

# Random Numbers

Rahul Ramachandran

## CONTENTS

1	Uniform Random Numbers	1
2	Central Limit Theorem	2
3	From Uniform to Other	4
4	Triangular Distributions	4
5	Maximal Likelihood	6
6	Gaussian to Other	9
7	Conditional Probability	10
8	Two Dimensions	10

**Abstract**—This manual provides a simple introduction to the generation of random numbers

## 1 UNIFORM RANDOM NUMBERS

Let  $U$  be a uniform random variable between 0 and 1.

1.1 Generate  $10^6$  samples of  $U$  using a C program and save into a file called uni.dat .

**Solution:** Download the following files and execute the C program.

```
wget https://github.com/gadepall/probability/
raw/master/manual/codes/exrand.c
wget https://github.com/gadepall/probability/
raw/master/manual/codes/coeffs.h
gcc exrand.c
./a.out
```

1.2 Load the uni.dat file into python and plot the empirical CDF of  $U$  using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr(U \leq x) \quad (1.1)$$

**Solution:**

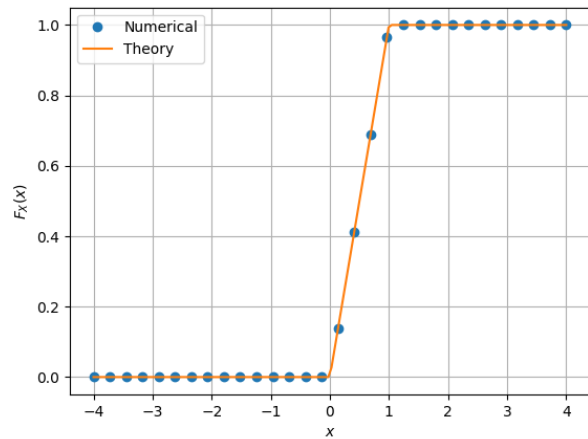


Fig. 1.2: The CDF of  $U$

```
wget https://github.com/gadepall/probability/
raw/master/manual/codes/cdf_plot.py
python3 cdf_plot.py
```

The following code plots Fig. 1.2

1.3 Find a theoretical expression for  $F_U(x)$ .

**Solution:**  $U$  is given by

$$U(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ 1, & x \in (0, 1) \\ 0, & x \in (1, \infty) \end{cases} \quad (1.2)$$

Therefore, we have:

$$F_U(x) = \int_0^x U(x) dx \quad (1.3)$$

Computing the integral, we get:

$$F_U(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ x, & x \in (0, 1) \\ 1, & x \in (1, \infty) \end{cases} \quad (1.4)$$

1.4 The mean of  $U$  is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^N U_i \quad (1.5)$$

and its variance as

$$\text{var}[U] = E[U - E[U]]^2 \quad (1.6)$$

Write a C program to find the mean and variance of  $U$ .

**Solution:** Add the following function to coeffs.h

```
double variance(char *str)
{
    int i=0,c;
    FILE *fp;
    double x, temp=0.0;

    fp = fopen(str,"r");
    //get numbers from file
    while(fscanf(fp,"%lf",&x)!=EOF)
    {
        //Count numbers in file
        i=i+1;
        //Add all numbers in file
        temp = temp+x*x;
    }
    double mn = mean(str);
    fclose(fp);
    temp = temp/(i-1);
    return temp - mn*mn ;
}
```

Following the steps mentioned below gives the required result:

```
gcc exrand.c
./a.out
mean = 0.500031
variance = 0.083247
```

1.5 Verify your result theoretically given that

$$E[U^k] = \int_{-\infty}^{\infty} x^k dF_U(x) \quad (1.7)$$

**Solution:** Since

$$dF_U(x) = p_U(x)dx \quad (1.8)$$

we have:

$$E[U^k] = \int_{-\infty}^{\infty} x^k p_U(x)dx \quad (1.9)$$

Also,

$$p_U(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ 1, & x \in (0, 1) \\ 0, & x \in (1, \infty) \end{cases} \quad (1.10)$$

Therefore, from Equations 1.9 and 1.10, we have:

$$E[U^2] = \int_{-\infty}^{\infty} x^2 p_U(x)dx \quad (1.11)$$

$$= \int_0^1 x^2 dx \quad (1.12)$$

$$= \frac{1}{3} \quad (1.13)$$

Similarly,

$$E[U] = \int_{-\infty}^{\infty} x p_U(x)dx \quad (1.14)$$

$$= \int_0^1 x dx \quad (1.15)$$

$$= \frac{1}{2} \quad (1.16)$$

Therefore, the mean is  $\frac{1}{2}$ , and the variance equals:

$$E[U^2] - E[U]^2 = \frac{1}{3} - \left(\frac{1}{2}\right)^2 \quad (1.17)$$

$$= \frac{1}{12} \quad (1.18)$$

## 2 CENTRAL LIMIT THEOREM

2.1 Generate  $10^6$  samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \quad (2.1)$$

using a C program, where  $U_i, i = 1, 2, \dots, 12$  are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat

**Solution:** Add the following line to **exrand.c** and execute the code:

```
gaussian("gau.dat", 1000000);
gcc exrand.c
./a.out
```

2.2 Load gau.dat in python and plot the empirical CDF of  $X$  using the samples in gau.dat. What properties does a CDF have?

