# Specification and Comparison of Seasonal ARIMA models on the Unemployment Rate in the USA

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## I. Purpose

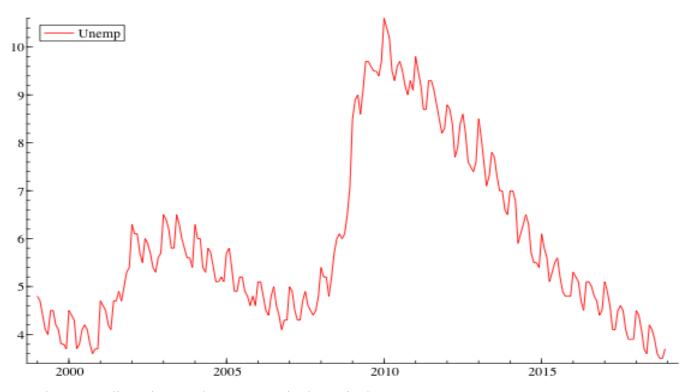
Seasonal Autoregressive Integrated Moving Average models are a way to model univariate data that exhibit seasonality and cycles. Seasonality refers to a pattern in the data that repeats after a set number of time steps. The season itself id the point/time at which the pattern repeats. The pattern is also called the cycle.

## II. Data

Our data comes from the (United States) Bureau of Labor Statistics. It's a simple record of the monthly unemployment rate. The data is not seasonally adjusted, since that's the goal of this exercise. Initially, I download the period from 1900-2018. There were many interesting things to observe in this series. However, the final timespan chosen to model was from December 2009 until December 2018.

## A. Initial Dataset 1990-2018

The first thing we'll take a look at is the full dataset. In Graph 01, we have the raw time-series. We can see what looks like three large multi-year trends. One trough finishes around late 2000. This is the start of the dot com bust. The second finished in late 2007. This is the start of the Great Recession. The final dip is for the end of 2018, with a low of 3.5 seen in November that year. Unfortunately, the pattern implies the rate will be on the rise again.



Graph 01: Unadjusted unemployment rate in the United States 1990 – 2018

The Bureau of Labor Statistics holds data going back to 1948. In the Jan-1948 – Jan-2019 dataset, the minimum value is a rate of 2.4 percent, in October 1952. May 1969 was 2.9%, and the lowest since has been 3.5 percent in (October and) November 2018. Another implication that an increase could be on the horizon.

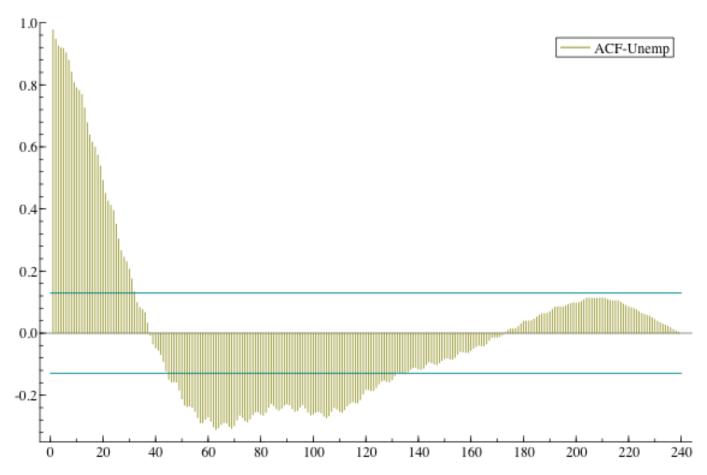
It's also useful to mention here that the data does not appear to be stationary. You could argue that it's somewhat mean-reverting (to 4%), but given the height and time span of the trend beyond 2008, that doesn't seem too accurate. Since the variance changes over time, the data is also heteroscedastic.

Our second graph (Graph 02) is the ACF of the unemployment rate. It uses a lag length equal to the full duration. We can see the autocorrelation between our final observation in 2018, and all months in previous years. There are very distinct "regime" changes, corresponding with what appeared to be our 3 large trends on Graph 01.

Regime 1: Dec-2018 to Dec-2015. Back towards Dec-2015, all unemployment rates have positive auto-correlation with Dec-2018.

Regime 2: Dec-2015 to Aug-2004. The observations here are all have negative auto-correlation with Regime 1.

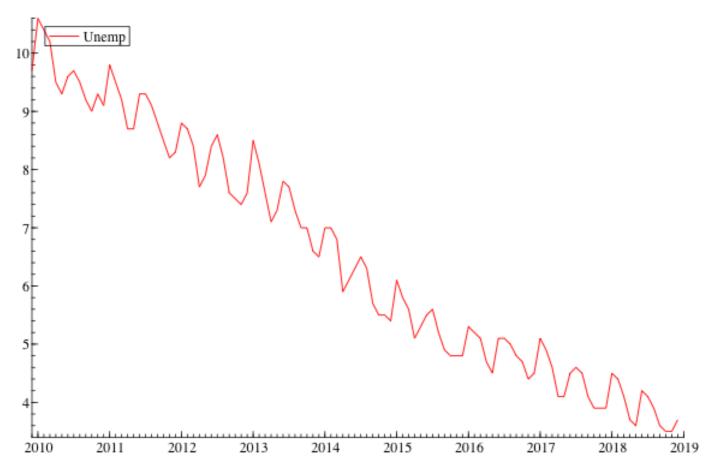
Regime 3: Aug-2004 to Jan-1999: This grouping correlates well with each other, and have less than statistically significant correlation with our Dec-2018 observation.



Graph 02: ACF of US unemployment rate 1990-2018 Shows periods of significant auto-correlation (positive or negative) with most recent data.

# B. Truncated Dataset 2009-2018

Back in Graph 01, there is a clear downward trend in the data starting around mid-2009. Therefore, for simplicity, the set used for modeling will be from 2009-12 to 2018-12 (9 years). The shortened data is plotted in Graph 03.

