

① See Lecture 2

② Note that

$$\begin{aligned}
 E(|X_a|; |X_a| > K) &= E(X_a; X_a > K) \\
 &= \int_K^\infty x(a-1)x^{-a} I_{[1,\infty)}(x) dx \\
 &= \int_K^\infty (a-1)x^{1-a} dx \quad (\text{for } K \geq 1) \\
 &= \left. \frac{(a-1)x^{2-a}}{2-a} \right|_K^\infty \\
 &= \frac{a-1}{a-2} K^{2-a} \quad \text{for } a > 2 \text{ (otherwise, it is infinite)}
 \end{aligned}$$

Let  $a^* = \inf A$ . If  $a^* > 2$ , then we have

$$\begin{aligned}
 E(X_a; X_a > K) &= \frac{a-1}{a-2} K^{2-a} = \left(1 + \frac{1}{a-2}\right) \frac{1}{K^{a-2}} \\
 &\leq \left(1 + \frac{1}{a^*-2}\right) \frac{1}{K^{a^*-2}} \quad \text{for all } K \geq 1, a \in A
 \end{aligned}$$

So if we choose  $K > \left(\frac{\varepsilon}{1 + \frac{1}{a^*-2}}\right)^{-\frac{1}{a^*-2}}$ , then

$$E(X_a; X_a > K) < \varepsilon \quad \text{for all } a \in A$$

Thus the family is uniformly integrable if  $a^* = \inf A > 2$ .

On the other hand, if  $a^* \leq 2$ , then either  $a \leq 2$  for some  $a \in A$  (so that  $E(X_a) = \infty$ ), or there is a sequence  $a_1 \geq a_2 \geq \dots$  of elements in  $A$  such that  $a_n \rightarrow 2$ . For these, we have

$$\lim_{n \rightarrow \infty} E(X_{a_n}; |X_{a_n}| > K) = \lim_{n \rightarrow \infty} \left(1 + \frac{1}{a_n-2}\right) \frac{1}{K^{a_n-2}} = \infty$$

for every choice of  $K$ . Thus there is no choice of  $K$  such that  $E(|X_a|; |X_a| > K) < \varepsilon$  for all  $a \in A$  if  $\varepsilon$  is given.

So the family is not uniformly integrable if  $a^* = \inf A \leq 2$ .

③ We can write

$$\begin{aligned} & \int_{-\infty}^{\infty} (P(X < x \leq Y) - P(Y < x \leq X)) dx \\ &= \int_{-\infty}^{\infty} \int_{\Omega} (I_{\{X < x \leq Y\}} - I_{\{Y < x \leq X\}}) dP dx \end{aligned}$$

Now note first that

$$\begin{aligned} & \int_{-\infty}^{\infty} \int_{\Omega} |I_{\{X < x \leq Y\}} - I_{\{Y < x \leq X\}}| dP dx \\ &= \int_{-\infty}^{\infty} \int_{\Omega} |I_{\{X < x \leq Y\}} + I_{\{Y < x \leq X\}}| dx dP \\ &\leq \int_{-\infty}^{\infty} \int_{\Omega} (I_{\{X < x \leq Y\}} + I_{\{Y < x \leq X\}}) dx dP \\ &= \int_{\Omega} ((Y-X)I_{\{X < Y\}} + (X-Y)I_{\{Y < X\}}) dP \\ &= \int_{\Omega} |Y-X| dP < \infty \quad \text{since } X, Y \text{ are integrable} \end{aligned}$$