**Solution:**

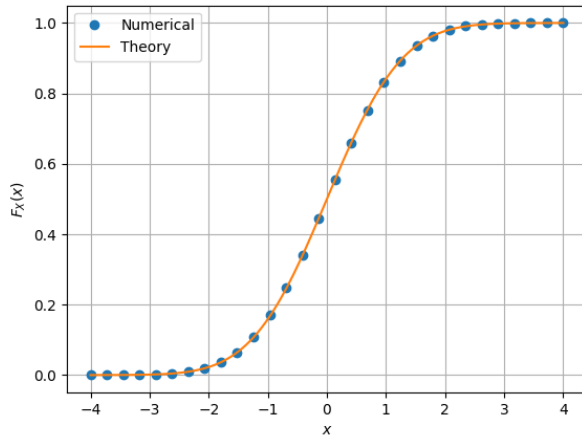


Fig. 2.2: The CDF of  $X$

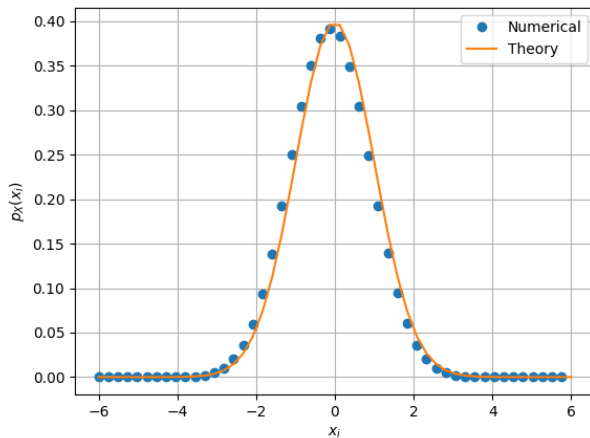


Fig. 2.3: The PDF of  $X$

The CDF of  $X$  is plotted in Fig. 2.2

- 2.3 Load `gau.dat` in python and plot the empirical PDF of  $X$  using the samples in `gau.dat`. The PDF of  $X$  is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \quad (2.2)$$

What properties does the PDF have?

**Solution:** The PDF of  $X$  is plotted in Fig. 2.3 using the code below

```
wget https://github.com/gadepall/probability/
raw/master/manual/codes/pdf_plot.py
python3 pdf_plot.py
```

To find the CDF theoretically, consider

- 2.4 Find the mean and variance of  $X$  by writing a C program.

**Solution:** Use the main and variance functions in `coeffs.h`, and execute the code below

```
gcc exrand.c
./a.out
```

We get

```
mean = 0.000685
variance = 1.000025
```

- 2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (2.3)$$

repeat the above exercise theoretically.

**Solution:** We have:

$$E[X] = \int_{-\infty}^{\infty} \frac{x}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \quad (2.4)$$

$$= -\frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) \Big|_{-\infty}^{\infty} \quad (2.5)$$

$$= 0 \quad (2.6)$$

Also,

$$E[X^2] = \int_{-\infty}^{\infty} \frac{x^2}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \quad (2.7)$$

$$= -\frac{x}{\sqrt{2\pi}} e^{\left(-\frac{x^2}{2}\right)} \Big|_{-\infty}^{\infty} + \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{\left(-\frac{x^2}{2}\right)} dx \quad (2.8)$$

$$= 0 + \frac{1}{\sqrt{2\pi}} \times \sqrt{2\pi} \quad (2.9)$$

$$= 1 \quad (2.10)$$

Hence,

$$\text{var}(X) = E[X^2] - E[X]^2 \quad (2.11)$$

$$= 1 \quad (2.12)$$

Therefore, the mean is 0 and the variance is 1. Running the empirical code in `./codes/exrand.c`, we get mean = 0.000685 and variance = 1.000025, which closely matches the theoretical values.

- 2.6 Find the theoretical CDF of  $X$

**Solution:** To find the theoretical CDF, consider:

$$Q_X(x) = \int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \quad (2.13)$$

$$= \frac{\operatorname{erfc}(\frac{x}{\sqrt{2}})}{2} \quad (2.14)$$

The CDF is then:

$$F_X(x) = 1 - Q_X(x) \quad (2.15)$$

$$= 1 - \frac{\operatorname{erfc}(\frac{x}{\sqrt{2}})}{2} \quad (2.16)$$

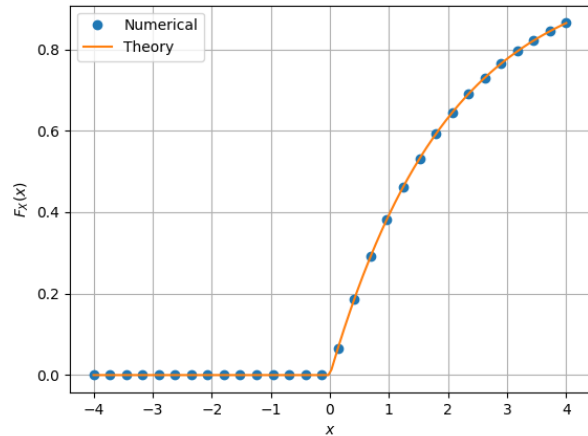


Fig. 3.2: The CDF of  $V$

### 3 FROM UNIFORM TO OTHER

#### 3.1 Generate samples of

$$V = -2 \ln(1 - U) \quad (3.1)$$

and plot its CDF. **Solution:**

Add the following function to **coeffs.h**:

```
void logarithmic(char *str){
    int i=0,c;
    FILE *fp, *fp2;
    double x, temp=0.0;

    fp = fopen("uni.dat","r");
    fp2 = fopen(str, "w");
    //get numbers from file
    while(fscanf(fp,"%lf",&x)!=EOF)
    {
        temp = -2*log(1-x);
        fprintf(fp2,"%lf\n",temp);
    }

    fclose(fp);
    fclose(fp2);

    return ;
}
```

Using this function in **exrand.c** prints the numbers in **log.dat**

#### 3.2 Find a theoretical expression for $F_V(x)$ .

**Solution:** We have:

$$F_V(x) = \Pr(V \leq x) \quad (3.2)$$

$$= \Pr(-2 \ln(1 - U) \leq x) \quad (3.3)$$

$$= \Pr\left(1 - U \geq \exp\left(-\frac{x}{2}\right)\right) \quad (3.4)$$

$$= \Pr\left(U \leq 1 - \exp\left(-\frac{x}{2}\right)\right) \quad (3.5)$$

$$= F_U\left(1 - \exp\left(-\frac{x}{2}\right)\right) \quad (3.6)$$

Therefore,

$$F_V(x) = \begin{cases} 0, & 1 - \exp\left(-\frac{x}{2}\right) \in (-\infty, 0) \\ 1 - \exp\left(-\frac{x}{2}\right), & 1 - \exp\left(-\frac{x}{2}\right) \in (0, 1) \\ 1, & 1 - \exp\left(-\frac{x}{2}\right) \in (1, \infty) \end{cases} \quad (3.7)$$

From this we get:

$$F_V(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ 1 - \exp\left(-\frac{x}{2}\right), & x \in (0, \infty) \end{cases} \quad (3.8)$$

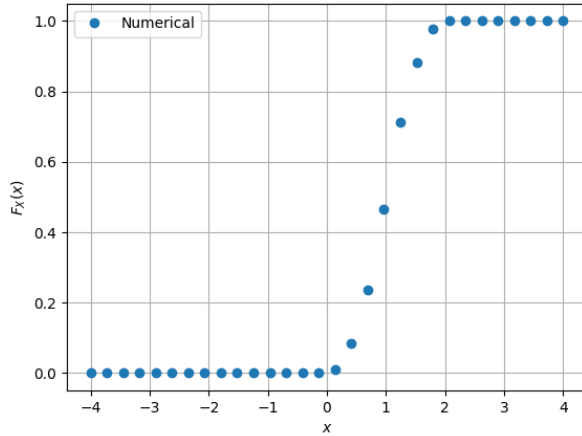
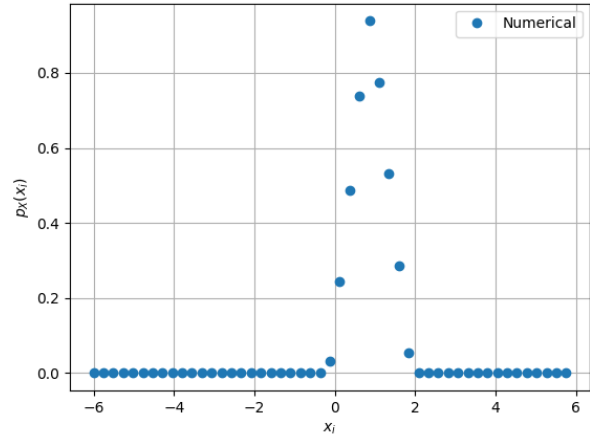
The CDF of  $V$  is plotted in Fig. 3.2

### 4 TRIANGULAR DISTRIBUTIONS

#### 4.1 Generate

$$T = U_1 + U_2 \quad (4.1)$$

**Solution:** Use the function 'triangular' in **exrand.c** and execute the following code:

Fig. 4.2: The CDF of  $T$ Fig. 4.3: The PDF of  $T$ 

```
wget https://github.com/Rahuboy/AI1110/blob
/main/RandomNumbers/codes/coeffs.h
wget https://github.com/Rahuboy/AI1110/blob
/main/RandomNumbers/codes/exrand.c
gcc exrand.c
./a.out
```

