

APPENDIX

A. Proof of Theorem 1

Proof. Now we consider a $GBA(U_{B1})$ where poison values exist on both sides. Let V_{B1}^L denote the set of poison values on $[D_L, O]$, and V_{B1}^R denote those on $[O, D_R]$.

Without loss of generality, consider the case where the mean of the left side is larger. Let

$$C_0 = \Sigma(V_{B1}^L - O) + \Sigma(V_{B1}^R - O) < 0. \quad (11)$$

Let us choose the largest poison value \mathbf{y}_r in V_{B1}^R , and an arbitrary subset of poison values \mathcal{Y}_L from V_{B1}^L , such that the following two formulas are satisfied concurrently:

$$\Sigma(\mathcal{Y}_L - O) + \mathbf{y}_r - O \leq 0, \quad (12)$$

and for any $\mathbf{y}_1 \in \mathcal{Y}_L$,

$$\Sigma(\mathcal{Y}_L' - O) + \mathbf{y}_r - O > 0, \quad (13)$$

where \mathcal{Y}_L' is the set of elements in \mathcal{Y}_L excluding \mathbf{y}_1 .

This is always available, because

$$\Sigma(O - V_{B1}^L) > \Sigma(V_{B1}^R - O) \geq \mathbf{y}_r - O.$$

Add $\mathbf{y}_1 - O$ on both sides of Equ. 13, we then have

$$\begin{aligned} \Sigma(\mathcal{Y}_L' - O) + \mathbf{y}_1 - O + \mathbf{y}_r - O &> \mathbf{y}_1 - O \\ \Sigma(\mathcal{Y}_L - O) + \mathbf{y}_r - O &> \mathbf{y}_1 - O \end{aligned} \quad (14)$$

Let

$$\mathbf{y}_1' - O = \Sigma(\mathcal{Y}_L - O) + \mathbf{y}_r - O \quad (15)$$

According to Equ. 12 and Equ. 14, we have

$$\mathbf{y}_1 - O < \mathbf{y}_1' - O \leq 0$$

Apparently, $\mathbf{y}_1' \in [D_L, O]$. Let us remove \mathbf{y}_r from V_{B1}^R , \mathcal{Y}_L from \mathcal{Y}_L , and add \mathbf{y}_1' into V_{B1}^L . Such an operation eliminates a poison value on the right hand without changing C_0 according to Equ. 11 and Equ. 15. Hence, the generated distribution can be reduced to the initial one.

Repeating the same operation until V_{B1}^R is empty, we finally obtain a $V_{B2}' = V_{B1}^L$. This is achievable, as both V_{B1}^L and V_{B1}^R are finite, and $C_0 < 0$. We finally obtain a $BBA(U_{B2})$ reporting V_{B2}' where poison values are on the left hand. \square

B. Proof of Theorem 2

Proof. The true mean O can be obtained by removing the effect of Y from V' :

$$O = \frac{\sum_{v'_i \in V'} v'_i - \sum_{v'_j \in V'_B} v'_j}{1 - \gamma}. \quad (16)$$

Likewise, we can obtain O' from removing the effect of T :

$$O' = \frac{\sum_{v'_i \in V'} v'_i - \sum_{v'_j \in T} v'_j}{1 - \gamma_{sup}}. \quad (17)$$

Since the values in T are the largest γ_{sup} in Y , so we have

$$\sum_{v'_i \in T} v'_i \geq \sum_{v'_j \in V'_B} v'_j$$

Compare Equ. 16 and Equ. 17, we have $O' \leq O$. \square

C. Proof of Theorem 5

Proof. Let $\mathcal{B} = \{B_{y_1}, \dots, B_{y_t}\}$, and $\bar{\mathcal{B}} = \{B_{y_{t+1}}, \dots, B_{y_{d'}}\}$. We start from the state where all buckets hold non-zero values, i.e., $B_{y_i} \neq 0, i \in \{1, \dots, d'\}$, and reconstruct the frequency histogram for poison values in $[-C, C]$. The likelihood estimator in Equ. 1 becomes

