

PGM DM 01 MVA 2019/2020

RAMI Hamza, MOKADIRI Mohammed Amine

ENS Paris Saclay, Cachan, Paris

November 22, 2019

1 Learning in discrete graphical models

Consider the following model: z and x are discrete variables taking respectively M and K different values with $p(z = m) = \pi_m$ and $p(x = k|z = m) = \theta_{mk}$

Let's take an i.i.d. sample of observations (Z_1, Z_2, \dots, Z_n) , we have : $P(Z_i = m) = \pi_m$, with $m \in [m_i, 1 \leq i \leq M]$.

The likelihood for (z_1, \dots, z_n) can be written as follow :

$$l_{(z_1, z_2, \dots, z_n)}(\pi) = \prod_{k=1}^n p(z_k|\pi) = \prod_{i=1}^M \pi_{m_i}^{\sum_{k=1}^n \mathbb{1}_{z_k=m_i}} \quad (1)$$

And the Log-likelihood is :

$$L_{(z_1, z_2, \dots, z_n)}(\pi) = \sum_{i=1}^M \left(\sum_{k=1}^n \mathbb{1}_{z_k=m_i} \right) \log(\pi_{m_i})$$

maximizing the loglikelihood can be solved by solving the problem
$$\begin{cases} \max L(\pi) \\ s.t : \\ \sum \pi_{m_i} = 1 \end{cases}$$

we consider the Lagrange multiplier : $h_\lambda(\pi) = L(\pi) + \lambda(1 - \sum_{i=1}^M \pi_{m_i})$

$$\frac{\partial h_\lambda}{\partial \pi_{m_i}}(\pi) = \frac{1}{\pi_{m_i}} \sum_{k=1}^n \mathbb{1}_{z_k=m_i} - \lambda = 0$$

$\pi_{m_i} \lambda = \sum_{k=1}^n \mathbb{1}_{z_k=m_i} \implies \sum_{i=1}^M \pi_{m_i} = \sum_{i=1}^M \sum_{k=1}^n \mathbb{1}_{z_k=m_i} = \sum_{k=1}^n \sum_{i=1}^M \mathbb{1}_{z_k=m_i}$
since $\sum_{i=1}^M \pi_{m_i} = 1$ and $\sum_{i=1}^M \mathbb{1}_{z_k=m_i} = 1$, So finally $\lambda = n$
and

$$\hat{\pi}_{m_i} = \frac{1}{n} \sum_{k=1}^n \mathbb{1}_{z_k=m_i}$$

Let's take an i.i.d. sample of observations (X_1, X_2, \dots, X_n) , we have : $P(X_i = k|z = m) = \theta_{mk}$, with $m \in [m_i, 1 \leq i \leq M]$ and $k \in [k_i, 1 \leq i \leq K]$.

The likelihood for (x_1, \dots, x_n) can be written as follow :

$$l_{(x_1, x_2, \dots, x_n)}(\theta) = \prod_{i=1}^n p(x_i|z, \theta, \pi) = \prod_{k,m} \theta_{mk}^{\sum_{i=1}^n \mathbb{1}_{x_i=k, z=m}} \quad (2)$$

The loglikelihood can be written as :

$$L_{(x_1, x_2, \dots, x_n)}(\theta) = \sum_{k,m} \left(\sum_{i=1}^n \mathbb{1}_{x_i=k, z=m} \right) \log(\theta_{mk}) \quad (3)$$

maximizing the loglikelihood can be solved by solving the problem
$$\begin{cases} \max L(\theta) \\ s.t : \\ \sum \theta_{mk} \pi_m = 1 \end{cases}$$

Setting the derivatives to zero as the previous question, we find :

$$\hat{\theta}_{mk} = \frac{\sum_{i=1}^n \mathbb{1}_{x_i=k, z=m}}{\hat{\pi}_m \sum_{k,m} (\sum_{i=1}^n \mathbb{1}_{x_i=k, z=m})} \quad (4)$$

2 Linear classification

2.1 Generative model (LDA)

a- Let us consider an i.i.d sample of observations $\{(x_i, y_i)\}_{1 \leq i \leq n}$ with $x_i \in \mathbb{R}^2, y_i \in \{0, 1\}$

Suppose that $y \sim \text{Bernoulli}(\pi)$ & $x|y = i \sim \mathcal{N}(\mu_i, \Sigma)$

The likelihood can be written as follow:

$$\begin{aligned} l(\pi, \mu_0, \mu_1, \Sigma) &= \prod_{i=1}^n p(x_i, y_i) \\ &= \prod_{i=1}^n p(x_i|y_i) p(y_i) \\ &= \prod_{i=1}^n \frac{\pi^{y_i} (1-\pi)^{1-y_i}}{2\pi \det(\Sigma)^{1/2}} \exp\left(-\frac{1}{2}(x_i - \mu_{y_i})^T \Sigma^{-1} (x_i - \mu_{y_i})\right) \end{aligned}$$

The Log-likelihood can be written as follow:

$$\begin{aligned} L(\pi, \mu_0, \mu_1, \Sigma) &= \sum_{i=1}^n \log(p(x_i, y_i)) \\ &= \sum_{i=1}^n -\frac{1}{2}(x_i - \mu_{y_i})^T \Sigma^{-1} (x_i - \mu_{y_i}) \sum_{i=1}^n y_i \log(\pi) + \sum_{i=1}^n (1 - y_i) \log(1 - \pi) - \frac{n}{2} \log(\det(\Sigma)) + C^{te} \end{aligned}$$

Setting the derivatives to zeros we get :

$$\hat{\pi} = \frac{\sum_{i=1}^n y_i}{n} \quad (5)$$

$$\hat{\mu}_k = \frac{\sum_{i=1}^n x_i \mathbb{1}_{y_i=k}}{\sum_{i=1}^n \mathbb{1}_{y_i=k}} \quad (6)$$

$$\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (x_i - \mu_k i)(x_i - \mu_k i)^T \text{ with } k i = \mathbb{1}_{y_i=1} \quad (7)$$

b- the form of $p(y = 1|x)$

$$\begin{aligned} p(y = 1|x) &= \frac{p(x|y = 1)p(y = 1)}{p(x|y = 1)p(y = 1) + p(x|y = 0)p(y = 0)} \\ &= \frac{1}{1 + \frac{(1-\pi)f_0(x)}{\pi f_1(x)}} \\ &= \sigma(w^T x + b) \end{aligned}$$

with :

$$w = \Sigma^{-1}(\mu_1 - \mu_0), b = \log\left(\frac{\pi}{1-\pi}\right) - \frac{1}{2}(\mu_1 + \mu_0)^T \Sigma^{-1}(\mu_1 - \mu_0)$$