4.2 Find the CDF of  $T$ .

**Solution:**

```
wget https://github.com/Rahuboy/AI1110/
blob/main/RandomNumbers/codes/
cdf_plot.py
python3 cdf_plot.py
```

The above code plots Fig. 4.2

4.3 Find the PDF of  $T$ .

```
wget https://github.com/Rahuboy/AI1110/
blob/main/RandomNumbers/codes/
pdf_plot.py
python3 pdf_plot.py
```

The above code plots Fig. 4.3

4.4 Find the theoretical expressions for the PDF and CDF of  $T$ .

**Solution:** When

$$Z = X + Y \quad (4.2)$$

where  $X$ ,  $Y$  and  $Z$  are random variables, we

have:

$$p_Z(t) = (p_X * p_Y)(t) \quad (4.3)$$

$$= \int_{-\infty}^{\infty} p_X(\tau) p_Y(t - \tau) d\tau \quad (4.4)$$

Here,  $p_X(t) = p_Y(t) = p_U(t)$ . Therefore:

$$p_T(t) = \int_{-\infty}^{\infty} p_U(\tau) p_U(t - \tau) d\tau \quad (4.5)$$

$$= \int_0^1 p_U(t - \tau) d\tau \quad (4.6)$$

When  $t < 0$  and  $t > 2$ , the integral evaluates to 0. When  $0 < t < 1$ :

$$p_T(t) = \int_0^1 p_U(t - \tau) d\tau \quad (4.7)$$

$$= \int_0^t p_U(t - \tau) d\tau \quad (4.8)$$

$$= \int_0^t 1 d\tau \quad (4.9)$$

$$= t \quad (4.10)$$

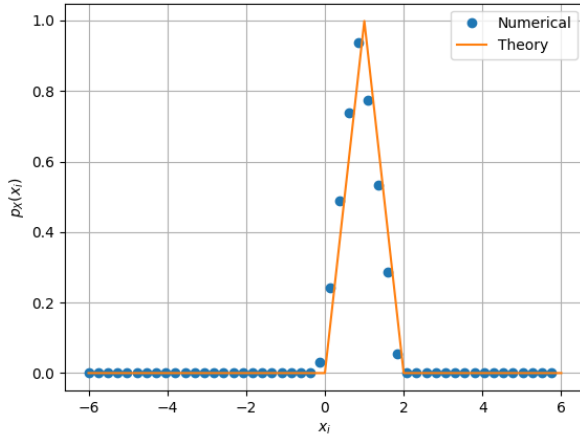
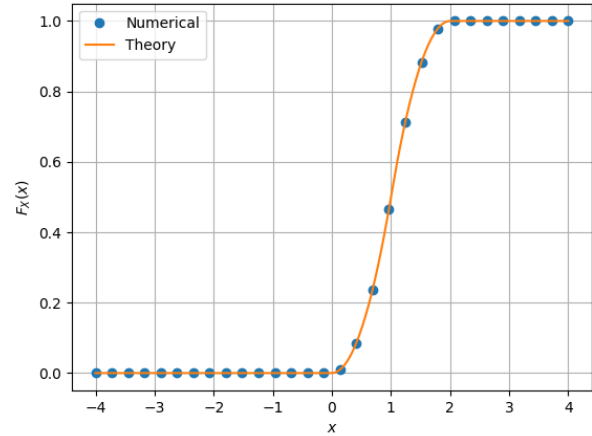
when  $1 < t < 2$ :

$$p_T(t) = \int_0^1 p_U(t - \tau) d\tau \quad (4.11)$$

$$= \int_{t-1}^1 p_U(t - \tau) d\tau \quad (4.12)$$

$$= \int_{t-1}^1 1 d\tau \quad (4.13)$$

$$= 2 - t \quad (4.14)$$

Fig. 4.5: The PDF of  $T$ Fig. 4.6: The CDF of  $T$ 

Therefore, we have:

$$p_T(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ x, & x \in (0, 1) \\ 2 - x, & x \in (1, 2) \\ 0, & x \in (2, \infty) \end{cases} \quad (4.15)$$

To find the CDF, we use:

$$F_T(x) = \int_{-\infty}^x p_T(t) dt \quad (4.16)$$

We get:

$$F_T(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ \frac{x^2}{2}, & x \in (0, 1) \\ -\frac{x^2}{2} + 2x - 1, & x \in (1, 2) \\ 1, & x \in (2, \infty) \end{cases} \quad (4.17)$$

