

Optimal Criterion

Payoff Matrix

		Response	
		Yes	No
Stimulus	S+N	V_{S+N}^{Yes}	V_{S+N}^{No}
	N	V_N^{Yes}	V_N^{No}

Optimal Criterion

		Response	
Stimulus	S+N	Yes	No
	N	V_N^{Yes}	V_N^{No}

$$E(Yes | x) = V_{S+N}^{Yes} p(S + N | x) + V_N^{Yes} p(N | x)$$

$$E(No | x) = V_{S+N}^{No} p(S + N | x) + V_N^{No} p(N | x)$$

Say yes if $E(Yes | x) \geq E(No | x)$

Optimal Criterion

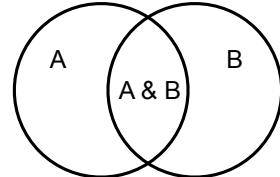
$$E(Yes | x) = V_{S+N}^{Yes} p(S + N | x) + V_N^{Yes} p(N | x)$$

$$E(No | x) = V_{S+N}^{No} p(S + N | x) + V_N^{No} p(N | x)$$

Say yes if $E(Yes | x) \geq E(No | x)$

$$\text{Say yes if } \frac{p(S + N | x)}{p(N | x)} \geq \frac{V_N^{No} - V_N^{Yes}}{V_{S+N}^{Yes} - V_{S+N}^{No}} = \frac{V(\text{Correct} | N)}{V(\text{Correct} | S + N)}$$

Aside: Bayes' Rule



$$p(A | B) = \text{probability of } A \text{ given that } B \text{ is asserted to be true} = \frac{p(A \& B)}{p(B)}$$

$$p(A \& B) = p(B)p(A | B)$$

$$= p(A)p(B | A)$$

$$\Rightarrow p(A | B) = \frac{p(B | A)p(A)}{p(B)}$$

Apply Bayes' Rule

$$p(S + N | x) = \frac{p(x | S + N)p(S + N)}{p(x)}$$

Prior

Likelihood

Posterior

Nuisance normalizing term

$$p(N | x) = \frac{p(x | N)p(N)}{p(x)}, \text{ hence}$$

$$\frac{p(S + N | x)}{p(N | x)} = \left(\frac{p(x | S + N)}{p(x | N)} \right) \left(\frac{p(S + N)}{p(N)} \right)$$

Prior odds

Posterior odds

Likelihood ratio

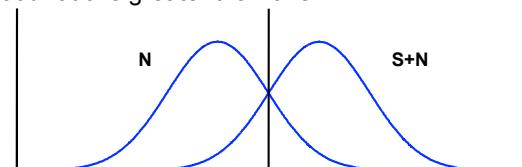
Prior odds

Optimal Criterion

$$\text{Say yes if } \frac{p(S + N | x)}{p(N | x)} \geq \frac{V(\text{Correct} | N)}{V(\text{Correct} | S + N)}$$

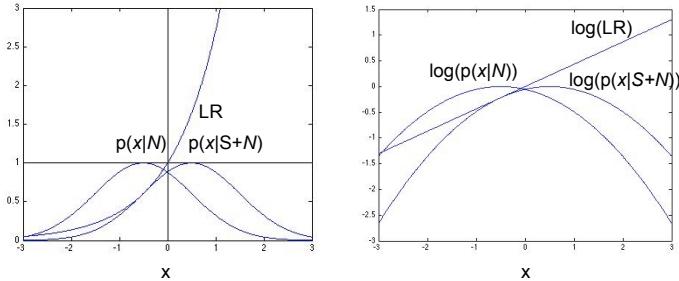
$$\text{i.e., if } \frac{p(x | S + N)}{p(x | N)} \geq \frac{p(N)}{p(S + N)} \frac{V(\text{Correct} | N)}{V(\text{Correct} | S + N)} = \beta$$

Example, if equal priors and equal payoffs, say yes if the likelihood ratio is greater than one:

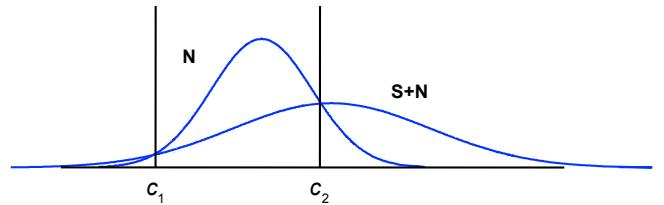


Likelihood Ratio

Optimal binary decisions are a function of the likelihood ratio (it is a sufficient statistic, i.e., no more information is needed or useful).



