

1 Interest Rate Only

This section follows [3].

1.1 Dynamics

The instantaneous risk-free rate is $r_t = X_t^1 + X_t^2$. The state variables are level X_t^1 , slope X_t^2 , and curvature X_t^3 , which follow a system of SDEs ($d\mathbf{X}_t = K^Q(-\mathbf{X}_t)dt + \Sigma d\mathbf{W}_t^Q$):

- Independent factor AFNS:

$$\begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \end{pmatrix} = - \begin{pmatrix} 0 & 0 & 0 \\ 0 & \lambda & -\lambda \\ 0 & 0 & \lambda \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \end{pmatrix} dt + \begin{pmatrix} \sigma_{11} & 0 & 0 \\ 0 & \sigma_{22} & 0 \\ 0 & 0 & \sigma_{33} \end{pmatrix} \begin{pmatrix} dW_{1t}^Q \\ dW_{2t}^Q \\ dW_{3t}^Q \end{pmatrix}, \lambda > 0.$$

- Correlated factor AFNS:

$$\begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \end{pmatrix} = - \begin{pmatrix} 0 & 0 & 0 \\ 0 & \lambda & -\lambda \\ 0 & 0 & \lambda \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \end{pmatrix} dt + \begin{pmatrix} \sigma_{11} & 0 & 0 \\ \sigma_{21} & \sigma_{22} & 0 \\ \sigma_{31} & \sigma_{32} & \sigma_{33} \end{pmatrix} \begin{pmatrix} dW_{1t}^Q \\ dW_{2t}^Q \\ dW_{3t}^Q \end{pmatrix}, \lambda > 0.$$

The zero-coupon bond yield is:

$$y(t, T) = X_t^1 + \frac{1 - e^{-\lambda(T-t)}}{\lambda(T-t)} X_t^2 + \left(\frac{1 - e^{-\lambda(T-t)}}{\lambda(T-t)} - e^{-\lambda(T-t)} \right) X_t^3 - \frac{A(t, T)}{T-t}, \quad (1)$$

where $-\frac{A(t, T)}{T-t}$ is a yield-adjustment term:

$$A(t, T) = \frac{1}{2} \sum_{i=1}^3 \int_t^T \left(\Sigma^\top \mathbf{B}(s, T) \mathbf{B}(s, T)^\top \Sigma \right)_{i,i} ds.$$

Assume an affine risk premium, the P-dynamics $d\mathbf{X}_t = K^P(\boldsymbol{\theta}^P - \mathbf{X}_t)dt + \Sigma d\mathbf{W}_t^P$ for

- An independent factor AFNS model are:

$$\begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \end{pmatrix} = \begin{pmatrix} \kappa_{11}^P & 0 & 0 \\ 0 & \kappa_{22}^P & 0 \\ 0 & 0 & \kappa_{33}^P \end{pmatrix} \left(\begin{pmatrix} \theta_1^P \\ \theta_2^P \\ \theta_3^P \end{pmatrix} - \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \end{pmatrix} \right) dt + \begin{pmatrix} \sigma_{11} & 0 & 0 \\ 0 & \sigma_{22} & 0 \\ 0 & 0 & \sigma_{33} \end{pmatrix} \begin{pmatrix} dW_{1t}^P \\ dW_{2t}^P \\ dW_{3t}^P \end{pmatrix}.$$

- A correlated-factor AFNS model are:

$$\begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \end{pmatrix} = \begin{pmatrix} \kappa_{11}^P & \kappa_{12}^P & \kappa_{13}^P \\ \kappa_{21}^P & \kappa_{22}^P & \kappa_{23}^P \\ \kappa_{31}^P & \kappa_{32}^P & \kappa_{33}^P \end{pmatrix} \left(\begin{pmatrix} \theta_1^P \\ \theta_2^P \\ \theta_3^P \end{pmatrix} - \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \end{pmatrix} \right) dt + \begin{pmatrix} \sigma_{11} & 0 & 0 \\ \sigma_{21} & \sigma_{22} & 0 \\ \sigma_{31} & \sigma_{32} & \sigma_{33} \end{pmatrix} \begin{pmatrix} dW_{1t}^P \\ dW_{2t}^P \\ dW_{3t}^P \end{pmatrix}.$$

