## Inferens 1, F2

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2 november 2022

## Today

- 7.1: Statistical inference, introduction
- 7.2: Estimation
  - 7.2.1: Properties of estimates
  - 7.2.2: Asymptotic properties

In WMS: 8.1-4

#### Example 1:

- In an opinion poll, 1000 randomly selected voters are asked about their political sympathies.
- Let X be the number of these voters who sympathize with the party P.
- Because the number of voters is so large, we may assume that  $X \sim \text{Bin}(1000, p)$ .
- Say that 100 of the selected voters sympathize with the party P.
- Which is your guess on the value of *p*?



http://www.fjellfotografen.se/Djur/Däggdjur/Älgar/Älgtjur i skog/@toj-03020 Example 2:

- A researcher measures the withers height (mankhöjd) of five randomly selected moose (älgtjurar).
- The researcher assumes that the withers height in cm is  $N(\mu, \sigma^2)$ .
- The measurements are 200 185 210 190 190
- Which is your guess on the value of  $\mu$ ? 190 195 197.5

- Probability:
  - Model with *known* parameters
  - $\rightarrow$  (probabilities for) random samples.
- Statistical inference:
  - Model with unknown parameters, observed random sample
  - $\rightarrow$  make inference on (e.g. estimate) parameter values.

### Definition (7.1)

 $x_1, x_2, ..., x_n$  is a *sample* (stickprov) from the random variable X with distribution F, if  $x_1, x_2, ..., x_n$  are observations of  $X_1, X_2, ..., X_n$ , which all have distribution F.

If, in addition, the variables  $X_1, X_2, ..., X_n$  are independent, we have a random sample (slumpmässigt stickprov) from X (from F).

#### Example 1:

- In an opinion poll, 1000 randomly selected voters are asked about their political sympathies.
- Say that 100 of the selected voters sympathize with the party P.
- This is a random sample of size n=1 where  $x_1=100$  is an observation of the random variable  $X_1 \sim \text{Bin}(1000, p)$ .
- Alternatively: A random sample  $u_1, ..., u_n$ , as observations of  $U_1, ..., U_n$ , n = 1000, where each  $U_i \sim \text{Be}(p)$  and where the  $U_i$  are independent. We observe

$$\sum_{i=1}^n u_i = 100.$$

#### Example 2:

- A researcher measures the withers height (mankhöjd) of five randomly selected moose (älgtjurar).
- The researcher assumes that the withers height in cm is  $N(\mu, \sigma^2)$ , where it is known that  $\sigma^2 = 100$ .
- The measurements are 200 185 210 190 190
- $(x_1, x_2, x_3, x_4, x_5) = (200, 185, 210, 190, 190)$  is a random sample from  $X \sim N(\mu, 100)$ , i.e.
- $x_1=200$  is an observation of the random variable  $X_1\sim N(\mu,100)$ ,  $x_2=185$  is an observation of the random variable  $X_2\sim N(\mu,100)$ ,  $x_3=210$  is an observation of the random variable  $X_3\sim N(\mu,100)$ ,  $x_4=190$  is an observation of the random variable  $X_4\sim N(\mu,100)$ ,  $x_5=190$  is an observation of the random variable  $X_5\sim N(\mu,100)$ , where  $X_1,X_2,X_3,X_4,X_5$  are independent (and distributed as  $X_5$ ).

### Estimation

- Suppose we have one unknown parameter  $\theta$ .
- Write

$$\mathbf{x} = (x_1, x_2, ..., x_n)$$
  
 $\mathbf{X} = (X_1, X_2, ..., X_n)$ 

### Definition (7.2)

An estimate (skattning)  $\theta^* = \theta^*(\mathbf{x})$  is a function of the sample  $\mathbf{x}$ .

The estimate is an observation of the *estimator*  $\theta^*(X)$ .



