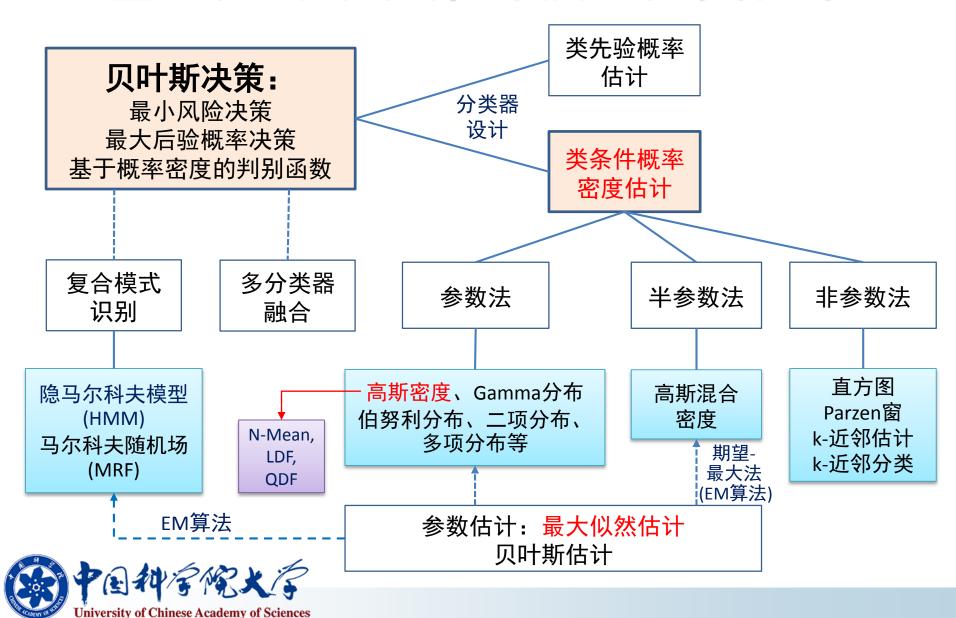
第3章:参数估计(续)

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基于贝叶斯决策的模式分类框架



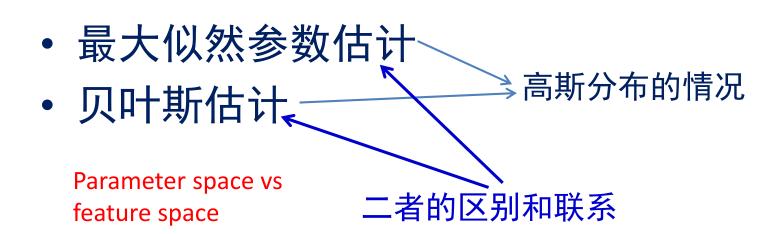
公式太多,怎么办?

- 注重宏观思维
 - 先整体,后局部,再回到整体
 - 理解概念最重要!
 - 特征空间、符号、公式的物理意义,形成直觉
 - 高维空间物理意义如何理解: 简化到低维, 再推广到高维
 - 注重不同方法之间的区别和联系(共性)
 - 理解概念的基础上再去了解细节和数学证明
 - 对主要的方法理解其原理、过程和结论
 - 复杂的数学证明过程可忽略,记住结论即可
 - 简单的情况要清楚细节,如高斯密度函数的最大似然估计求解过程
 - 高斯混合密度的最大似然估计(EM算法)了解主要步骤(E-step, M-step)
 - 低维空间和简单模型能写出详细过程, 高维或复杂模型则不要求
 - 数学分析(形式化)和证明的能力对创新研究很重要,但不可能(没有精力)把所有细节都搞懂
 - 善于利用已有概念、原理和结论,理解和会用是基础



上次课主要内容回顾

- 离散变量贝叶斯决策
- 复合模式分类





提纲

- 第3章:参数估计 (贝叶斯分类的参数法、半参数法)
 - -特征维数问题
 - 期望最大法
 - 隐马尔可夫模型



特征维数问题

- 统计模式分类
 - 特征空间划分
 - 贝叶斯决策:最小风险规则, MAP
- 增加特征有什么好处
 - 判别性: 类别间有差异的特征有助于分类
- 带来什么问题
 - 计算
 - 存储
 - 泛化性能,Overfitting