The measurement equation is:

$$\mathbf{y}_t = B\mathbf{X}_t + \mathbf{A} + \boldsymbol{\epsilon}_t \iff$$

$$\begin{pmatrix} y_t(\tau_1) \\ y_t(\tau_2) \\ \vdots \\ y_t(\tau_N) \end{pmatrix} = \begin{pmatrix} 1 & \frac{1-e^{-\lambda\tau_1}}{\lambda\tau_1} & \frac{1-e^{-\lambda\tau_1}}{\lambda\tau_1} - e^{-\lambda\tau_1} \\ 1 & \frac{1-e^{-\lambda\tau_2}}{\lambda\tau_2} & \frac{1-e^{-\lambda\tau_2}}{\lambda\tau_2} - e^{-\lambda\tau_2} \\ \vdots & \vdots & \vdots \\ 1 & \frac{1-e^{-\lambda\tau_N}}{\lambda\tau_N} & \frac{1-e^{-\lambda\tau_N}}{\lambda\tau_N} - e^{-\lambda\tau_N} \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \end{pmatrix} - \begin{pmatrix} \frac{A(\tau_1)}{\tau_1} \\ \frac{A(\tau_2)}{\tau_2} \\ \vdots \\ \frac{A(\tau_N)}{\tau_N} \end{pmatrix} + \begin{pmatrix} \epsilon_t(\tau_1) \\ \epsilon_t(\tau_2) \\ \vdots \\ \epsilon_t(\tau_N) \end{pmatrix},$$

where the measurement errors $\epsilon_t(\tau_i)$ are iid noises.

1.2 Estimation

The state transition equation is:

$$\mathbf{X}_t = [I - \exp(-K^P \Delta t)] \boldsymbol{\theta}^P + \exp(-K^P \Delta t) \mathbf{X}_{t-1} + \boldsymbol{\eta}_t,$$

with the error structure

$$\begin{pmatrix} \boldsymbol{\eta}_t \\ \boldsymbol{\epsilon}_t \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} \mathbf{0} \\ \mathbf{0} \end{pmatrix}, \begin{pmatrix} Q & 0 \\ 0 & H \end{pmatrix} \right],$$

where Δt is the time between observations, $H = \text{diag}(\sigma_\epsilon^2(\tau_1), \sigma_\epsilon^2(\tau_2), \dots, \sigma_\epsilon^2(\tau_N))$, and

$$Q = \int_0^{\Delta t} e^{-K^P s} \Sigma \Sigma^\top e^{-(K^P)^\top s} ds.$$

We apply the Kalman filtering to estimate the parameters:

1. Initialise the filter at the unconditional mean and variance of the state variables under the P-measure:

$$\mathbf{X}_0 = \boldsymbol{\theta}^P, \Sigma_0 = \int_0^\infty e^{-K^P s} \Sigma \Sigma^\top e^{-(K^P)^\top s} ds.$$

2. Let the information available at time t be $\mathbf{Y}_t = (y_1, \dots, y_t)$. The prediction step is:

$$\begin{aligned} \mathbf{X}_{t|t-1} &= [I - \exp(-K^P \Delta t)] \boldsymbol{\theta}^P + \exp(-K^P \Delta t) \mathbf{X}_{t-1}, \\ \Sigma_{t|t-1} &= \exp(-K^P \Delta t) \Sigma_{t-1} \exp(-K^P \Delta t)^\top + \int_0^{\Delta t} e^{-K^P s} \Sigma \Sigma^\top e^{-(K^P)^\top s} ds. \end{aligned}$$

We utilise the analytical formulas in [4].