So  $I_{\{X < x \leq Y\}} - I_{\{Y < x \leq X\}}$  is integrable, and Fubini's theorem applies. Thus

$$\begin{aligned} & \int_{-\infty}^{\infty} (P(X < x \leq Y) - P(Y < x \leq X)) dx \\ &= \int_{-\infty}^{\infty} \int_{\Omega} (I_{\{X < x \leq Y\}} - I_{\{Y < x \leq X\}}) dP dx \\ &= \int_{\Omega} \int_{-\infty}^{\infty} (I_{\{X < x \leq Y\}} - I_{\{Y < x \leq X\}}) dx dP \\ &= \int_{\Omega} ((Y-X)I_{\{X < Y\}} - (X-Y)I_{\{Y < X\}}) dP \\ &= \int_{\Omega} (Y-X) dP = EY - EX \end{aligned}$$

④ See Lecture 4

⑤ (a) A stopping time with respect to a filtration  $\mathcal{F}_n$  is a random variable  $T$  with values in  $\{0, 1, \dots, \infty\}$  such that

$$\{T = n\} \in \mathcal{F}_n$$

$\tau_1$  is not a stopping time:

$\tau_1 = \sup \{n: P_n \leq 10\}$  is the last time that  $P_n \leq 10$ . Since the possible values of  $P_n$  are  $1, 2, 4, 8, \dots$ , we can also write

$$\{\tau_1 = n\} = \{P_n = 8\} \cap \{P_{n+1} = 16\} = \{P_n = 8\} \cap \{X_{n+1} = 2\}$$

which is not in  $\mathcal{F}_n$ , since it involves  $X_{n+1}$ .

$\tau_2$  is a stopping time:

$\tau_2 = \inf \{n: P_n > 10\}$  is the first time that  $P_n > 10$ . We have

$$\{\tau_2 = n\} = \{P_{n-1} = 8\} \cap \{P_n = 16\} = \{P_{n-1} = 8\} \cap \{X_n = 2\} \in \mathcal{F}_n$$

$\tau_3$  is a stopping time:

$$\tau_3 = \inf \{n: P_n = P_{n-10}\} = \inf \{n: X_n = X_{n-1} = \dots = X_{n-9} = 1\}$$

is the first time that ten consecutive  $X_i$  are equal to 1. We have

$$\{\tau_3 = n\} = \{X_n = X_{n-1} = \dots = X_{n-9} = 1\} \cap \{X_j X_{j-1} \dots X_{j-9} \neq 1 \text{ for all } j < n\}$$

This also lies in  $\mathcal{F}_n$ .

(b) We have

$$\begin{aligned} E(P_n | \mathcal{F}_{n-1}) &= E(P_{n-1} X_n | \mathcal{F}_{n-1}) \\ &= P_{n-1} E(X_n) = \frac{3}{2} P_{n-1} \end{aligned}$$

Thus for  $c = \frac{2}{3}$ ,

$$E(c^{-n} P_n | \mathcal{F}_{n-1}) = c^{-n} E(P_n | \mathcal{F}_{n-1}) = c^{-n} \cdot \frac{3}{2} P_{n-1} = c^{-(n-1)} P_{n-1},$$

meaning that  $c^{-n} P_n$  is a martingale.

Note finally that

$$\begin{aligned} \ln(c^{-n} P_n) &= -n \ln c + \ln P_n \\ &= -n \ln c + \ln(X_1 X_2 \dots X_n) \\ &= (\ln X_1 - \ln c) + (\ln X_2 - \ln c) + \dots + (\ln X_n - \ln c) \end{aligned}$$

$$\begin{aligned}\text{Since } E(\ln X_i - \ln c) &= \left(\frac{1}{2} \ln 2 + \frac{1}{2} \ln 1\right) - \ln c \\ &= \frac{1}{2} \ln 2 - \ln \frac{3}{2} = \frac{1}{2} (\ln 2 - \ln \frac{9}{4}) \\ &= \frac{1}{2} \ln \frac{8}{9} < 0\end{aligned}$$

it follows from the strong law of large numbers that

$$\frac{1}{n} \ln(c^{-n} P_n) \xrightarrow{\text{a.s.}} \frac{1}{2} \ln \frac{8}{9} < 0$$

Thus  $\ln(c^{-n} P_n) \xrightarrow{\text{a.s.}} -\infty$  and  $c^{-n} P_n \xrightarrow{\text{a.s.}} 0$ .