分类错误率与特征的关系

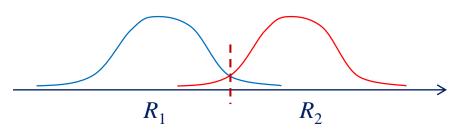
- 二类高斯分布
 - $-p(\mathbf{x}|\omega_i)^{\sim}N(\mu_i,\Sigma), j=1,2$,等协方差矩阵
 - Bayes error rate

$$P(e) = \frac{1}{\sqrt{2\pi}} \int_{r/2}^{\infty} e^{-u^2/2} du \qquad r^2 = (\mu_1 - \mu_2)^t \mathbf{\Sigma}^{-1} (\mu_1 - \mu_2)$$

$$P(error) = P(\mathbf{x} \in \mathcal{R}_2, \omega_1) + P(\mathbf{x} \in \mathcal{R}_1, \omega_2)$$

$$= P(\mathbf{x} \in \mathcal{R}_2 | \omega_1) P(\omega_1) + P(\mathbf{x} \in \mathcal{R}_1 | \omega_2) P(\omega_2)$$

$$= \int_{\mathcal{R}_2} p(\mathbf{x} | \omega_1) P(\omega_1) d\mathbf{x} + \int_{\mathcal{R}_1} p(\mathbf{x} | \omega_2) P(\omega_2) d\mathbf{x}.$$



• 二类高斯分布

- Conditionally independent case $\Sigma = diag(\sigma_1^2, ..., \sigma_d^2)$
 - 每一维二类均值之间距离反映区分度,决定错误率
 - 特征增加有助于减小错误率(r²增大)

$$r^2 = \sum_{i=1}^d \left(\frac{\mu_{i1} - \mu_{i2}}{\sigma_i}\right)^2$$
 $P(e) = \frac{1}{\sqrt{2\pi}} \int_{r/2}^{\infty} e^{-u^2/2} du$

- 特征维数决定可分性的例子
 - 3D空间完全可分
 - 2D和1D投影空间有重叠

然而,增加特征也可能导致分类性能更差, 因为有模型估计误差(wrong model)



计算复杂度

- 最大似然估计
 - 高斯分布, d维特征, n个样本
 - 参数估计的复杂度, 主要由Σ决定

$$g(\mathbf{x}) = -\frac{1}{2} (\mathbf{x} - \hat{\boldsymbol{\mu}})^t \widehat{\boldsymbol{\Sigma}}^{-1} (\mathbf{x} - \hat{\boldsymbol{\mu}}) - \underbrace{\frac{O(1)}{d} \ln 2\pi}_{O(n)} - \underbrace{\frac{O(d^2n)}{1} \ln |\widehat{\boldsymbol{\Sigma}}|}_{O(n)} + \underbrace{\frac{O(n)}{1} \ln |\widehat{$$

• 参数存储复杂度

$$c(d+d(d+1)/2)$$

- 分类复杂度?
 - 计算逆矩阵比较复杂,一般为 $O(d^3)$



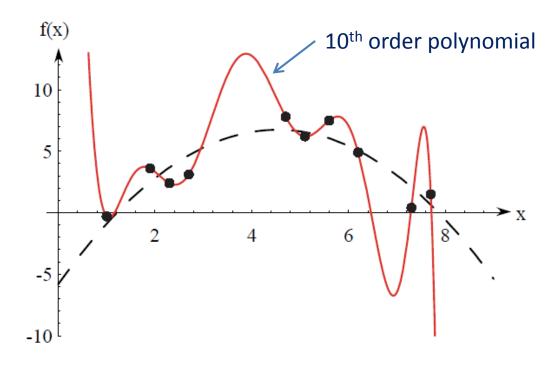
过拟合(Overfitting)

- Overfitting
 - 特征维数高、训练样本少导致模型参数估计不准确
 - 比如协方差矩阵需要样本数在d以上
- 克服办法
 - 特征降维: 特征提取(变换)、特征选择
 - 参数共享/平滑
 - 共享协方差矩阵Σ₀
 - Shrinkage (a.k.a. Regularized Discriminant Analysis)

$$\Sigma_i(\alpha) = \frac{(1-\alpha)n_i\Sigma_i + \alpha n\Sigma}{(1-\alpha)n_i + \alpha n}$$
$$\Sigma(\beta) = (1-\beta)\Sigma + \beta I$$



• 过拟合的例子



$$f(x) = ax^2 + bx + c + \epsilon$$
 where $p(\epsilon) \sim N(0, \sigma^2)$



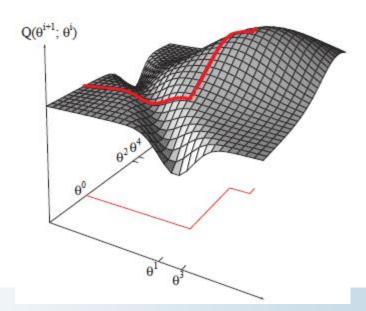
期望-最大法(EM)