3. Update

$$\begin{aligned} \mathbf{X}_t &= \mathbf{X}_{t|t-1} + \Sigma_{t|t-1} B^\top \text{Cov}(\mathbf{v}_t)^{-1} \mathbf{v}_t, \\ \Sigma_t &= \Sigma_{t|t-1} - \Sigma_{t|t-1} B^\top \text{Cov}(\mathbf{v}_t)^{-1} B \Sigma_{t|t-1}, \end{aligned}$$

where

$$\begin{aligned} \mathbf{v}_t &= \mathbf{y}_t - \mathbf{A} - B\mathbf{X}_{t|t-1}, \\ \text{Cov}(\mathbf{v}_t) &= B\Sigma_{t|t-1}B^\top + H. \end{aligned}$$

4. The log-likelihood is:

$$\log \ell(y_1, \dots, y_T) = \sum_{t=1}^T \left(-\frac{N}{2} \log(2\pi) - \frac{1}{2} \log \det(\text{Cov}(\mathbf{v}_t)) - \frac{1}{2} \mathbf{v}_t^\top \text{Cov}(\mathbf{v}_t)^{-1} \mathbf{v}_t \right),$$

where N is the number of observed yields.

5. Compute MLE by the Nelder-Mead approach.

We use the monthly RBA data from July 1992 to May 2021, maturing every quarter until 10 years. We adapt the R codes from [1].

1.3 Results

1.3.1 Independent-factor AFNS

Assuming K^P is diagonalisable, the estimated parameters are:

$$\hat{K}^P, \hat{\Sigma}, \hat{\theta}^P, \hat{\lambda}$$

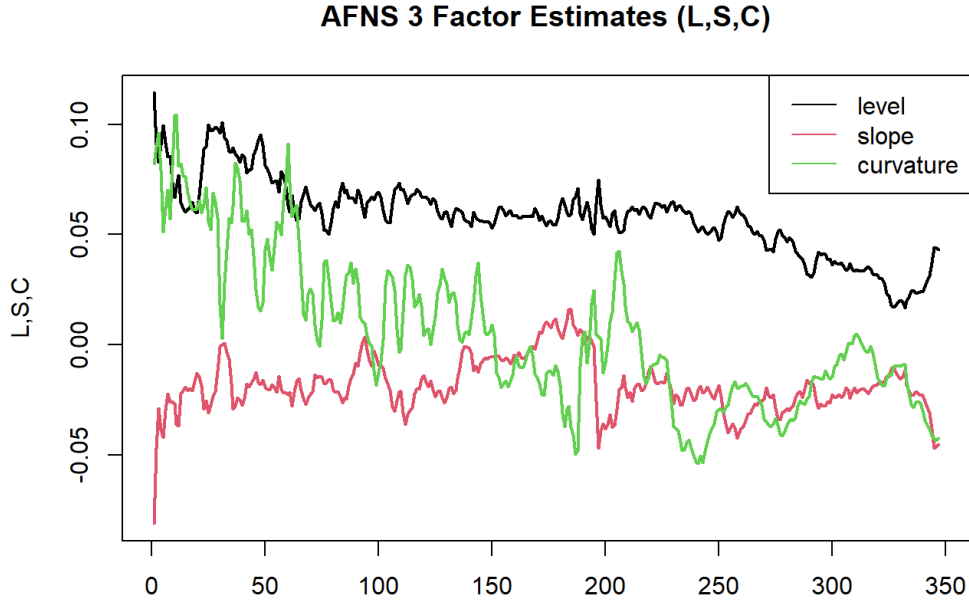


Figure 1: Fitted state variables for the independent factor AFNS.

