Supplementary Methods

A novel analytical framework to quantify co-gradient and counter-gradient variation M.A. Albecker, G.C. Trussell, and K.E. Lotterhos

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Supplemental Methods. 1. Proof for Cov_{GE}

The following equation is the standard approach to measure the covariance (Cov) between a pair of variables (here shown as x and y).

(1)
$$Sample Cov(x,y) = \frac{\sum (x - \overline{x})(y - \overline{y})}{N - 1}$$

(2)
$$Population Cov(x,y) = \frac{\sum (x - \mu_x)(y - \mu_y)}{N}$$

In which N is the total number of samples, \bar{x} is the sample mean of x, μ_x is the population mean of x, while \bar{y} is the sample mean of y and μ_y is the population mean of y.

The standard Cov(X,Y) equation cannot be directly applied to calculating Cov(G,E), because in Cov(G,E) we do not want to measure the covariance among every combination of genotypic and environmental effects, but only among the pairs of genotypic and environmental effects sourced from the same native environment. Therefore, we added the indicator variable (I) to ensure that only those data in which genotypes (i) are correctly matched to the environments (j) to which they were sourced (or are native to) are used to estimate Covariance. $I_{ij} = 1$ when genotype is correctly matched with the environment to which it is native, and $I_{ij} = 0$ otherwise (See Fig. 1 in the main text). In keeping, we take the covariance between genotypic and environmental effects on phenotypes, across all genotype-environment pairs $(n_{gen} = n_{env})$. This gives:

(3)
$$Population Cov(G, E) = \frac{1}{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (I_{ij})} (\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (\mu_i - \mu)(\mu_j - \mu)(I_{ij}))$$

In this equation, μ_i is the population mean phenotype of the *i*th genotype, and μ_j is the population mean phenotype of the *j*th environment, and μ is the overall population mean phenotype.

To allow for comparisons across taxa, study systems, and environmental gradients, we sought to standardize Cov(G, E). The typical approach to standardize covariance is correlation, which standardizes by the product of the standard deviation of the x and y variable $(\sigma_x^*\sigma_y)$. Here, we standardize by the variance in group means (instead of variance in all the data) because Cov(G,E) is based on covariance between the group means. We use σ_{μ_i} to represent the population standard deviation in the group means for genotypic effects, and σ_{μ_j} to represent the standard deviation in the group means for the environmental effects.

However, the group means for environmental effects and the group means for genotypic effects are not independent of each other, because they are both calculated from the phenotypic means.

For example, the following standardization produces inaccurate results:

(4)
$$Cor(G, E) = \frac{1}{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (I_{ij})} \left(\frac{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (\mu_i - \mu)(\mu_j - \mu)(I_{ij})}{(\sigma_{\mu_i} * \sigma_{\mu_j})} \right)$$

Equation 4 produces inaccurate results in simulations (e.g., does not produce Cov(G,E) estimates between -1 and 1) because the standardization by $\sigma_{\mu_i} * \sigma_{\mu_j}$ assumes that G and E are independent, but in this case, they are not. Thus, we multiply this equation by a correction term k:

(5)
$$\left(\frac{1}{\sum_{i=1}^{n_{gen}} \sum_{i=1}^{n_{gen}} (I_{ij})} \frac{\sum_{j=1}^{n_{env}} (\mu_i - \mu) (\mu_j - \mu) (I_{ij})}{(\sigma_{\mu_i} * \sigma_{\mu_j})} \right) * k$$

When the correction factor k is equal to the following, the result is constrained between -1 and 1:

(6)
$$k = \max(\frac{\sigma_{\mu_i}}{\sigma_{\mu_j}}, \frac{\sigma_{\mu_j}}{\sigma_{\mu_i}})$$

For example, if $\sigma_{\mu_i} > \sigma_{\mu_j}$ and we apply $\frac{\sigma_{\mu_j}}{\sigma_{\mu_i}}$ to the above equation, the σ_{μ_j} terms cancel each other and the correction becomes:

(7)
$$\sigma_{\mu_i} * \sigma_{\mu_i} = \sigma_{\mu_i}^2$$

Which when applied to the correlation equation, can be written as:

(8)
$$\frac{1}{\sum_{i=1}^{n_{gen}} \sum_{i=1}^{n_{gen}} (I_{ij})} \left(\frac{\sum_{j=1}^{n_{env}} (\mu_i - \mu)(\mu_j - \mu)(I_{ij})}{\sigma_{\mu_i}^2} \right)$$

