# Analyzing Flow-Cytometry Count Data with Regression Mixtures

Amit Meir University of Washington

Joint work with **Raphael Gottardo** and **Greg Finak** *Hutchinson Cancer Research Center* 

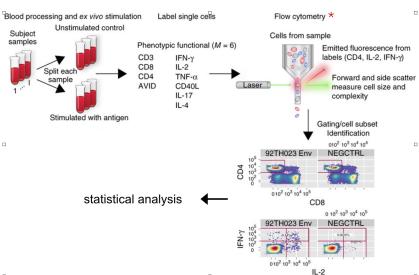
February 23, 2017

#### **Outline**

- Introduction to Flow-Cytometry
- Motivation
  - Existing methods
  - Why a regression model?
- Models:
  - A Marginal Model
  - A Joint HMRF model
- Data analysis
  - RV144 HIV Vaccine Trial
  - Controlled Human Malaria Infection Study.
- Computation (time permitting)



#### The RV144 HIV Vaccine Trial



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PTID	Subset	stim	count	parentcount
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P1003	CD154	${\tt nonstim}$	31	28099
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P1003	IFNg,IL2,CD154	stim	0	23524
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0

0

stim

23524

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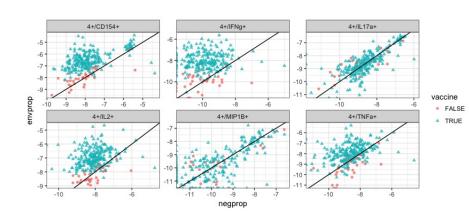
P1003

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#### The RV144 HIV Vaccine Trial

- 262 Subjects
  - 226 Cases
  - 36 Controls
- 2 Types of stimulus
  - HIV antigen
  - Negative control
- 6 types of cytokines.

# Marginal Counts for RV144



#### Motivation: COMPASS

# ANALYSIS



# COMPASS identifies T-cell subsets correlated with clinical outcomes

 $\label{limited} Lin^1, Greg Finak^1, Kevin Ushey^1, Chetan Seshadri^2, Thomas R Hawn^2, Nicole Frahm^1, Thomas J Scriba^3, Hassan Mahomed^3, Willem Hanekom^3, Pierre-Alexandre Bart^4, Giuseppe Pantaleo^4, Georgia D Tomaras^5, Supachai Rerks-Ngarm^6, Jaranit Kaewkungwal^7, Sorachai Nitayaphan^8, Punnee Pitisuttithum^9, Nelson L Michael^{10}, Jerome H Kim^{10}, Merlin L Robb^{11}, Robert J O'Connell^{12}, Nicos Karasavvas^{12}, Peter Gilbert^1, Stephen C De Rosa^{1,13}, M Juliana McElrath^{1,2,13} & Raphael Gottardo^1$ 

Or in general, with immune response



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#### Or in general, with immune response.



# How do current methods work? (Approximately)

Current models are baseline/stimulation models.

• Unstimulated blood sample are compared stimulated ones.

For the unstimulated sample of the ith subject out of n, we sample a count proportion:

$$p_{i0} \sim {\sf Dirichlet}(lpha_0, eta_0),$$
  $y_{i0} \sim {\sf Multinomial}(N_{i0}, p_{i0}).$ 

Let  $k_i \in \{0, 1\}^p$  indicate in which subsets i responds:

$$k_{ij} \sim \mathsf{Ber}(w_j),$$
  $p_{i1, au=0} \sim \delta(p_{i0, au=0}), \quad p_{i1, au=1}|p_{i0, au=0} \propto \mathsf{Dirichlet}(lpha_1,eta_1)$   $y_{i1} \sim \mathsf{Multinomial}(N_{i1},p_{i1})$ 



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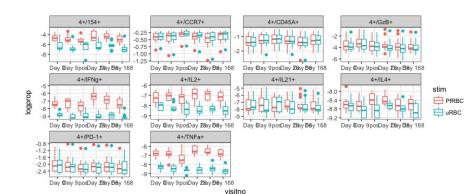
# Controlled Human Malaria Infection Study

- 9 subjects were infected with Malaria.
  - +3 controls.

