

① (a) See Lecture 2 and Lecture 4

(b) Note that for every n , we have

$$\begin{aligned} P(X_n = Y_n) &= \sum_{i=1}^n P(X_n = i \text{ and } Y_n = i) \\ &= \sum_{i=1}^n P(X_n = i)P(Y_n = i) \\ &= n \cdot \left(\frac{1}{n}\right)^2 = \frac{1}{n} \end{aligned}$$

and similarly

$$\begin{aligned} P(X_n = Y_n = Z_n) &= \sum_{i=1}^n P(X_n = i)P(Y_n = i)P(Z_n = i) \\ &= n \cdot \left(\frac{1}{n}\right)^3 = \frac{1}{n^2} \end{aligned}$$

Since $\sum_{n=1}^{\infty} \frac{1}{n} = \infty$, but $\sum_{n=1}^{\infty} \frac{1}{n^2} < \infty$, it follows from the Borel-Cantelli lemmas that

$$P(X_n = Y_n \text{ infinitely often}) = 1, \text{ but}$$

$$P(X_n = Y_n = Z_n \text{ infinitely often}) = 0$$

② Note that

$$\begin{aligned} P(X_i = X_{i+1} = \dots = X_{i+n-1}) &= P(X_i = X_{i+1} = \dots = X_{i+n-1} = 1) \\ &\quad + P(X_i = X_{i+1} = \dots = X_{i+n-1} = -1) \\ &= \left(\frac{1}{2}\right)^n + \left(\frac{1}{2}\right)^n = 2^{1-n} \end{aligned}$$

So if Z_N is the number of times that we have n consecutive identical values among the first N , we have

$$\begin{aligned} E(Z_N) &= E\left(\sum_{i=1}^{N-n+1} I_{\{X_i = X_{i+1} = \dots = X_{i+n-1}\}}\right) \\ &= \sum_{i=1}^{N-n+1} P(X_i = X_{i+1} = \dots = X_{i+n-1}) \\ &= (N-n+1)2^{1-n} \end{aligned}$$

If $N = \lfloor c^n \rfloor$, where $c < 2$, then

$$E(Z_N) = (\lfloor c^n \rfloor - n + 1)2^{1-n} \leq c^n \cdot 2^{1-n} = 2\left(\frac{c}{2}\right)^n \rightarrow 0 \text{ as } n \rightarrow \infty$$

Since $P(Z_N > 0) = P(Z_N \geq 1) \leq E(Z_N)$ by the Markov inequality, the first part follows.

For the second part, consider the $\lfloor \frac{N}{n} \rfloor$ disjoint n -tuples

$$X_1, X_2, \dots, X_n$$

$$X_{n+1}, X_{n+2}, \dots, X_{2n}$$

$$X_{\lfloor \frac{N}{n} \rfloor n - n + 1}, X_{\lfloor \frac{N}{n} \rfloor n - n + 2}, \dots, X_{\lfloor \frac{N}{n} \rfloor n}$$

Since they are disjoint, the events $X_1 = X_2 = \dots = X_n$, $X_{n+1} = X_{n+2} = \dots = X_{2n}$ etc are independent. It follows that

$$\begin{aligned} P(Z_N = 0) &\leq P(\{X_1, X_2, \dots, X_n \text{ not all equal}\} \cap \{X_{n+1}, X_{n+2}, \dots, X_{2n} \text{ not all equal}\} \cap \dots) \\ &= \prod_{i=1}^{\lfloor \frac{N}{n} \rfloor} (1 - P(X_{(i-1)n+1} = X_{(i-1)n+2} = \dots = X_{in})) \\ &= (1 - 2^{1-n})^{\lfloor \frac{N}{n} \rfloor} \\ &= (1 - 2^{1-n})^{2^{n-1} (\lfloor \frac{N}{n} \rfloor \cdot 2^{1-n})} \end{aligned}$$

Since $\lfloor \frac{N}{n} \rfloor \cdot 2^{1-n} = \lfloor \frac{c}{n} \rfloor \cdot 2^{1-n} \rightarrow \infty$ for $c > 2$ and $\lim_{n \rightarrow \infty} (1 - 2^{1-n})^{2^{n-1}} = \frac{1}{e}$, it follows that $P(Z_N = 0) \rightarrow 0$, which concludes the second part.

③ See Lecture 4

④ (a) A stopping time with respect to a filtration \mathcal{F}_n is a random variable T with values in $\{0, 1, \dots, \infty\}$ such that

$$\{T = n\} \in \mathcal{F}_n$$

T_1 is a stopping time:

$$\{\tau_1 = n\} = \{S_n = 10\} \cap \{S_1 \neq 10\} \cap \{S_2 \neq 10\} \dots \cap \{S_{n-1} \neq 10\}$$

can be determined from S_1, S_2, \dots, S_n and is thus in \mathcal{F}_n .

T_2 is not a stopping time:

$$\{\tau_2 = n\} = \{S_n = 10\} \cap \{S_{n+1} \neq 10\} \cap \{S_{n+2} \neq 10\} \dots$$

cannot be determined from S_1, S_2, \dots, S_n alone.