Thus, the Cov(G,E) is standardized by the variance in group means for genotypic effects or the variance in group means for environmental effects, whichever is greater. This standardization preserves the relationship between the genotypic and environmental effects in the data, while constraining Cov(G,E) to be between -1 and 1. Thus, the population and estimated measures for standardized covariance (which we call Cov_{GE}) are as follows:

(9)
$$Population Cov_{GE} = \frac{1}{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (I_{ij})} \left(\frac{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (\mu_i - \mu)(\mu_j - \mu)(I_{ij})}{max(\sigma_{\mu_i}^2, \sigma_{\mu_j}^2)} \right)$$

(10)
$$Sample \, \hat{Cov}_{GE} = \frac{1}{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (I_{ij}) - 1} \left(\frac{\left(\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (\bar{y}_i - \bar{y})(\bar{y}_j - \bar{y})I_{ij}\right)}{max(s_{\bar{y}_i}^2, s_{\bar{y}_j}^2)} \right)$$

For the sample estimate, \bar{y}_i is the estimated marginal mean phenotype for ith genotype, \bar{y}_j is the estimated marginal mean phenotype for jth environment, and \bar{y} is the estimated overall mean phenotype. $s_{\bar{y}_i}^2$ is the variance in the group mean phenotype for genotypic effects, while $s_{\bar{y}_j}^2$ is the variance in the group mean phenotype for environmental effects.

Supplemental Methods. 2. POPULATION EQUATIONS

We used simulated data to measure sampling estimates of \hat{Cov}_{GE} and $\bar{\Delta}_{GxE}$, which are shown in the main document. However, to validate these measures, we also calculated population parameters based on observed \hat{Cov}_{GE} when $\epsilon_{ijk} = 0$ (Eqn. 4 in main text) and used the following equations to measure population parameters of \hat{Cov}_{GE} and $\bar{\Delta}_{GxE}$. See Supplemental Figure S2 below for simulation results showing agreement between sampling estimates and population parameter.

(11)
$$Population Cov_{GE} = \frac{1}{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (I_{ij})} \left(\frac{\sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} (\mu_i - \mu)(\mu_j - \mu)(I_{ij})}{max(\sigma_{\mu_i}^2, \sigma_{\mu_j}^2)} \right)$$

In this equation, n_{gen} is the number of genotypes and n_{env} is the number of environments. μ_i is the mean phenotype for environment j, and μ is the overall mean phenotype. I_{ij} is the indicator variable which equals one when genotype i is correctly matched with the environment j to which it is native, and zero otherwise. Standard deviation of genotypic group means is indicated by σ_{μ_i} and standard deviation of environmental group means is σ_{μ_i} .

(12)
$$Population \,\bar{\Delta}_{GxE} = \frac{1}{(n_{gen})(n_{env})} \sum_{i=1}^{n_{gen}} \sum_{j=1}^{n_{env}} |\mu_{ij} - \mu_i - \mu_j + \mu|$$

In this equation, n_{gen} is the number of genotypes and n_{env} is the number of environments. μ_{ij} is the mean phenotype for genotype i in environment j, μ_i is the mean phenotype for each genotype, μ_j is the mean phenotype for each environment, and μ is the overall mean phenotype. There is no indicator variable, as all combinations of genotype and environment are used to calculate $\bar{\Delta}_{GxE}$.

Supplemental Methods. 3. Imbalanced Designs

For common garden experimental designs, there may be imbalance where there are unequal sample sizes or unequal numbers of genotypes for each native environment. Estimated marginal means reduce bias that may arise from unequal sample sizes within groups, thus do not affect \hat{Cov}_{GE} . However, because environment and genotype are treated as factors in the ANOVA to extract estimated marginal means, imbalanced numbers of genotypes (example shown in Supplemental Figure S7B below) introduce bias in the environmental means (\bar{y}_j) or μ and overall mean phenotype (\bar{y}) or μ (See B in the table below) In fact, both studies analyzed in the main text have imbalanced numbers of genotypes. To correct for the bias, a second ANOVA model with native environment substituted for genotype can then be used to extract estimated marginal means to calculate unbiased (\bar{y}_j) or μ and (\bar{y}) or μ . The original model should be used to generate estimated marginal means for (\bar{y}) or μ . As shown in column C in the table below, this approach returns μ and μ to the values in balanced designs (A).

