

# Central Limit Theorem in Complete Feedback Games

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# Lady tasting tea experiment ( 1920)



Ronald Fisher



Milk first or  
Tea first



Muriel Bristol

## 5. Statement of Experiment

A LADY declares that by tasting a cup of tea made with milk she can discriminate whether the milk or the tea infusion was first added to the cup. We will consider the problem of designing an experiment by means of which this assertion can be tested. For this purpose

The Design of Experiments: Chapter II



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$\#\{\text{Number of correct guesses}\} \sim \text{Hypergeometric}(N = 8, K = 4, n = 4) .$

- $N = 8$  Population size
- $K = 4$  Number of success states
- $n = 4$  Number of draws



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## Optimal Strategy and Score

Knowing that  $x$  cups of Type  $T$  and  $y$  cups of type  $M$  are remaining, she should guess the type corresponding to  $\max(T, M)$ .





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Knowing that  $x$  cups of Type  $T$  and  $y$  cups of type  $M$  are remaining, she should guess the type corresponding to  $\max(T, M)$ . With this strategy, she can make  $373/70 = 5.3$  correct guesses (on average). ( **Diaconis and Graham '81**)



## Complete feedback games: Setup

- Consider a well-shuffled deck of cards of  $n$ -types:  $1, 2, \dots, n$ .
- Number of cards of type  $i$ :  $m$ .
- Total number of cards:  $n \times m$ .



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## Zener Cards

$$n = 5, \quad m = 5.$$



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## Rule of the game

- The player guesses the **type of the topmost card** in the deck.
- The player is **shown** the card and the **card is removed** from the deck.
- The game **continues till the deck is exhausted**.



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# Complete feedback games: General Setup

- Consider a well-shuffled deck of cards of  $n$ -types:  $1, 2, \dots, n$ .
- Number of cards of type  $i$ :  $m_i$ .
- Total number of cards:  $\sum_{i=1}^n m_i$ .

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## The simplest case: $m = 1$

- The probability of guessing the first card correctly:  $\frac{1}{n}$ .
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$$\mathbb{E}[T_{1,n}] = 1 + \frac{1}{2} + \dots + \frac{1}{n} \sim \log(n) .$$

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## Gaussian fluctuation

$$\frac{T_{1,n} - \log(n)}{\log(n)} \xrightarrow{n \rightarrow \infty} \mathcal{N}(0, 1) .$$



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  - Hospital may rule out some subjects on the grounds of medical conditions.
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  - Fixed  $n$  and large  $m$ :

$$\mathbb{E}[T_{n,m}] \sim m + \frac{\pi}{2}M_n\sqrt{m} + o_n(\sqrt{m}),$$

where  $M_n$  is the expected value of  $n$  i.i.d. standard Gaussian.





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- Similar asymptotics for fixed  $n$  and  $\mathbf{m} = (m_1, \dots, m_n)$ .



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- Diaconis and Graham (1981):
  - When  $n = 2$  and  $m_1 = m_2 = m$ , the total score  $T_{2,m}$  satisfies Rayleigh limit law:

$$\mathbb{P}\left(\frac{T_{2,m} - m}{\sqrt{m/2}} \leq x\right) \rightarrow 1 - e^{\frac{-x^2}{2}} \quad x \geq 0.$$

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- This case behaves **very differently**. In particular,

$$\mathbb{P} (T_{2,(m_1,m_2)} - \max\{m_1, m_2\} = k) \rightarrow \gamma(1 - \gamma)^k, \quad \text{where } \gamma = \frac{2|p - q|}{1 + |p - q|}.$$



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  - More precise understanding of  $T_{2,(m_1,m_2)}$ .



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  - In particular, Rayleigh limit law holds as long as  $m_1 - m_2 = o(m_1)$ .
  - Phase transition for the limit law of  $T_{2,(m_1,m_2)}$ .