4.5 Verify your result for the PDF through a plot.

**Solution:** Execute the following code:

```
python3 pdf_plot.py
```

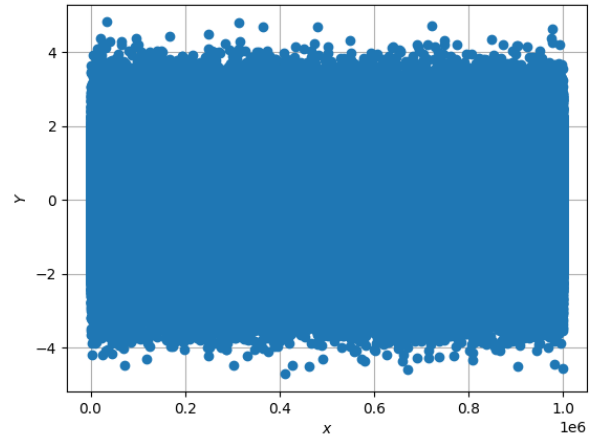
The theoretical PDF is plotted in Fig. 4.5

4.6 Verify your result for the CDF through a plot.

**Solution:** Execute the following code:

```
python3 cdf_plot.py
```

The theoretical CDF is plotted in Fig. 4.6

Fig. 5.2: The Plot of  $Y$ 

where  $A = 5$  dB,  $X_1 \{1, -1\}$ , is Bernoulli and  $N \sim \mathcal{N}(0, 1)$ .

**Solution:** Use the functions 'bernoulli' and 'maxlike' in **exrand.c**:

```
gcc exrand.c
./a.out
```

5.2 Plot  $Y$ .

$Y$  is plotted in Fig. 5.2

5.3 Guess how to estimate  $X$  from  $Y$ .

**Solution:** To estimate  $X$  from  $Y$ , we define the following function:

$$\text{sgn}(y) = \begin{cases} -1, & y \in (-\infty, 0] \\ 1, & y \in (0, \infty) \end{cases} \quad (5.2)$$

## 5 MAXIMAL LIKELIHOOD

5.1 Generate

$$Y = AX + N, \quad (5.1)$$

Using `sgny`, we can operate on  $Y$  to find corresponding values of  $X$ .

5.4 Find

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1) \quad (5.3)$$

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1) \quad (5.4)$$

**Solution:** Use the function "maxlike\_proberr" in **exrand.c** to find the respective probabilities:

```
gcc exrand.c
./a.out
P_(e|0) = 0.312414
P_(e|1) = 0.310985
```

5.5 Find  $P_e$ .

**Solution:** Assume a general value of  $A$ . Our estimation function predicts that the data points above the  $x$  axis correspond to  $X = 1$ , and the data points below the  $x$ -axis correspond to  $X = -1$ . This isn't always the case, as  $Y = AX + N$ , and the  $N$  causes some spill-over. We have:

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1) \quad (5.5)$$

$$= \Pr(AX + N < 0|X = 1) \quad (5.6)$$

$$= \Pr(N < -A) \quad (5.7)$$

$$= \int_{-\infty}^{-A} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \quad (5.8)$$

$$= \int_A^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \quad (5.9)$$

$$= Q_N(A) \quad (5.10)$$

where  $Q_N$  is the  $Q$ -function of the normal distribution.

Similarly,

$$P_{e|1} = Q_N(A) \quad (5.11)$$

Therefore,

$$P_e = P_{e|0} \times \Pr(X = 1) + P_{e|1} \times \Pr(X = -1) \quad (5.12)$$

$$= \frac{1}{2} P_{e|0} + \frac{1}{2} P_{e|1} \quad (5.13)$$

$$= \frac{1}{2} Q_N(A) + \frac{1}{2} Q_N(A) \quad (5.14)$$

$$= Q_N(A) \quad (5.15)$$

5.6 Verify by plotting the theoretical  $P_e$ .

**Solution:** The graph of  $P_e$  is plotted in Fig. 5.6

5.7 Now, consider a threshold  $\delta$  while estimating

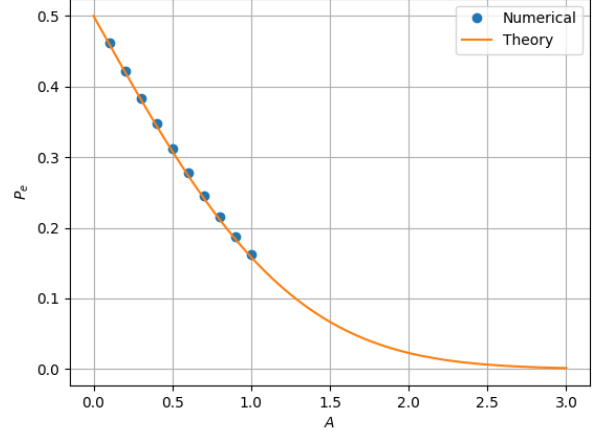


Fig. 5.6: The Plot of  $P_e$

$X$  from  $Y$ . Find the value of  $\delta$  that maximizes the theoretical  $P_e$ .