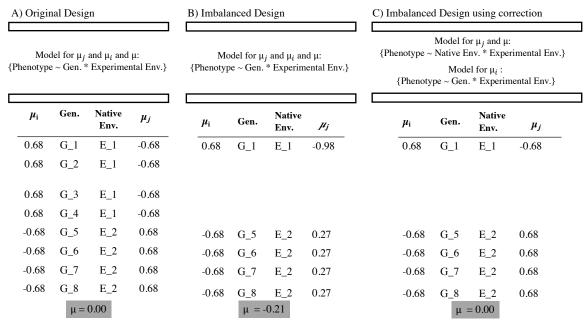


TABLE 1. (A) shows example data for a strong negative Cov_{GE} pattern for a Common-garden design. When the design is balanced, four genotypes native to two different environments were measured across two experimental environments (Supplemental Figure S7 below for depiction of corresponding sample data, Fig. S7A shows data for panels A in the above table, Fig. S7B shows data for panels B and C in above table). When the design is balanced, the mean phenotypes for genotype (μ_i) , environment (μ_j) , and overall (μ) are calculated by extracting estimated marginal means from a single model. When the design is unbalanced (as shown in B), continuing to use the same model to generate estimated marginal means produces bias in environmental means (μ_j) , and overall means (μ) . As shown in (C), the bias is corrected by conducting a secondary ANOVA with native environment replacing genotype to calculate environmental means (μ_j) , and overall means (μ) . Genotypic means (μ) are calculated using the original ANOVA model.

Supplemental Methods. 4. VARIANCE PARTITIONING

In the main manuscript, we focus on estimating the effect size of Cov_{GE} and $\bar{\Delta}_{GxE}$ and their significance. The effect size provides information about the strength of the pattern, which is a distinct type of information from the percent of variation in the phenotype explained by the following components: (i) genetic effects on phenotype, (ii) environmental effects on phenotypes, (iii) genetic x environment interactions, (iv) covariance between genetic and environmental effects on phenotypes, and (iv) residual error. Both Falconer (1989) and Conover and Schultz (1995) have previously discussed Cov_{GE} in this more traditional sense as the percent of variation in the phenotype explained by different variance components:

$$(13) V_P = V_G + V_E + V_{GxE} + xCov_{GE}$$

In a 2x2 reciprocal transplant or common garden design with 2 genotypes in 2 environments, x = 2. This factor of x = 2 does not extend to more complex designs, as explained below.

Here, we show how to extend sums of squares (SS) calculations from a traditional analysis of variance to incorporate $SS_{Cov_{GE}}$, which can in turn be used to understand the percent of variation in phenotypes explained by different components. These calculations assume a fully factorial reciprocal transplant design. We do not advocate that these SS be used to test the significance of the variance components with a traditional F-test, because the presence of Cov_{GE} likely violates the assumption of independence among samples and complicates calculations of degrees of freedom (violations that are addressed in the main manuscript by using bootstrap and permutation for hypothesis testing). It is, however, useful to compare the percent of variation explained by different components to their effect sizes, because it furthers understanding of the relative influence of genetic differentiation and plasticity on the evolved patterns in the population.

In a reciprocal transplant experiment, there are g genotypes transplanted into e environmental patches, for a total of $g * e = n_{ge}$ genotype-environment combinations. In a fully factorial reciprocal transplant experiment, g = e, g is the number of levels of genotypes from i = 1, 2...g, and e is the number of levels of environments from j = 1, 2...e.

Assuming the equal sample sizes r (k = 1, 2, ... r) within each genotype-environment combination, the following sums of squares can be estimated as:

(14)
$$V_G = SS_G = re \sum_{i=1}^{g} (\bar{y}_i - \bar{y})^2$$

(15)
$$V_E = SS_E = rg \sum_{j=1}^{e} (\bar{y}_j - \bar{y})^2$$

(16)
$$V_{GxE} = SS_{GE} = r \sum_{i=1}^{g} \sum_{j=1}^{e} (y_{ij} - y_{ij} - y_{ij} + \bar{y})^{2}$$

(17)
$$V_{Cov_{GE}} = SS_{Cov_{GE}} = |xr \sum_{i=1}^{g} \sum_{j=1}^{e} (\bar{y}_i - \bar{y})(\bar{y}_j - \bar{y})I_{ij}|$$

(18)
$$V_{error} = SS_{Error} = \sum_{i=1}^{g} \sum_{j=1}^{e} \sum_{k=1}^{r} (y_{ijk} - \bar{y}_{ij})^2$$

where $x = \frac{ge}{(\sum_{i=1}^g \sum_{j=1}^e I_{ij})}$, and I_{ij} is an indicator variable that is 1 when the genotype i originated from environment j and 0 otherwise. In a 2x2 reciprocal transplant design, g = 2, e = 2, and $\sum_{i=1}^g \sum_{j=1}^e I_{ij} = 2$, so x = 2 as is assumed in Eq. S1. However in a 4x4 reciprocal transplant design, g = 4, e = 4, and $\sum_{i=1}^g \sum_{j=1}^e I_{ij} = 4$, so x = 4. The factor x ensures that the $SS_{Cov_{GE}}$ scales appropriately with the SS of the other components with the size of the experiment. Finally, since Cov_{GE} can be negative under countergradient scenarios, we take the absolute value for partitioning variance.