- Blood samples were collected at 6 time points.
  - Day 0, day 9, blood parasitemia, Day 28, Day 56, Day 168.
- Two types of stimulation:
  - Infected/uninfected blood-cells.

- 53 cell subsets.
  - (10 types of cytokines in 8 cell-types)

# Controlled Human Malaria Infection Study



# Motivation - A Regression Model

- We want to be able to include covariates:
  - Batch effects.
  - · Other covariates such as age, gender...

- Longitudal data.
- More than one stimulation.
- Explicit dependence model:
  - For the observed proportions.
  - For response/non-response.

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#### Motivation - Unique Challanges

- Dependence
  - Within sample between cell subsets.
  - Within subject / across time.

- Heterogenous treatment effect
- Over-dispersed Binomial counts

# A Marginal Model - Single Subset

Indexing: i-subject, t- stimulation/time-point.

 $Over dispersion \Rightarrow Beta \ Binomial \ Model.$ 

$$egin{aligned} & \mathsf{logit}(\mu_{\mathit{it}}) = X_{\mathit{it}}eta, \ & p_{\mathit{it}} \sim \mathsf{Beta}(\textit{M}\mu, \textit{M}(\mathsf{1}-\mu)), \ & y_{\mathit{it}} \sim \mathsf{Binom}(\textit{N}_{\mathit{it}}, p_{\mathit{it}}) \end{aligned}$$

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Dependence ⇒ 'random' subject baseline:

$$u_i \sim \textit{N}(0, \sigma^2) \ ext{logit}(\mu_{it}) = \textit{X}_{it} eta + 
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# A Marginal Model - Single Subset

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Non-response ⇒ Mixture-Model:

$$k \sim \textit{Ber}(\theta),$$

$$logit(\mu_{itk}) = X_{it}\beta + T_{it}\tau_k + \nu_i,$$

- T a matrix of covariates related to the treatment.
- $\tau_k$  equals 0 if k = 0 or  $\tau \neq 0$  if k = 1.

# A Marginal Model - Recap

Indexing: i-subject, t- stimulation, k- cluster.

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Model can be estimated via an EM algorithm



#### Wellness of Fit Evaluation

#### How do we evaluate the model?

- We fit the model without information regarding the true treatment allocation.
- The model should be able to discriminate between vaccinees and placebos.
- We use three type of figures:
  - Scatter plots w/classification information.
  - Receiver-Operator Curves.
  - False Detection Rates.

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#### Marginal Model - Results

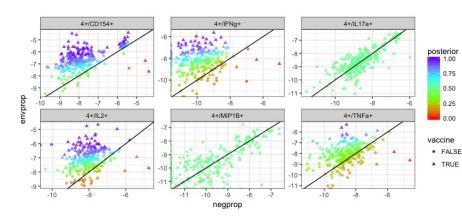


Figure: Posterior Probabilities for RV144 dataset - Independence Model



# Comparison w/ MIMOSA - Finak et al. (2013)

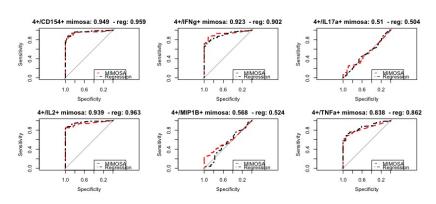


Figure: Comparison with MIMOSA (univariate COMPASS)

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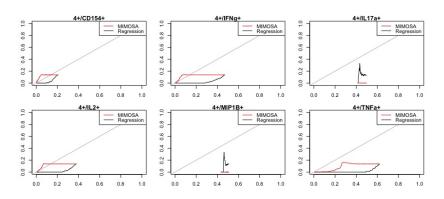


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# Subject-Response Model

Performing analysis for each cell-subset at a time doesn't use all of the information available.