**Solution:** To estimate  $X$  from  $Y$ , we now consider the following:

$$X = \begin{cases} 1, & Y > \delta \\ -1, & Y < \delta \end{cases} \quad (5.16)$$

Therefore,

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1) \quad (5.17)$$

$$= \Pr(AX + N < \delta|X = 1) \quad (5.18)$$

$$= \Pr(N < \delta - A) \quad (5.19)$$

$$= \int_{-\infty}^{\delta-A} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \quad (5.20)$$

$$= \int_{A-\delta}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \quad (5.21)$$

$$= Q_N(A - \delta) \quad (5.22)$$

Where  $Q_N$  is the  $Q$ -function of the normal distribution. Similarly,

$$P_{e|1} = Q_N(A + \delta) \quad (5.23)$$

Therefore,

$$P_e = P_{e|0} \Pr(X = 1) + P_{e|1} \Pr(X = -1) \quad (5.24)$$

$$= \frac{1}{2} (Q_N(A - \delta) + Q_N(A + \delta)) \quad (5.25)$$

$$(5.26)$$

To minimise  $P_e$ , we differentiate the above

equation wrt  $\delta$ :

$$0 = \frac{d}{d\delta} \left( \frac{1}{2} (Q_N(A - \delta) + Q_N(A + \delta)) \right) \quad (5.27)$$

$$= \frac{1}{2} \left( \frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta-A)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} \right) \quad (5.28)$$

Therefore,

$$(\delta - A)^2 = (\delta + A)^2 \quad (5.29)$$

$$\implies \delta = 0 \quad (5.30)$$

5.8 Repeat the above exercise when

$$p_X(0) = p \quad (5.31)$$

**Solution:** Using Eq. (5.24), we have:

$$P_e = P_{e|0}p + P_{e|1}(1-p) \quad (5.32)$$

$$= pQ_N(A - \delta) + (1-p)Q_N(A + \delta) \quad (5.33)$$

Differentiating as before, we get:

$$0 = p \frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta-A)^2}{2}} - (1-p) \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} \quad (5.34)$$

Taking ln on both sides we have:

$$\ln p - \frac{(\delta - A)^2}{2} = \ln(1-p) + \frac{(\delta + A)^2}{2} \quad (5.35)$$

$$\implies \delta = \frac{1}{2A} \ln \frac{1-p}{p} \quad (5.36)$$

5.9 Repeat the above exercise using the MAP criterion.

**Solution:** Assume that  $\Pr(X = -1) = p$ , and  $\Pr(X = 1) = (1-p)$ . Then, using the Law of Total Probability, we have:

$$p_Y(y) = p_{Y|X=-1}(y|-1) \Pr(X = -1) + p_{Y|X=1}(y|1) \Pr(X = 1) \quad (5.37)$$

$$= p \times p_{(-A+N)}(y) + (1-p) \times p_{(A+N)}(y) \quad (5.38)$$

where  $p_Y(y)$  is the pdf of  $Y$ . Now,  $p_{(-A+N)}$  is just the pdf of a shifted normal distribution, and therefore:

$$p_Y(y) = p \frac{e^{-\frac{(y+A)^2}{2}}}{\sqrt{2\pi}} + (1-p) \frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}} \quad (5.39)$$

To use the MAP criterion, we must find  $p_{X|Y}(x|y)$ . To do this, we use the Theorem of Conditional Probability:

$$p_{X|Y}(x|y) = \frac{p_{Y|X}(y|x) \times p_X(x)}{p_Y(y)} \quad (5.40)$$

When  $X = 1$ , we have:

$$p_{X|Y}(1|y) = \frac{p_{Y|X}(y|1) \times p_X(1)}{p_Y(y)} \quad (5.41)$$

$$= \frac{(1-p) \frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}}{p \frac{e^{-\frac{(y+A)^2}{2}}}{\sqrt{2\pi}} + (1-p) \frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}} \quad (5.42)$$

$$= \frac{(1-p) e^{2yA}}{p + (1-p) e^{2yA}} \quad (5.43)$$

Similarly, when  $X = -1$ , we get:

$$p_{X|Y}(-1|y) = \frac{p}{p + (1-p) e^{2yA}} \quad (5.44)$$

Therefore, when  $p_{X|Y}(1|y) > p_{X|Y}(-1|y)$ , we have:

$$\frac{(1-p) e^{2yA}}{p + (1-p) e^{2yA}} > \frac{p}{p + (1-p) e^{2yA}} \quad (5.45)$$

$$e^{2yA} > \frac{p}{(1-p)} \quad (5.46)$$

$$y > \frac{1}{2A} \ln \frac{p}{(1-p)} \quad (5.47)$$

Therefore, when Eq. (5.47), we can assert that  $X = 1$ , and  $X = -1$  otherwise. Now, consider when  $p = \frac{1}{2}$ . We have:

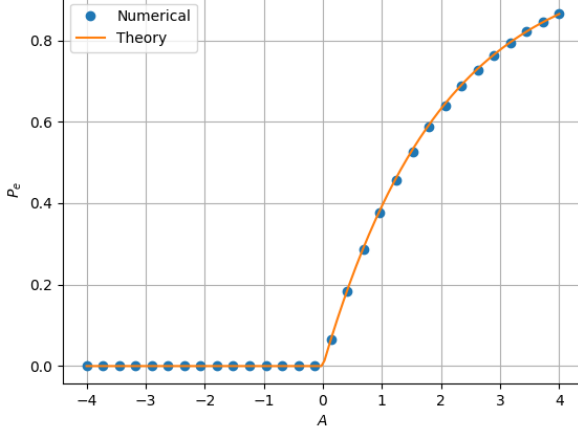
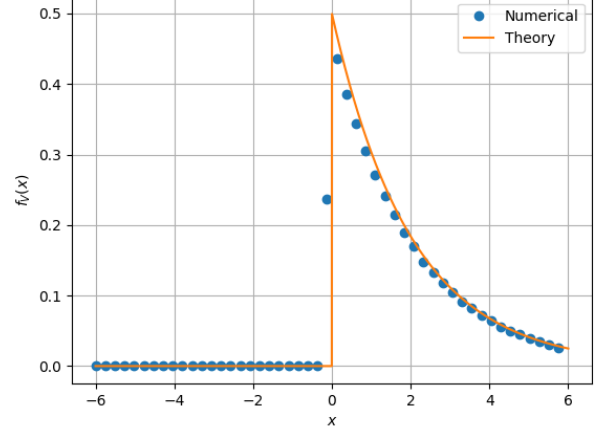
$$y > \frac{1}{2A} \ln \frac{p}{(1-p)} \quad (5.48)$$

$$= \frac{1}{2A} \ln 1 \quad (5.49)$$

$$= 0 \quad (5.50)$$

Therefore, when  $y > 0$ , we choose  $X = 1$ , and we choose  $X = -1$  otherwise.



Fig. 6.1: The CDF of  $V$ Fig. 6.1: The PDF of  $V$ 

## 6 GAUSSIAN TO OTHER

6.1 Let  $X_1 \sim \mathcal{N}(0, 1)$  and  $X_2 \sim \mathcal{N}(0, 1)$ . Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \quad (6.1)$$

**Solution:** Use the function "chi" in exrand.c and execute:

```
gcc exrand.c
./a.out
```

Define the functions "chi\_pdf" and "chi\_cdf" in **functions.py** and execute:

```
python3 cdf_plot.py
python3 pdf_plot.py
```

The graphs are plotted in Fig. 6.1

6.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \geq 0 \\ 0 & x < 0, \end{cases} \quad (6.2)$$

find  $\alpha$ .

**Solution:** We will assume that  $X_1$  and  $X_2$  are i.i.d. Let

$$X_1 = r \cos \theta \quad (6.3)$$

$$X_2 = r \sin \theta \quad (6.4)$$

$$(6.5)$$

The Jacobian Matrix is then defined as:

$$J = \begin{pmatrix} \frac{\partial x_1}{\partial r} & \frac{\partial x_1}{\partial \theta} \\ \frac{\partial x_2}{\partial r} & \frac{\partial x_2}{\partial \theta} \end{pmatrix} \quad (6.6)$$

$$J = \begin{pmatrix} \frac{\partial r \cos \theta}{\partial r} & \frac{\partial r \cos \theta}{\partial \theta} \\ \frac{\partial r \sin \theta}{\partial r} & \frac{\partial r \sin \theta}{\partial \theta} \end{pmatrix} \quad (6.7)$$

$$J = \begin{pmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{pmatrix} \quad (6.8)$$

$$\Rightarrow |J| = R \quad (6.9)$$

Then,

$$f_{X_1, X_2}(x_1, x_2) = f_{X_1}(x_1)f_{X_2}(x_2) \quad (6.10)$$

$$= \frac{1}{2\pi} e^{-\frac{(x_1^2 + x_2^2)}{2}} \quad (6.11)$$

$$= \frac{1}{2\pi} e^{-\frac{r^2}{2}} \quad (6.12)$$

Now, since

$$f_{R, \theta}(r, \theta) = |J|f_{X_1, X_2}(x_1, x_2) \quad (6.13)$$

we have:

$$f_{R, \theta}(r, \theta) = \frac{r}{2\pi} e^{-\frac{r^2}{2}} \quad (6.14)$$

Therefore,

$$f_R(r) = \int_0^{2\pi} f_{R, \theta}(r, \theta) d\theta \quad (6.15)$$

$$= \int_0^{2\pi} \frac{r}{2\pi} e^{-\frac{r^2}{2}} d\theta \quad (6.16)$$

$$= r e^{-\frac{r^2}{2}} \quad (6.17)$$

$$(6.18)$$

We then have:

$$F_R(r) = \Pr(R \leq r) \quad (6.19)$$

$$= \int_0^r f_R(r) dr = 1 - e^{-\frac{r^2}{2}} \quad (6.20)$$

$F_V(x)$  is given by:

$$F_V(x) = F_{X_1^2 + X_2^2}(x) \quad (6.21)$$

$$= F_{R^2}(x) \quad (6.22)$$

$$= \Pr(R^2 \leq x) \quad (6.23)$$

$$= \Pr(R \leq \sqrt{x}) \quad (6.24)$$

Therefore,

$$F_V(x) = \begin{cases} 0, & x < 0 \\ 1 - e^{-\frac{x}{2}}, & x \geq 0 \end{cases} \quad (6.25)$$

