Analysis of Time Series, L14

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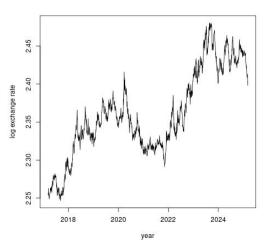
Today

5.2: Unit root testing

- DF test
- ADF test.
- With deterministic terms
- KPSS test (not in book)
- Application
- Menti

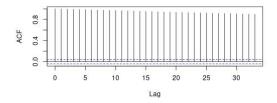


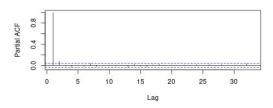
Log exchange rate, Skr/Euro, March 6 2017-March 5 2025:

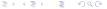




Log Skr/Euro, ACF (slowly decaying) and PACF (cutting off).







```
> y1=log(y)
> arima(y1,order=c(1,0,0))
Call:
arima(x = y1, order = c(1, 0, 0))
Coefficients:
         ar1
              intercept
      0.9979
                 2.3610
s.e. 0.0014
                 0.0347
sigma<sup>2</sup> estimated as 1.608e-05: log likelihood = 8558.44,
aic = -17110.87
```

Estimated AR parameter very close to 1.

```
> arima(y1,order=c(2,0,0))
Call:
```

Coefficients:

ar1 ar2 intercept 0.8848 0.1134 2.3610

s.e. 0.0217 0.0217 0.0389

arima(x = y1, order = c(2, 0, 0))

sigma^2 estimated as 1.587e-05: log likelihood = 8571.93, aic = -17135.86

Sum of estimated AR parameters very close to 1.



AR(1) process

$$x_t = \phi x_{t-1} + w_t, \quad t = 1, 2, ..., n.$$

- Causal if.f. $|\phi| < 1$.
- Stationary for a suitable choice of distribution for x_0 ($var(x_0) = \frac{\sigma_w^2}{1-\phi^2}$) if.f. $|\phi| < 1$.
- Test H_0 : $\phi = 1$ vs H_1 : $|\phi| < 1$.
- A natural test statistic is $\hat{\phi} 1$.
- The Dickey-Fuller (DF) test.



• Under H_0 : $\phi = 1$, it follows that (why?)

$$\hat{\phi} - 1 = \frac{\sum_{t=1}^{n} x_t x_{t-1}}{\sum_{t=1}^{n} x_{t-1}^2} - 1 = \frac{\sum_{t=1}^{n} w_t x_{t-1}}{\sum_{t=1}^{n} x_{t-1}^2}.$$

• Moreover (why?), assuming $x_0 = 0$,

$$\hat{\phi} - 1 = \frac{x_n^2 - \sum_{t=1}^n w_t^2}{2\sum_{t=1}^n x_{t-1}^2}.$$



Definition (5.1)

A continuous time process $\{W(t); t \geq 0\}$ is called a *standard Brownian motion* if it satisfies

- (i) W(0) = 0.
- (ii) For any $0 \le t_1 < t_2 < ... < t_n$ and integer n, $W(t_2) W(t_1)$, $W(t_3) W(t_2)$,..., $W(t_n) W(t_{n-1})$ are independent.
- (iii) $W(t + \Delta t) W(t) \sim N(0, \Delta t)$ for $\Delta t > 0$.

One may show that for a white noise process $\{w_t\}$, as $n \to \infty$,

$$\frac{1}{\sigma_w \sqrt{n}} \sum_{j=1}^{\lfloor nt \rfloor} w_j \xrightarrow{\mathcal{L}} W(t)$$

where |a| is the integer part of a.



• As $n \to \infty$,

$$\frac{1}{\sigma_w \sqrt{n}} \sum_{j=1}^{\lfloor nt \rfloor} w_j \xrightarrow{\mathcal{L}} W(t)$$

It follows that (why?)

$$n(\hat{\phi}-1) = n \frac{x_n^2 - \sum_{t=1}^n w_t^2}{2\sum_{t=1}^n x_{t-1}^2} \xrightarrow{\mathcal{L}} \frac{W(1)^2 - 1}{2\int_0^1 W(t)^2 dt}.$$

• No "standard" distribution!