T_3 is a stopping time:

$$\sup \{n : S_n = n\} = \sup \{n : X_1 = X_2 = \dots = X_n = 1\}$$

$$\Rightarrow \{\tau_3 = n\} = \{X_1 = X_2 = \dots = X_{n-1} = 1\} \cap \{X_n = -1\} \in \mathcal{F}_n$$

(b) See Lecture 11

$$\begin{aligned}(c) \quad E(S_n^2 - n | \mathcal{F}_{n-1}) &= E((S_{n-1} + X_n)^2 - n | \mathcal{F}_{n-1}) \\&= E(S_{n-1}^2 + 2X_n S_{n-1} + X_n^2 - n | \mathcal{F}_{n-1}) \\&= S_{n-1}^2 + 2S_{n-1} E(X_n | \mathcal{F}_{n-1}) + E(X_n^2 - n | \mathcal{F}_{n-1}) \\&= S_{n-1}^2 + 2S_{n-1} E(X_n) + 1 - n \quad \leftarrow \text{always } -1 \\&= S_{n-1}^2 - (n-1) \quad \leftarrow 0\end{aligned}$$

showing that $S_n^2 - n$ is a martingale.

By the optional stopping theorem, we have

$$E(S_\tau^2 - \tau) = E(S_0 - 0) = 0$$

$\stackrel{||}{=} \text{by definition}$

$$\Rightarrow 100 - E(\tau) = 0 \Rightarrow E(\tau) = 100$$

⑤(a) We can express X_n as $X_n = Y_n \cdot X_{n-1}$, where Y_n is uniform on $[1, 2]$.
Note that

$$\begin{aligned}E(c^n X_n | \mathcal{F}_{n-1}) &= c^n E(Y_n X_{n-1} | \mathcal{F}_{n-1}) \\&= c^n E(Y_n) X_{n-1} \\&= c \cdot \frac{3}{2} \cdot c^{n-1} X_{n-1}\end{aligned}$$

So $c^n X_n$ is a martingale if $c = \frac{2}{3}$, a supermartingale if $c < \frac{2}{3}$ and a submartingale if $c > \frac{2}{3}$.

(b) We have

$$\begin{aligned}\ln(c^n X_n) &= n \ln c + \ln(Y_1 \cdots Y_n) \\&= \ln(cY_1) + \ln(cY_2) + \cdots + \ln(cY_n)\end{aligned}$$

$$\begin{aligned}\text{Now } E(\ln(cY_i)) &= \int_1^2 \ln(cx) dx = \ln c + x(\ln x - 1) \Big|_1^2 \\&= \ln c + 2\ln 2 - 1 = \ln\left(\frac{4c}{e}\right)\end{aligned}$$

If $c > \frac{e}{4}$, it follows from the strong law of large numbers that

$\frac{\ln(c^n X_n)}{n} \rightarrow \ln\left(\frac{4c}{e}\right) > 0 \Rightarrow \ln(c^n X_n) \rightarrow \infty$
 and thus $c^n X_n \rightarrow \infty$. Likewise, for $c < \frac{e}{4}$, we have

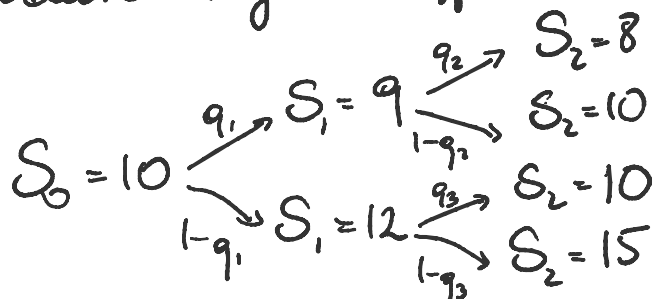
$$\frac{\ln(c^n X_n)}{n} \rightarrow \ln\left(\frac{4c}{e}\right) < 0 \Rightarrow \ln(c^n X_n) \rightarrow -\infty$$

and thus $c^n X_n \rightarrow 0$. This proves the statement with $c_0 = \frac{e}{4}$.

⑥ See Lecture 12

⑦ (a), (b): See Lectures 19 and 20

(c) A martingale measure is determined by the probabilities of the different transitions:



For $\bar{S}_t = S_t$ to be a martingale, we need

- $9q_1 + 12(1-q_1) = 10 \Leftrightarrow 2 = 3q_1 \Leftrightarrow q_1 = \frac{2}{3}$
- $8q_2 + 10(1-q_2) = 9 \Leftrightarrow 1 = 2q_2 \Leftrightarrow q_2 = \frac{1}{2}$
- $10q_3 + 15(1-q_3) = 12 \Leftrightarrow 3 = 5q_3 \Leftrightarrow q_3 = \frac{3}{5}$

Since there is a martingale measure, the model is viable.
 Since it is also unique, the model is complete.

⑧ See Lecture 18