

# Weighted Least Squares

Rahul Ramachandran  
cs21btech11049

Rishit D  
cs21btech11053

October 1, 2023

## Contents

<b>1 Solution</b>	<b>1</b>
1.1 Likelihood and Prior	1
1.2 ML and MAP Estimation	1
1.3 Error Function and Solution	2

## 1 Solution

### 1.1 Likelihood and Prior

In a homoscedastic setting, the likelihood is given by:

$$p(t_n|x_n, \mathbf{w}) \sim \mathcal{N}(t_n|\mathbf{w}^T\phi(x_n), \beta^{-1})$$

We arrive at this by assuming that the noise is Gaussian and independent of the input:

$$t_n = y(x_n, \mathbf{w}) + \epsilon$$

In a heteroscedastic setting, we instead assume that the noise depends on the input. In particular, we will assume that the standard deviation of the noise is given by  $\frac{\sigma^2}{r_n}$ , where  $\sigma^2 = \frac{1}{\beta}$ . Therefore, the likelihood is given by:

$$p(t_n|x_n, \mathbf{w}) \sim \mathcal{N}(t_n|\mathbf{w}^T\phi(x_n), (\beta r_n)^{-1})$$

The prior is given by (**check this**):

$$p(\mathbf{w}|\alpha) \sim \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha^{-1}\mathbf{I})$$

### 1.2 ML and MAP Estimation

The likelihood is given by:

$$p(\mathbf{t}|\mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n|\mathbf{w}^T\phi(x_n), (\beta r_n)^{-1})$$

Here, we assume that the  $t_i$ s are independent. The above expression simplifies to:

$$p(\mathbf{t}|\mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n|\mathbf{w}^T\phi(x_n), (\beta r_n)^{-1}) \tag{1}$$

$$= \prod_{n=1}^N \frac{1}{\sqrt{2\pi(\beta r_n)^{-1}}} \exp\left(-\frac{1}{2}(\beta r_n)(t_n - \mathbf{w}^T\phi(x_n))^2\right) \tag{2}$$

For the MAP estimate, we consider:

$$p(\mathbf{w}|\mathbf{x}, \mathbf{t}, \alpha, \beta) \propto p(\mathbf{t}|\mathbf{w}, \beta)p(\mathbf{w}|\alpha)$$

Therefore, we have:

$$\begin{aligned} p(\mathbf{w}|\mathbf{x}, \mathbf{t}, \alpha, \beta) &\propto p(\mathbf{t}|\mathbf{w}, \beta)p(\mathbf{w}|\alpha) \\ &= \left( \prod_{n=1}^N \frac{1}{\sqrt{2\pi(\beta r_n)^{-1}}} \exp \left( -\frac{1}{2}(\beta r_n)(t_n - \mathbf{w}^T \phi(x_n))^2 \right) \right) \left( \frac{\alpha}{2\pi} \right)^{(M+1)/2} \exp \left( -\frac{\alpha}{2} \mathbf{w}^T \mathbf{w} \right) \end{aligned}$$

### 1.3 Error Function and Solution

To maximise the likelihood, we take log on both sides of 2:

$$\begin{aligned} \log p(\mathbf{t}|\mathbf{w}, \beta) &= \sum_{n=1}^N \log \frac{1}{\sqrt{2\pi(\beta r_n)^{-1}}} \exp \left( -\frac{1}{2}(\beta r_n)(t_n - \mathbf{w}^T \phi(x_n))^2 \right) \\ &= \sum_{n=1}^N \left( \log \frac{1}{\sqrt{2\pi(\beta r_n)^{-1}}} + \log \exp \left( -\frac{1}{2}(\beta r_n)(t_n - \mathbf{w}^T \phi(x_n))^2 \right) \right) \\ &= \sum_{n=1}^N \left( \log \frac{1}{\sqrt{2\pi(\beta r_n)^{-1}}} - \frac{1}{2}(\beta r_n)(t_n - \mathbf{w}^T \phi(x_n))^2 \right) \\ &= -\frac{\beta}{2} \sum_{n=1}^N r_n (t_n - \mathbf{w}^T \phi(x_n))^2 + \sum_{n=1}^N \log \frac{1}{\sqrt{2\pi(\beta r_n)^{-1}}} \\ &= -\beta E_D(\mathbf{w}) + \sum_{n=1}^N \log \frac{1}{\sqrt{2\pi(\beta r_n)^{-1}}} \end{aligned}$$

Therefore, we have to minimise the error function:

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N r_n (t_n - \mathbf{w}^T \phi(x_n))^2$$

To do this, we rewrite  $E_D$  as  $(T - \Phi \mathbf{w})^T R (T - \Phi \mathbf{w})$ , where:

$$\Phi = \begin{pmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \cdots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \cdots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \cdots & \phi_{M-1}(\mathbf{x}_N) \end{pmatrix} \quad \mathbf{t} = \begin{pmatrix} t_1 \\ t_2 \\ \vdots \\ t_N \end{pmatrix} \quad \mathbf{R} = \begin{pmatrix} r_1 & 0 & \cdots & 0 \\ 0 & r_2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & r_N \end{pmatrix}$$

Therefore, we have:

$$\begin{aligned} \frac{\partial E_D}{\partial \mathbf{w}} &= \frac{\partial}{\partial \mathbf{w}} (T - \Phi \mathbf{w})^T R (T - \Phi \mathbf{w}) \\ &= -2(T - \Phi \mathbf{w})^T R \Phi \\ &= 0 \\ \implies \Phi^T R (T - \Phi \mathbf{w}) &= 0 \\ \implies \Phi^T R T &= \Phi^T R \Phi \mathbf{w} \\ \implies \mathbf{w} &= (\Phi^T R \Phi)^{-1} \Phi^T R T \end{aligned}$$