- 数据缺失情况下的参数估计
 - Good features, missing/bad features $\mathbf{x}_k = \{\mathbf{x}_{kg}, \mathbf{x}_{kb}\}$

$$\mathcal{D} = \{\mathbf{x}_1, ..., \mathbf{x}_n\} = \mathcal{D}_g \cup \mathcal{D}_b$$

- 已知参数值θⁱ情况下估计新参数值θ
 - 对缺失数据求期望(marginalize)

$$\max Q(\boldsymbol{\theta}; \; \boldsymbol{\theta}^i) = \mathcal{E}_{\mathcal{D}_b}[\ln p(\mathcal{D}_g, \mathcal{D}_b; \; \boldsymbol{\theta}) | \mathcal{D}_g; \; \boldsymbol{\theta}^i]$$



Expectation-Maximization (EM)

Algorithm 1 (Expectation-Maximization)

```
1 begin initialize \theta^{0}, T, i = 0
2 do i \leftarrow i + 1
3 E step: compute Q(\theta; \theta^{i})
5 M step: \theta^{i+1} \leftarrow \arg \max_{\theta} Q(\theta; \theta^{i})
6 until Q(\theta^{i+1}; \theta^{i}) - Q(\theta^{i}; \theta^{i-1}) \leq T
7 return \hat{\theta} \leftarrow \theta^{i+1}
8 end
```

The EM algorithm guarantees that the log-likelihood of good data increases monotonically.



Example: EM for a 2D Gaussian

$$\mathcal{D} \, = \, \{\mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3, \mathbf{x}_4\} \, = \, \left\{ \binom{0}{2}, \binom{1}{0}, \binom{2}{2}, \binom{*}{4} \right\}$$

parameters
$$\theta=\begin{pmatrix} \mu_1\\ \mu_2\\ \sigma_1^2\\ \sigma_2^2 \end{pmatrix}$$
 initially $\theta^0=\begin{pmatrix} 0\\ 0\\ 1\\ 1 \end{pmatrix}$

initially
$$\theta^0 = \begin{pmatrix} 0 \\ 0 \\ 1 \\ 1 \end{pmatrix}$$

$$Q(\theta; \theta^0) = \mathcal{E}_{x_{41}}[\ln p(\mathbf{x}_g, \mathbf{x}_b; \theta | \theta^0; \mathcal{D}_g)]$$

$$= \int_{-\infty}^{\infty} \left[\sum_{k=1}^{3} \ln p(\mathbf{x}_{k}|\boldsymbol{\theta}) + \ln p(\mathbf{x}_{4}|\boldsymbol{\theta}) \right] p(x_{41}|\boldsymbol{\theta}^{0}; \ x_{42} = 4) \ dx_{41}$$

$$= \sum_{k=1}^{3} \left[\ln p(\mathbf{x}_{k}|\boldsymbol{\theta})\right] + \int_{-\infty}^{\infty} \ln p\left(\binom{x_{41}}{4}\right|\boldsymbol{\theta}\right) \frac{p\left(\binom{x_{41}}{4}|\boldsymbol{\theta}^{0}\right)}{\left(\int_{-\infty}^{\infty} p\left(\binom{x'_{41}}{4}|\boldsymbol{\theta}^{0}\right) dx'_{41}\right)} dx_{41}$$

 $Q(\theta; \ \theta^{0}) = \sum_{k=1}^{3} \left[\ln p(\mathbf{x}_{k}|\theta) \right] + \frac{1}{K} \int \ln p\left(\begin{pmatrix} x_{41} \\ 4 \end{pmatrix} \middle| \theta \right) \frac{1}{2\pi \left| \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \right|} \exp\left[-\frac{1}{2} (x_{41}^{2} + 4^{2}) \right] dx_{41}$

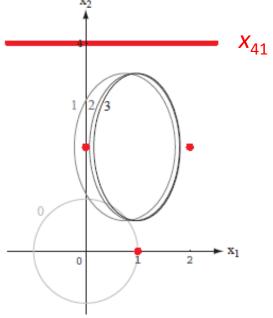
$$= \sum_{k=1}^{3} \left[\ln p(\mathbf{x}_{k}|\boldsymbol{\theta}) \right] - \frac{1 + \mu_{1}^{2}}{2\sigma_{1}^{2}} - \frac{(4 - \mu_{2})^{2}}{2\sigma_{2}^{2}} - \ln (2\pi\sigma_{1}\sigma_{2}).$$



$$\max \sum_{k=1}^{3} [\ln p(\mathbf{x}_{k}|\boldsymbol{\theta})] - \frac{1 + \mu_{1}^{2}}{2\sigma_{1}^{2}} - \frac{(4 - \mu_{2})^{2}}{2\sigma_{2}^{2}} - \ln (2\pi\sigma_{1}\sigma_{2})$$

$$\boldsymbol{\theta}^{1} = \begin{pmatrix} 0.75 \\ 2.0 \\ 0.938 \\ 2.0 \end{pmatrix}$$