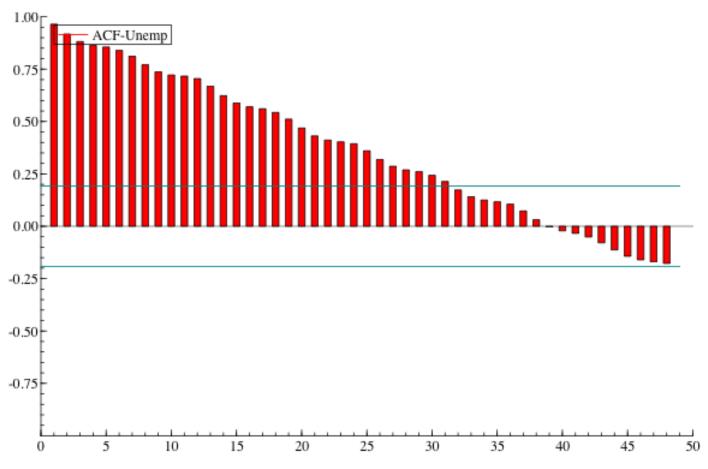


Graph 03: Nine-year trend of declining unemployment rate, with seasonal hikes

## III. Determining the model Part I

## A. Stationarity

The ACF plot of our 9-year period, is just a portion of what we saw in Graph 01. It doesn't help much. There's a slow decay, with statistical significance for 30 or so lags.



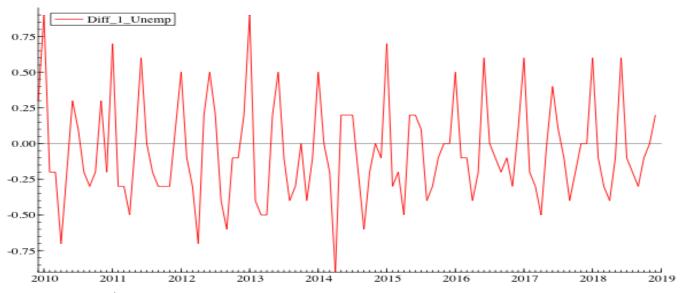
Graph 04: ACF of US unemployment rate, Dec-2009 to Dec-2018

We'd like to figure out when the data is, or is most stationary. From the outset, our base data from 2009 on (Graph 03) is not stationary, but could arguably be called trend stationary. However, we will still give consideration to the first and second differences, to see if they pain a better picture. Moreover, these latter series will give us some intuition about our later model.

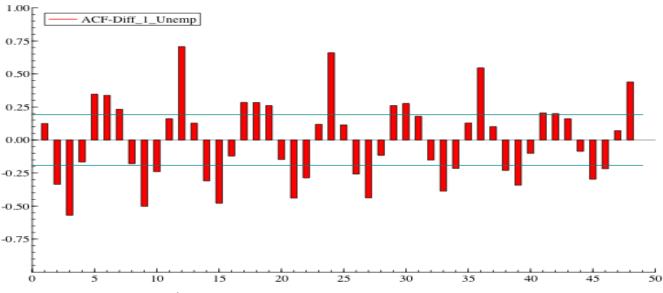
Our visual analysis of the autocorrelation functions didn't correspond well with initial beliefs. Moreover, the plots were somewhat misleading. What follows is a rapid description of these plots and assumed implications, so that we quickly get to the numerical analysis/results.

# A1. Auto-correlation function plots

The time-series and ACF plots for both 1<sup>st</sup> and 2<sup>nd</sup> differences of the unemployment rate data are below. They exhibit some useful, some strange, and some misleading behavior.



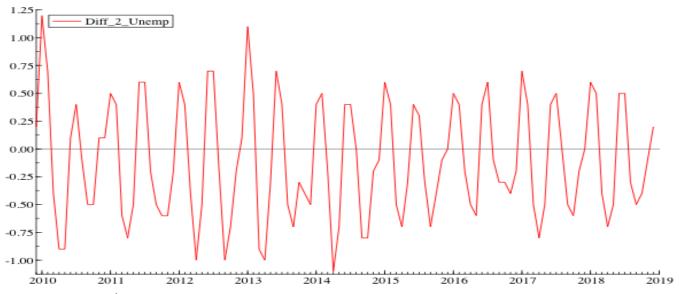
Graph 05a: 1st differenced unemployment rate



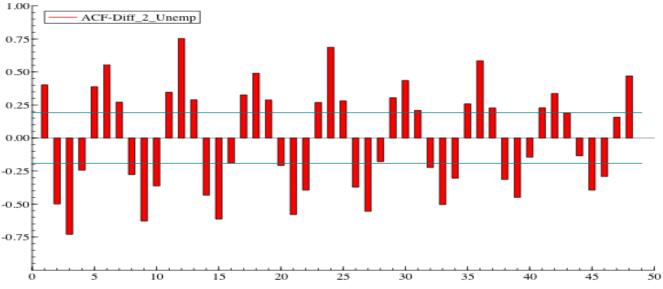
Graph 05b: ACF plot, 1st differenced unemployment rate

When using the 1st differences, our ACF for this 9-year period (Graph 05b), shows clear signs of yearly seasonality. There are prominent spikes in autocorrelation at multiples of 12 lags. Additionally, it shows what might be quarterly seasonality as well. In the differenced timeseries (Graph 05a) we see a pattern that, if not centered around 0, seems to be stationary.

Switching to the 2<sup>nd</sup> differenced time-series. The time-series (Graph 06a) seems less stationary than the 1<sup>st</sup> differenced series. From 2009-2013, and then from 2015-2018 there seems to be slight trends upward. The plot does not appear to be centered at 0 either. Moving to the PACF of this series, we again see evidence of a 12-month season, but also potentially at 6 and 3 months. Moreover, the two months on either side a 3-month multiple are also statistically significant. It would not be better to model with this series.



Graph 06a: 2<sup>nd</sup> differenced unemployment rate



Graph 06b: ACF plot, 2<sup>nd</sup> differenced unemployment rate

# A2. ADF Tests

Since our raw data has an obvious downward trend, we would want to run the ADF on this data assuming "trend", but can also check for trend and constant. On the other hand, we've calculated 1st and 2nd differences of the series. These series, visually, do not suggest a trend. Though there might, be a slight one that we cannot detect visually. The goal of our ADF testing is to find the setup that yields the lowest Information Criteria.