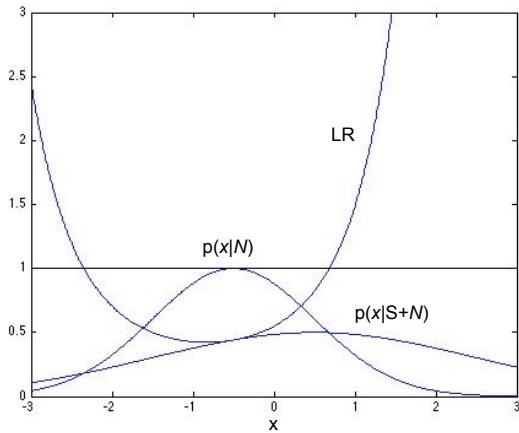
Gaussian Unequal Variance



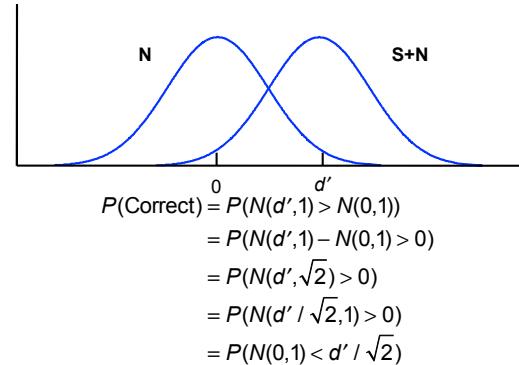
SDT (suboptimal): Say yes if $x \geq c_2$

Optimal strategy: Say yes if $x \geq c_2$ or $x \leq c_1$

Gaussian Unequal Variance

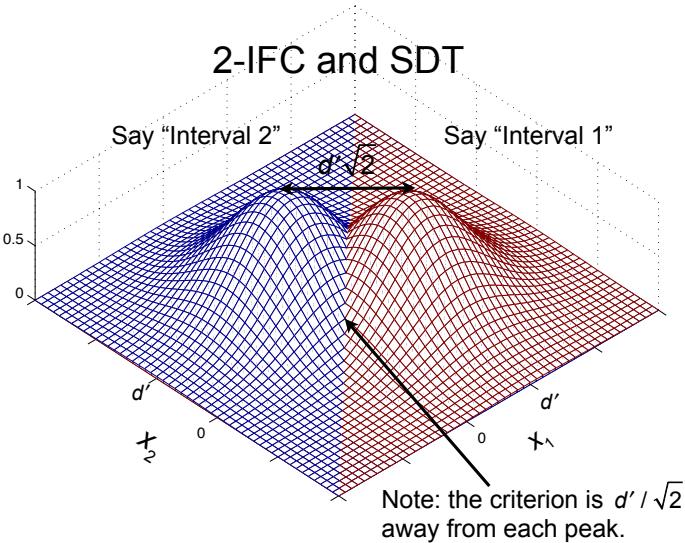


2-IFC and SDT

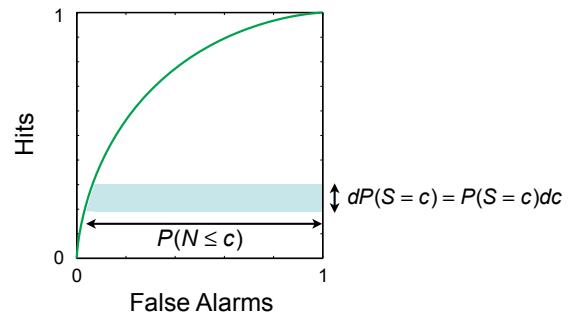


Where $N(\mu, \sigma)$ means a random variable with mean μ and SD σ .