After 3 iterations $\mu = \begin{pmatrix} 1.0 \\ 2.0 \end{pmatrix}$, and $\Sigma = \begin{pmatrix} 0.667 & 0 \\ 0 & 2.0 \end{pmatrix}$



 x_{41} unknown

What if $x_4 = (1,4)^T$



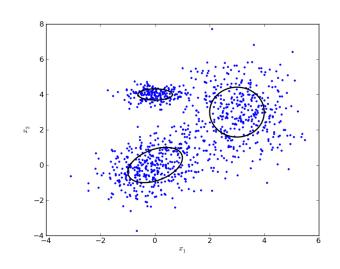
- EM for Gaussian mixture
 - 参数型概率密度函数,可以表示复杂的分布

$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k p(\mathbf{x} \mid \theta_k)$$

subject to
$$\sum_{k=1}^{K} \pi_k = 1$$

Gaussian component

$$p(\mathbf{x} \mid \theta_k) = \mathcal{N}(\mathbf{x} \mid \mu_k, \Sigma_k)$$



- 参数估计: Maximum Likelihood (ML)

$$\max LL = \log \prod_{n=1}^{N} p(\mathbf{x}_n) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k p(\mathbf{x}_n \mid \theta_k)$$
$$\nabla_{\pi_k} LL = 0, \quad \nabla_{\mu_k} LL = 0, \quad \nabla_{\Sigma_k} LL = 0$$

不能解析求解



EM Algorithm for Gaussian mixture

Incomplete data X, complete data {X,Z}

missing
$$z_{nk} \in \{0,1\}, k = 1,...,K$$

Expectation of complete data log-likelihood

$$Q(\Theta, \Theta^{old}) = \sum_{\mathbf{Z}} [\log p(\mathbf{X}, \mathbf{Z} | \Theta)] p(\mathbf{Z} | \mathbf{X}, \Theta^{old})$$

- 1. Choose an initial set of parameters for Θ^{old}
- 2. Do

E-step: Evaluate p(Z|X, Θold)

M-step: Update parameters

$$\Theta^{new} = \arg\max_{\Theta} Q(\Theta, \Theta^{old})$$

If convergence condition is not satisfied

$$\Theta^{old} \leftarrow \Theta^{new}$$

3. End



EM Algorithm for Gaussian mixture

E-step
$$p(\mathbf{X}, \mathbf{Z} \mid \Theta) = \prod_{n=1}^{N} \prod_{k=1}^{K} \pi_k^{z_{nk}} \mathcal{N}(\mathbf{x}_n \mid \mu_k, \Sigma_k)^{z_{nk}}$$

$$Q(\Theta, \Theta^{old}) = \mathbb{E}_{\mathbf{Z}}[\log p(\mathbf{X}, \mathbf{Z} \mid \Theta)] = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \{\log \pi_k + \log \mathcal{N}(\mathbf{x} \mid \mu_k, \Sigma_k)\}$$

$$\gamma(z_{nk}) = P(z_{nk} = 1 \mid \mathbf{x}_n) = \frac{\pi_k N(\mathbf{x}_n \mid \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j N(\mathbf{x}_n \mid \mu_j, \Sigma_j)}$$

M-step
$$\nabla_{\pi_k} Q = 0$$
, $\nabla_{\mu_k} Q = 0$, $\nabla_{\Sigma_k} Q = 0$

$$\mu_k^{new} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{X}_n$$

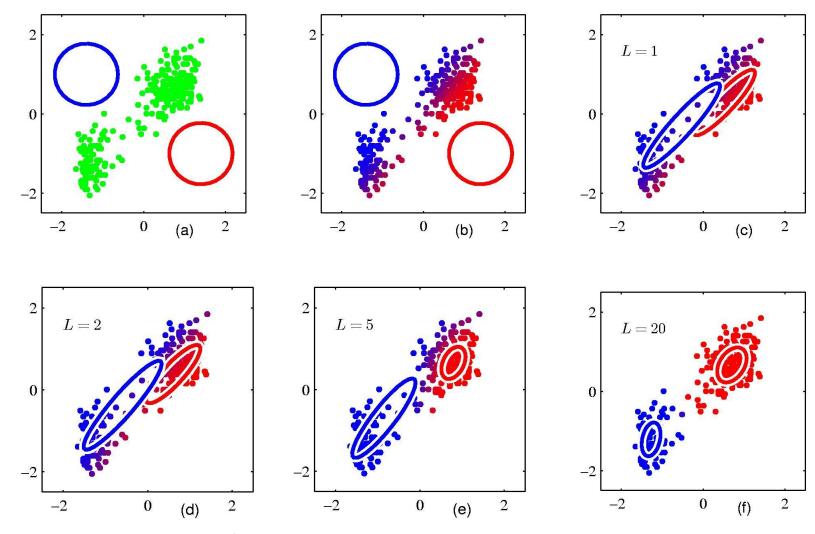
$$\sum_k^{new} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{X}_n - \mu_k^{new}) (\mathbf{X}_n - \mu_k^{new})^T$$