### Estimation

#### Example 2:

- $(x_1, x_2, x_3, x_4, x_5) = (200, 185, 210, 190, 190)$  is a random sample from  $X \sim N(\mu, 100)$ , i.e.
- $x_1=200$  is an observation of the random variable  $X_1\sim N(\mu,100)$ ,  $\vdots$   $x_5=190$  is an observation of the random variable  $X_5\sim N(\mu,100)$ ,
  - where  $X_1, X_2, X_3, X_4, X_5$  are independent and distributed as X.
- ullet Estimate  $\mu$  by the sample mean.
- Estimate

$$\mu^*(x_1, x_2, x_3, x_4, x_5) = \overline{x} = \frac{x_1 + x_2 + x_3 + x_4 + x_5}{5} = 195$$

Estimator

$$\mu^*(X_1, X_2, X_3, X_4, X_5) = \overline{X} = \frac{X_1 + X_2 + X_3 + X_4 + X_5}{5} \sim N(\mu, 20)$$

### Estimation

#### Example 1:

- In an opinion poll, 1000 randomly selected voters are asked about their political sympathies.
- We have one observation x = 100 of the random variable  $X \sim \text{Bin}(1000, p)$ .
- Estimate  $p^* = x/1000 = 100/1000 = 0.1$ .
- Estimator  $p^*(X) = X/1000$ .



- $X = (X_1, ..., X_n)$
- Estimator  $\theta^*(X)$
- $\theta^* \theta = [E\{\theta^*(\mathbf{X})\} \theta] + [\theta^* E\{\theta^*(\mathbf{X})\}]$ = systematic error + random error

### Definition (7.3)

An estimate  $\theta^*$  is said to be *unbiased* (väntevärdesriktig) if it satisfies  $E\left\{\theta^*(\mathbf{X})\right\} = \theta$ 

i.e. it has no systematic error.



Example 2:

$$\mu^*(\mathbf{X}) = \overline{X} = \frac{X_1 + X_2 + X_3 + X_4 + X_5}{5} \sim N(\mu, 20)$$

Is  $\mu^*(\mathbf{x})$  an unbiased estimate of  $\mu$ ?

Example 1:

$$p^*(X) = \frac{X}{1000}$$

where  $X \sim \text{Bin}(1000, p)$ .

Is  $p^*(x)$  an unbiased estimate of p?

If we have more than one unbiased estimate, which is the best one?

### Definition (7.4)

If  $\theta_1^*$  and  $\theta_2^*$  are unbiased estimates of  $\theta$  and

$$V\left\{\theta_1^*(\mathbf{X})\right\} \le V\left\{\theta_2^*(\mathbf{X})\right\}$$

for all  $\theta$  with strict inequality for some  $\theta$ , we say that  $\theta_1^*$  is *more efficient* than  $\theta_2^*$ .

#### Example 2:

- $(x_1, x_2, x_3, x_4, x_5) = (200, 185, 210, 190, 190)$  is a random sample from  $X \sim N(\mu, 100)$ .
- Some possible estimates:

$$\overline{x} = \frac{x_1 + x_2 + x_3 + x_4 + x_5}{5} = \frac{200 + 185 + 210 + 190 + 190}{5} = 195,$$

$$\text{median}(x_1, x_2, x_3, x_4, x_5) = 190,$$

$$\text{midrange} = \frac{\min x_i + \max x_i}{2} = \frac{185 + 210}{2} = 197.5.$$

- It may be shown that all of these estimates are unbiased.
- Which one is most efficient?

#### Example 2:

- Suppose that  $X_1, X_2, ..., X_n$  are independent  $N(\mu, \sigma^2)$ , n = 5.
- It may be shown that
  - $V(\overline{X}) = \sigma^2/5 = 0.2\sigma^2$ , (most efficient!)
  - $V(\text{median}) \approx 0.29\sigma^2$ ,
  - $V(\text{midrange}) \approx 0.23\sigma^2$ .
- For n arbitrary,
  - $V(\overline{X}) = \sigma^2/n$ ,
  - $V(\text{median}) \approx (\pi/2)\sigma^2/n$ , n large,
  - $V(\text{midrange}) \approx (\pi^2/24)\sigma^2/(\ln n)$ , n large.
- For n = 5, the median (as a random variable), Y say, has density function

$$f(y) = 30f_X(y)F_X(y)^2\{1 - F_X(y)\}^2$$

and the minimum  $Y_1$  and maximum  $Y_2$  have the simultaneous density

$$f(y_1, y_2) = 20f_X(y_1)f_X(y_2)\{F_X(y_2) - F_X(y_1)\}^3.$$

#### Example 2':

- A researcher randomly selects four moose and observes the withers heights  $x_1, ..., x_4$ , assumed to be a random sample from  $N(\mu, 100)$ .
- Another researcher randomly selects five other moose and observes the withers heights  $y_1, ..., y_5$ , assumed to be a random sample from  $N(\mu, 100)$ .
- ullet Based on this, the following two estimates of  $\mu$  are proposed:

$$\begin{split} \mu_1^* &= \frac{\overline{x} + \overline{y}}{2} \\ \mu_2^* &= \frac{4 * \overline{x} + 5 * \overline{y}}{9} \end{split}$$

- Are the estimates unbiased?
- 2 If so, which one is most efficient?

