### Analysis of Time Series, L8

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# Today

- 3.7: Integrated models
- 3.8: Building ARIMA models

Recall:

### Definition (2.4)

The backshift operator is defined by

$$Bx_t = x_{t-1}.$$

For  $k = 1, 2, ..., B^k x_t = x_{t-k}$ .

### Definition (2.5)

Differences of order d are defined by

$$\nabla^d x_t = (1 - B)^d x_t.$$

### Special cases:

- $\nabla^1 x_t = \nabla x_t = (1 B)x_t = x_t x_{t-1}$
- $\nabla^2 x_t = (1 B)^2 x_t = x_t 2x_{t-1} + x_{t-2}$

- Two types of trends:
  - **1** Deterministic (example:  $x_t = \beta_0 + \beta_1 t + w_t$ )
  - 2 Stochastic (example:  $x_t = x_{t-1} + w_t$ )
- May be removed by differencing.
- Watch out for overdifferencing!
- Example:  $x_t = x_{t-1} + w_t \Rightarrow \nabla x_t = w_t \Rightarrow \nabla^2 x_t = \nabla w_t = w_t w_{t-1}$ .  $\operatorname{corr}(\nabla x_{t+1}, \nabla x_t) = 0$ ,  $\operatorname{corr}(\nabla^2 x_{t+1}, \nabla^2 x_t) = -1/2 \neq 0$ .
- Too much differencing may introduce extra autocorrelations and non invertibility!

### Example:

Let

$$x_t = \beta_0 + \beta_1 t + \beta_2 t^2 + y_t,$$

where  $\beta_2 \neq 0$  and  $y_t$  is stationary.

• How many differences are required to make  $x_t$  stationary?



### Definition (3.11)

A process  $\{x_t\}$  is said to be ARIMA(p, d, q) if

$$\nabla^d x_t = (1 - B)^d x_t$$

is ARMA(p, q). We may write the model as

$$\phi(B)(1-B)^d x_t = \theta(B) w_t.$$

If  $E(\nabla^d x_t) = \mu$ , we write the model as

$$\phi(B)(1-B)^d x_t = \delta + \theta(B)w_t,$$

where

$$\delta = \mu(1 - \phi_1 - \dots - \phi_p).$$

Prediction (based on infinite past):

- Write the process on AR form to obtain the forecast.
- Write the process on MA form to obtain the prediction error.
- Example: Forecasting IMA(1,1)

$$\nabla x_t = w_t - \lambda w_{t-1}.$$

Leads to the Holt and Winter method:

$$\tilde{x}_{n+1} = (1 - \lambda)x_n + \lambda \tilde{x}_n.$$

Why?



- Oheck if the series is stationary. If not:
  - Remove trends by differencing.
  - Make the variance constant by transformations.
- Identify an ARMA model by looking at
  - ACF
  - PACF
  - Information criteria (AIC, BIC...)
- Estimate the model.
- Check the model fit by performing residual diagnostics.
- If the model does not fit well, start over from 1.

- Make the variance constant!
- Box-Cox transformation (chap.2, p.59)

$$y_t = \begin{cases} \frac{x_t^{\lambda} - 1}{\lambda}, & \lambda \neq 0, \\ \log x_t, & \lambda = 0. \end{cases}$$

Modified (not in book):

$$\tilde{y}_t = \operatorname{gm}(x)^{1-\lambda} y_t,$$

where 
$$gm(x) = (\prod_{i=1}^{n} x_i)^{1/n}$$
.

By Taylor expansion (why?),

$$\frac{x_t^{\lambda}-1}{\lambda}=\log x_t+O(\lambda).$$



Model identification via ACF and PACF:

	AR(p)	MA(q)	ARMA(p,q)
ACF	Tails off	Cuts off after lag $q$	Tails off
PACF	Cuts off after lag $p$	Tails off	Tails off

Let  $\hat{\sigma}_k^2 = SSE_k/n$ , where  $SSE_k$  is the residual sum of squares and k is the number of parameters in the model.