2-IFC and SDT

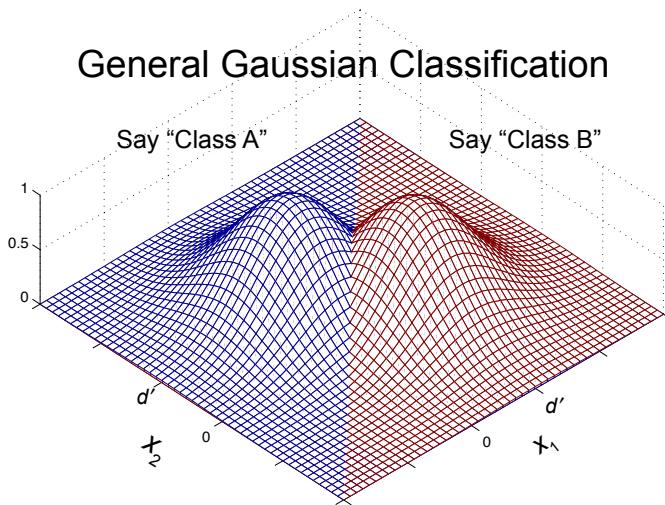


2-IFC and the Yes-No ROC



$$\begin{aligned} P(\text{Correct in 2-IFC}) &= \int P(S = c)P(N \leq c)dc \\ &= \text{Area under the ROC} \end{aligned}$$

General Gaussian Classification



General Gaussian Classification

The d -dimensional Gaussian distribution:

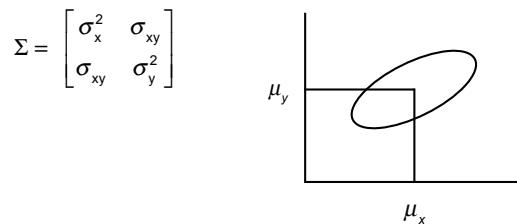
$$p(\bar{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (\bar{x} - \bar{\mu})^T \Sigma^{-1} (\bar{x} - \bar{\mu}) \right]$$

Σ = the covariance matrix

General Gaussian Classification

In 2 dimensions: Mahalanobis distance

$$p(\bar{x}) = \frac{1}{2\pi |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (\bar{x} - \bar{\mu})^T \Sigma^{-1} (\bar{x} - \bar{\mu}) \right]$$



General Gaussian Classification

Say "Category A" if

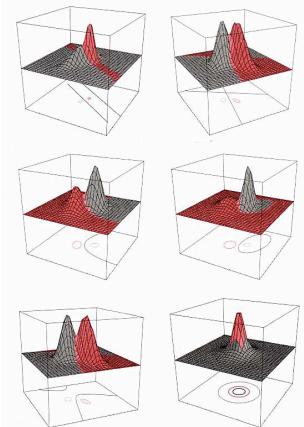
$$\frac{1}{(2\pi)^{d/2} |\Sigma_A|^{1/2}} \exp \left[-\frac{1}{2} (\bar{x} - \bar{\mu}_A)^T \Sigma_A^{-1} (\bar{x} - \bar{\mu}_A) \right] > \frac{1}{(2\pi)^{d/2} |\Sigma_B|^{1/2}} \exp \left[-\frac{1}{2} (\bar{x} - \bar{\mu}_B)^T \Sigma_B^{-1} (\bar{x} - \bar{\mu}_B) \right]$$

Take log and simplify. Say "Category A" if

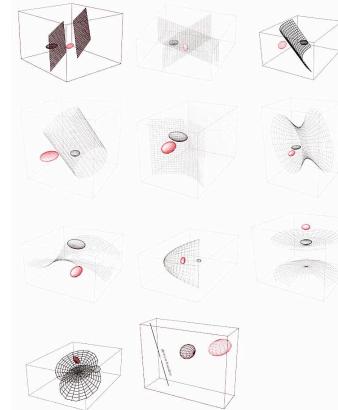
$$(\bar{x} - \bar{\mu}_B)^T \Sigma_B^{-1} (\bar{x} - \bar{\mu}_B) - (\bar{x} - \bar{\mu}_A)^T \Sigma_A^{-1} (\bar{x} - \bar{\mu}_A) \\ = MD(\bar{x}, \bar{\mu}_B) - MD(\bar{x}, \bar{\mu}_A) > C$$

which is a quadratic in \bar{x} .