An alternative is the t test

$$\hat{t} = \frac{\phi - 1}{\sqrt{s^2 / \sum_{t=2}^{n} x_{t-1}^2}},$$

where

$$s^2 = \frac{1}{n-1} \sum_{t=2}^{n} (x_t - \hat{\phi} x_{t-1})^2.$$

• One may show that, as $n \to \infty$,

$$\hat{t} \xrightarrow{\mathcal{L}} \frac{W(1)^2 - 1}{2\sqrt{\int W(t)^2 dt}}.$$



AR(1):
$$x_t = \phi x_{t-1} + w_t \Leftrightarrow \nabla x_t = \gamma x_{t-1} + w_t, \ \gamma = \phi - 1.$$

Extension to AR(2):

•

$$x_t = \phi_1 x_{t-1} + \phi_2 x_{t-2} + w_t$$

Equivalent to (why?)

$$\nabla x_t = \gamma x_{t-1} + \psi_1 \nabla x_{t-1} + w_t,$$

where

$$\gamma = \phi_1 + \phi_2 - 1, \quad \psi_1 = -\phi_2.$$



Extension to AR(p):

•

$$x_t = \phi_1 x_{t-1} + \dots + \phi_p x_{t-p} + w_t$$

Equivalent to

$$\nabla x_{t} = \gamma x_{t-1} + \psi_{1} \nabla x_{t-1} + \dots + \psi_{p-1} \nabla x_{t-(p-1)} + w_{t}$$

where

$$\gamma = \sum_{j=1}^{p} \phi_j - 1, \quad \psi_j = -\sum_{i=j+1}^{p} \phi_i.$$

- The Augmented Dickey-Fuller (ADF) test: Test H_0 : $\gamma = 0$ vs H_1 : $\gamma < 0$.
- Modified test statistics, but the same limit distributions as before.



With deterministic terms

Incorporating deterministic terms:

•
$$x_t = \beta_0 + \phi x_{t-1} + w_t$$

• Test
$$H_0$$
: $(\beta_0, \phi) = (0, 1)$ vs H_1 : $\neg H_0$.

•
$$x_t = \beta_0 + \beta_1 t + \phi x_{t-1} + w_t$$

- Test H_0 : $(\beta_1, \phi) = (0, 1)$ vs H_1 : $\neg H_0$.
- Modified test statistics, other limit distributions.

KPSS test

- Test H_0 : stationarity vs H_1 : non stationarity.
- Model

$$x_t = r_t + \varepsilon_t,$$

$$r_t = r_{t-1} + u_t,$$

where $\{\varepsilon_t\}$ and $\{u_t\}$ are independent white noise sequences, t=1,2,...,n.

- Test H_0 : $var(u_t) = \sigma_u^2 = 0$ vs H_1 : $\sigma_u^2 > 0$.
- Under H_0 , $x_t = r_0 + \varepsilon_t$ is stationary.
- Let $S_t = \sum_{j=1}^t e_j$, $e_j = (x_j \bar{x})$ and estimate its variance by

$$s^{2}(I) = \frac{1}{n} \sum_{t=1}^{n} e_{t}^{2} + \frac{2}{n} \sum_{s=1}^{I} \left(1 - \frac{s}{I+1}\right) \sum_{t=s+1}^{n} e_{t} e_{t-s}.$$



KPSS test

• Let $S_t = \sum_{j=1}^t e_j$, $e_j = (x_j - \bar{x})$ and estimate its variance by

$$s^{2}(I) = \frac{1}{n} \sum_{t=1}^{n} e_{t}^{2} + \frac{2}{n} \sum_{s=1}^{I} \left(1 - \frac{s}{I+1} \right) \sum_{t=s+1}^{n} e_{t} e_{t-s}.$$

Test statistic

$$\hat{\eta} = \frac{\sum_{t=1}^n S_t^2}{n^2 s^2(I)}.$$

• As $n, l \to \infty$ such that $l/\sqrt{n} \to 0$,

$$\hat{\eta} \stackrel{d}{\rightarrow} \int_0^1 \left\{ W(r) - rW(1) \right\}^2 dr,$$

where W(r) is the standard Brownian motion.



Unit root tests for log Skr/Euro in R:

```
> v1 = log(v)
> library(tseries)
> adf.test(y1)
Augmented Dickey-Fuller Test
data: y1
Dickey-Fuller = -2.6273, Lag order = 12, p-value = 0.3128
alternative hypothesis: stationary
> kpss.test(v1)
KPSS Test for Level Stationarity
data: y1
KPSS Level = 15.517, Truncation lag parameter = 8, p-value = 0.01
Warning message:
In kpss.test(y1): p-value smaller than printed p-value
```

The adf test does not reject non stationarity and the KPSS test rejects stationarity.

News of today

- Testing for non stationarity (unit root)
- Test statistics
- Limit distributions (non standard)
- Extensions:
 - Allowing for autocorrelation (ADF test)
 - Allowing for deterministic terms
- Testing the null of stationarity (KPSS test)