Q1.

$$\begin{split} \widehat{\beta}_{0} &= \overline{Y} - \widehat{\beta}_{1} \overline{X} \\ \widehat{\beta}_{1} &= \frac{S_{XY}}{S_{XX}} = \frac{\sum (x_{i} - \overline{X})(y_{i} - \overline{Y})}{S_{XX}} = \frac{\sum (x_{i}y_{i} - x_{i}\overline{Y} - \overline{X}y_{i} + \overline{X}\overline{Y})}{S_{XX}} \\ &= \frac{\sum (x_{i}y_{i}) - \sum (x_{i}\overline{Y}) - \sum (\overline{X}y_{i}) + \sum (\overline{X}\overline{Y})}{S_{XX}} \\ &= \frac{\sum (x_{i}y_{i}) - n\overline{X}\overline{Y} - n\overline{X}\overline{Y} + n\overline{X}\overline{Y}}{S_{XX}} = \frac{\sum (x_{i}y_{i}) - n\overline{X}\overline{Y}}{S_{XX}} = \frac{\sum (x_{i} - \overline{X})y_{i}}{S_{XX}} \\ Var(\widehat{\beta}_{1}) &= Var\left(\frac{\sum (x_{i} - \overline{X})y_{i}}{S_{XX}}\right) = \frac{\left[\sum (x_{i} - \overline{X})\right]^{2}}{\left(S_{XX}\right)^{2}} Var(y_{i}) = \frac{\left[\sum (x_{i} - \overline{X})\right]^{2}}{\left[\sum (x_{i} - \overline{X})^{2}\right]^{2}} Var(y_{i}) = \frac{1}{S_{XX}} \sigma^{2} \end{split}$$

(a).

$$\begin{aligned} \text{cov}\big(\hat{\beta}_{0}, \hat{\beta}_{1}\big) &= \text{cov}\big(\overline{Y} - \hat{\beta}_{1}\overline{X}, \hat{\beta}_{1}\big) = \text{cov}\big(\overline{Y}, \hat{\beta}_{1}\big) - \text{cov}\big(\hat{\beta}_{1}\overline{X}, \hat{\beta}_{1}\big) = 0 - \overline{X}\text{cov}\big(\hat{\beta}_{1}, \hat{\beta}_{1}\big) \\ &= -\overline{X}\text{Var}\big(\hat{\beta}_{1}\big) = \frac{-\overline{X}\sigma^{2}}{S_{XX}} \end{aligned}$$

(b).

$$\begin{aligned} & \operatorname{cov}(\overline{y}, \widehat{\beta}_{1}) = \operatorname{cov}\left(\frac{\sum y_{i}}{n}, \frac{\sum (x_{i} - \overline{X})y_{i}}{S_{XX}}\right) = \frac{1}{nS_{XX}} \operatorname{cov}\left(\sum y_{i}, \sum (x_{i} - \overline{X})y_{i}\right) \\ &= \frac{1}{nS_{XX}} \operatorname{cov}(y_{1} + y_{2} + \dots + y_{n}, (x_{1} - \overline{X})y_{1} + (x_{2} - \overline{X})y_{2} + \dots + (x_{n} - \overline{X})y_{n}) \\ &= \frac{1}{nS_{XX}} \left[\operatorname{cov}(y_{1}, (x_{1} - \overline{X})y_{1}) + \operatorname{cov}(y_{2}, (x_{2} - \overline{X})y_{2}) + \dots + \operatorname{cov}(y_{n}, (x_{n} - \overline{X})y_{n})\right] \\ &= \frac{\sum \operatorname{cov}(y_{i}, (x_{i} - \overline{X})y_{i})}{nS_{XX}} \Rightarrow \frac{\sum (x_{i} - \overline{X})\operatorname{cov}(y_{i}, y_{i})}{nS_{XX}} = \frac{(n\overline{X} - n\overline{X})\sigma^{2}}{nS_{XX}} = \frac{0 \times \sigma^{2}}{nS_{XX}} = 0 \end{aligned}$$

Q2.