### Example 3 (stratification):

We are interested in the proportion p of Swedish citizens that last year have traveled by plane in connection with work. To find out, we want to take a sample consisting of n=1000 people. Consider the following two ways to do this:

 Randomly draw 1000 people, of which x traveled by plane, and use the estimate

$$p_1^* = \frac{x}{n}.$$

 Randomly draw 500 men and 500 women, of which y and z, respectively, traveled by plane, and use the estimate

$$p_2^* = \frac{y+z}{n}.$$

- Are the estimates unbiased?
- 2 If so, which one is most efficient?



How can we assign a numerical value to the dispersion of an estimate?

### Definition (7.6)

The standard error (medelfelet) of the estimate  $\theta^*$  is an estimate of the standard deviation  $D\{\theta^*(\mathbf{X})\}$ . It is denoted by  $d\{\theta^*(\mathbf{x})\} = d(\theta^*)$ .

#### Example 1:

- We have one observation x = 100 of the random variable  $X \sim \text{Bin}(1000, p)$ .
- Estimate  $p^* = x/1000 = 100/1000 = 0.1$ .
- Calculate the standard error of this estimate.

Answer: 0.0094

#### Example 2:

- $(x_1, x_2, x_3, x_4, x_5) = (200, 185, 210, 190, 190)$  is a random sample from  $X \sim N(\mu, 100)$ .
- Estimate

$$\overline{x} = \frac{x_1 + x_2 + x_3 + x_4 + x_5}{5} = \frac{200 + 185 + 210 + 190 + 190}{5} = 195$$

Calculate the standard error of this estimate.

Answer: 4.47

The accuracy of an estimate should improve as the sample size increases.

### Definition (s.270)

The bias (väntevärdesfelet) for the estimate  $\theta^*$  is defined as

$$B(\theta^*) = E\left\{\theta^*(\mathbf{X})\right\} - \theta.$$

Observe: An unbiased estimate has bias zero.

### Definition (7.7)

If the bias  $B(\theta_n^*)$  tends to zero as  $n \to \infty$  for all  $\theta$ , the estimate  $\theta_n^*$  is said to be asymptotically unbiased.

Observe: An unbiased estimate is always asymptotically unbiased.



#### Example 4:

- Let  $x_1, ..., x_n$  be a random sample from  $N(\mu, \sigma^2)$ , where  $\mu$  is unknown. The object is to estimate  $\sigma^2$ .
- The estimate  $\sigma_n^{2^*} = \frac{1}{n} \sum_{i=1}^n (x_i \bar{x})^2$  is *not* unbiased, but it is asymptotically unbiased.
- The estimate  $s_n^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i \bar{x})^2$  is unbiased for  $\sigma^2$ . Why?
- Is  $s_n$  unbiased for  $\sigma$ ?

#### Definition (7.9)

The estimate  $\theta_n^*$  is said to be *consistent* for  $\theta$  if the corresponding estimator converges to  $\theta$  in probability for all  $\theta$ .

#### Definition (7.8)

The estimator  $\theta_n^*(\mathbf{X})$  converges to  $\theta$  in probability if for all  $\varepsilon > 0$ ,

$$P(|\theta_n^*(\mathbf{X}) - \theta| > \varepsilon) \to 0$$

as  $n \to \infty$  for all  $\theta$ .



### Theorem (Sats 7.2)

If the estimate  $\theta_n^*$  is asymptotically unbiased and  $V\{\theta_n^*(\mathbf{X})\}\to 0$  as  $n\to\infty$  for all  $\theta$ , then it is consistent.

#### Example 2":

- Let  $x_1, ..., x_n$  be a random sample from  $N(\mu, \sigma^2)$ , where  $\sigma^2$  is known.
- $\bullet$  Estimate  $\mu$  by

$$\mu_n^* = \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i.$$

- Show that the estimate is unbiased.
- Calculate the variance for the corresponding estimator.
- Show that the estimate is consistent.

# News of today

- Random sample
- Estimate (skattning)
- Estimator
- Unbiased (väntevärdesriktig)
- More efficient
- Standard error (medelfel)
- Asymptotically unbiased
- Consistent

Problem: 7.2.2