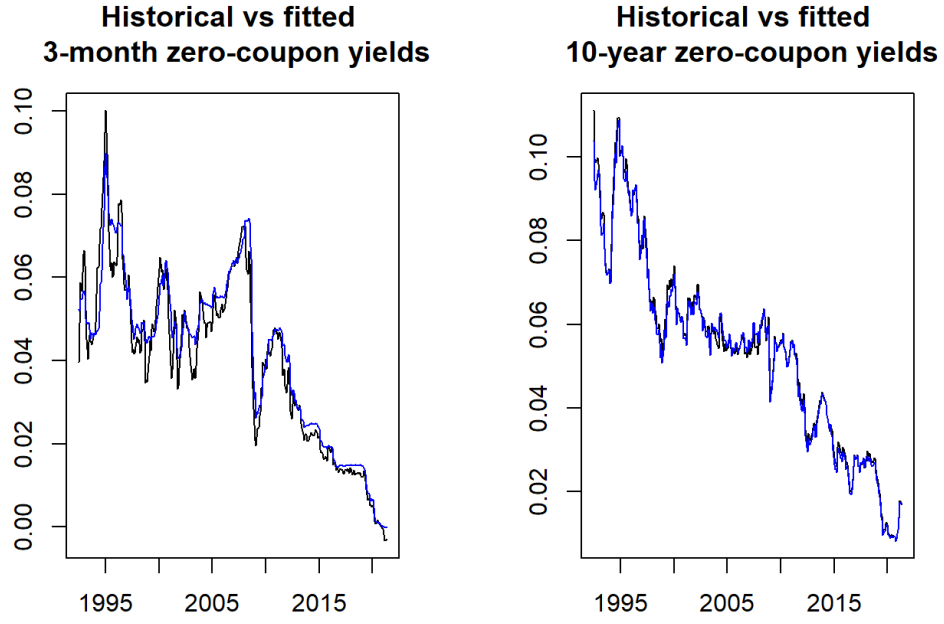


Figure 2: Fitted zero-coupon bond rates for the independent factor AFNS.

1.3.2 Correlated-factor AFNS

Assuming K^P is diagonalisable, the estimated parameters are:

$$\hat{K}^P, \hat{\Sigma}, \hat{\theta}^P, \hat{\lambda}$$

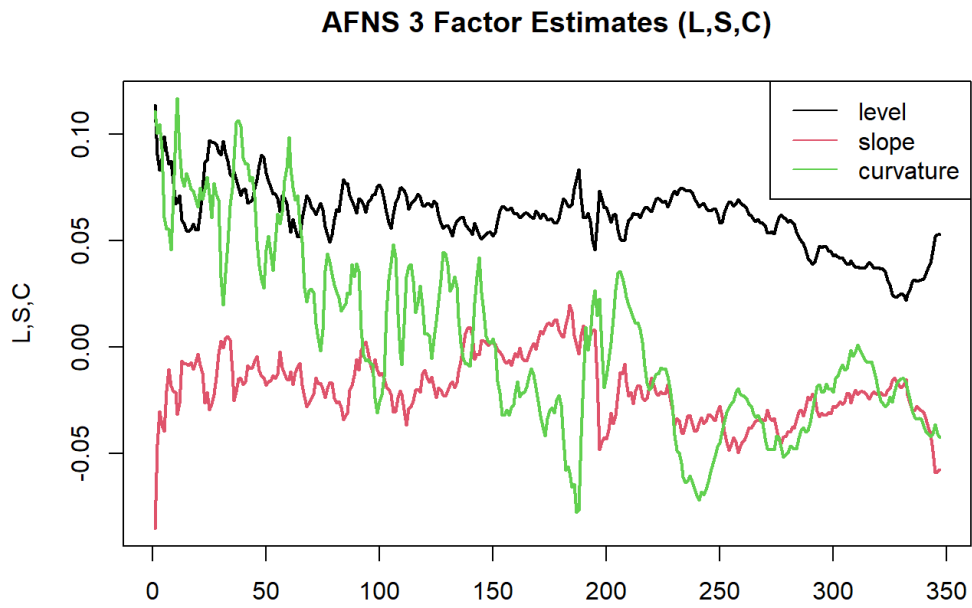


Figure 3: Fitted state variables for the correlated factor AFNS.

