Lecture 10 Multiple Linear Regression

STAT 512 Spring 2011

Background Reading

KNNL: 6.1-6.5

Topic Overview

• Multiple Linear Regression Model

Data for Multiple Regression

- Y_i is the response variable (as usual)
- $X_{i1}, X_{i2}, \dots X_{i,p-1}$ are the p-1 explanatory variables for cases $i=1,2,\dots,n$
- Example In HW #2, you considered predicting freshman GPA based on ACT scores. Perhaps we consider high school GPA and an intelligence test as well. Then for this problem, we would be using p = 4.

Multicollinearity

- Predictor variables are often correlated to each other.
- If predictor variables are highly correlated, they will be "fighting" to explain the same part of the variation in the response variable.
- Caution: Using highly correlated predictor variables in the same model will not lead to useful parameter estimates. Want to be careful of this.

Multiple Regression Model

$$Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \dots + \beta_{p-1} X_{i,p-1} + \varepsilon_i$$

- i = 1, 2, ..., n observations
- Assumptions exactly as before:

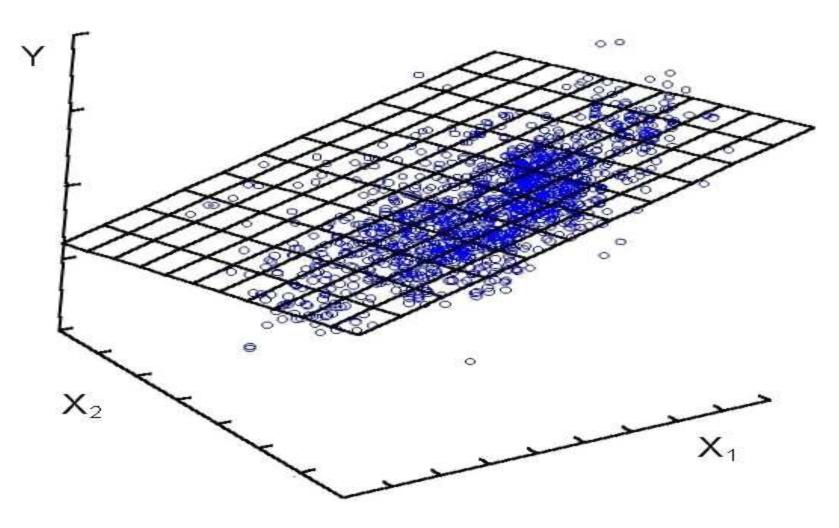
$$arepsilon_{i}\overset{iid}{\sim}N\left(0,\sigma^{2}
ight)$$

- Y_i is the value of the response variable for the i^{th} case.
- X_{ik} is the value of the k^{th} explanatory variable for the i^{th} case.

Multiple Regression Model (2)

- β_0 is the intercept (think multidimensional).
- $\beta_1, \beta_2, \dots, \beta_{p-1}$ are the regression (slope) coefficients for the explanatory variables.
- Parameters as usual include all of the β 's as well as σ^2 . These need to be estimated from the data.

Regression Plane/Surface



Model in Matrix Form

$$\mathbf{Y} = \mathbf{X} \boldsymbol{\beta} + \boldsymbol{\varepsilon}$$
 $n \times 1 = n \times p \boldsymbol{\beta} + \boldsymbol{\varepsilon}$

$$\mathbf{\varepsilon} \sim \mathbf{N}\left(\mathbf{0}, \sigma^2 \mathbf{I}_{n \times n}\right)$$

$$\mathbf{Y} \sim N(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I})$$

Design matrix X

$$\mathbf{X}_{n imes p} = egin{bmatrix} 1 & X_{11} & X_{12} & \cdots & X_{1,p-1} \ 1 & X_{21} & X_{22} & \cdots & X_{2,p-1} \ dots & dots & dots & dots \ 1 & X_{n1} & X_{n2} & \cdots & X_{n,p-1} \end{bmatrix}$$

Coefficient matrix β

$$oldsymbol{eta}_{p imes 1} = egin{bmatrix} eta_0 \ eta_1 \ dots \ eta_{p-1} \end{bmatrix}$$

Least Squares Solution

- Minimize distances between point and response surface
- Find b to minimize

$$SSE = (\mathbf{Y} - \mathbf{Xb})'(\mathbf{Y} - \mathbf{Xb})$$

• Obtain normal equations as before:

$$X'Xb = X'Y$$

• Least Squares Solution as before:

$$\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$

Fitted Values / Residuals

• Fitted (predicted) values for the mean of Y are

$$\hat{\mathbf{Y}} = \mathbf{X}\mathbf{b} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} = \mathbf{H}\mathbf{Y}$$

Residuals are

$$e = Y - \hat{Y} = Y - HY = (I - H)Y$$

• Note formulas are same as before, with hat matrix:

$$\mathbf{H} = \mathbf{X} (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'$$

"Linear" Regression Models

- The term *linear* here refers to the **parameters**, not the predictor variables.
- We can use *linear* regression models to deal with almost any "function" of a predictor variable (e.g. X^2 , $\log(X)$, etc.)
- We cannot use *linear* regression models to deal with nonlinear functions of the parameters (unless we can find a transformation that makes them linear).