- Random Effects are correlated, can be estimated better simultaneously.
- Correlation structure might be of interest in itself.
- Response is probably not independent across cell-subsets.
- We might be able to improve classification of response by looking at several cell-subsets at once.

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Indexing: i-subject, t- stimulation, j- subset, k- cluster.

Denote cluster (Response) by a  $k \in \{0, 1\}^p$  vector with 1 indicating a responsive subset.

We assume an Ising model for the dependence structure between subsets:

$$P(k) \propto \sum_{j=1}^{p} k_j \theta_j + \sum_{s \neq t} k_t k_s \theta_{st},$$

$$P(k_j = 1 | k_{-j}) = \theta_j + \sum_{t \neq j} k_t \theta_{tj}.$$

We can induce sparsity through an  $\ell_1$  penalty.

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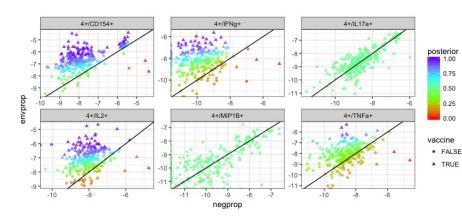


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#### HMRF Model - Results

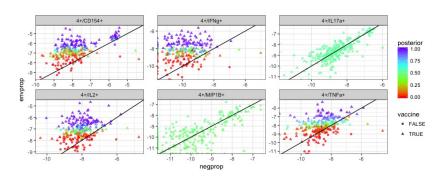


Figure: Scatter Plot for HMRF Modle Model

#### Subset-Response Model - Results

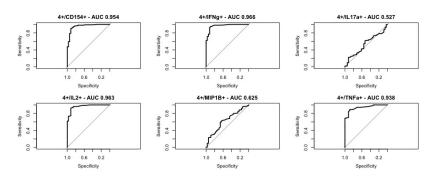


Figure: ROC for HMRF Modle

### Subset-Response Model - Results

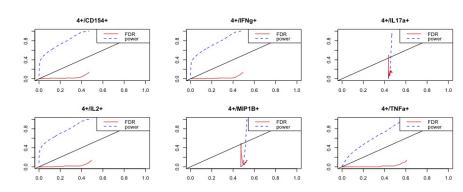


Figure: FDR for HMRF Modle

# Simulated Data Experiment

- Posterior probabilities are not well calibrated.
  - Might be due to true non-response.
- Is the optimization algorithm estimating the model properly?
- Does the model fit the data well?
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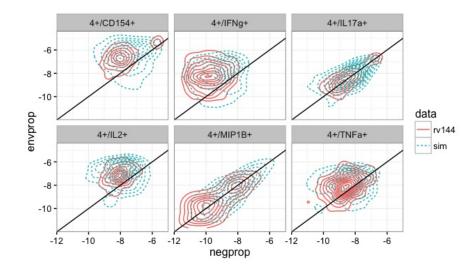


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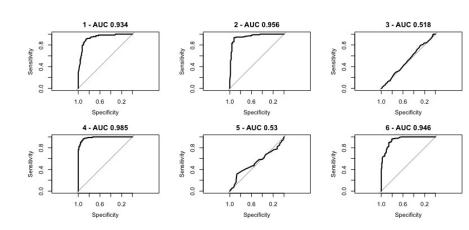
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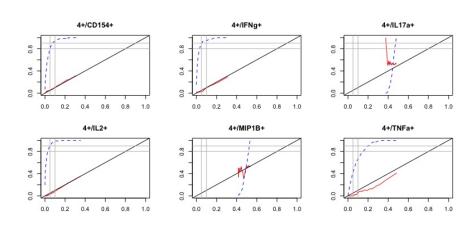
### How close are we to the distribution of the data?



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stim

0

0

23524

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P1003

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 So far we worked with marginal counts - can be obtained from bulk assays.

 single-cell measurements enable a more comprehensive understanding of cellular functions.