⑥ Note first that  $X_n$  is a martingale with respect to the natural filtration  $\mathcal{F}_n = \sigma(X_0, X_1, \dots, X_n)$ :

$$E(X_n | \mathcal{F}_{n-1}) = \frac{1}{2} \cdot \frac{X_{n-1}^2 + 2X_{n-1} - 1}{2} + \frac{1}{2} \cdot \frac{-X_{n-1}^2 + 2X_{n-1} + 1}{2} = X_{n-1}.$$

Moreover,  $X_n \in (-1, 1)$  for all  $n$ : this follows by induction since  $X_0 \in (-1, 1)$ , and

$$\frac{X_{n-1}^2 + 2X_{n-1} - 1}{2} = \frac{(X_{n-1} + 1)^2 - 2}{2} \in \left(\frac{0^2 - 2}{2}, \frac{2^2 - 2}{2}\right) = (-1, 1)$$

if  $X_{n-1} \in (-1, 1)$  as well as

$$\frac{-X_{n-1}^2 + 2X_{n-1} + 1}{2} = \frac{-(X_{n-1} - 1)^2 + 2}{2} \in \left(\frac{-2^2 + 2}{2}, \frac{-0^2 + 2}{2}\right) = (-1, 1)$$

if  $X_{n-1} \in (-1, 1)$ .

So  $X_n$  is a bounded martingale, thus convergent (almost surely) by the martingale convergence theorem.

A possible limit  $L$  needs to satisfy  $L = \lim_{n \rightarrow \infty} X_n = \lim_{n \rightarrow \infty} X_{n-1}$ , thus

$$L = \frac{L^2 + 2L - 1}{2} \quad \text{and} \quad L = \frac{-L^2 + 2L + 1}{2}$$

$$\Leftrightarrow L^2 = 1$$

$$\Leftrightarrow L = \pm 1$$

$$\Leftrightarrow L^2 = 1$$

$$\Leftrightarrow L = \pm 1$$

So the possible limits are  $\pm 1$ . Finally,  $E(X_n) = E(X_0) = a$  for all  $n$ , thus

$$E\left(\lim_{n \rightarrow \infty} X_n\right) = \lim_{n \rightarrow \infty} E(X_n) = a$$

and at the same time

$$\begin{aligned} E(\lim_{n \rightarrow \infty} X_n) &= P(\lim_{n \rightarrow \infty} X_n = 1) - P(\lim_{n \rightarrow \infty} X_n = -1) \\ &= P(\lim_{n \rightarrow \infty} X_n = 1) - (1 - P(\lim_{n \rightarrow \infty} X_n = 1)) \\ &= 2P(\lim_{n \rightarrow \infty} X_n = 1) - 1, \end{aligned}$$

thus  $P(\lim_{n \rightarrow \infty} X_n = 1) = \frac{a+1}{2}$  and  $P(\lim_{n \rightarrow \infty} X_n = -1) = \frac{1-a}{2}$ .

⑦ See Lectures 15 and 16

⑧ (a) See Lectures 16 and 17

(b) See Lectures 19 and 20

(c) For example, consider the following simple model with a single time step:

$$\begin{array}{l} S_0 = 2 \quad \begin{array}{l} \nearrow_{q_1} S_1 = 1 \\ \searrow_{q_2} S_1 = 2 \\ \searrow_{q_3} S_1 = 3 \end{array} \end{array}$$

Assume also that the risk-free rate is 0. Now both  $q_1 = q_2 = q_3 = \frac{1}{3}$  and  $q_1 = q_3 = \frac{1}{4}, q_2 = \frac{1}{2}$  (and infinitely many others) define equivalent martingale measures.

Since there is a martingale measure, but not a unique one, the model is viable, but not complete.