Table 01 shows the output of the Information Criteria for ADF test on each time-series. In bold, we can see, across all series, what the lowest results of the Information Criteria were. They all occur in the base time-series, under a test that assumes neither an intercept, nor a trend. This was unexpected.

For our <u>raw</u> data, cons No Intercept No T		
	riteria (to be minimized)	
Akaike		-0.503884
Schwarz		
	_	
Intercept and No T	rend	
	riteria (to be minimized)	
	-0.452532 Shibata	-0.488265
Schwarz	-0.078564 Hannan-Quinn	-0.301368
Intercept and Tren	d	
	riteria (to be minimized)	
	-0.432930 Shibata	-0.473496
Schwarz		
Schwarz	-0.032230 Haiman-Quim	-0.270707
	d data, considering 12 lags.	
No Intercept and n		
	riteria (to be minimized)	
Akaike	<u>-0.406170</u> Shibata	-0.437331
Schwarz	-0.058915 Hannan-Quinn	-0.265804
Intercept and No T	rend	
	riteria (to be minimized)	
Akaike		-0.467641
Schwarz	-0.057940 Hannan-Quinn	
Schwarz	-0.03/940 Haillian-Quilli	-0.280744
Intercept and Time	e Trend	
Information C	riteria (to be minimized)	
Akaike	-0.428850 Shibata	-0.469416
Schwarz	-0.028171 Hannan-Quinn	-0.266889
For our 2nd difference	ed data, considering 12 lags.	
No Intercept no Tr		
-	riteria (to be minimized)	
	,	-0.421350
Akaike	-0.390189 Shibata	
Schwarz	-0.042934 Hannan-Quinn	-0.249823
Intercept no Trend		
	riteria (to be minimized)	
Akaike	-0.404545 Shibata	-0.440278
Schwarz	-0.030577 Hannan-Quinn	
Intercent and T	.1	
Intercept and Tren		
	riteria (to be minimized)	0.441045
Akaike	-0.401276 Shibata	-0.441842
Schwarz	-0.000597 Hannan-Quinn	-0.239315

Table 01: Abbreviated ADF Test results

Not only was this the best results, but based on the ADF statistic and critical values for each of the 9 tests, only the raw data with no assumed intercept nor trend, was stationary. In every other test, the ADF statistic was not less than the critical value at any given alpha (1%, 5%, 10%). Full results for the ADF tests are in Appendix A. We will keep the results of the ADF test in mind when using the X12 Arima module on OxMetrics.

## A. PACF plots

Non-stationary differenced data is not what we would expect. The PACF of each time-series will be used later on for some intuition on a decent model. For now, the plot for all three series is presented here (Graph 07), showing a lag of up to 48 months. Comparatively, the first plot behaves better. While the correlations taper off a bit faster in the last two plots, the last two plots are also quite complicated within the first 12 lags. Whereas the PACF of the raw unemployment rate, has seasonal spikes at lags 13 and 25; and, outside of lags 1 and 4, no other statistically significant lags prior to lag 13.



Graph 07: PACF. Unemployment rate (top), 1<sup>st</sup> difference (middle), 2<sup>nd</sup> difference (bottom) Consecutively significant lags, and significant lags at irregular times (e.g. 7,11) are problematic.

# B. Interpretations

There is strong evidence that differencing, either first or second is quite useless. By useless, I mean that after both 1st or 2nd differencing, our PACF indicates a series that is less stationary then the raw data. In fact, the PACF's for the differenced data implies non-stationarity.

In general, *the ACF controls the MA* - moving average order, and *the PACF controls the AR* - autoregressive order. Although both graphs help with guessing an ARMA specification. By "control", we mean that we can read the order from the graph, from the last point at which the lags are statistically significant.

This brings our attention back to Graph 07 above. The interpretations are:

#### - PACF raw

Lag 2 is almost significant. Lag 4 is significant. Lag 13, and 25 are significant. However, 13+12 = 25 We can say that this is a Seasonal effect. In short, we'd consider AR up to 4 lags.

#### - PACF 1st difference

This is a mess. We get partial autocorrelation at lag 19, 12,9,8,7,4,3,2. How would we interpret this? You'll notice the ACF of this series is also quite interesting

## - PACF 2nd difference

Lags 18,11,10,8,7,4,2,1 are all significant.

Lag 24 was on the border, but Lag 12 was near zero partial-autocorrelation

I'd say an AR order of 18 is a bit excessive. There's no good pattern here.

Maybe, the series is non-stationary

Our ADF tests confirmed our suspicions about Graph 07. The null hypotheses could not be rejected for neither the 1st differenced series, nor the 2nd differenced series. As such, I would think it strange to arbitrarily throw in 1st and 2nd differencing into consideration of our SARIMA model. We can look, but I expect to find nothing useful. Basically, a non-seasonal differencing parameter of d = 0 should produce our best model.

## V. Model Specification, Comparison, and Selection

# A. Specification

There are 7 parameters for a Seasonal ARIMA model: 3 pairs of autoregressive, moving average, and integration parameters, and one final parameter to denote the length of the season.

ARIMA 
$$(p, d, q) * (P, D, Q)_s$$

The X12 Arima module in OXMetrics, as a setting for automatic model selection. In our previous discussion about the ACF and PACF plots, we decided that there was seasonality, with S = 12. It also made sense that the non-seasonal auto-regressive order of up to p = 4 could work. In total, 7 models were run, including one that used a log transformation, and a model generated my OxMetrics using the "Automatic ARIMA" option. The different specifications are listed below. In **bold** is the model I've selected. Full results for all of the model is in the Appendix.

