

1. For Brownian motion  $B = \{B_t\}_{t \geq 0}$  and in terms of the Brownian transition density  $p_t(x) = (2\pi t)^{-1/2} e^{-x^2/2t}$ ,  $t > 0$ ,  $x \in \mathbb{R}$ , the joint distribution of the pair  $(B_s, B_t)$ ,  $s < t$ , has the joint density function  $p_s(x)p_{t-s}(y-x) dx dy$ , and the conditional distribution of  $B_s$  given that  $B_t = b$ , has density

$$\frac{p_s(x)p_{t-s}(b-x)}{p_t(b)} = \frac{1}{\sqrt{2\pi s(t-s)/t}} \exp \left\{ -\frac{(x - sb/t)^2}{2s(t-s)/t} \right\}$$

Let  $Y = \{Y_t\}_{0 \leq t \leq 1}$  be the Brownian Bridge derived from  $B$  by putting  $Y_t = B_t$  conditioned on  $B_1 = 0$ . It follows that  $Y$  is Gaussian with density function

$$f_{Y_t}(x) = \frac{1}{\sqrt{2\pi t(1-t)}} \exp \left\{ -\frac{x^2}{2t(1-t)} \right\}.$$

In particular,  $E(Y_t) = 0$ ,  $\text{Var}(Y_t) = t(1-t)$ ,  $0 \leq t \leq 1$ . Similarly as above, the two-dimensional joint density of  $(Y_s, Y_t)$ ,  $0 < s < t < 1$ , is obtained as

$$\frac{p_s(x)p_{t-s}(y)p_{1-t}(0)}{p_1(0)} = \frac{1}{2\pi \sqrt{s(t-s)(1-t)}} \exp \left\{ -\frac{t(1-t)x^2 - 2s(1-t)xy + s(1-s)y^2}{2s(t-s)(1-t)} \right\}.$$

By comparing with the bivariate Gaussian distribution it is seen that the covariance and the correlation function of  $Y$  are given by

$$C(s, t) = s(1-t), \quad \rho(s, t) = \sqrt{\frac{s(1-t)}{t(1-s)}}.$$

It can be shown that  $\{Y_t\}_{0 \leq t \leq 1}$  has the same distribution as the process  $\{\tilde{Y}_t\}_{0 \leq t \leq 1}$  defined by  $\tilde{Y}_t = B_t - tB_1$ . This is frequently used as an alternative definition of the Brownian bridge, which yields additional methods for deriving the joint density, the covariance, etc.

2. We know that  $m_A = E(A_t) > 0$  is independent of  $t$  and that  $\text{Cov}(A_s, A_t)$  is a function  $C_A(t-s)$  which only depends on  $|t-s|$ . By the independence of  $A$  and  $\Phi$ ,

$$E(X_t) = E(A_t)E[\sin(\omega t + \Phi)] = \frac{E(A_t)}{2\pi} \int_0^{2\pi} \sin(\omega t + y) dy = \frac{E(A_t)}{2\pi} [-\cos(\omega t + y)]_0^{2\pi} = 0,$$

so that  $m(t) = 0$  and hence the mean is independent of  $t$ . The covariance function is therefore

$$C_X(s, t) = E[X_s X_t] = E[A_s A_t] \mathbb{E}[\sin(\omega s + \Phi) \sin(\omega t + \Phi)].$$

Here,  $E[A_s A_t] = C_A(t - s) + m^2$  and, using a suitable trigonometric formula,

$$\begin{aligned}\mathbb{E}[\sin(\omega s + \Phi) \sin(\omega t + \Phi)] &= \frac{1}{2\pi} \int_0^{2\pi} \sin(\omega s + y) \sin(\omega t + y) dy \\ &= \frac{1}{2\pi} \int_0^{2\pi} \frac{1}{2} (\cos(\omega|t - s|) - \cos(\omega(s + t) + 2y)) dy = \frac{1}{2} \cos \omega|t - s|.\end{aligned}$$

It follows that  $X$  is stationary, since

$$C_X(s, t) = \frac{1}{2}(C_A(t - s) + m^2) \cos \omega|t - s|,$$

which is a function that only depends on the difference of  $s$  and  $t$ .

3. Clearly,  $Y_n \in \mathcal{F}_n$  and  $E[|Y_n|] \leq 1$  for all  $n \geq 1$ , so  $Y$  is adapted and integrable. Moreover,

$$E[Y_n | \mathcal{F}_{n-1}] = (-1)^n E[\cos(\pi X_{n-1} + \pi Z_n) | \mathcal{F}_{n-1}].$$

By independence,

$$\begin{aligned}E[\cos(\pi X_{n-1} + \pi Z_n) | \mathcal{F}_{n-1}] &= \frac{1}{2} E[\cos(\pi X_{n-1} + \pi) | \mathcal{F}_{n-1}] + \frac{1}{2} E[\cos(\pi X_{n-1} - \pi) | \mathcal{F}_{n-1}] \\ &= \frac{1}{2} (\cos(\pi X_{n-1} + \pi) + \cos(\pi X_{n-1} - \pi)) \\ &= -\frac{1}{2} (\cos(\pi X_{n-1}) + \cos(\pi X_{n-1})),\end{aligned}$$

and therefore  $E[Y_n | \mathcal{F}_{n-1}] = Y_{n-1}$ , as required.