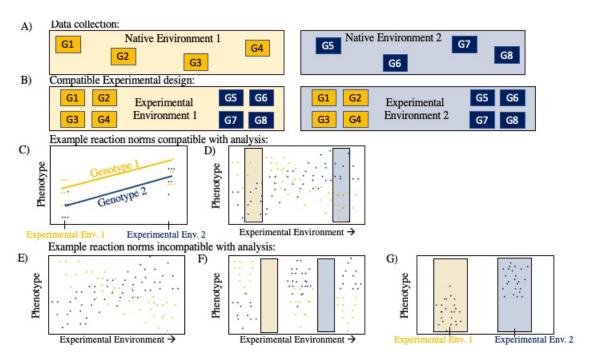
The percent of variation explained by each component (Comp) can then be estimated as

(19)
$$\eta_{Comp}^2 = \frac{SS_{Comp}}{SS_G + SS_E + SS_{GxE} + SS_{CovGE} + SS_{Error}}$$

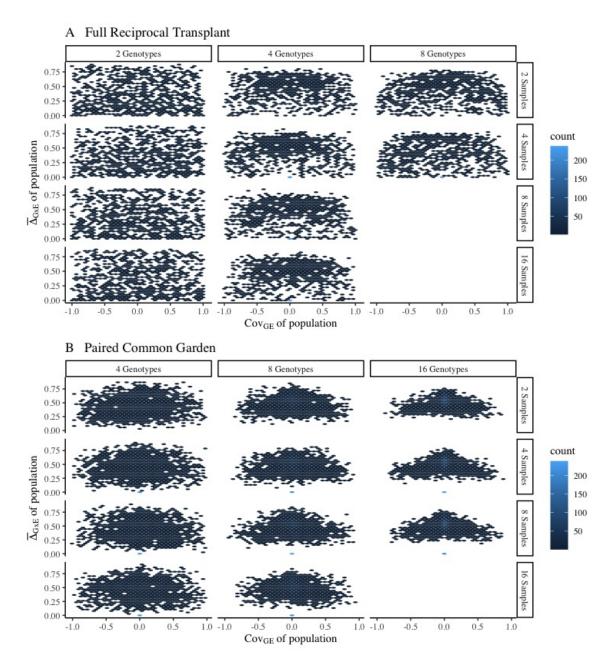
We recognize that this is a crude approach. However, it provides a reasonable way to compare to the percent of variation explained by different components to their effect size estimates in the main text. For example, when $|Cov_{GE}|$ is maximized, equal amounts of non-residual variance are explained by V_G , V_E , and $V_{Cov_{GE}}$ for the fully factorial reciprocal transplant experiment with an arbitrary number of populations (Supplementary. Fig 6a). For this reason, when residual variation is minimized the maximum percent of phenotypic variance explained by $V_{Cov_{GE}}$ is rarely greater than 1/3 of the total phenotypic variance (Supplementary Fig. S5, Fig.S6a).

The comparison of effect sizes and variance components is also useful for understanding the relative influences of V_G and V_E on a particular value of $|Cov_{GE}|$; for instance, similarly intermediate values of $|Cov_{GE}|$ can be driven by higher V_E and lower V_G (Supplementary Fig. 6B), or lower V_E and higher V_G (Supplementary. Fig. S6c).