 $\begin{array}{lll} \text{1. Generic:} & S_{12}\text{ARIMA} \ (1,1,0) \ (1,1,0) \\ \text{2. Generic:} & S_{12}\text{ARIMA} \ (0,1,0) \ (1,1,0) \\ \text{3. 2}^{\text{nd}} \ \text{differencing A:} & S_{12}\text{ARIMA} \ (0,2,0) \ (1,1,0) \\ \text{4. 2}^{\text{nd}} \ \text{differencing B:} & S_{12}\text{ARIMA} \ (3,2,0) \ (1,1,0) \\ \text{5. Intuitive:} & S_{12}\text{ARIMA} \ (4,0,0) \ (2,1,0) \\ \text{6. Auto-optimized:} & S_{12}\text{ARIMA} \ (3,1,1) \ (0,1,1) \end{array}$ 

For comparison, we will look at how forecasts from model #5 and #6 match up with the real data. Two quick notes on determining the models.

- 1. The "vanilla" (0,1,1) (0,1,1) model is the first assumption in automatic model selection. Since, OxMetrics did not consider it to be the best models, there's not real reason to compare it. However, I did compare the forecasts against Model-6 and consistently, each value was slightly greater than the forecasted points in Model-6.
- 2. The data implies an additive model: The seasonal pattern neither expands nor shrinks over time. When using the option for additive outliers, the results returned "No AO outliers identified". As such, none of the above listed models was made using *Outlier Detection* options.

# B. Comparison

Model-6 is the best model according to the system. Compared to Model-5, it has higher Log likelihood, and low Information Criteria. There's also a lower variance, and std. err of variance. However, if we look into the result output for Model-6, it actually has issues with statistical significance of Non-Seasonal Lag 1 and Lag 2. Dividing Estimate by Standard Error, we get a value less than 2 is absolute value. Lag three is ok, but not by much. This lack of significance is also the problem with the Non-seasonal MA component. Additionally, AR Lag 1 Standard Error is larger than in Model-5. Essentially, these problems occur because of having d = 1. Previously we saw that  $1^{st}$  differencing the data was a bit useless. Table 02 has Model-5 and Model-6 output for parameters.

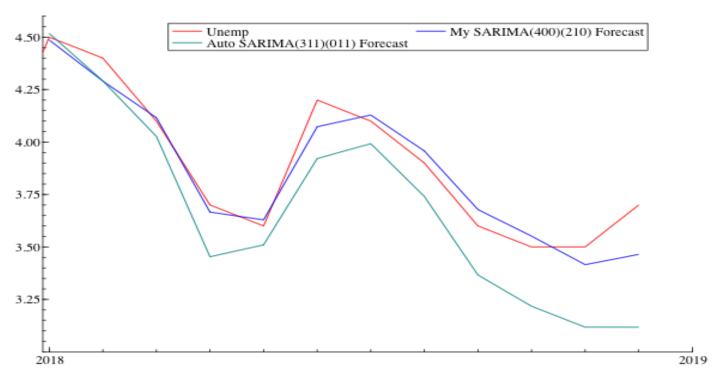
Model-6 ARIMA Model Parameter	: (3, 1, 1) (0, 1, 1) Estimate	Standard Errors	
Nonseasonal AR  Lag 1  Lag 2  Lag 3	.0345 1370 2284		
Nonseasonal MA Lag 1 Seasonal MA Lag 12	.3076 .9998	.22122 .11226	
Variance SE of Var			
Model-5 ARIMA Model Parameter		Standard Errors	
Nonseasonal AR  Lag 1  Lag 2  Lag 3  Lag 4	.7399 .0732 0569 .2389		
Seasonal AR Lag 12 Lag 24	7079 2459	.10691 .10940	
Variance SE of Var	. 27893 . 42786		

Table 02: Estimated parameter comparison, Model-6 and Model-5

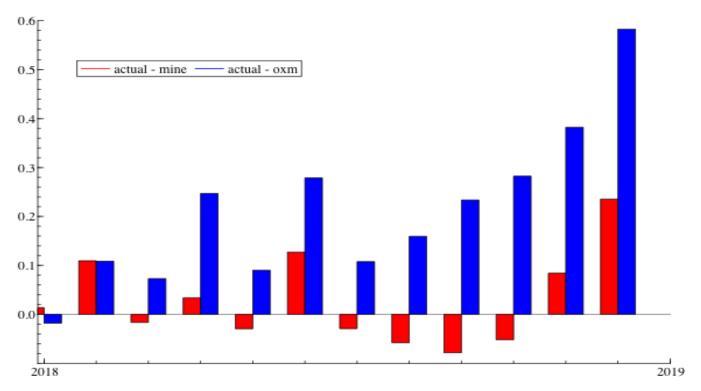
## C. Selection

While the Likelihood Statistics declare Model-5 the winner, I disagree. The parameters used don't make sense. If we look back at Graph 05b, ACF of 1<sup>st</sup> differences, and Graph 07 of the PACF's, what exactly motivates the reason to have p = 3 and q = 1? Nothing really. On the other hand, we discussed sound reasons for the Model-5 parameters. However, since the ACF of the raw series (Graph 04) wasn't too nice, admittedly there was no goo way to decide on the MA order. Something like 30 would be too many lags to include.

Visually, my select model, Model-5 is a clear winner, based on the forecast. Model-6 plays it safe. For the entire forecast is underestimates the unemployment rate. Realistically, that's bad. Moreover, in general it's further away from the actual values (except for the 1<sup>st</sup> estimate), than Model-5. It seems Model-5 is penalized since it slightly overestimates the unemployment are for various points. Graph 08 shows the forecasts plotted with the actual values. Graph 08 is a simple bar plot of (Actual Value – Model Forecasted Value) to further illustrate the reason for my selection.



Graph 08: Actual Unemployment rate, and Forecasts: My/Model-5, Auto/Model-6



Graph 09: Forecast Error, red – Model-5 (Mine), blue – Model-6 (OxMetrics)

To briefly conclude, while Model-6 did have better results for matching the sample data, my selected Model-5 was superior in the forecast, and would be the one to go with.