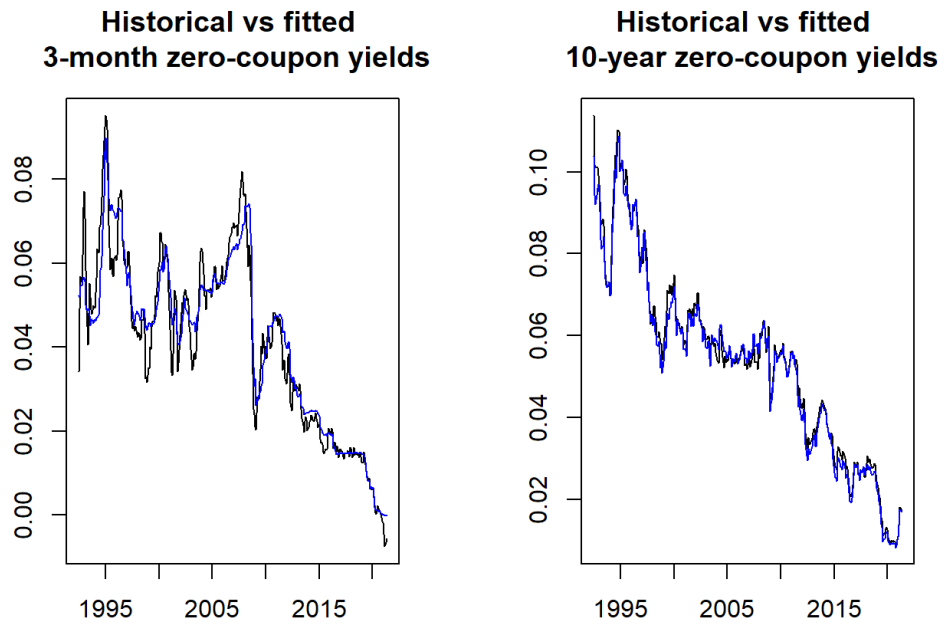


Figure 4: Fitted state variables for the correlated factor AFNS.

1.4 Testing

1.5 Simulation

Fix the time step $h > 0$. Discretise the SDEs under the P measure:

$$\mathbf{X}_{t+h} - \mathbf{X}_t = K^P(\boldsymbol{\theta}^P - \mathbf{X}_t)h + \Sigma\sqrt{h}\mathbf{Z}, \mathbf{Z} \sim \mathcal{N}(\mathbf{0}, I).$$

This gives a trajectory of the state variables: $\mathbf{X}_0, \mathbf{X}_h, \mathbf{X}_{2h}, \dots$. We use it with formula (1) to simulate the yields of a maturity- τ month zero-coupon bond.

Simulated 3-month zero-coupon yield Simulated 10-year zero-coupon yield

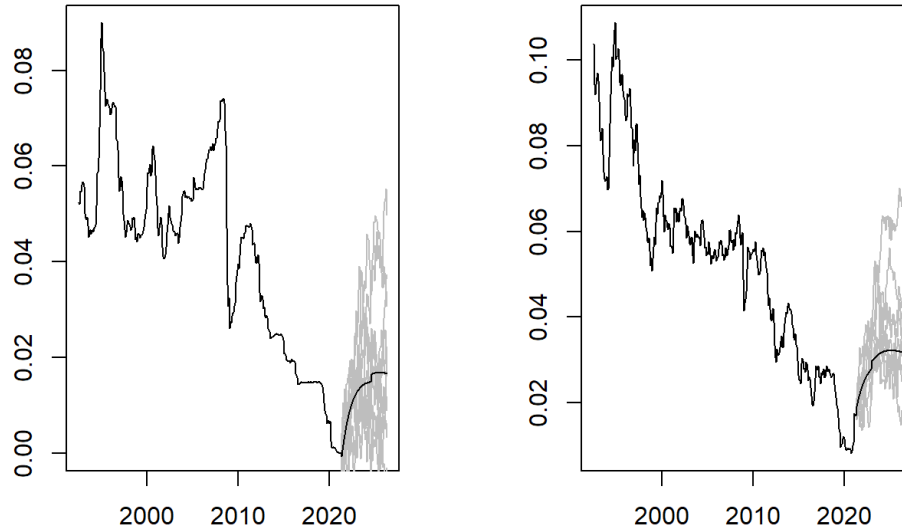


Figure 5: Simulated zero-coupon bond rates for the independent factor AFNS.

2 Interest Rate, Home Value Indexes, Stock Prices

This section is based on [2]. Granger causality between these variables are proven in "VAR_Framework.pdf".

Assume that the interest rate follows AFNS as above, both the home value indexes and stock prices follow the GBM. Denote the additional state variables as: logged home value index NSW $X_t^4 = \log(H_t)$ and logged stock price $X_t^5 = \log(S_t)$. The new state variables are correlated with X_t^1, X_t^2, X_t^3 .