$$\pi_k^{new} = \frac{N_k}{N}$$

$$N_k = \sum_{n=1}^N \gamma(z_{nk})$$
Property of Chinese Academy of Sciences



An example, from (C.M. Bishop, *Pattern Recognition and Machine Learning*, 2006. Figure 9.8)





Break



隐马尔可夫模型

- Sequential (Temporal) Pattern
 - Variable length
 - Distortion



- Ambiguous boundary between primitives (symbols)
- Bayesian Classification
 - Sequence of patterns (observations) $\mathbf{O} = O_1 O_2 \cdots O_T$
 - Sequence class (states) $\mathbf{q} = q_1 q_2 \cdots q_T$
 - Posterior probability $P(\mathbf{q} \mid \mathbf{O}) = \frac{p(\mathbf{O} \mid \mathbf{q})P(\mathbf{q})}{p(\mathbf{O})}$
- Hidden Markov Model (HMM)
 - Model $p(\mathbf{O}|\mathbf{q}), p(\mathbf{O},\mathbf{q})$



Markov Chain

Sequence of States (classes)

$$P(q_1 q_2 \cdots q_T) = P(q_1) P(q_2 | q_1) P(q_3 | q_1 q_2) \cdots P(q_T | q_1 \cdots q_{T-1})$$

$$q_t \in \{S_1, ..., S_N\}$$

First-Order Markov chain

$$P(q_{t} = S_{j} / q_{t-1} = S_{i}, q_{t-2} = S_{k}, \dots) = P(q_{t} = S_{j} / q_{t-1} = S_{i})$$

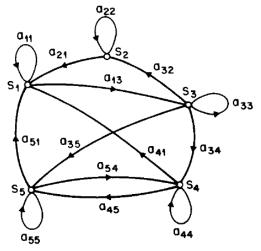
$$P(q_{1}q_{2} \dots q_{T}) = 0$$

$$P(q_1)P(q_2 | q_1)P(q_3 | q_2)\cdots P(q_T | q_{T-1})$$

State transition probabilities

$$a_{ij} = P(q_t = S_j | q_{t-1} = S_i), \quad 1 \le i, j \le N$$

$$\sum_{j=1}^{N} a_{ij} = 1$$





State duration (self-transition)

$$O = \{S_i, S_i, S_i, ..., S_i, S_j \neq S_i\}$$
1 2 3 d d+1
$$P(O | \text{Model}, q_1 = S_i) = (a_{ii})^{d-1} (1 - a_{ii}) = p_i(d)$$

Expected duration of specific state

$$\overline{d}_{i} = \sum_{d=1}^{\infty} d p_{i}(d)$$

$$= \sum_{d=1}^{\infty} d (a_{ii})^{d-1} (1 - a_{ii}) = \frac{1}{1 - a_{ii}}$$

Example: Transition of Weather

State 1: rain or (snow)
State 2: cloudy
A =
$$\{a_{ij}\}$$
 = $\begin{bmatrix} 0.4 & 0.3 & 0.3 \\ 0.2 & 0.6 & 0.2 \\ 0.1 & 0.1 & 0.8 \end{bmatrix}$

Expected number of days for sunny and cloudy?



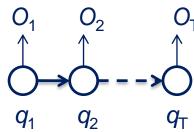
Hidden Markov Model (HMM)

- Markov Chain: States are Observable
- Hidden States: An Example
 - Imagine you are in a un-windowed room, cannot see the weather outside. Instead, you can guess the weather from the temperature and humidity in room
 - Observations: temperature, humidity
 - Hidden states: weather
 - Hidden Markov Model (HMM): Doubly embedded
 stochastic process

$$P(O_1,O_2,...,O_T)$$

 $P(q_1,q_2,...,q_T)$

Infer states from observations



$$q_i \in \{S_1, S_2, \dots, S_N\}$$



Elements of an HMM

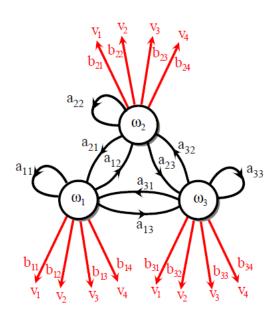
$$\lambda = (\mathbf{A}, \mathbf{B}, \boldsymbol{\pi})$$