# VI. Appendix

# A. Full output for Augmented Dickey-Fuller tests

# A1. ADF Test, Unemployment rate, base data

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TESTS:
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ADF Test with 12 lags

No intercept and no time trend

H0: Unemp is I(1)

ADF Statistics: -2.59152

Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993)

1% 5% 10% -2.56572 -1.94093 -1.61663

## **OLS** Results

	Coefficient t-value			
y_1	-0.018579 -2.5915			
dy_1	-0.350456 -3.6357			
dy_2	-0.338199 -3.3752			
$dy_3$	-0.384948 -3.6617			
dy_4	-0.217074 -1.9255			
dy_5	0.045419 0.39564			
dy_6	-0.021181 -0.18361			
dy_7	0.078827  0.68828			
dy_8	0.113341 0.99000			
dy_9	-0.079062 -0.70322			
dy_10	-0.160711 -1.5420			
dy_11	-0.111135 -1.0967			
dy_12	0.401754 4.3490			
RSS	2.645402			
OBS	96.000000			
Information Criteria (to be minimized)				
Akaike	-0.472724 Shibata	-0.503884		

#### TESTS:

Schwarz

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ADF Test with 12 lags Intercept and no time trend H0: Unemp is I(1)

ADF Statistics: -1.37578

Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993)

-0.125469 Hannan-Quinn

-0.332358

1% 5% 10% -3.4323 -2.86228 -2.56721

## **OLS** Results

 $\begin{array}{cccc} & & Coefficient & t\text{-value} \\ y\_1 & & -0.014992 & -1.3758 \\ dy\_1 & & -0.360416 & -3.6227 \end{array}$ 

```
dy 2
             -0.350053 -3.3577
dy_3
             -0.398539 -3.6202
dy 4
             -0.231399 -1.9627
dy 5
             0.030670 \quad 0.25525
dy 6
             -0.034920 -0.29081
dy_7
             0.065913
                       0.55486
dy 8
             0.102076 0.86596
dy 9
             -0.087828 -0.76551
dy 10
             -0.167509 -1.5822
dy_11
             -0.116883 -1.1384
dy 12
                         4.1956
              0.394923
Constant
              -0.031933 -0.43884
RSS
             2.639203
OBS
             96.000000
Information Criteria (to be minimized)
             -0.452532 Shibata
Akaike
                                       -0.488265
Schwarz
             -0.078564 Hannan-Quinn
                                       -0.301368
```

## TESTS:

\_\_\_\_\_

ADF Test with 12 lags Intercept and time trend H0: Unemp is I(1)

ADF Statistics: -0.675754

Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993)

1% 5% 10% -3.96104 -3.41127 -3.12748

O E D Trebu	100			
	Coefficient t-value			
y_1	-0.056336 -0.67575			
dy_1	-0.318105 -2.4296			
dy_2	-0.311039 -2.3820			
dy_3	-0.363116 -2.7651			
dy_4	-0.199370 -1.4808			
dy_5	0.063472  0.46205			
dy_6	-0.001222 -0.0088470			
dy_7	0.097626 0.72246			
dy_8	0.134413  0.99630			
dy_9	-0.058246 -0.44963			
dy_10	-0.144612 -1.2489			
dy_11	-0.098970 -0.90645			
dy_12	0.406073 4.1798			
Constant	0.243573			
Trend	-0.002750 -0.50026			
RSS	2.631074			
OBS	96.000000			
Information Criteria (to be minimized)				
Akaike	-0.432930 Shibata	-0.473496		
Schwarz	-0.032250 Hannan-Quinn	-0.270969		

#### **A2.** ADF Test, 1<sup>st</sup> differences of Unemployment rate TESTS: ADF Test with 12 lags No intercept and no time trend H0: Diff 1 Unemp is I(1) ADF Statistics: -1.02122 Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993) 1% 5% 10% -2.56572 -1.94093 -1.61663 **OLS** Results Coefficient t-value -0.300457 -1.0212 y\_1 dy 1 -1.012798 -3.3453 dy 2 -1.250332 -4.0815 dy 3 -1.518229 -4.8091 dy 4 -1.592918 -4.7660 dy 5 -1.399044 -3.9490 dy 6 -1.263128 -3.5638 dy 7 -1.013631 -2.9774 dy 8 -0.736959 -2.3734 dy 9 -0.649565 -2.5285 dy\_10 -0.647637 -3.2961 dy 11 -0.616194 -4.2783 dy 12 -0.095734 -0.96922 RSS 2.827454 OBS 96.000000 Information Criteria (to be minimized) Akaike -0.406170 Shibata -0.437331 Schwarz -0.058915 Hannan-Quinn -0.265804 TESTS: ADF Test with 12 lags Intercept and no time trend H0: Diff 1 Unemp is I(1) ADF Statistics: -2.25524 Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993) 1% 5% 10% -3.4323 -2.86228 -2.56721 **OLS** Results

	Coefficient t-	-value
y_1	-1.770243	-2.2552
dy_1	0.399206	0.52418
dy_2	0.061829	0.086153
dy_3	-0.321524	-0.47970
dy_4	-0.529396	-0.85140
dy_5	-0.474992	-0.82486

```
dy 6
            -0.480121 -0.92002
dy_7
            -0.376743 -0.81862
dy 8
            -0.236780 -0.60205
dy 9
            -0.278672 -0.89213
dy 10
             -0.396337 -1.7247
dy_11
             -0.469576 -2.9517
dy 12
             -0.042677 -0.42457
Constant
             -0.101437 -2.0139
RSS
             2.694201
OBS
             96.000000
Information Criteria (to be minimized)
            -0.431907 Shibata
Akaike
                                       -0.467641
            -0.057940 Hannan-Quinn
                                      -0.280744
Schwarz
```

## TESTS:

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ADF Test with 12 lags Intercept and time trend H0: Diff\_1\_Unemp is I(1)

ADF Statistics: -2.57428

Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993)

