

# Inference 2, 2023, lecture 10

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# Today

## Chap. 5. Testing hypotheses

- Test problems
- P value
- Decision rules
- The Neyman-Pearson test and lemma

# Test problems

- Sample  $\mathbf{X} = (X_1, \dots, X_n)$ , statistical model  $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$ .
- Let  $\Theta = \Theta_0 \cup \Theta_1$  with  $\Theta_0 \cap \Theta_1 = \emptyset$ .
- Test  $H_0: \theta \in \Theta_0$  vs  $H_1: \theta \in \Theta_1$ .
- Two-sided test:  $H_0: \theta = \theta_0$ ,  $H_1: \theta \neq \theta_0$ .
- One-sided tests for  $\Theta \subseteq \mathcal{R}$ :
  - Test  $H_0: \theta \geq \theta_0$  vs  $H_1: \theta < \theta_0$
  - Test  $H_0: \theta \leq \theta_0$  vs  $H_1: \theta > \theta_0$
- Alternatively:
  - For  $\Theta = (-\infty, \theta_0]$ , test  $H_0: \theta = \theta_0$  vs  $H_1: \theta < \theta_0$ .
  - For  $\Theta = [\theta_0, \infty)$ , test  $H_0: \theta = \theta_0$  vs  $H_1: \theta > \theta_0$ .

# P value

- Base the hypothesis test on a suitable **test statistic**  $T(\mathbf{X})$ .
- Suppose that  $H_0$  is true.
- If the observed value  $T(\mathbf{x}) = t_{\text{obs}}$  is “improbable” (“extreme”), then we have evidence against  $H_0$ .

## Definition (5.1)

The **p value** corresponding to an observed  $T(\mathbf{x}) = t_{\text{obs}}$  is the probability of  $T(\mathbf{X})$  lying at or beyond  $t_{\text{obs}}$  in the “direction of the more extreme values” of the alternative, computed from the null distribution.

# P value

Example 1:

- Let  $\mathbf{X} = (X_1, \dots, X_n)$  where the  $X_i$  are independent uniform on  $[0, \theta]$ .
- We have  $n = 4$  and observations  $\mathbf{x} = (0.74, 1.99, 0.42, 1.08)$ .
- Let  $\Theta = [2, \infty)$ .
- Test  $H_0: \theta = 2$  vs  $H_1: \theta > 2$ .
- Take  $T(\mathbf{X}) = \max_i X_i$ .
- Compute the p value.

# P value

- For one-sided tests, i.e.
  - test  $H_0: \theta \geq \theta_0$  vs  $H_1: \theta < \theta_0$ ,
  - or test  $H_0: \theta \leq \theta_0$  vs  $H_1: \theta > \theta_0$ ,

the P value is given by either  $P_0^T(T \leq t_{\text{obs}})$  or  $P_0^T(T \geq t_{\text{obs}})$ , depending on the direction of the alternative and the test statistic  $T$ .

- For the two-sided test:  $H_0: \theta = \theta_0$ ,  $H_1: \theta \neq \theta_0$ , the p value is (most often) given by

$$2 \min\{P_0^T(T \leq t_{\text{obs}}), P_0^T(T \geq t_{\text{obs}})\}.$$

# P value

- Suppose that the test statistic  $T$  has a continuous distribution.
- Suppose the p value is given by  $P_0(T \leq t_{\text{obs}})$ , where  $P_0$  is the probability calculated under  $H_0$ .
- The p value may be regarded as a random variable which is uniformly distributed on  $[0, 1]$ . (Why?)

# Decision rules

We now shift focus from the parameter space  $\Theta$  to the sample space  $\mathcal{X}$ .

## Definition (5.2)

A (nonrandomized) **test** is a statistic  $\varphi : \mathcal{X} \rightarrow \{0, 1\}$  such that

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } \mathbf{x} \in C_1 \quad (\text{reject } H_0) \\ 0 & \text{if } \mathbf{x} \in C_0 \quad (\text{do not reject } H_0) \end{cases}$$

where  $\mathcal{X} = C_1 \cup C_0$ , with  $C_1 \cap C_0 = \emptyset$ .

$C_1$  is called the **critical region**.

Consider example 1. If we want to reject with probability 0.05 under  $H_0$ , which is the critical region?



# Decision rules

Example 2:

- Let  $\mathbf{X} = (X_1, \dots, X_n)$  where the  $X_i$  are independent discrete uniform on  $[1, 2, \dots, \theta]$ .
  - We have  $n = 4$  and observations  $\mathbf{x} = (52, 99, 35, 12)$ .
  - Let  $\Theta = [100, \infty)$ .
  - Consider the test  $H_0: \theta = 100$  vs  $H_1: \theta > 100$ .
- 1 Calculate the p value.
  - 2 If we want to reject with probability *at most* 0.05 under  $H_0$ , which is the critical region?

# Decision rules

## Definition (5.3)

A **randomized test**  $\varphi$  is a step function on  $C_1$ ,  $C_=\$  and  $C_0$  to  $[0, 1]$ , where  $\mathcal{X} = C_1 \cup C_=\cup C_0$  and  $C_1$ ,  $C_=\$  and  $C_0$  are mutually disjoint:

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } \mathbf{x} \in C_1 & (\text{reject } H_0) \\ \gamma & \text{if } \mathbf{x} \in C_= & (\text{reject } H_0 \text{ with probability } \gamma) \\ 0 & \text{if } \mathbf{x} \in C_0 & (\text{do not reject } H_0) \end{cases}$$

*Observe:* Randomized tests are only of theoretical interest! They are never performed in practice. What would be the main problem if they were?

Example 2':

What should  $C_=\$  and  $\gamma$  be in example 2 if we want to reject with probability 0.05 under  $H_0$ ?