2.1 Dynamics

The Q-dynamics are:

$$dX_t^4 = \left(X_t^1 + X_t^2 - \frac{\sigma_4^2}{2} \right) dt + \sigma_4 \sum_{i=1}^3 \rho_{4i} dW_{it}^Q + \sigma_4 \sqrt{1 - \rho_4^2} dW_{4t}^Q,$$

$$dX_t^5 = \left(X_t^1 + X_t^2 - \frac{\sigma_5^2}{2} \right) dt + \sigma_5 \sum_{i=1}^4 \rho_{5i} dW_{it}^Q + \sigma_5 \sqrt{1 - \rho_5^2} dW_{5t}^Q.$$

The standard Brownian motions W_{4t}^Q and W_{5t}^Q are correlated with other BMs by:

$$\begin{aligned} \text{Cov} \left(dW_{4t}^Q, dW_{it}^Q \right) &= \rho_{4i} dt, i = 1, 2, 3, \\ \text{Cov} \left(dW_{5t}^Q, dW_{it}^Q \right) &= \rho_{5i} dt, i = 1, 2, 3, 4. \end{aligned}$$

The ρ 's are correlation coefficients which take value from -1 to 1. Let $\rho_4^2 = \sum_{i=1}^3 \rho_{4i}^2 \leq 1$, and $\rho_5^2 = \sum_{i=1}^4 \rho_{5i}^2 \leq 1$.

The state variables $\mathbf{X}_t = (X_t^1, X_t^2, X_t^3, X_t^4, X_t^5)$ follow the system of SDEs $d\mathbf{X}_t = (\boldsymbol{\theta}^Q - K^Q \mathbf{X}_t)dt + \Sigma d\mathbf{W}_t^Q$:

- Independent factor:

$$\begin{aligned} \begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \\ dX_t^4 \\ dX_t^5 \end{pmatrix} &= \left[\begin{pmatrix} 0 \\ 0 \\ 0 \\ -\frac{\sigma_4^2}{2} \\ -\frac{\sigma_5^2}{2} \end{pmatrix} - \begin{pmatrix} 0 & 0 & 0 & 0 & 0 \\ 0 & \lambda & -\lambda & 0 & 0 \\ 0 & 0 & \lambda & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \\ X_t^4 \\ X_t^5 \end{pmatrix} \right] dt + \\ &\quad \begin{pmatrix} \sigma_{11} & 0 & 0 & 0 & 0 \\ 0 & \sigma_{22} & 0 & 0 & 0 \\ 0 & 0 & \sigma_{33} & 0 & 0 \\ \sigma_4 \rho_{41} & \sigma_4 \rho_{42} & \sigma_4 \rho_{43} & \sigma_4 \sqrt{1 - \rho_4^2} & 0 \\ \sigma_5 \rho_{51} & \sigma_5 \rho_{52} & \sigma_5 \rho_{53} & \sigma_5 \rho_{54} & \sigma_5 \sqrt{1 - \rho_5^2} \end{pmatrix} \begin{pmatrix} dW_{1t}^Q \\ dW_{2t}^Q \\ dW_{3t}^Q \\ dW_{4t}^Q \\ dW_{5t}^Q \end{pmatrix} \end{aligned}$$

- Correlated factor:

$$\begin{aligned} \begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \\ dX_t^4 \\ dX_t^5 \end{pmatrix} &= \left[\begin{pmatrix} 0 \\ 0 \\ 0 \\ -\frac{\sigma_4^2}{2} \\ -\frac{\sigma_5^2}{2} \end{pmatrix} - \begin{pmatrix} 0 & 0 & 0 & 0 & 0 \\ 0 & \lambda & -\lambda & 0 & 0 \\ 0 & 0 & \lambda & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \\ X_t^4 \\ X_t^5 \end{pmatrix} \right] dt + \\ &\quad \begin{pmatrix} \sigma_{11} & 0 & 0 & 0 & 0 \\ \sigma_{21} & \sigma_{22} & 0 & 0 & 0 \\ \sigma_{31} & \sigma_{32} & \sigma_{33} & 0 & 0 \\ \sigma_4 \rho_{41} & \sigma_4 \rho_{42} & \sigma_4 \rho_{43} & \sigma_4 \sqrt{1 - \rho_4^2} & 0 \\ \sigma_5 \rho_{51} & \sigma_5 \rho_{52} & \sigma_5 \rho_{53} & \sigma_5 \rho_{54} & \sigma_5 \sqrt{1 - \rho_5^2} \end{pmatrix} \begin{pmatrix} dW_{1t}^Q \\ dW_{2t}^Q \\ dW_{3t}^Q \\ dW_{4t}^Q \\ dW_{5t}^Q \end{pmatrix} \end{aligned}$$