1% 5% 10% -3.96104 -3.41127 -3.12748

OLS Results				
Coefficient t-value				
-2.173735 -2.5743				
0.783221 0.95858				
0.422973 0.54950				
0.016064 0.022343				
-0.220430 -0.33109				
-0.201247 -0.32827				
-0.244675 -0.44314				
-0.182460 -0.37739				
-0.086386 -0.21099				
-0.171578 -0.53204				
-0.328503 -1.3971				
-0.438096 -2.7307				
-0.035465 -0.35355				
-0.125208 -2.3370				
0.000918 1.2672				
2.641830				
96.000000				
Information Criteria (to be minimized)				
-0.428850 Shibata	-0.469416			
-0.028171 Hannan-Quinn	-0.266889			
	Coefficient t-value -2.173735 -2.5743 0.783221 0.95858 0.422973 0.54950 0.016064 0.022343 -0.220430 -0.33109 -0.201247 -0.32827 -0.244675 -0.44314 -0.182460 -0.37739 -0.086386 -0.21099 -0.171578 -0.53204 -0.328503 -1.3971 -0.438096 -2.7307 -0.035465 -0.35355 -0.125208 -2.3370 0.000918 1.2672 2.641830 96.000000 on Criteria (to be minimized) -0.428850 Shibata			

#### **A3.** ADF Test, 2<sup>nd</sup> differences of Unemployment rate TESTS: -----ADF Test with 12 lags No intercept and no time trend H0: Diff 2 Unemp is I(1) ADF Statistics: -0.899516 Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993) 1% 5% 10% -2.56572 -1.94093 -1.61663 **OLS** Results Coefficient t-value y 1 -0.133898 -0.89952 dy 1 -0.228020 -1.2994 dy 2 -1.163495 -6.7481 dy\_3 -0.547854 -2.8433 dy 4 -1.224532 -6.2468 dy 5 -0.390890 -1.8521 $dy_6$ -1.081098 -5.1140 dy 7 -0.194561 -0.97044 dy 8 -0.764772 -3.9291 dy\_9 -0.151056 -0.95956 dy 10 -0.715038 -4.9028 dy 11 -0.111763 -1.0962 dy 12 -0.142546 -1.4345 RSS 2.873003 96.000000 OBS Information Criteria (to be minimized) Akaike -0.390189 Shibata -0.421350 -0.042934 Hannan-Quinn Schwarz -0.249823 TESTS: \_\_\_\_\_ ADF Test with 12 lags Intercept and no time trend H0: Diff 2 Unemp is I(1) ADF Statistics: -1.96477 Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993) 1% 5% 10% -3.4323 -2.86228 -2.56721

0 20 1100	*****	
	Coefficient t	-value
y_1	-0.810843	-1.9648
$dy_1$	0.401915	1.0086
dy_2	-0.550203	-1.4157
dy 3	-0.018181	-0.050965

```
dy 4
             -0.727988 -2.1238
dy_5
             0.015211 0.048849
dy 6
             -0.702305 -2.3390
dy 7
             0.086287 0.33895
dy 8
             -0.509190 -2.1115
dy_9
             0.016780 0.091933
dy 10
             -0.574866 -3.4904
dy 11
             -0.053048 -0.49993
dy 12
             -0.096995 -0.95547
Constant
              -0.092627 -1.7555
RSS
             2.768939
OBS
             96.000000
Information Criteria (to be minimized)
Akaike
             -0.404545 Shibata
                                       -0.440278
             -0.030577 Hannan-Quinn
Schwarz
                                       -0.253381
TESTS:
ADF Test with 12 lags
Intercept and time trend
H0: Diff 2 Unemp is I(1)
ADF Statistics: -2.30016
Asymptotic critical values, Davidson, R. and MacKinnon, J. (1993)
```

1% 5% 10% -3.96104 -3.41127 -3.12748

	Coefficient t-value	
y_1	-1.020002 -2.3002	
dy_1	0.592292 1.3942	
dy_2	-0.361769 -0.87147	
dy_3	0.148931 0.39256	
dy_4	-0.569207 -1.5636	
dy_5	0.146768  0.44832	
dy_6	-0.582443 -1.8553	
dy_7	0.176847 0.67078	
dy_8	-0.435091 -1.7588	
dy_9	0.063915  0.34422	
dy_10	-0.542450 -3.2655	
dy_11	-0.044401 -0.41905	
dy_12	-0.091249 -0.90117	
Constant	-0.117287 -2.0907	
Trend	0.000926 1.2603	
RSS	2.715690	
OBS	96.000000	
Informatio	on Criteria (to be minimized)	
Akaike	-0.401276 Shibata	-0.441842
Schwarz	-0.000597 Hannan-Quinn	-0.239315

## B. Full results for SARIMA models

# B1. $S_{12}$ ARIMA (1,1,0) (1,1,0)

MODEL DEFINITION for Unemp

Transformation: No transformation

ARIMA Model: (1 1 0)(1 1 0)

regARIMA Model Span: 2009.Dec to 2017.Dec

#### MODEL ESTIMATION/EVALUATION

Estimation converged in 5 ARMA iterations, 16 function evaluations.

ARIMA Model: (1 1 0)(1 1 0)

Nonseasonal differences: 1
Seasonal differences: 1

Parameter	Estimate	Standard Errors
Nonseasonal AR		
Lag 1	1999	. 10544
Seasonal AR		
Lag 12	5831	. 08458
Variance	.31957E-01	
SE of Var	.49311E-02	

-----

#### Likelihood Statistics

Number of observations (nobs)	97
Effective number of observations (nefobs)	84
Number of parameters estimated (np)	3
Log likelihood (L)	22.9167
AIC	-39.8334
AICC (F-corrected-AIC)	-39.5334
Hannan Quinn	-36.9019
BIC	-32.5410

Roots of ARIMA Model

Root	Real	Imaginary	Modulus	Frequency
	-5.0024	. 0000	5.0024	. 5000
Seasonal AR Root 1	-1.7150	. 0000	1.7150	. 5000