# Decision rules

## Definition (5.4)

For the (randomized) test problem  $H_0: \theta \in \Theta_0$  vs  $H_1: \theta \in \Theta_1$  we define

- ① The **error of type I**: reject  $H_0$ , but  $\theta \in \Theta_0$ .
- ② The **error of type II**: do not reject  $H_0$ , but  $\theta \in \Theta_1$ .
- ③ The function defined on  $\Theta$  by

$$\pi(\theta) = E_{\theta}\{\varphi(\mathbf{X})\} = P_{\theta}(C_1) + \gamma P_{\theta}(C_2)$$

is called the **power function** of the test  $\varphi$ .

*Note:*

- The power function gives the probability of rejecting  $H_0$  when the parameter equals  $\theta$ .
- If  $\theta \in \Theta_1$ , this is one minus the probability of a type II error for this particular  $\theta$ .

# Decision rules

Calculate the power for

- 1  $\theta = 3$  in example 1.
- 2  $\theta = 150$  in example 2.

# The Neyman-Pearson test

- Consider testing a simple  $H_0: \theta = \theta_0$  vs a simple  $H_1: \theta = \theta_1$ .
- In other words, we test  $H_0: P_0$  vs  $H_1: P_1$ .

## Definition (5.5)

For testing a simple  $H_0$  vs a simple  $H_1$ ,  
the **size** of the type I error is  $\alpha = P_0(\text{reject } H_0)$ .

The size of the type II error is  $\beta = P_1(\text{do not reject } H_0)$ .

In example 1 with  $H_0: \theta = 2$ ,  $H_1: \theta = 3$ , calculate  $\alpha$  and  $\beta$  when the critical regions are

- 1  $C_1 = \{t_{\text{obs}} > 1.975\}$
- 2  $C_1 = \{t_{\text{obs}} > 1.9\}$

# The Neyman-Pearson test

The set of all possible  $\{\alpha(\varphi), \beta(\varphi)\}$ , where  $\varphi$  is a test, is called the set of  **$\alpha\beta$ -representations**.

## Theorem (5.1)

*The set of  $\alpha\beta$ -representations*

- ① *is convex,*
- ② *is included in the closed unit square,*
- ③ *includes the points  $(0, 1)$  and  $(1, 0)$ .*

# The Neyman-Pearson test

We want to find a test  $\varphi$  which is optimal in the sense that, given a pre-assigned **significance level**  $\alpha = E_0(\varphi)$ , the power  $1 - \beta = E_1(\varphi)$  is as large as possible.

## Definition (5.6)

A test  $\varphi^*$  is called **most powerful** (MP) of size  $\alpha$  if.f.  $E_0\{\varphi^*(\mathbf{X})\} = \alpha$  and  $E_1\{\varphi^*(\mathbf{X})\} \geq E_1\{\varphi(\mathbf{X})\}$  for all tests  $\varphi$  with  $E_0\{\varphi(\mathbf{X})\} \leq \alpha$ .

# The Neyman-Pearson test

## Definition (5.7)

For real numbers  $k \geq 0$  and  $\gamma \in [0, 1]$ , tests of the form

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } p_0(\mathbf{x}) < kp_1(\mathbf{x}), \\ \gamma & \text{if } p_0(\mathbf{x}) = kp_1(\mathbf{x}), \\ 0 & \text{if } p_0(\mathbf{x}) > kp_1(\mathbf{x}) \end{cases}$$

are called **Neyman-Pearson (NP) tests**.

*Note:* For non randomized tests, this corresponds to rejecting if.f. **the likelihood ratio**

$$\Lambda^*(\mathbf{x}) = \frac{p_0(\mathbf{x})}{p_1(\mathbf{x})} < k.$$

The NP test is also called the likelihood ratio (LR) test.



# The Neyman-Pearson test

The Neyman-Pearson lemma tells us that NP tests, given the significance level (size), are optimal in terms of power.

## Theorem (5.2)

### The Neyman-Pearson lemma.

*For all  $\alpha \in [0, 1]$  there exist  $\gamma = \gamma(\alpha)$  and  $k = k(\alpha)$  such that the NP test with  $\gamma(\alpha)$ ,  $k(\alpha)$  has size  $\alpha$  and is most powerful of size  $\alpha$ .*

*Especially, for  $\alpha > 0$ ,*

- ① *If there exists a  $k = k(\alpha)$  such that  $P_0\{p_0(\mathbf{X}) < kp_1(\mathbf{X})\} = \alpha$ , then  $\gamma = \gamma(\alpha) = 0$ . (The non randomized case.)*
- ② *Otherwise take  $k = k(\alpha)$  with  $P_0\{p_0(\mathbf{X}) < kp_1(\mathbf{X})\} < \alpha < P_0\{p_0(\mathbf{X}) \leq kp_1(\mathbf{X})\}$  and*

$$\gamma = \gamma(\alpha) = \frac{\alpha - P_0\{p_0(\mathbf{X}) < kp_1(\mathbf{X})\}}{P_0\{p_0(\mathbf{X}) = kp_1(\mathbf{X})\}}.$$

# News of today

- Tests, one-sided and two-sided
- P value
- In the continuous case, the P value can be seen as a uniformly distributed random variable.
- Test (function)
  - non randomized (continuous case)
  - randomized (discrete case)
- Critical region
- Errors of type I and II
- Power function
- Testing a one point  $H_0$  vs a one point  $H_1$ :
  - The set of  $\alpha\beta$  representations.
  - Most powerful test
  - Neyman-Pearson (likelihood ratio) tests
  - ...are optimal in terms of power (**Neyman-Pearson lemma**)