

## A PROOF OF THEOREM 5.2

PROOF. Let  $\mathcal{B} = \{B_{y_1}, \dots, B_{y_t}\}$ , and  $\bar{\mathcal{B}} = \{B_{y_{t+1}}, \dots, B_{y_{d'}}\}$ . We start from the state where all buckets hold non-zero values, i.e.,  $B_{y_i} \neq 0, i \in \{1, \dots, d'\}$ , and reconstruct the frequency histogram for poison values in  $[-C, C]$ . The likelihood estimator in Equ. 1 becomes

$$\begin{aligned} l(F) &= \sum_{i=1}^N \ln \left( \sum_{k=1}^d \hat{x}_k \Pr[v'_i | v_i \in B_{x_k}] + \sum_{j=1}^{d'} \hat{y}_j \Pr[v'_i | v_i \in B_{y_j}] \right) \\ &= \sum_{t=1}^{d'} n_t \ln \left( \sum_{k=1}^d \hat{x}_k M_{b_t x_k} + \sum_{j=1}^{d'} \hat{y}_j M_{b_t y_j} \right). \end{aligned}$$

Note that  $\sum_{k=1}^d \hat{x}_k + \sum_{j=1}^{d'} \hat{y}_j = 1$ , we employ the Lagrangian multiplier method to derive the extremum. The Lagrangian function can be written as:

$$L(F) = l(F) + \lambda \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^{d'} \hat{y}_j - 1 \right). \quad (12)$$

Let all first-order partial derivatives of  $L$  w.r.t.  $\hat{x}_k$  and  $\hat{y}_j$  equal zero

$$\begin{aligned} \frac{\partial L(F)}{\partial \hat{x}_k} &= \sum_{t=1}^{d'} n_t \frac{M_{b_t x_k}}{\sum_{k=1}^d \hat{x}_k M_{b_t x_k} + \sum_{j=1}^{d'} \hat{y}_j} + \lambda = 0, \quad k \in \{1, \dots, d\} \\ \frac{\partial L(F)}{\partial \hat{y}_j} &= \sum_{t=1}^{d'} n_t \frac{M_{b_t y_j}}{\sum_{k=1}^d \hat{x}_k M_{b_t x_k} + \sum_{j=1}^{d'} \hat{y}_j} + \lambda = 0, \quad j \in \{1, \dots, d'\}, \end{aligned}$$

we have:

$$\hat{x}_k = 0, \quad k \in \{1, \dots, d\}, \quad \hat{y}_j = \frac{n_j}{N}, \quad j \in \{1, \dots, d'\}, \quad \lambda = -N.$$

This result shows all collected values converge to poison values if no bucket is suppressed.

When we suppress the bucket  $B_{y_{d'}}$  (by setting  $\hat{y}_{d'} = 0$ ) and carry out EMF, the collected values in  $B'_{b_{d'}}$  can only converge to the normal values. The collected values in  $B'_{b_j}$  ( $j \in \{1, \dots, t\}$ ) will converge to both the poison values in  $B_{y_j}$  ( $j \in \{1, \dots, t\}$ ) and the normal values, so:

$$\frac{n_{d'}}{N} + \sum_{j=1}^t \frac{n_j}{N} \leq \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0}.$$

Then suppress the bucket  $B_{y_{d'-1}}$ , similarly, we have:

$$\frac{n_{d'} + n_{d'-1}}{N} + \sum_{j=1}^t \frac{n_j}{N} \leq \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0, y_{d'-1}=0}.$$

Since the collected values in  $B'_{b_{d'-1}}$  can only converge to the normal values and the decrement of collected values in  $B'_{b_j}$  ( $j \in \{1, \dots, t\}$ ) will also converge to the normal values, we have:

$$\left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0} \leq \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0, y_{d'-1}=0}. \quad (13)$$

Suppress all buckets in  $\bar{\mathcal{B}}$ . Similar to Equ. 13, we have:

$$\begin{aligned} \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0} &\leq \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0, y_{d'-1}=0} \leq \\ &\dots \leq \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0, \dots, y_{t+1}=0}, \end{aligned}$$

and we infer that:

$$\frac{\sum_{i=t+1}^{d'} n_i}{N} + \sum_{j=1}^t \frac{n_j}{N} = \left( \sum_{k=1}^d \hat{x}_k + \sum_{j=1}^t \hat{y}_j \right) \Big|_{y_{d'}=0, \dots, y_{t+1}=0} = 1,$$

which is the optimal case where none of the collected values will converge to buckets in  $\bar{\mathcal{B}}$ .