FORECASTING

Origin 2017.Dec Number 12

## Confidence intervals with coverage probability ( .95000)

Date	Lower	Forecast	Upper
2018.Jan	4.10	4.45	4.80
2018.Feb	3.86	4.31	4.76
2018.Mar	3.59	4.12	4.66
2018.Apr	3.07	3.68	4.29
2018.May	2.89	3.57	4.24
2018.Jun	3.34	4.08	4.82
2018.Jul	3.33	4.12	4.92
2018.Aug	3.18	4.02	4.87
2018.Sep	2.85	3.74	4.63
2018.0ct	2.66	3.60	4.54
2018.Nov	2.44	3.42	4.41
2018.Dec	2.45	3.48	4.51

-----

# B2. $S_{12}$ ARIMA (0,1,0) (1,1,0)

MODEL DEFINITION for Unemp

Transformation: No transformation

ARIMA Model: (0 1 0)(1 1 0)

regARIMA Model Span: 2009.Dec to 2017.Dec

#### MODEL ESTIMATION/EVALUATION

Estimation converged in 4 ARMA iterations, 9 function evaluations.

ARIMA Model: (0 1 0)(1 1 0) Nonseasonal differences: 1 Seasonal differences: 1

Parameter	Estimate	Standard Errors
Seasonal AR Lag 12	5783	.08550
Variance SE of Var	.33308E-01 .51396E-02	

\_\_\_\_\_

#### Likelihood Statistics

\_\_\_\_\_\_ Number of observations (nobs) Effective number of observations (nefobs) 84 Number of parameters estimated (np) 2 21.2479 Log likelihood (L) AIC -38.4957 AICC (F-corrected-AIC) -38.3476 Hannan Quinn -36.5414 BIC -33.6341

#### Roots of ARIMA Model

Root	Real	Imaginary	Modulus	Frequency
Seasonal AR Root 1	-1.7291	. 0000	1.7291	. 5000

#### FORECASTING

Origin 2017.Dec Number 12

Date	Lower	Forecast	Upper
2018.Jan	4.08	4.44	4.80
2018.Feb	3.79	4.30	4.81
2018.Mar	3.50	4.12	4.74
2018.Apr	2.96	3.67	4.39
2018.May	2.76	3.56	4.36
2018.Jun	3.20	4.07	4.95
2018.Jul	3.17	4.12	5.06
2018.Aug	3.00	4.02	5.03
2018.Sep	2.66	3.73	4.80
2018.Oct	2.46	3.59	4.72
2018.Nov	2.23	3.42	4.60
2018.Dec	2.23	3.47	4.71

# B3. S<sub>12</sub>ARIMA (0,2,0) (1,1,0)

MODEL DEFINITION for Unemp

Transformation: No transformation

ARIMA Model: (0 2 0)(1 1 0)

regARIMA Model Span: 2009.Dec to 2017.Dec

#### MODEL ESTIMATION/EVALUATION

Estimation converged in 4 ARMA iterations, 9 function evaluations.

ARIMA Model: (0 2 0)(1 1 0) Nonseasonal differences: 2 Seasonal differences: 1

Parameter	Estimate	Standard Errors
Seasonal AR Lag 12	5758	. 08653
Variance SE of Var	.79383E-01 .12323E-01	

#### Likelihood Statistics

\_\_\_\_\_\_ Number of observations (nobs) Effective number of observations (nefobs) 83 Number of parameters estimated (np) 2 Log likelihood (L) -15.0496 AIC 34.0993 AICC (F-corrected-AIC) 34.2493 Hannan Quinn 36.0428 BIC 38.9370

#### Roots of ARIMA Model

Root	Real	Imaginary	Modulus	Frequency
Seasonal AR				
Root 1	-1.7367	. 0000	1.7367	. 5000

#### FORECASTING

Origin 2017.Dec Number 12

Date	Lower	Forecast	Upper
2018. Jan 2018. Feb 2018. Mar 2018. Apr 2018. May 2018. Jun 2018. Jul 2018. Aug	3.848 2.980 1.922 .478 750 -1.450 -2.716 -4.211	4.400 4.215 3.988 3.503 3.345 3.818 3.818 3.676	4.952 5.450 6.054 6.528 7.441 9.086 10.352 11.563
2018.Sep 2018.Oct	-5.974 -7.672	3.349 3.164	12.671 13.999
2018.Nov 2018.Dec	-7.672 -9.473 -11.115	2.949 2.964	15.370 17.043
2010.060		2.504	17.045

## B4. S<sub>12</sub>ARIMA (3,2,0) (1,1,0)

MODEL DEFINITION for Unemp

Transformation: No transformation ARIMA Model: (3 2 0)(1 1 0)

regARIMA Model Span: 2009.Dec to 2017.Dec

#### MODEL ESTIMATION/EVALUATION

Estimation converged in 6 ARMA iterations, 31 function evaluations.

ARIMA Model: (3 2 0)(1 1 0)

Nonseasonal differences: 2

Seasonal differences: 1

Parameter	Estimate	Standard Errors
Nonseasonal AR		
Lag 1	8126	. 10192
Lag 2	5376	.12463
Lag 3	3256	. 10295
Seasonal AR		
Lag 12	5885	. 08506
Variance	.45044E-01	
SE of Var	.69922E-02	

Likelihood Statistics

Number of observations (nobs)	97
Effective number of observations (nefobs)	83
Number of parameters estimated (np)	5
Log likelihood (L)	7.9039
AIC	-5.8078
AICC (F-corrected-AIC)	-5.0286
Hannan Quinn	9490
BIC	6.2864

-----

Roots of ARIMA Mode	el Real	Imaginary	Modulus	Frequency
Nonseasonal AR				
Root 1	1165	1.4670	1.4716	. 2626
Root 2	1165	-1.4670	1.4716	2626
Root 3	-1.4184	. 0000	1.4184	. 5000
Seasonal AR				
Root 1	-1.6992	. 0000	1.6992	. 5000

#### FORECASTING

Origin 2017.Dec Number 12

Date	Lower	Forecast	Upper
2018.Jan	3.984	4.400	4.816
2018.Feb	3.588	4.234	4.879
2018.Mar	3.170	4.067	4.964
2018.Apr	2.426	3.600	4.773
2018.May	1.940	3.463	4.985
2018.Jun	2.087	3.964	5.842
2018.Jul	1.740	3.997	6.253
2018.Aug	1.220	3.878	6.536
2018.Sep	. 493	3.580	6.668
2018.Oct	110	3.424	6.959
2018.Nov	768	3.234	7.235
2018.Dec	-1.211	3.276	7.764

## **B5.** S<sub>12</sub>ARIMA (4,0,0) (2,1,0) *Author's Selected Model*

MODEL DEFINITION for Unemp

Transformation: No transformation

ARIMA Model: (4 0 0)(2 1 0) regARIMA Model Span: 2009.Dec to 2017.Dec

#### MODEL ESTIMATION/EVALUATION

Estimation converged in 16 ARMA iterations, 115 function evaluations.