$$\begin{aligned} l(F) &= \sum_{i=1}^N \ln \left(\sum_{k=1}^d \hat{x}_k \Pr[v'_i | v_i \in B_{x_k}] + \sum_{j=1}^{d'} \hat{y}_j \Pr[v'_i | v_i \in B_{y_j}] \right) \\ &= \sum_{t=1}^{d'} c_t \ln \left(\sum_{k=1}^d \hat{x}_k M_{b_t x_k} + \sum_{j=1}^{d'} \hat{y}_j M_{b_t y_j} \right). \end{aligned}$$

Note that $\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^{d'} \hat{y}_j = 1$, we employ the Lagrangian multiplier method to derive the extremum. The Lagrangian function can be written as:

$$L(F) = l(F) + \lambda \left(\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^{d'} \hat{y}_j - 1 \right).$$

Let all first-order partial derivatives of L w.r.t. \hat{x}_k and \hat{y}_j equal zero

$$\begin{aligned} \frac{\partial L(F)}{\partial \hat{x}_k} &= \sum_{t=1}^{d'} c_t \frac{M_{b_t x_k}}{\sum_{k=1}^d \hat{x}_k M_{b_t x_k} + \hat{y}_t} + \lambda = 0, \quad k \in \{1, \dots, d\} \\ \frac{\partial L(F)}{\partial \hat{y}_j} &= \sum_{t=1}^{d'} c_t \frac{M_{b_t y_j}}{\sum_{k=1}^d \hat{x}_k M_{b_t x_k} + \hat{y}_t} + \lambda = 0, \quad j \in \{1, \dots, d'\}, \end{aligned}$$

we have:

$$\hat{x}_k = 0, \quad k \in \{1, \dots, d\}, \quad \hat{y}_j = \frac{c_j}{N}, \quad j \in \{1, \dots, d'\}, \quad \lambda = -N.$$

This result shows all collected values converge to poison values if no bucket is suppressed, and we can obtain:

$$\left(\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_i \neq 0, i \in \{1, \dots, d'\}} = \sum_{j=1}^t \frac{c_j}{N}.$$

When we suppress the bucket $B_{y_{d'}}$ (by setting $\hat{y}_{d'} = 0$) and carry out EMF, the collected values in $B'_{b_{d'}}$ can only converge to $B_{x_k} (k \in \{1, \dots, d\})$, but not $B_{y_j} (j \in \{1, \dots, d'\})$. Hence, every \hat{x}_k will increase. Therefore, suppressing $B_{y_{d'}}$ leads to the increase of all \hat{x}_k , which in turn results in the decrease of all \hat{y}_j . However, since the decrease of $\hat{y}_j (j \in \{1, \dots, t\})$ is a part of increment of \hat{x} , we can figure out $(\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j) \Big|_{y_i \neq 0, i \in \{1, \dots, d'\}} \leq$

$$\begin{aligned} &(\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j) \Big|_{y_{d'}=0} \\ &(\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j) \Big|_{y_i \neq 0, i \in \{1, \dots, d'\}} \leq (\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j) \Big|_{y_{d'}=0} \\ &\leq \dots \leq (\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j) \Big|_{y_{d'}=0, \dots, y_{t+1}=0}. \end{aligned}$$

When the number of suppressed buckets in $\bar{\mathcal{B}}$ increases, the corresponding interference of $\bar{\mathcal{B}}$ decreases. Therefore, the collected values more accurately converge to the buckets that

they should belong to, and thus achieve a better convergence result.

After suppressing all buckets in \overline{B} , all collected values will convergence to normal values and poison values in $B_{y_j}(j \in 1, \dots, t)$ and we can infer that $(\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j) \Big|_{y_{d'}=0, \dots, y_{t+1}=0} = 1$, which is the optimal case where none of the collected values will converge to buckets in \overline{B} . \square