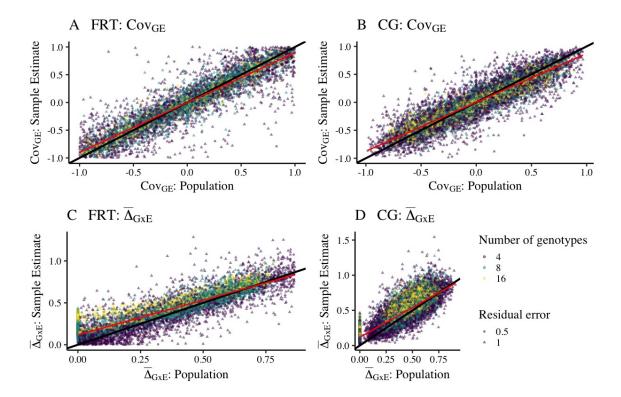
Supplemental Figures. FIGURES



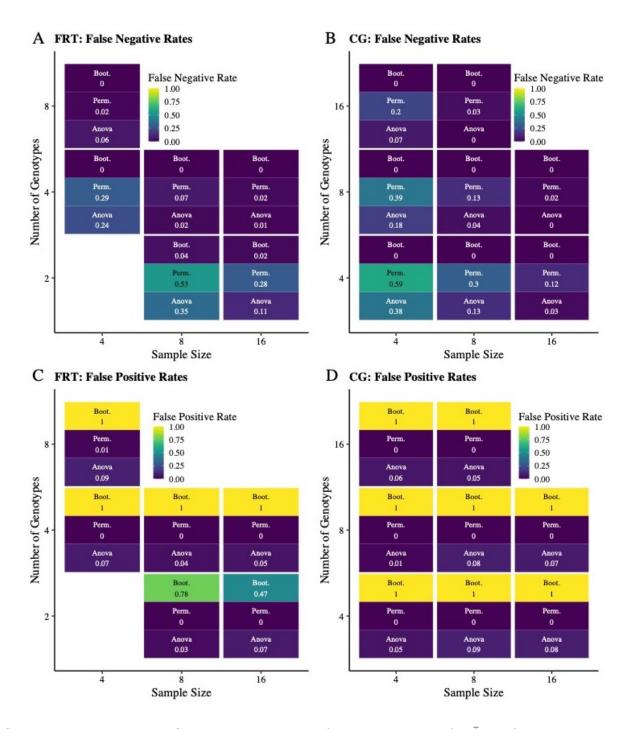
SUPPLEMENTARY FIGURE S1. Examples of experimental designs that are compatible or incompatible with Cov_{GE} . Plot A shows how phenotypic data are collected. Each box represents a different environment (orange, blue). Four genotypes are shown within each environment, with boxes representing the genotypes from which individuals (samples) are collected. Plot B shows an example of a fully factorial experimental design in which genotypes native to each environment are exposed to both environmental conditions. Plots C and D indicate phenotypic data that may result from the experimental design in panel B that are compatible with the analysis presented herein. Plot C shows categorical phenotypic data collected in either Environment 1 (orange), or Environment 2 (blue). Plot D shows phenotypic collected from an experimental design that used a continuous environmental gradient. However, to be analyzed, continuous phenotypic data would need to be binned. Note that in panel D, there are uneven numbers of samples in each bin. This is okay, because estimated marginal means reduce bias from unequal sample sizes. Plots E, F, and G show designs that are not incompatible with the Cov_{GE} equation. Panel E shows phenotypic data along a continuous axis but there is no information on which environment each genotype is native to. Because our analysis requires that information on the native environment of each genotype is known, this design would remain incompatible even if environment was categorical (like in plot C). Plot F demonstrates phenotypic data in which the native environment is known, but was not included as a treatment in the experiment. Indeed, not only is information about the native environment required, but it also must be included as a treatment. Plot G shows an instance where a single genotype is tested across multiple environments. This design is incompatible because the Cov_{GE} calculation requires at least 2 genotypes from at least 2 environments (but note that 2x2 designs are often underpowered - see Fig. 4 in the main text).



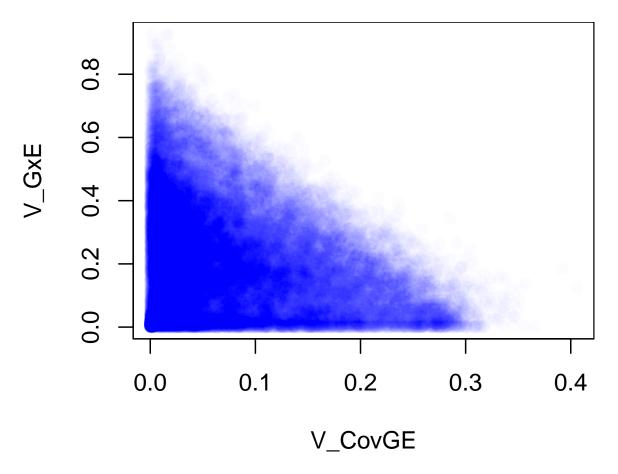
SUPPLEMENTARY FIGURE S2. Coverage of parameter space of Cov_{GE} and $\bar{\Delta}_{GxE}$ for full reciprocal transplant (A) and paired common garden designs (B). Hexagons are colored according to the density of observations in each bin.