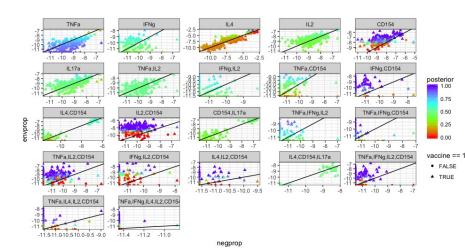
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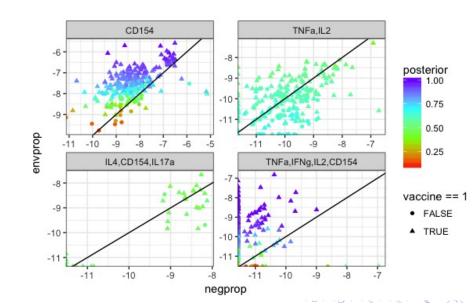
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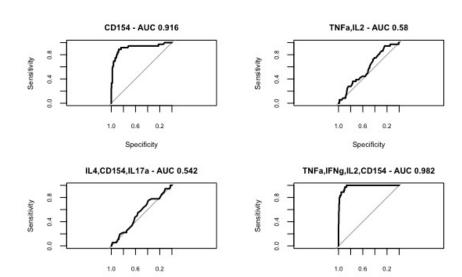
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- 262 Subjects
  - 226 Cases
  - 36 Controls
- 2 Types of stimulus
  - HIV antigen
  - Negative control
- 23 types of cells with non-negligible counts.
  - (At least two instances of *count* ≥ 5)

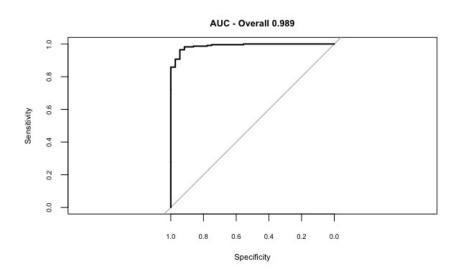






Specificity

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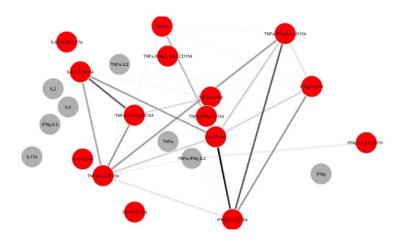


Figure: Estimated Ising Model - Red marks CD154

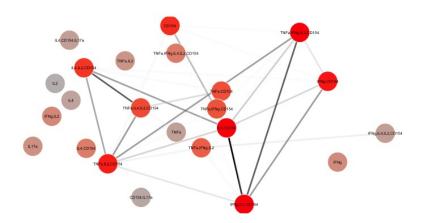
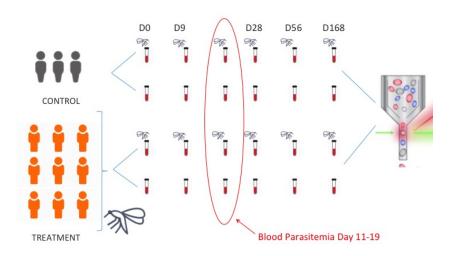


Figure: Estimated Ising Model - Red marks AUC



- 9 subjects were infected with Malaria.
  - +3 controls.

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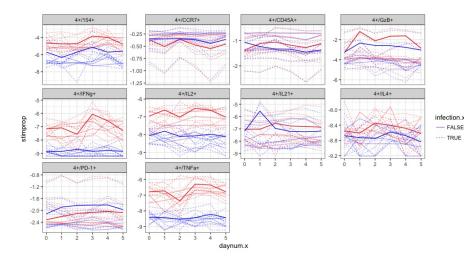
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- Individuals who experience malaria infections develop immunity.
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  - Even at day 0!
  - What is the profile of the immune response?
- The immunity is not long lived.
  - We might expect to see a rise in response during experiment.
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#### malaraFourPlusSmoothed



# FDR Adjusted p-values for CHMI Study

Standard errors for significance tests computed using Jackknife.