- N: number of states in the model, $S=\{S_1, S_2, ..., S_N\}$
- M: number of observation symbols, $V=\{v_1, v_2, ..., v_M\}$
- State transition probability distribution $A = \{a_{ij}\}$

$$a_{ij} = P(q_{t+1} = S_j | q_t = S_i), \quad 1 \le i, j \le N$$

- Observation symbol (emission) probability distribution $B=\{b_j(k)\}$ $b_j(k)=P(v_k \text{ at } t \mid q_t=S_j), \ 1 \le j \le N, \ 1 \le k \le M$
- Initial state distribution $\pi = {\pi_i}$

$$\pi_i = P(q_1 = S_i), 1 \le i \le N$$





Three Basic Problems of HMM

– Problem 1 (Evaluation):

How to efficiently compute the probability of observation sequence $P(O|\lambda)$

– Problem 2 (Decoding):

How to choose the best state sequence responding to an observation sequence

– Problem 3 (Training):

How to estimate the model parameters

Evaluation Problem

- Given model $\lambda = (\mathbf{A}, \mathbf{B}, \pi)$ and observation sequence $O = O_1 O_2 ... O_T$, compute $P(O \mid \lambda)$
 - Direct computation

$$P(O \mid \lambda) = \sum_{all \ Q} P(O \mid Q, \lambda) P(Q \mid \lambda)$$

$$= \sum_{q_1, q_2, \dots, q_T} \pi_{q_1} b_{q_1}(O_1) a_{q_1 q_2} b_{q_2}(O_2) \cdots a_{q_{T-1} q_T} b_{q_T}(O_T)$$

Conditional independence
$$P(O \mid Q, \lambda) = \prod_{t=1}^{T} P(O_t \mid q_t, \lambda) \qquad P(Q \mid \lambda) = \pi_{q_1} a_{q_1 q_2} a_{q_2 q_3} \cdots a_{q_{T-1} q_T} \\ = b_{q_1}(O_1) b_{q_2}(O_2) \cdots b_{q_T}(O_T) \qquad \text{Markov chain of states}$$

– Complexity: O(2TN^T)!



Evaluation: Forward Procedure

Define forward variable

$$\alpha_t(i) = P(O_1 O_2 \cdots O_t, q_t = S_i \mid \lambda)$$

- Initialization $\alpha_1(i) = \pi_i b_i(O_1), 1 \le i \le N$
- Induction

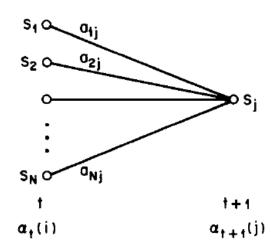
$$\alpha_{t+1}(j) = \left[\sum_{i=1}^{N} \alpha_{t}(i) a_{ij}\right] b_{j}(O_{t+1}),$$

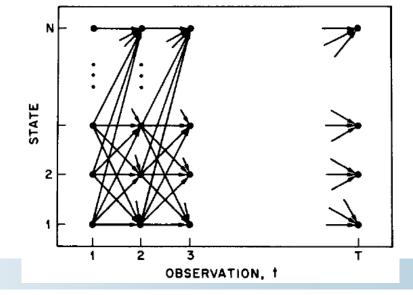
$$1 \le t \le T - 1, \qquad 1 \le j \le N.$$

Termination

$$P(O \mid \lambda) = \sum_{i=1}^{N} \alpha_{T}(i)$$

Complexity: O(TN²)







Evaluation: Backward Procedure

Define backward variable

$$\beta_{t}(i) = P(O_{t+1}, ..., O_T | q_t = S_i, \lambda)$$

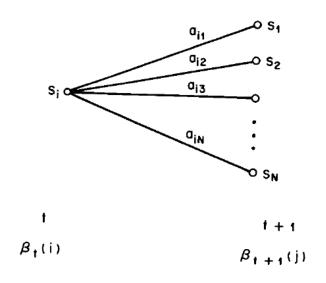
Initialization

$$\beta_T(i) = 1, \quad 1 \le i \le N$$

Induction

$$\beta_{t}(i) = \sum_{j=1}^{N} a_{ij} b_{j}(O_{t+1}) \beta_{t+1}(j),$$

$$1 \le t \le T - 1, \quad 1 \le i \le N$$



- Termination

$$P(O \mid \lambda) = \sum_{i=1}^{N} \pi_i b_i(O_1) \beta_1(i) = \sum_{i=1}^{N} \alpha_1(i) \beta_1(i)$$

– Complexity?