$$\begin{split} SS_R &= \sum_{i=1}^n (\hat{y}_i - \overline{y})^2 = \sum_{i=1}^n \big( \hat{y}_i{}^2 - 2\hat{y}_i \overline{y} + \overline{y}^2 \big) \\ &= \sum_{i=1}^n \hat{y}_i{}^2 - 2\sum_{i=1}^n \hat{y}_i \overline{y} + \sum_{i=1}^n \overline{y}^2 = \sum_{i=1}^n \hat{y}_i{}^2 - 2\overline{y}\sum_{i=1}^n \hat{y}_i + n\overline{y}^2 \end{split}$$

$$= \sum_{i=1}^n {\hat{y}_i}^2 - 2 \bar{y}(n\bar{y}) + n\bar{y}^2 = \sum_{i=1}^n {\hat{y}_i}^2 - 2n\bar{y}^2 + n\bar{y}^2 = \sum_{i=1}^n {\hat{y}_i}^2 - n\bar{y}^2$$

Q3.

(a).

$$H = X(X^{T}X)^{-1}X^{T}$$

$$HH = [X(X^{T}X)^{-1}X^{T}][X(X^{T}X)^{-1}X^{T}] = X(X^{T}X)^{-1}X^{T}X(X^{T}X)^{-1}X^{T}$$

$$= X[(X^{T}X)^{-1}(X^{T}X)](X^{T}X)^{-1}X^{T} = X(X^{T}X)^{-1}X^{T} \quad \because [(X^{T}X)^{-1}(X^{T}X)] = 1$$

$$= X(X^{T}X)^{-1}X^{T} = H$$

$$(1 - H)^{2} = (1 - H)(1 - H) = I - 2H + H^{2} = I - 2H + H \quad \therefore H^{2} = HH = 1$$

$$= 1 - H$$

According to (a) and (b), H is idempotent.

(b).

Define 
$$\hat{Y} = X(X^TX)^{-1}X^TY = HY$$

$$Var(\hat{Y}) = Var(HY) = HH \times Var(Y) = H \times Var(Y)$$

$$\therefore Y = x\beta + \varepsilon , x\beta \text{ is a constant } , \varepsilon \sim N(0, \sigma^2)$$

$$\therefore Var(Y) = Var(x\beta + \varepsilon) = Var(x\beta + \varepsilon) = Var(\varepsilon) = \sigma^2$$

$$\Rightarrow Var(\hat{Y}) = H \times Var(Y) = H\sigma^2$$

Q4.

$$R^{2} = \frac{SS_{Regression}}{SS_{Total}} = \frac{SS_{R}}{SS_{TO}}$$

$$SS_{R} = \sum_{i=1}^{n} (\hat{y}_{i} - \bar{y})^{2}, \hat{y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}x_{i}$$

$$SS_{TO} = \sum_{i=1}^{n} (y_{i} - \bar{y})^{2}, y_{i} = \beta_{0} + \beta_{1}x_{i} + \varepsilon_{i}$$

Although we estimate the  $\beta_0$  and  $\beta_1$  with a close or equal estimator  $\hat{\beta}_0$  and  $\hat{\beta}_1$ , we still cannot estimate the error correctly. Thus  $y_i$  will always larger than  $\hat{y}_i$ , by  $\sum_{i=1}^n \varepsilon_i^2$ . Hence,  $(SS_{TO} - SS_R)$  must be a positive number smaller than  $SS_{TO}$ , leading  $R^2$  smaller than 1.

$X_i$	VIF	$R_j^2$
CYLINDERS	10.738	0.907
DISPLACEMENT	21.837	0.954
HORSEPOWER	9.944	0.899
WEIGHT	10.831	0.908
ACCELERATION	2.626	0.619
MODEL YEAR	1.245	0.197
ORIGIN	1.772	0.436

While building a multiple liner regression model, we need to check whether model has collinearity(共線性) problem with, which leads to certain variables increase predictive power between each other, and reduce of the model accuracy.

In the chart we can see that cylinders, displacement, weight have VIF value larger than 10, namely, they are highly correlate to each other (not independent), and makes them have better predictive power than the other variables. We can combine them into one variable, or deleting two of them instead.

## DA HW 04

March 28, 2021

## 0.0.1 Q5

```
[7]: import pandas as pd from statsmodels.stats.outliers_influence import variance_inflation_factor import statsmodels.api as sm
```

```
Variables
                        VIF
0
          const 763.557531
1
         origin
                   1.772386
                   1.244952
2
    model year
3 acceleration
                   2.625806
4
         weight
                 10.831260
5
    horsepower
                 9.943693
  displacement
                  21.836792
7
      cylinders
                 10.737535
```