	4+	4+/CXCR5+	56+dim	56+hi	8+	8+/CXCR5+	NK T cells	PD-1+
154+	0.029	0.004			0.103	0.75	0.006	0.024
CCR7+	0.649	0.996			0.596	0.51		
CD45A+	0.575	0.307			0.543	0.54		
IFNg+	0.001	0.006	0.065	0.146	0.001		0.052	0.097
IL2+	0	0.005			0.119	0.56	0.321	0.052
IL21+	0.676	0.649	0.751	0.589	0.649		0.71	
IL4+	0.12	0.543		0.751	0.649		0.583	
TNFa+	0	0.001	0.261	0.309	0.276		0.053	0.09
GzB+	0.583		0.511	0.001	0.589		0.596	
PD-1+	0.751				0.596	0.83		

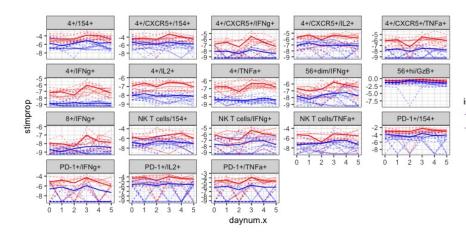


Figure: Significant Subsets

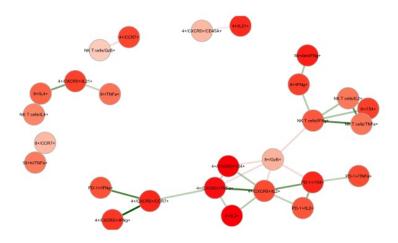


Figure: Estimated Graph - Red marks probability of response

# Computation - A Difficult Likelihood

With cluster assignments and random effects known:

$$f(y_i|\nu_i,k_i) = \prod_j \prod_t f(y_{ijt}|\nu_{ij},k_{ij})$$

The log-likelihood of the data is given by:

$$\ell(\beta, \tau, \theta, \Sigma) = \sum_{i=1}^{n} \log \left( \sum_{k \in \{0,1\}^{J}} P_{\theta}(k) \int_{\mathbb{R}^{J}} f_{\beta, \tau}(y_{i} | \nu_{i}, k_{i}) \varphi(\nu_{i}; 0, \Sigma) d\nu_{i} \right)$$

The log-likelihood is intractable for large J.

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The integrals are replaced with conditional expectations:

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We can use sampling to approximate the intractable integrals

$$k_{i1}^*, ..., k_{iM}^* \sim P_{\theta}(k|y_i)$$

$$\nu_{i1}^* \sim f(\nu_i|y_i, k_{i1}), ..., \nu_{iM}^* \sim f(\nu_i|y_i, k_{iM})$$

$$(\beta, \tau, \theta, \Sigma) = \arg\max \sum_{i=1}^n \frac{1}{M} \sum_{m=1}^M \sum_{i=1}^M \sum_{t=1}^m \log f(y_{ijt}|\nu_{im}^*, k_{im}^*)$$

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$$k_{i1}^*,...,k_{iM}^* \sim P_{\theta}(k|y_i)$$
 $u_{i1}^* \sim f(\nu_i|y_i,k_{i1}),...,\nu_{iM}^* \sim f(\nu_i|y_i,k_{iM})$ 
 $(\beta,\tau,\theta,\Sigma) = \arg\max\sum_{i=1}^n \frac{1}{M} \sum_{m=1}^M \sum_i \sum_t \log f(y_{ijt}|\nu_{im}^*,k_{im}^*)$ 

# Computation - Stochastic EM Algorithm

We alternate between a stochastic E-step and an M-step.

#### S - Step

- k\* Gibbs sampler.
- $\nu^* | \mathbf{k}^*$  component-wise MH algorithm.

#### M - Step

- $\beta$ ,  $\tau$  glm, glmnet for sparsity or gamlss for BB.
- θ Pseudo-likelihood or isingFit for sparsity.
- Σ Option for sparse estimation via PSDE package.

# Thank you!

Questions?

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