The measure change is facilitated by the Girsanov's Theorem: $d\mathbf{W}_t^Q = d\mathbf{W}_t^P + \Gamma_t dt$. Assume an essentially affine risk premium:

$$\begin{aligned}\Gamma_t &= \gamma^0 + \gamma^1 \mathbf{X}_t \\ &= \begin{pmatrix} \gamma_1^0 \\ \gamma_2^0 \\ \gamma_3^0 \\ \gamma_4^0 \\ \gamma_5^0 \end{pmatrix} + \begin{pmatrix} \gamma_{11}^1 & \gamma_{12}^1 & \gamma_{13}^1 & \gamma_{14}^1 & \gamma_{15}^1 \\ \gamma_{21}^1 & \gamma_{22}^1 & \gamma_{23}^1 & \gamma_{24}^1 & \gamma_{25}^1 \\ \gamma_{31}^1 & \gamma_{32}^1 & \gamma_{33}^1 & \gamma_{34}^1 & \gamma_{35}^1 \\ \gamma_{41}^1 & \gamma_{42}^1 & \gamma_{43}^1 & \gamma_{44}^1 & \gamma_{45}^1 \\ \gamma_{51}^1 & \gamma_{52}^1 & \gamma_{53}^1 & \gamma_{54}^1 & \gamma_{55}^1 \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \\ X_t^4 \\ X_t^5 \end{pmatrix}.\end{aligned}$$

Substituting it into the Q-dynamics yields:

$$\begin{aligned}d\mathbf{X}_t &= (\boldsymbol{\theta}^Q - K^Q \mathbf{X}_t)dt + \Sigma [d\mathbf{W}_t^P + (\gamma^0 + \gamma^1 \mathbf{X}_t)dt] \\ &= [(\boldsymbol{\theta}^Q + \Sigma \gamma^0) - (K^Q - \Sigma \gamma^1) \mathbf{X}_t] dt + \Sigma d\mathbf{W}_t^P.\end{aligned}$$

We can make further assumptions to γ^0, γ^1 such that the P-dynamics preserve the structure of the corresponding Q-dynamics:

- Independent factor:

$$\begin{aligned}\begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \\ dX_t^4 \\ dX_t^5 \end{pmatrix} &= \left[\begin{pmatrix} \theta_1^P \\ \theta_2^P \\ \theta_3^P \\ \theta_4^P \\ \theta_5^P \end{pmatrix} - \begin{pmatrix} \kappa_{11}^P & 0 & 0 & 0 & 0 \\ 0 & \kappa_{22}^P & 0 & 0 & 0 \\ 0 & 0 & \kappa_{33}^P & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \\ X_t^4 \\ X_t^5 \end{pmatrix} \right] dt + \\ &\quad \begin{pmatrix} \sigma_{11} & 0 & 0 & 0 & 0 \\ 0 & \sigma_{22} & 0 & 0 & 0 \\ 0 & 0 & \sigma_{33} & 0 & 0 \\ \sigma_4 \rho_{41} & \sigma_4 \rho_{42} & \sigma_4 \rho_{43} & \sigma_4 \sqrt{1 - \rho_4^2} & 0 \\ \sigma_5 \rho_{51} & \sigma_5 \rho_{52} & \sigma_5 \rho_{53} & \sigma_5 \rho_{54} & \sigma_5 \sqrt{1 - \rho_5^2} \end{pmatrix} \begin{pmatrix} dW_{1t}^P \\ dW_{2t}^P \\ dW_{3t}^P \\ dW_{4t}^P \\ dW_{5t}^P \end{pmatrix}\end{aligned}$$