Decoding Problem

- This is Pattern Recognition
- Optimal Sequence of States

$$\max_{q_1q_2\cdots q_T} P(q_1q_2\cdots q_T \mid O, \lambda) = \max_{q_1q_2\cdots q_T} P(q_1q_2\cdots q_T, O \mid \lambda)$$

- Viterbi Algorithm (DP: dynamic programming)
 - Define variable

$$\delta_t(i) = \max_{q_1, q_2, \dots, q_{t-1}} P(q_1 q_2 \dots q_t = i, O_1 O_2 \dots O_t / \lambda)$$

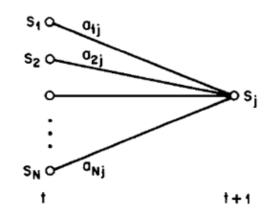
- DP

$$\delta_{t+1}(j) = \left[\max_{i} \delta_{t}(i) a_{ij} \right] \cdot b_{j}(O_{t+1})$$

Initialization

$$\delta_1(i) = \pi_i b_i(O_1), \quad 1 \le i \le N$$

$$\psi_1(i) = 0.$$



$$\delta_t(i) = \max_{q_1, q_2, \dots, q_{t-1}} P(q_1 q_2 \cdots q_t = i, O_1 O_2 \cdots O_t \mid \lambda)$$

Viterbi Algorithm (Cont.)

Recursion

$$\delta_{t}(j) = \left[\max_{1 \leq i \leq N} \delta_{t-1}(i) a_{ij}\right] b_{j}(O_{t}),$$

$$\psi_{t}(j) = \arg\max_{1 \leq i \leq N} \delta_{t-1}(i) a_{ij},$$

$$2 \leq t \leq T, \quad 1 \leq j \leq N$$

- Termination

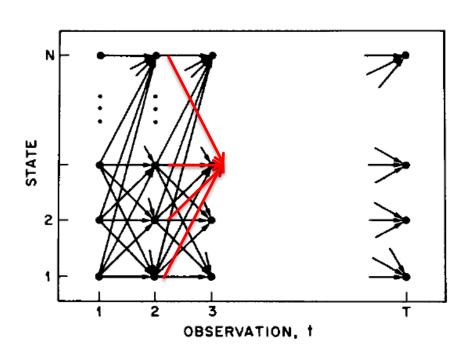
$$P^* = \max_{1 \le i \le N} \delta_T(i)$$

$$q_T^* = \arg\max_{1 \le i \le N} \delta_T(i)$$

Backtracking

$$q_t^* = \psi_{t+1}(q_{t+1}^*),$$

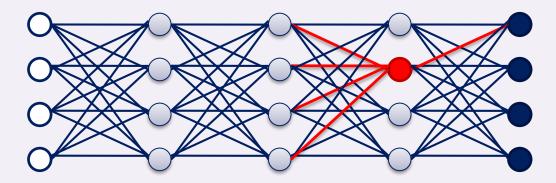
 $1 \le t \le T - 1$



Complexity: $O(TN^2)$



- Appendix: Dynamic Programming (DP) Principle (Bellman Principle of Optimality)
 - The best path through a particular, intermediate place is the best way from start to it, followed by the best way from it to the goal.
 - Implication: from multiple ways reaching an intermediate place, only retain the best one
 - Often used in sequence matching and HMMs



Training Problem

- $\max P(O | \lambda)$ Maximum Likelihood (ML)
- Baum-Welch Algorithm (EM)

$$\max_{\overline{\lambda}} Q(\lambda, \overline{\lambda}) = \sum_{\mathcal{Q}} [\log P(Q, O \,|\, \overline{\lambda})] P(Q, O \,|\, \lambda)$$
 - Define variable
$$\xi_t(i, j) = P(q_t = S_i, q_{t+1} = S_j \,|\, O, \lambda)$$

$$\xi_{t}(i, j) = P(q_{t} = S_{i}, q_{t+1} = S_{j} \mid O, \lambda)$$

$$\xi_{t}(i,j) = \frac{\alpha_{t}(i) a_{ij} b_{j}(O_{t+1}) \beta_{t+1}(j)}{P(O|\lambda)} P(O, q_{t} = S_{i}, q_{t+1} = S_{j} | \lambda)$$

$$= \frac{\alpha_{t}(i) a_{ij} b_{j}(O_{t+1}) \beta_{t+1}(j)}{\sum_{i=1}^{N} \sum_{j=1}^{N} \alpha_{t}(i) a_{ij} b_{j}(O_{t+1}) \beta_{t+1}(j)} \sum_{i=1}^{N} \frac{\alpha_{t}(i) a_{ij} b_{j}(O_{t+1}) \beta_{t+1}(j)}{\sum_{i=1}^{N} \sum_{j=1}^{N} \alpha_{t}(i) a_{ij} b_{j}(O_{t+1}) \beta_{t+1}(j)}$$

|a_{ij}b_j(0₁₊₁)

Define probability

$$\gamma_{t}(i) = P(q_{t} = S_{i} \mid O, \lambda) = \frac{\alpha_{t}(i)\beta_{t}(i)}{P(O \mid \lambda)} = \frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{i=1}^{N} \alpha_{t}(i)\beta_{t}(i)} = \sum_{j=1}^{N} \xi_{t}(i, j)$$

Baum-Welch Algorithm (Cont.)