When the number of suppressed buckets in  $\bar{\mathcal{B}}$  increases, the corresponding interference of  $\bar{\mathcal{B}}$  decreases. Therefore, the collected values more accurately converge to the buckets that they should belong to, and thus achieve a better convergence result.  $\square$

## B PROOF OF THEOREM 5.3

PROOF. Let  $v_{tj}$  denote the  $j$ -th value in group  $G_t$ ,  $v'_{tj}$  denote the perturbed  $v_{tj}$ , and  $M_t$  denote the mean value of  $v'_{tj}$ . The variance of  $\tilde{M}$ , which is a linear combination of  $M_t$ , can be written as:

$$\begin{aligned} \text{Var}(\tilde{M}) &= \text{Var} \left( \sum_{t=1}^h w_t M_t \right) = \sum_{t=1}^h w_t^2 \text{Var}(M_t) \\ &= \sum_{t=1}^h w_t^2 \text{Var} \left( \frac{\sum_{j=1}^{n'} v'_{tj}}{n'} \right) = \sum_{t=1}^h \frac{w_t^2}{n'^2} \sum_{j=1}^{n'} \text{Var}(v'_{tj}), \end{aligned} \quad (14)$$

where  $\sum w_t = 1$ , and  $n'$  is the number of normal users in each group.

Since  $\text{Var}(v'_{tj})$  in Equ. 14 relies on the input of each user, we consider the worst-case at the maximum variance, i.e., all inputs  $v_{tj}$  are either 1 or -1. The worst-case variance  $\text{Var}_{\text{worst}}(v'_{tj})$  can be expressed as:

$$\begin{aligned} \text{Var}_{\text{worst}}(v'_{tj}) &= \frac{v_{tj}^2}{e^{\epsilon_t/2} - 1} + \frac{e^{\epsilon_t/2} + 3}{3(e^{\epsilon_t/2} - 1)^2} \Big|_{v_{tj}=\pm 1} \\ &= \frac{1}{e^{\epsilon_t/2} - 1} + \frac{e^{\epsilon_t/2} + 3}{3(e^{\epsilon_t/2} - 1)^2}. \end{aligned}$$

Let  $B_t = n' \text{Var}_{\text{worst}}(v'_{tj})$ . Equ. 14 can be rewritten as:

$$\text{Var}(\tilde{M}) = \sum_{t=1}^h \frac{w_t^2}{n'^2} B_t. \quad (15)$$

We regard the variance as a function of  $w_t$ , and the minimal variance is the extreme point of Equ. 15. By the Lagrangian method, we have:

$$\mathcal{L} = \sum_{t=1}^h \frac{w_t^2}{n'^2} B_t + C_0 \left( 1 - \sum_{t=1}^h w_t \right).$$

The first partial derivatives of  $\mathcal{L}$  w.r.t.  $w_t$  is:

$$\frac{\partial \mathcal{L}}{\partial w_t} = \frac{2w_t}{n'^2} B_t - C_0. \quad (16)$$

Let  $\frac{\partial \mathcal{L}}{\partial w_t} = \frac{2w_t}{n'^2} B_t - C_0 = 0$ , then we have  $w_t = \frac{C_0 n'^2}{2B_t}$ . Through the restriction  $\sum_{t=1}^h w_t = 1$ , we figure out

$$C_0 = \frac{2}{n'^2 \sum_{t=1}^h \frac{1}{B_t}}, \quad w_t = \frac{1}{B_t \sum_{i=1}^h \frac{1}{B_i}}.$$

And the final minimal variance of  $\tilde{M}$  is:

$$\text{Var}(\tilde{M})_{\min} = \left[ \sum_{t=1}^h \frac{n'^2}{B_t} \right]^{-1}.$$

$\square$