General Gaussian Classification



General Gaussian Classification

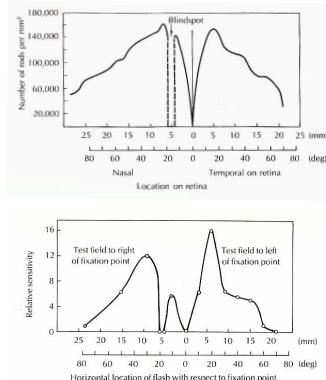


Application: Absolute Threshold

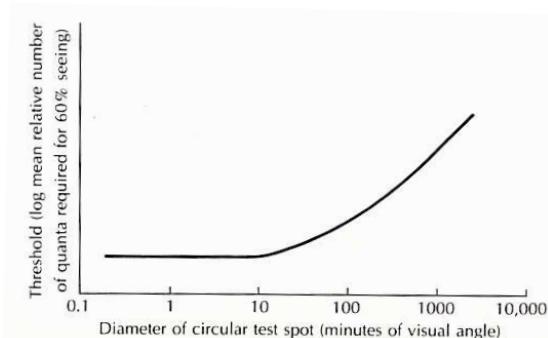
Question: What is the minimal number of photons required for a stimulus to be visible under the best of viewing conditions?

Hecht, Schlaer & Pirenne (1942)

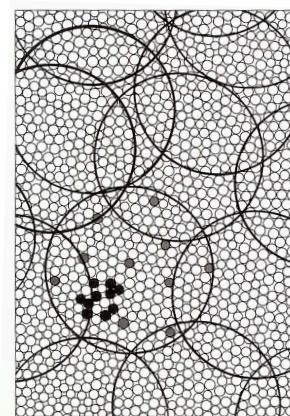
Sensitivity by Eccentricity



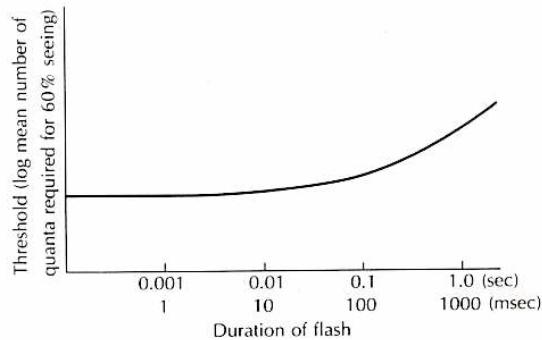
Spatial Summation



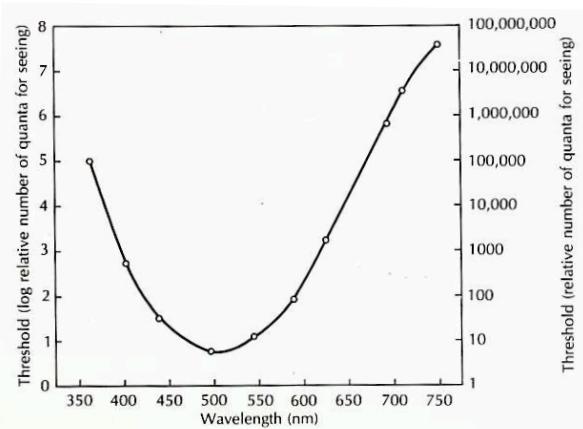
Spatial Summation



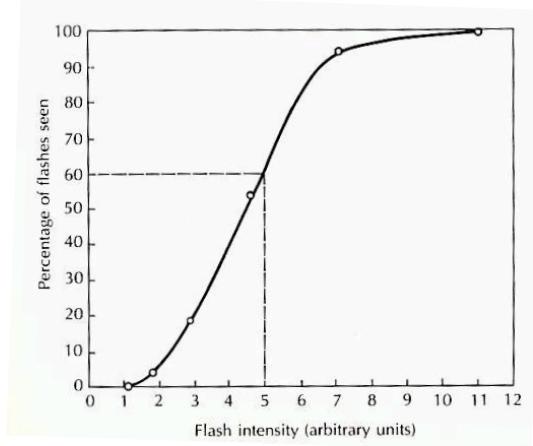
Temporal Summation



Choice of Wavelength



The Data



Poisson Distribution

Counting distribution for processes consisting of “events” that occur randomly at a fixed rate λ events/sec.

