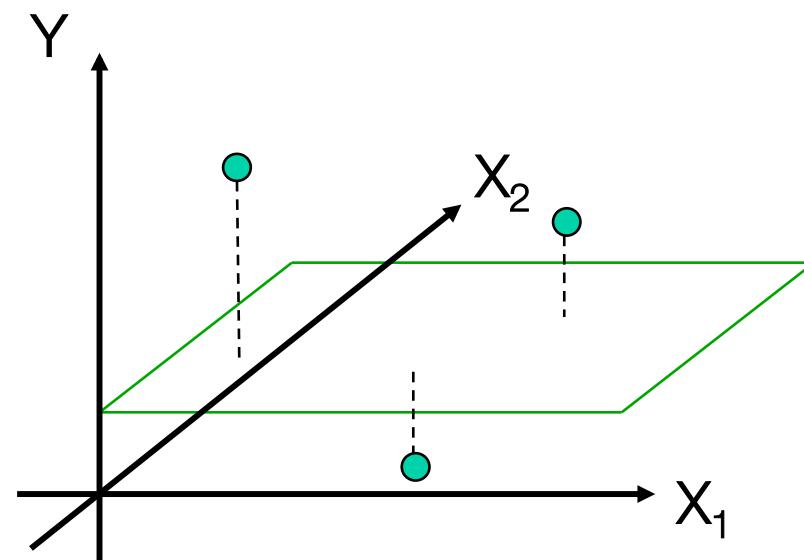
Sums of squared errors (SSE)

Least Squares Estimation

One explanatory variable X .



Two explanatory variables X_1 and X_2 .



Least Squares Estimation

- The least squares criterion:

$$Q = (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})$$

- Taking the derivative with respect to $\boldsymbol{\beta}$ and setting it to 0 gives us the normal equations:

$$\mathbf{X}'\mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}'\mathbf{Y}$$

- The ordinary least squares (OLS) estimators are given by

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$



OLS Solution

- Ordinary least squares solution:

$$\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

- Properties:

$$E(\hat{\beta}) = \beta \quad \text{Var}(\hat{\beta}) = \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1}$$

- Any other unbiased estimator of β will have a larger variance than the OLS solution.
 - Best linear unbiased estimator (BLUE)

Estimation

- If ϵ is i.i.d., then Ordinary Least Square (OLS) estimate is optimal

$$\text{model} \quad \mathbf{Y} = \mathbf{X}\beta + \epsilon \quad \rightarrow \quad \text{estimate} \quad \hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$

- If $\text{Var}(\epsilon) = V\sigma^2 \neq I\sigma^2$, then Generalized Least Squares (GLS) estimate is optimal

$$\text{model} \quad \mathbf{Y} = \mathbf{X}\beta + \epsilon \quad \rightarrow \quad \text{estimate} \quad \hat{\beta} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{Y}$$



Residuals

model estimate

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon} \quad \longrightarrow \quad \hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{Y}$$

fitted values residuals

$$\begin{aligned} \hat{\mathbf{Y}} &= \mathbf{X}\hat{\boldsymbol{\beta}} \\ \mathbf{r} &= \mathbf{Y} - \hat{\mathbf{Y}} \\ &= (\mathbf{I} - \mathbf{X}(\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1})\mathbf{Y} \\ &= \mathbf{R}\mathbf{Y} \end{aligned}$$



Estimating the Variance

- Even if we assume $\boldsymbol{\epsilon}$ is i.i.d., we still need to estimate the residual variance, σ^2 .
- Our estimate:

$$\hat{\sigma}^2 = \frac{\mathbf{r}^T \mathbf{r}}{tr(\mathbf{RV})}$$

- For OLS:

$$\hat{\sigma}^2 = \frac{\mathbf{r}^T \mathbf{r}}{N - p}$$

- Estimating $\mathbf{V} \neq \mathbf{I}$ more difficult.



A geometric interpretation of the GLM

Think of y as a vector in a space with one dimension per observation

Minimize sum of squared errors (Q):



$$Q = (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})$$

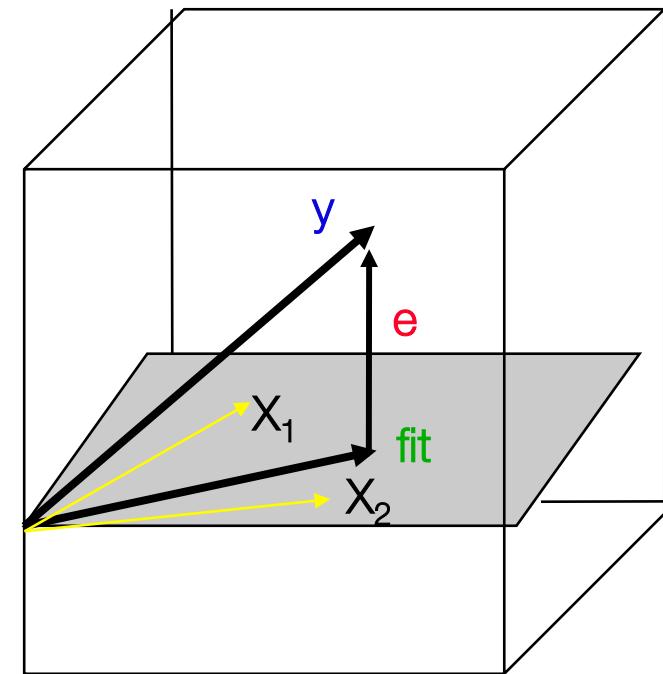
Solve for beta:

Project data (y) onto subspace of X

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$$

$$\hat{\boldsymbol{\beta}} = \mathbf{X}^{-1}\mathbf{y}$$

Projection matrix



- (pseudoinverse of X) * y

End of Module



@fMRIstats