SUPPLEMENTARY FIGURE S3. Agreement between population measures and sample estimates of Cov_{GE} and $\bar{\Delta}_{GxE}$ for Full Reciprocal Transplant (A, C) and paired Common Garden designs (B, D). The black line falls along a 1:1 line while the red line reflects the pattern of the data. Point colors indicate the number of genotypes, while point shapes indicate the level of residual variation. As expected, sample estimates deviate more from the population measure in situations with low sample sizes and higher residual variation.

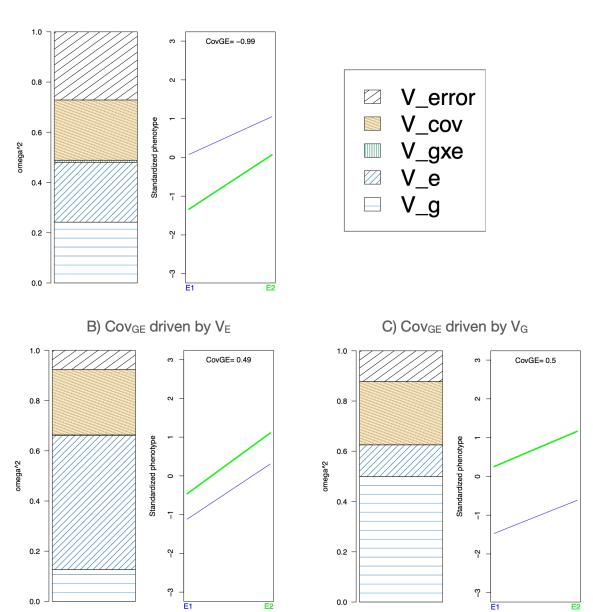


SUPPLEMENTARY FIGURE S4. Heat map showing false negative rates for $\bar{\Delta}_{GxE}$ for Full Reciprocal Transplant (A) and paired Common Garden designs (B) and false positive rates for full reciprocal transplant (C) and paired Common Garden designs (D). Tiles are split to show rates for bootstrapping (top), permutation (middle) and ANOVA (bottom) approaches. This figure complements Fig. 4 in the main text and is meant to demonstrate that bootstrapping produces unreliable confidence intervals for $\bar{\Delta}_{GxE}$.

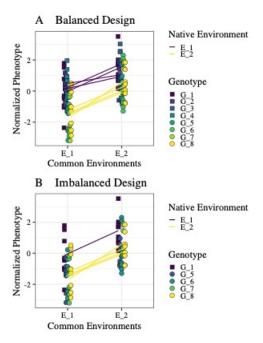


Supplementary Figure S5. Comparison of the percent of variance explained by V_{GxE} and $V_{Cov_{GE}}$ across all simulations.

A) Max abs(Cov_{GE})



SUPPLEMENTARY FIGURE S6. Comparison of the percent of variance explained by $V_{Cov_{GE}}$ with the effect size of Cov_{GE} for three example simulations from a 2 genotype x 2 environment reciprocal transplant design. In each subpanel, the stacked barplot on the left of each panel shows the proportion of variance explained by each component corresponding to the legend (See Supplemental Materials 4), while the plot on the right shows the mean standardized reaction norms for each genotype (the effect size of Cov_{GE} is shown at the top). A) When $|Cov_{GE}| = 1$ (i.e., maximal Cov_{GE}), equal proportions of phenotypic variance is explained by V_G , V_E , and $V_{Cov_{GE}}$. B) An intermediate effect size of $Cov_{GE} = 0.5$ arising from larger V_E compared to V_G . C) An intermediate effect size of $Cov_{GE} = 0.5$ arising from larger V_G compared to V_E .



SUPPLEMENTARY FIGURE S7. (A) An example of normalized phenotypic data for a balanced common garden design, meaning that both native environments have equal numbers of genotypes (in this case, 4 genotypes each). Points show phenotypic data colored according to genotype, while reaction norms are colored according to the environment to which the different genotypes are native. Data point shapes correspond to the native environment as well. (B) Normalized phenotypic for an imbalanced design in which there are 4 genotypes native to Environment 2, but only 1 genotype native to Environment 1. Imbalanced designs can generate bias in Cov_{GE} unless corrected. See Supplemental Methods 3 for details on the correction.