The time until the next event occurs follows an exponential distribution:

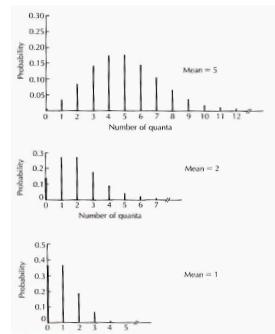
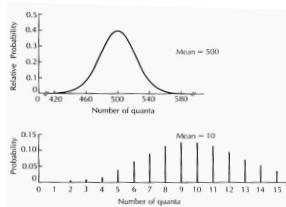
$$p(t) = \lambda e^{-\lambda t}$$

Since λ is the rate, then over a period of τ seconds, the expected count is $\lambda \tau$. The distribution of event count is Poisson:

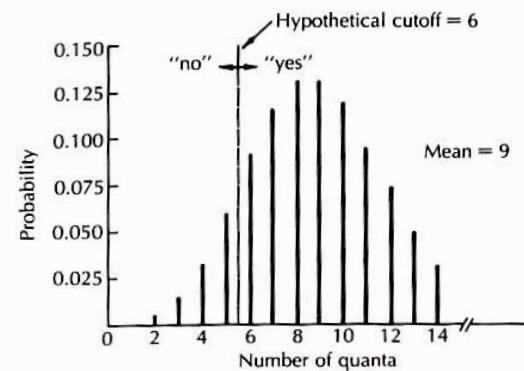
$$p(k) = \frac{e^{-\lambda \tau} (\lambda \tau)^k}{k!}$$

Stimulus Variability (Poisson)

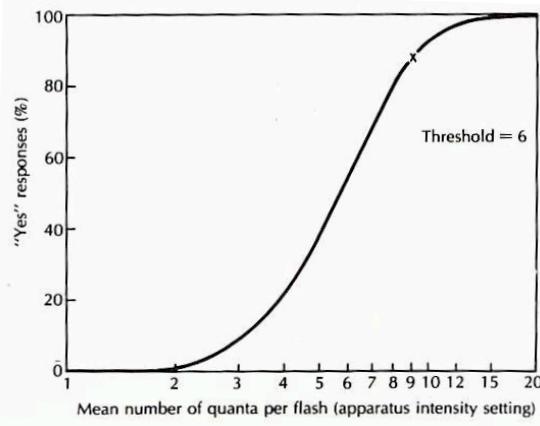
$$p(k) = \frac{e^{-\lambda \tau} (\lambda \tau)^k}{k!}$$



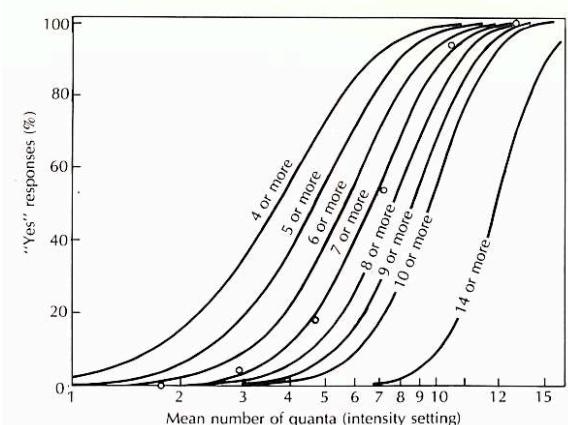
Model with no Subject Variability



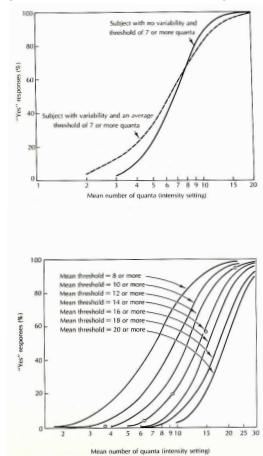
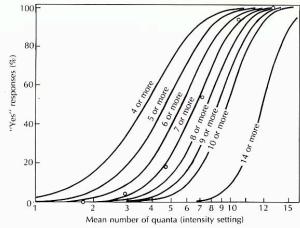
Model with no Subject Variability



Model with no Subject Variability



Model with Subject Variability



Bayes classifier: ideal observer

Question: What is the minimal contrast required to carry out various detection and discrimination tasks?