Definition (Akaike Information Criterion)

$$AIC = \log(\hat{\sigma}_k^2) + \frac{2k}{n}.$$

In R: AIC =  $-2\log(L) + 2k$ .

Definition (Bayesian Information Criterion)

$$\mathsf{BIC} = \log(\hat{\sigma}_k^2) + \frac{k \log(n)}{n}.$$

- BIC is recommended for large samples, AIC for small samples.
- It is *not recommended* to compare AIC, BIC between ARIMA models with different *d*, or for different transformations.

- Test  $H_0$ : The residuals are white noise vs  $H_1$ :  $\neg H_0$ .
- Ljung-Box-Pierce Q statistic (p.139)

$$Q = n(n+2) \sum_{h=1}^{H} \frac{\hat{\rho}_{e}^{2}(h)}{n-h},$$

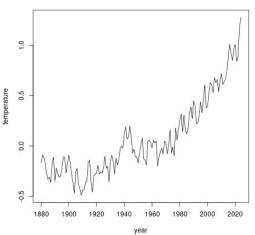
where  $\hat{\rho}_e(h)$  are the estimated autocorrelations of the residuals.

• Asymptotically,  $Q \sim \chi^2_{H-p-q}$ .

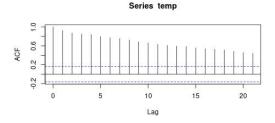


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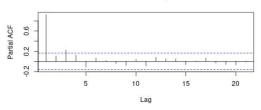


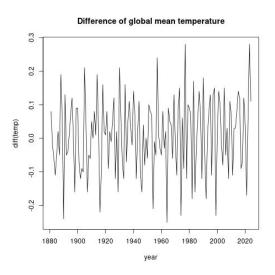


Global mean temperature, ACF and PACF (typical signs of a trend):

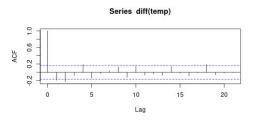


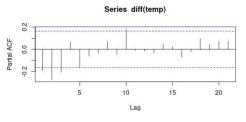
### Series temp





Difference of global mean temperature, ACF (cuts off after lag 2) and PACF (tails off?):





```
In R, try MA(2) for differences (ARIMA(0,1,2)):
> arima(temp, order=c(0,1,2))
Call:
arima(x = temp, order = c(0, 1, 2))
Coefficients:
          ma1
                   ma2
      -0.3010 -0.2118
s.e. 0.0783 0.0702
sigma^2 estimated as 0.01106: log likelihood = 119.84,
aic = -233.68
```

Among many models, ARIMA(3,1,1) gave the smallest AIC.

```
> arima(temp,order=c(3,1,1))
```

### Call:

```
arima(x = temp, order = c(3, 1, 1))
```

### Coefficients:

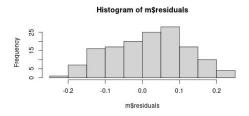
```
sigma^2 estimated as 0.01053: log likelihood = 123.27, aic = -236.55
```

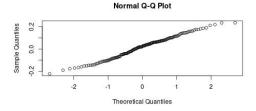
Check: all coefficients are significant, i.e. outside the two times s.e. bound.

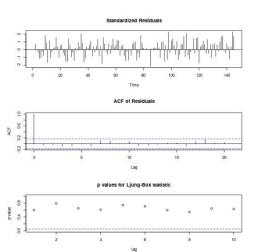
```
Residual diagnostics, ARIMA(3,1,1):
```

```
> m=arima(temp,order=c(3,1,1))
```

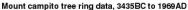
- > par(mfrow=c(2,1))
- > hist(m\$residuals)
- > qqnorm(m\$residuals)
- > dev.off()
- > tsdiag(m)

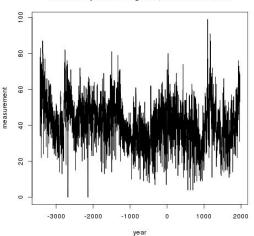




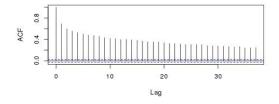


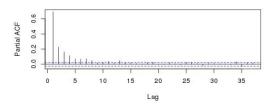
No sign of autocorrelations in the residuals!