Reestimation formulas

 $\bar{\pi}_i$ = expected frequency (number of times) in state S_i at time (t=1)= $\gamma_1(i)$

$$\overline{a_{ij}} = \frac{\text{expected number of transitions from state } S_i \text{ to state } S_j}{\text{expected number of transitions from state } S_i}$$

$$= \frac{\sum_{t=1}^{T-1} \xi_t(i,j)}{\sum_{t=1}^{T-1} \gamma_t(i)} = \frac{\sum_{t=1}^{T-1} \alpha_t(i) a_{ij} b_j(O_{t+1}) \beta_{t+1}(j)}{\sum_{t=1}^{T-1} \alpha_t(i) \beta_t(i)}$$

$$\overline{b}_{j}(k) = \frac{\text{expected number of times in state } j \text{ and observing symbol } v_{k}}{\text{expected number of times in state } j}$$

$$= \frac{\sum_{t=1, \text{ s.t. } O_t = v_k}^{T-1} \gamma_t(j)}{\sum_{t=1}^{T-1} \gamma_t(j)} = \frac{\sum_{t=1, \text{ s.t. } O_t = v_k}^{T} \alpha_t(j) \beta_t(j)}{\sum_{t=1}^{T} \alpha_t(j) \beta_t(j)}$$

Proposition

Total Proposition

**Total Propo



Continuous Density HMM

- Handling Continuous Observations
 - Continuous features: vector \boldsymbol{O}_{T}
 - Discretization: vector quantization (VQ)
 - Each vector replaced with its closest codevector, which is viewed as a symbol
 - Small codebook: distortion
 - Large codebook: large data required in emission probability estimation
 - Continuous emission density: Gaussian mixture (GM)

$$b_{j}(O) = \sum_{m=1}^{M} c_{jm} \mathcal{N}(O; \mu_{jm}, U_{jm}), \quad 1 \leq j \leq N$$

Parameter Estimation of Continuous HMM (omitted)

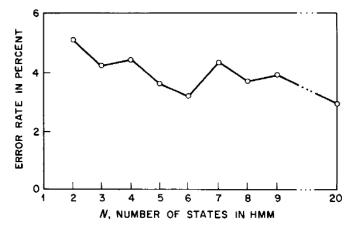


Application to Speech Recognition

- Isolated Word Recognition
 - Given HMM λ^{v} for each word in the vocabulary
 - Input observation sequence O, Bayes decision (assuming equal prior probabilities)

$$v^* = \arg \max_{1 \le v \le V} P(\lambda^v \mid O) = \arg \max_{1 \le v \le V} P(O \mid \lambda^v)$$

- Acoustic features (O_t) (details omitted)
- Vector quantization (discrete observation symbols)
- Choice of model parameters
 - Number of states: empirical (crossvalidation), can be equal for all word models
 - Number of components in GM



Extensions of HMM

- Hybrid HMM/Neural
 - HMM: parametric $b_j(O_t)=p(O_t|q_t=S_j)$, conditional independence
 - Neural: discriminative emission probability $p(q_t=S_i|O_t)$
 - Neural network outputs approximate posterior probabilities
 - Replace $p(O_t|q_t)$ with $p(q_t|O_t)/p(q_t)$

$$\frac{p(O_t \mid q_t)}{p(O_t)} = \frac{p(q_t \mid O_t)}{P(q_t)}$$

- ANN may input multiple frames to learn the correlation

$$p(q_t = S_i | ...O_{t-1}O_tO_{t+1}...)$$

Latest: deep neural networks



讨论

- 特征维数与过拟合
 - 克服过拟合的方法?
- 期望最大法(EM)
 - 对数似然度对缺失数据的期望
 - EM for Gaussian mixture
- 隐马尔可夫模型(HMM)
 - Three basic problems
 - Viterbi Algorithm
 - Extensions



下次课内容

- 第4章 非参数法
 - 密度估计
 - Parzen窗方法
 - K近邻估计
 - 最近邻规则
 - 距离度量
 - Reduced Coulomb Energy Network
 - Approximation by Series Expansion