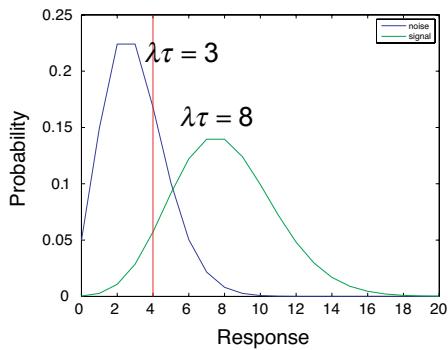
Geisler (1989)
See also Wandell Appendix 3

Examples:

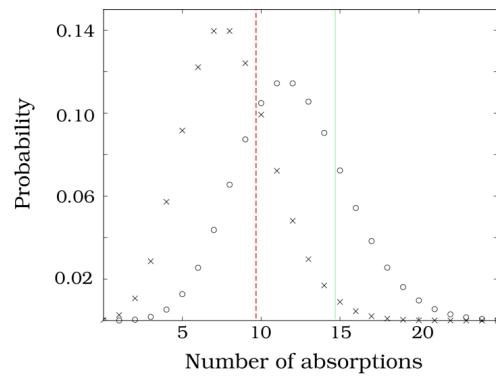
- detecting a light
- distinguishing two spatial patterns

Poisson Distribution SDT

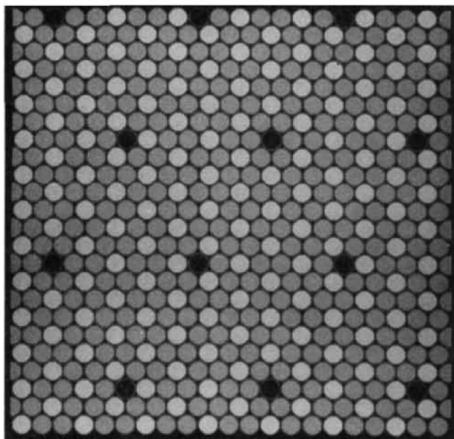
$$p(k) = \frac{e^{-\lambda\tau}(\lambda\tau)^k}{k!} \quad LR(k) \propto \left(\frac{\lambda_{S+N}}{\lambda_N} \right)^k$$



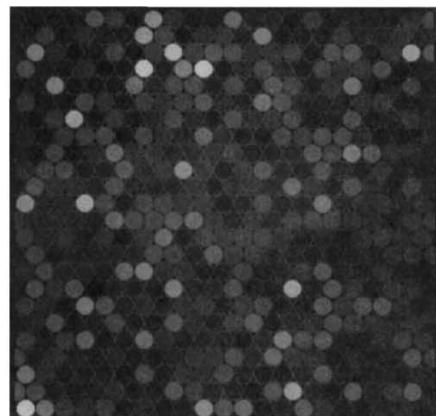
Optimal Criterion



Idealized Receptor Lattice



Poisson Statistics



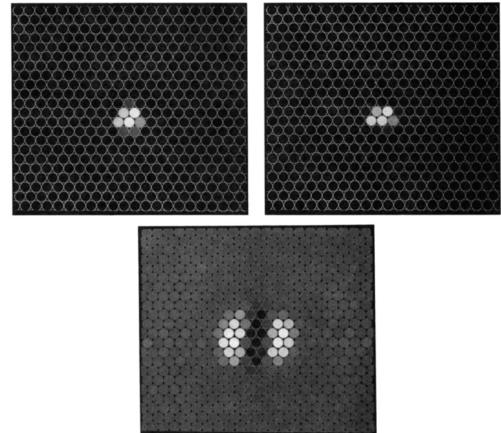
Ideal Discriminator

Given two known possible targets A and B with expected photon absorptions a_i and b_i and actual photon catches Z_i in receptor i , respectively, calculate the likelihood ratio $p(Z_i | a)/p(Z_i | b)$, take the log, do some algebra and discover the following quantity is monotonic in likelihood ratio:

$$Z = \sum_i Z_i \ln(a_i / b_i)$$

Decisions are thus based on an “ideal receptive field.”

Example: 2-Point Resolution



Bayes classifier for spatial discrimination