Comparing with Eq. 6.1 we get:

$$\alpha = \frac{1}{2} \quad (6.26)$$

6.3 Plot the CDF and PDF of

$$A = \sqrt{V} \quad (6.27)$$

**Solution:** Use the function "ray" in **exrand.c** and execute:

```
gcc exrand.c
./a.out
```

Add the functions "ray\_pdf" and "ray\_cdf" to **functions.py** and execute the below files:

```
python3 pdf_plot.py
python3 cdf_plot.py
```

The graphs are plotted in Fig. 6.3

## 7 CONDITIONAL PROBABILITY

7.1 Plot

$$P_e = \Pr(\hat{X} = -1 | X = 1) \quad (7.1)$$

for

$$Y = AX + N, \quad (7.2)$$

where  $A$  is Raleigh with  $E[A^2] = \gamma$ ,  $N \sim \mathcal{N}(0, 1)$ ,  $X \in (-1, 1)$  for  $0 \leq \gamma \leq 10$  dB.

7.2 Assuming that  $N$  is a constant, find an expression for  $P_e$ . Call this  $P_e(N)$

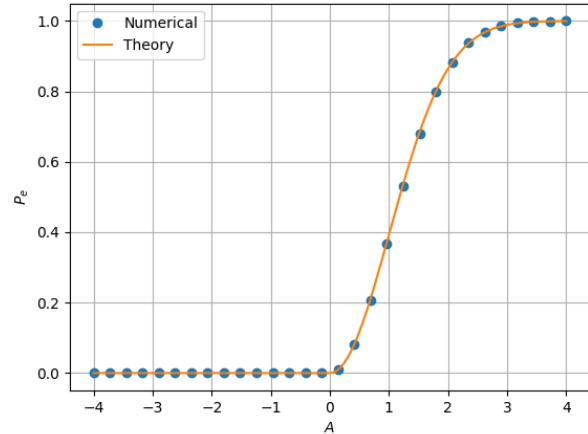


Fig. 6.3: The CDF of A

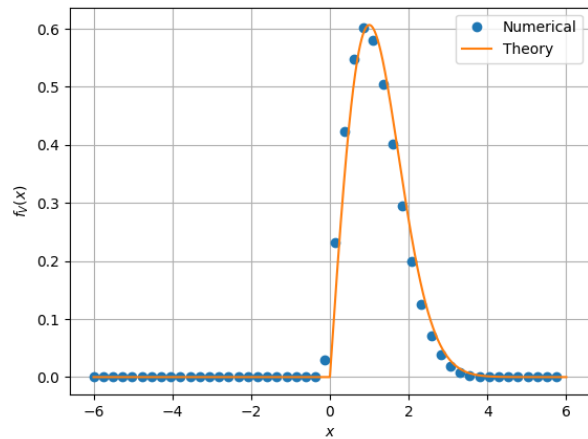


Fig. 6.3: The PDF of A

7.3 For a function  $g$ ,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) p_X(x) dx \quad (7.3)$$

Find  $P_e = E[P_e(N)]$ .

7.4 Plot  $P_e$  in problems 7.1 and 7.3 on the same graph w.r.t  $\gamma$ . Comment.

## 8 TWO DIMENSIONS

Let

$$\mathbf{y} = \mathbf{A}\mathbf{x} + \mathbf{n}, \quad (8.1)$$

where

$$x \in (\mathbf{s}_0, \mathbf{s}_1), \mathbf{s}_0 = \begin{pmatrix} 1 \\ 0 \end{pmatrix}, \mathbf{s}_1 = \begin{pmatrix} 0 \\ 1 \end{pmatrix} \quad (8.2)$$

$$\mathbf{n} = \begin{pmatrix} n_1 \\ n_2 \end{pmatrix}, n_1, n_2 \sim \mathcal{N}(0, 1). \quad (8.3)$$

8.1 Plot

$$\mathbf{y}|\mathbf{s}_0 \text{ and } \mathbf{y}|\mathbf{s}_1 \quad (8.4)$$

on the same graph using a scatter plot.

8.2 For the above problem, find a decision rule for detecting the symbols  $\mathbf{s}_0$  and  $\mathbf{s}_1$ .

8.3 Plot

$$P_e = \Pr(\hat{\mathbf{x}} = \mathbf{s}_1 | \mathbf{x} = \mathbf{s}_0) \quad (8.5)$$

with respect to the SNR from 0 to 10 dB.

8.4 Obtain an expression for  $P_e$ . Verify this by comparing the theory and simulation plots on the same graph.