Types of Predictors

- Continuous Predictors we are used to these.
- Qualitative Predictors
 - Two possible outcomes (e.g. male/female) represented by 0 or 1
- Polynomial Regression
 - Use squared or higher-ordered terms in regression model.
 - Typically always include lower order terms.

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \dots + \beta_{p-1} X_i^{p-1} + \varepsilon_i$$

Types of Predictors (2)

- Using Transformed Variables
 - Transform one or more X's
 - Transform Y
- Interaction Effects
 - Use Product of Predictor variables as an additional variable.
 - Each variable in the product included by itself as well.
 - $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1} X_{i2} + \varepsilon_i$
- More on some of these models later...

Analysis of Variance

Formulas for sums of squares(in matrix terms) are the same as before

$$SSR = \sum (\hat{Y}_i - \overline{Y})^2 = \mathbf{b}' \mathbf{X}' \mathbf{Y} - \left(\frac{1}{n}\right) \mathbf{Y}' \mathbf{J} \mathbf{Y}$$

$$SSE = \sum (\hat{Y}_i - \hat{Y}_i)^2 = \mathbf{e}' \mathbf{e} = \mathbf{Y}' \mathbf{Y} - \mathbf{b}' \mathbf{X}' \mathbf{Y}$$

$$SSTO = \sum (\hat{Y}_i - \overline{Y})^2 = \mathbf{Y}' \mathbf{Y} - \left(\frac{1}{n}\right) \mathbf{Y}' \mathbf{J} \mathbf{Y}$$

Analysis of Variance (2)

- Degrees of Freedom depend on the model
- Always n 1 total degrees of freedom
- Model degrees of freedom is equal to the number of terms in the model (p-1)
 - Each variable has at least one term
 - May be additional terms for squares, interactions, etc.
- Error degrees of freedom is difference between total and model degrees of freedom (n p).

Analysis of Variance (3)

 Mean Squares obtained by dividing SS by DF for each source.

• The mean square error (MSE) is still, always, and forever, the estimate of σ^2 .

ANOVA Table

Source	df	SS	MS	F
Regression (Model)	p-1	$\sum \Bigl(\hat{Y_i} - ar{Y}\Bigr)^2$	$\frac{SSR}{df_R}$	MSR MSE
Error	n-p	$\sum \left(Y_i - \hat{Y_i} ight)^2$	$\frac{SSE}{df_E}$	
Total	n-1	$\sum ig(Y_i - ar{Y}ig)^2$	$\frac{SSTO}{df_T}$	

F-test for model significance

- The ratio F = MSR / MSE is again used to test for a regression relationship.
- Difference from SLR
 - Null Hyp: $H_0: \beta_1 = \beta_2 = ... = \beta_{p-1} = 0$
 - Alt Hyp: H_a : at least one $\beta_k \neq 0$
- Tests model significance, not individual variables; gives no indication of which variable(s) in particular are important

F-test for model significance (2)

- Under null, has F-distribution with degrees of freedom p-1 and n-p.
- Reject if statistic is larger than critical value for $\alpha = 0.05$; or if p-value for test (given in SAS ANOVA table) is less than 0.05
- If reject, conclude at least one of the explanatory variables is important.
- If fail to reject, and sample size large enough (power), then none of the explanatory variables are useful.

Coefficient of Multiple Determination

•
$$R^2 = \frac{SSR}{SSTO} = 1 - \frac{SSE}{SSTO}$$

- Measures the percentage of variation explained by the variables in the model.
- Additional variables will make R^2 go up; so cannot really use R^2 to determine whether a variable should be added.

Adjusted R²

•
$$R_a^2 = 1 - \frac{MSE}{MSTO} = 1 - \left(\frac{n-1}{n-p}\right) \frac{SSE}{SSTO}$$

- Recall mean squares are SS adjusted by degrees of freedom
- R_a^2 can increase or decrease when a new variable is introduced into the model; depending on whether the decrease in SSE is offset by the lost degree of freedom.
- R_a^2 can be used to decide if variables are important in a model.

Inference for INDIVIDUAL Regression Coefficients

• We already have $\mathbf{b} \sim N(\boldsymbol{\beta}, \sigma^2(\mathbf{X}'\mathbf{X})^{-1})$, so define

$$\mathbf{s^2} \{ \mathbf{b} \} = MSE \times (\mathbf{X'X})^{-1}$$

• For individual b_k , the estimated variance is the k^{th} diagonal element of this matrix:

$$s^{2}\left\{b_{k}\right\} = \left[\mathbf{s^{2}}\left\{\mathbf{b}\right\}\right]_{k,k}$$

Note: k=0,1,...,p-1.

Confidence Intervals for β_k

- CI for β_k is $b_k \pm t_{crit}s\{b_k\}$
- Critical value comes from t-distribution with n-p degrees of freedom (DF for error)
- If CI includes zero, then we cannot reject $H_0: \beta_k = 0$ (i.e. that variable is not significant when added to the model containing all of the other variables.)

Significance Test for β_k

- Is known as a *variable-added-last* test; tests whether the k^{th} explanatory variable is important when added to all of the other variables in the model (i.e. it is a conditional test).
- Test statistic is as before: $t^* = b_k / s\{b_k\}$
- Compare to t-critical value on n p degrees of freedom.
- This is the test given in the model parameters section of SAS for PROC REG.

Upcoming in Lecture 11

• Case Study: Computer Science Student Data