## A Brief History-II

- Diaconis, Graham, He, Spiro (2022): For fixed  $m$  as  $n \rightarrow \infty$

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$$\mathbb{E}[T_{n,\mathbf{m}}] \sim H_{m^*} H_n + \sum_{j=1}^{m^*} \log(\gamma_j) + O\left(\log(n) \left(\frac{\log(n)}{n}\right)^{1/m^*}\right) ,$$

where  $m^* = \max_{i=1}^n m_i$  and  $\gamma_j = \left(\frac{1}{n} \sum_{i=1}^n \binom{m_i}{j}\right)^{1/j}$ .



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- This suggests a phase transition. Likely at around  $\log n \sim m$ .



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- Pretend that  $X_i$ s are independent and normally distributed with correct mean and variance. Then,

$$\begin{aligned} \mathbb{E}(T_{m,n}) &= \sum_{t=1}^{mn} \frac{\mathbb{E}(\max_i X_i(t))}{t} \\ &\approx m + \sqrt{2m \log n} \int_0^1 \sqrt{\frac{1-p}{p}} dp . \end{aligned}$$

- In the general regime the variance, and the fluctuations are not understood.



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## Theorem (Ottolini and T.'2023)

Assume that  $m_i \leq m$  for some  $m$  and  $\epsilon_n \geq \epsilon$  for some  $\epsilon > 0$  independent of  $n$ . Then,

$$\mathbb{E}[T_{\mathbf{m}^n, n}] \sim \text{Var}[T_{\mathbf{m}^n, n}] \sim \left(1 + \frac{1}{2} + \dots + \frac{1}{\mathbf{m}_{\max}^n}\right) \log n \quad \text{as } n \rightarrow \infty.$$



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And, there exists  $C(\epsilon, m) > 0$  such that

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P} \left( \frac{T_{\mathbf{m}^n, n} - \mathbb{E}[T_{\mathbf{m}^n, n}]}{\sqrt{\text{Var}[T_{\mathbf{m}^n, n}]}} \leq x \right) - \Phi(x) \right| \leq C(\epsilon, m) \frac{\log \log n}{\sqrt{\log n}}$$





## Proof Idea:

Define the following random variables:

$$T_j = \max\{t \in \{0, \dots, |\mathbf{m}|\} : \text{Among last } t \text{ cards, no card appears more than } j \text{ times}\}$$



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Example deck with 3 types and  $\mathbf{m} = (3, 3, 2)$

Listed from the last to the first:

$$(1, 2, 2, 1, 3, 1, 2, 3).$$

In this case,  $T_1 = 2, T_2 = 5$



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## Proof Idea contd...

### Lemma (Key Lemma)

For any deck  $\mathbf{m}$ , the optimal score  $S_{\mathbf{m}}$  can be written as

$$S_{\mathbf{m}} = \sum_{j=1}^m \sum_{s=1}^{\widetilde{W}_j} X_{j,s},$$

where the  $X_{j,s}$  are conditionally independent - given the  $\widetilde{W}_j$ 's - Bernoulli random variables with  $\mathbb{P}(X_{j,s} = 1) = \frac{1}{s}$ .



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### 1. CLT for the conditional score.

- Conditioned on  $\tilde{W}_j, j = 1, \dots, m$ , we prove CLT for the conditional score with suitable conditional mean  $\mu'_{\mathbf{m}}$  and variance  $\sigma'^2_{\mathbf{m}}$ .



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1. CLT for the conditional score.
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2. Removing the conditioning.
  - This requires understanding the behavior of  $T_j$ s and the dependence of  $\widetilde{W}_j$ s on  $T_j$ .





## Some other variants

- Partial Feedback Games



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  - Diaconis, Graham, He, Spiro (2022), Nie (2022): Expected score  $m + \Theta(\sqrt{m})$  uniformly in  $n$  for  $n \gg m \gg 1$ .



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- Decks that are not properly shuffled
  - Ciucu (1998): Dovetail shuffle
  - Liu (2021): Riffle-Shuffle with complete feedback
  - Kuba and Panholzer (2023): Limit law with one riffle-shuffle and no feedback



Thank you!