- Correlated factor:

$$\begin{aligned}\begin{pmatrix} dX_t^1 \\ dX_t^2 \\ dX_t^3 \\ dX_t^4 \\ dX_t^5 \end{pmatrix} &= \left[\begin{pmatrix} \theta_1^P \\ \theta_2^P \\ \theta_3^P \\ \theta_4^P \\ \theta_5^P \end{pmatrix} - \begin{pmatrix} \kappa_{11}^P & \kappa_{12}^P & \kappa_{13}^P & 0 & 0 \\ \kappa_{21}^P & \kappa_{22}^P & \kappa_{23}^P & 0 & 0 \\ \kappa_{31}^P & \kappa_{32}^P & \kappa_{33}^P & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \\ -1 & -1 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \\ X_t^4 \\ X_t^5 \end{pmatrix} \right] dt + \\ &\quad \begin{pmatrix} \sigma_{11} & 0 & 0 & 0 & 0 \\ \sigma_{21} & \sigma_{22} & 0 & 0 & 0 \\ \sigma_{31} & \sigma_{32} & \sigma_{33} & 0 & 0 \\ \sigma_4 \rho_{41} & \sigma_4 \rho_{42} & \sigma_4 \rho_{43} & \sigma_4 \sqrt{1 - \rho_4^2} & 0 \\ \sigma_5 \rho_{51} & \sigma_5 \rho_{52} & \sigma_5 \rho_{53} & \sigma_5 \rho_{54} & \sigma_5 \sqrt{1 - \rho_5^2} \end{pmatrix} \begin{pmatrix} dW_{1t}^P \\ dW_{2t}^P \\ dW_{3t}^P \\ dW_{4t}^P \\ dW_{5t}^P \end{pmatrix}\end{aligned}$$

The measurement equation is:

$$\begin{pmatrix} y_t(\tau_1) \\ y_t(\tau_2) \\ \vdots \\ y_t(\tau_N) \\ H_t \\ S_t \end{pmatrix} = \begin{pmatrix} 1 & \frac{1-e^{-\lambda\tau_1}}{\lambda\tau_1} & \frac{1-e^{-\lambda\tau_1}}{\lambda\tau_1} - e^{-\lambda\tau_1} & 0 & 0 \\ 1 & \frac{1-e^{-\lambda\tau_2}}{\lambda\tau_2} & \frac{1-e^{-\lambda\tau_2}}{\lambda\tau_2} - e^{-\lambda\tau_2} & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & \frac{1-e^{-\lambda\tau_N}}{\lambda\tau_N} & \frac{1-e^{-\lambda\tau_N}}{\lambda\tau_N} - e^{-\lambda\tau_N} & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} X_t^1 \\ X_t^2 \\ X_t^3 \\ X_t^4 \\ X_t^5 \end{pmatrix} - \begin{pmatrix} \frac{A(\tau_1)}{\tau_1} \\ \frac{A(\tau_2)}{\tau_2} \\ \vdots \\ \frac{A(\tau_N)}{\tau_N} \\ 0 \\ 0 \end{pmatrix} + \begin{pmatrix} \epsilon_t(\tau_1) \\ \epsilon_t(\tau_2) \\ \vdots \\ \epsilon_t(\tau_N) \\ \epsilon_H \\ \epsilon_S \end{pmatrix},$$

2.2 Estimation

2.3 Results

2.4 Testing

2.5 Simulation

References

- [1] Arbitrage-free nelson siegel model with r code. <https://www.r-bloggers.com/2021/05/arbitrage-free-nelson-siegel-model-with-r-code/>. Accessed: 2022-01-30.
- [2] Reverse mortgage : Long term care project. unpublished.
- [3] Jens HE Christensen, Francis X Diebold, and Glenn D Rudebusch. The affine arbitrage-free class of nelson–siegel term structure models. *Journal of Econometrics*, 164(1):4–20, 2011.
- [4] Jens HE Christensen, Jose A Lopez, and Glenn D Rudebusch. Analytical formulas for the second moment in affine models with stochastic volatility. 2015.