ARIMA I	Model:	(4 0	0)(2	Τ	<b>(</b> )	Seasonal differences:
Parameter					Estimate	Standard Errors

Nonseasonal AR			
Lag 1	. 7399	. 10279	
Lag 2	.0732	. 12824	
Lag 3	0569	.12626	
Lag 4	. 2389	. 10399	
Seasonal AR			
Lag 12	7079	. 10691	
Lag 24	2459	. 10940	
Variance	.27893E-01		
SE of Var	.42786E-02		

#### Likelihood Statistics

Number of observations (nobs) Effective number of observations (nefobs) 85

Number of parameters estimated (np) 7 Log likelihood (L) 26.8876 AIC -39.7752 AICC (F-corrected-AIC) -38.3207 Hannan Quinn -32.8977 -22.6767

#### Roots of ARIMA Model

Root		Real	Imaginary	Modulus	Frequency
Root	1 AR	1.0029	.0000	1.0029	. 0000
	1	-1.6489	.0000	1.6489	. 5000
	2	.4420	1.5283	1.5909	. 2052
Root Seasonal A	4 R	.4420	-1.5283	1.5909	2052
Root	1	-1.4391	1.4123	2.0164	.3765
Root		-1.4391	-1.4123	2.0164	3765

FORECASTING: Origin 2017.Dec Number 12

Date	Lower	Forecast	Upper
2018. Jan	4.16	4.49	4.81
2018. Feb	3.88	4.29	4.70
2018. Mar	3.66	4.12	4.57
2018. Apr	3.19	3.67	4.15
2018. Apr	3.11	3.63	4.14
2018.Jun	3.52	4.07	4.62
2018.Jul	3.54	4.13	4.72
2018.Aug	3.34	3.96	4.57
2018.Sep	3.03	3.68	4.32
2018.Oct	2.88	3.55	4.22
2018.Nov	2.72	3.42	4.11
2018.Dec	2.74	3.46	4.19
2010.060	2./4	3.40	7.13

# B6. S<sub>12</sub>ARIMA (3,1,1) (0,1,1) OxMetrics' Selected Model

Series Title- X-12-ARIMA run of Unemp Series Name- Unemp

-Period covered- 12th month,2009 to 12th month,2017
Automatic ARIMA Model Selection

Procedure based closely on TRAMO method of Gomez and Maravall (2000)

"Automatic Modeling Methods for Univariate Series",
A Course in Time Series

(Edited by D. Pena, G. C. Tiao, R. S. Tsay),
New York: J. Wiley and Sons

Maximum order for regular ARMA parameters: 2
Maximum order for seasonal ARMA parameters: 1
Maximum order for regular differencing: 2
Maximum order for seasonal differencing: 1

Results of Unit Root Test for identifying orders of differencing:

Regular difference order : 1 Seasonal difference order : 1

Mean is not significant.

Automatic model choice : (0 1 1)(0 1 1)

Final automatic model choice :  $(3\ 1\ 1)(0\ 1\ 1)$  End of automatic model selection procedure.

Estimation converged in 10 ARMA iterations, 61 function evaluations.

ARIMA Model: (3 1 1)(0 1 1)

Nonseasonal differences: 1

Seasonal differences: 1

Parameter	Estimate	Standard Errors	
Nonseasonal AR Lag 1 Lag 2 Lag 3	. 0345 1370 2284	.22571 .09815 .10202	
Nonseasonal MA Lag 1	. 3076	. 22122	
Seasonal MA Lag 12	. 9998	.11226	
Variance SE of Var	.19780E-01 .30521E-02		

## Likelihood Statistics

Number of observations (nobs)	97
Effective number of observations (nefobs)	84
Number of parameters estimated (np)	6
Log likelihood (L)	32.9541
AIC	-53.9083
AICC (F-corrected-AIC)	-52.8174
Hannan Quinn	-48.0453
BIC	-39.3234

Roots of ARIMA Mode Root	l Real	Imaginary	Modulus	Frequency
Nonseasonal AR				
Root 1	. 6484	1.3740	1.5193	.1798
Root 2	. 6484	-1.3740	1.5193	1798
Root 3	-1.8964	.0000	1.8964	. 5000
Nonseasonal MA				
Root 1	3.2510	. 0000	3.2510	.0000
Seasonal MA				
Root 1	1.0002	. 0000	1.0002	. 0000

## FORECASTING

Origin 2017.Dec Number 12

Date	Lower	Forecast	Upper
2018.Jan	4.24	4.52	4.79
2018.Feb	3.95	4.29	4.63
2018.Mar	3.65	4.03	4.40
2018.Apr	3.06	3.45	3.84
2018.May	3.10	3.51	3.92
2018.Jun	3.49	3.92	4.36
2018.Jul	3.53	3.99	4.45
2018.Aug	3.26	3.74	4.23
2018.Sep	2.86	3.37	3.87
2018.Oct	2.69	3.22	3.74
2018.Nov	2.57	3.12	3.66
2018.Dec	2.56	3.12	3.68

# VII. References

United States Department of Labor - Bureau of Labor Statistics https://beta.bls.gov/dataViewer/view/timeseries/LNU04000000