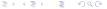




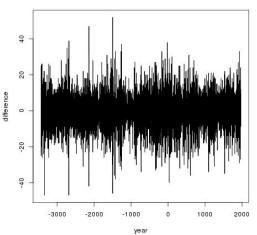
Mount Campito, ACF and PACF. Maybe not stationary?



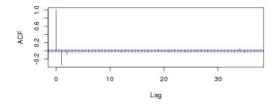


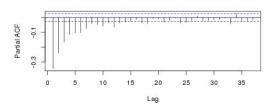


### Mount campito tree ring data, differences



### Mount Campito differences, ACF and PACF. MA(2)?





MA(2) for differences, estimation in R (intercept not significant):

```
> dy=diff(y)
> a=arima(dy,order=c(0,0,2));a
```

### Call:

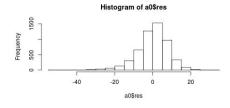
arima(x = dy, order = c(0, 0, 2))

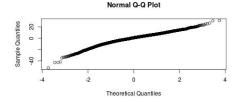
### Coefficients:

 $sigma^2$  estimated as 65.34: log likelihood = -18961.64, aic = 37931.27

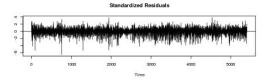
```
MA(2) without constant (note: AIC two units lower):
> a0=arima(dy,order=c(0,0,2),include.mean=FALSE);a0
Call:
arima(x = dy, order = c(0, 0, 2), include.mean = FALSE)
Coefficients:
          ma1
                   ma2
      -0.5449 -0.1921
s.e. 0.0130 0.0140
sigma^2 estimated as 65.34: log likelihood = -18961.64,
aic = 37929.27
```

- > par(mfrow=c(2,1))
- > hist(a0\$res)
- > qqnorm(a0\$res)

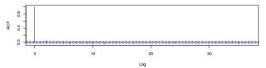




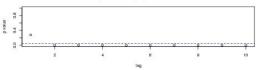
### > tsdiag(a0)



#### **ACF of Residuals**



#### p values for Ljung-Box statistic



Small p values, i.e. significant autocorrelations.

MountCampito: Find the ARMA(p, q) model for differences without constant with the smallest AIC:

р	q	AIC
0	0	39339.2
0	1	38107.1
1	0	38657.1
0	2	37929.3
1	1	37885.3
2	0	38341.7
0	3	37901.2
1	2	37854.4
2	1	37872.9
3	0	38191.7
0	4	37887.8
1	3	37818.9
2	2	37815.9
3	1	37851.8
4	0	38123.2
2	3	37826.7
3	2	37818.0
3	3	37819.8

### Try ARMA(2,2):

```
> a1=arima(dy,order=c(2,0,2),include.mean=FALSE);a1
```

### Call:

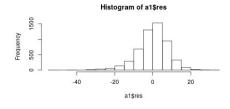
arima(x = dy, order = c(2, 0, 2), include.mean = FALSE)

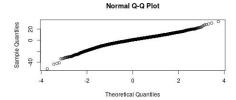
### Coefficients:

```
sigma<sup>2</sup> estimated as 63.93: log likelihood = -18902.95, aic = 37815.9
```

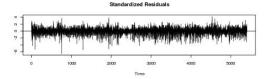
Observe: The invertibility condition  $\theta_2 - \theta_1 < 1$  is not satisfied!

- > par(mfrow=c(2,1))
- > hist(a1\$res)
- > qqnorm(a1\$res)

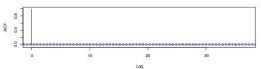




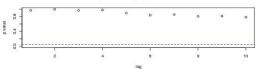
### > tsdiag(a1)



#### **ACF of Residuals**

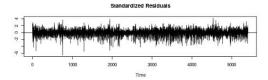


#### p values for Ljung-Box statistic

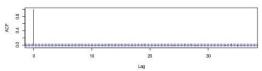


Better on autocorrelations, but try more lags for Ljung-Box!

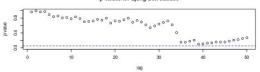
### > tsdiag(a1,50)



#### **ACF of Residuals**

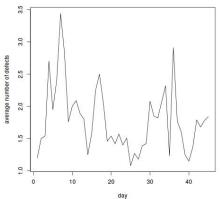


#### p values for Ljung-Box statistic

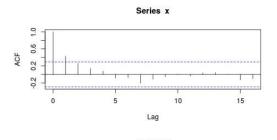


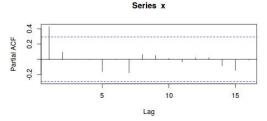
Ljung-Box a bit suspicious here (p values for high lags close to 0.05)...

Daily average number of defects per truck found in the final inspection at the end of the assembly line of a truck manufacturing plant. (Looks stationary.)



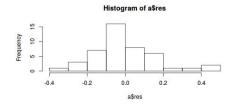
Trucks, ACF (tails off) and PACF (cuts off after lag 1). AR(1)?

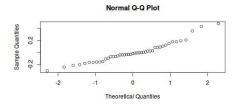




```
Try AR(1):
> a=arima(x,order=c(1,0,0));a
Call:
arima(x = x, order = c(1, 0, 0))
Coefficients:
         ar1
              intercept
      0.4322
                 1.7799
s.e. 0.1340
                 0.1189
sigma^2 estimated as 0.2118: log likelihood = -29.04,
aic = 64.07
Significant coefficients!
```

- > par(mfrow=c(2,1))
- > hist(a\$res)
- > qqnorm(a\$res)

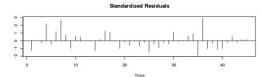




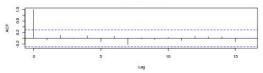
Maybe some outliers to the right?



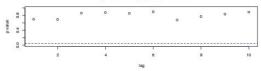
### > tsdiag(a)



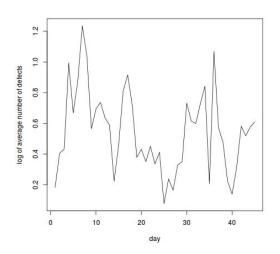
#### **ACF of Residuals**



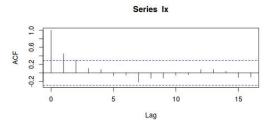
#### p values for Ljung-Box statistic



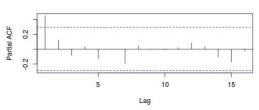
Take logarithms (to get a smaller effect of outliers):



### Trucks in logs, ACF and PACF



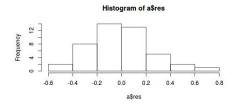
### Series Ix

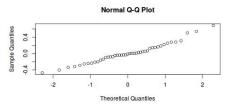


```
Try AR(1):
> lx = log(x)
> a=arima(lx,order=c(1,0,0));a
Call:
arima(x = lx, order = c(1, 0, 0))
Coefficients:
              intercept
         ar1
      0.4582
                  0.5391
s.e. 0.1330
                 0.0641
sigma<sup>2</sup> estimated as 0.05625: log likelihood = 0.78,
aic = 4.43
```

Do not compare AIC to the previous value!

- > par(mfrow=c(2,1))
- > hist(a\$res)
- > qqnorm(a\$res)

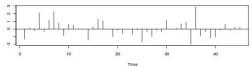




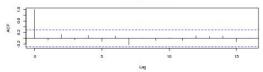
Less pronounced outliers now.

### > tsdiag(a)

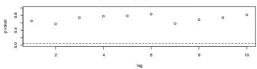
#### Standardized Residuals



### ACF of Residuals



#### p values for Ljung-Box statistic



# News of today

- Removing trends by differencing
- Model building:
  - Transformation
  - Identification
  - Estimation
  - Diagnostics