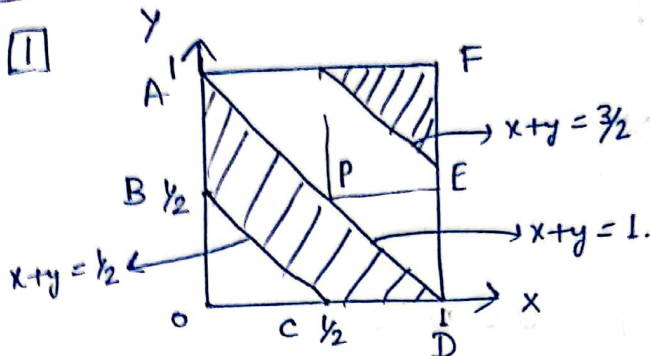


5.1



Since $f_{x,y}(x,y) = \begin{cases} c & \text{in Shaded region} \\ 0 & \text{otherwise} \end{cases}$

We know, Area under curve = 1 $\left[\because \iint f_{x,y}(x,y) = 1 \right]$

$\Rightarrow c = \frac{1}{\text{Area of Shaded region}}$

$\Rightarrow c = \frac{1}{\text{Area}(ABCD) + \text{Area}(EFCD)}$

$\Rightarrow c = \frac{1}{\frac{1}{2} \times \frac{1}{2} + \frac{1}{2} \times 1}$

$\Rightarrow \boxed{c = 2}$

(2)

Marginal PDF of X = Integration of $f_{x,y}(x,y)$ w.r.t to y

$\Rightarrow f_x(x) = \int_{-\infty}^{\infty} f_{x,y}(x,y) dy$

\Rightarrow When $0 \leq x < \frac{1}{2}$, Area of Shaded region, is given by the area b/w curve $x+y = \frac{1}{2}$ & $x+y = 1$

So, $\left| \int_{y=\frac{1}{2}-x}^{1-x} 2 dy \right|$ when $0 < x < \frac{1}{2}$

$= 2 \left[1-x - \frac{1}{2} + x \right] = 1$

When $\frac{1}{2} < x < 1$, Area of shaded region is given by the sum of area b/w curve $xy=1$ & $y=0$ and $xy=\frac{3}{2}$ & $y=1$

$$\text{So, } \left| \int_0^{1-x} 2dy \right| + \left| \int_{\frac{3}{2}-x}^1 2dy \right| \quad \frac{1}{2} \leq x < 1$$

$$= 2(1-x) + 2(1-\frac{3}{2}+x)$$

$$= 2-2x + 2-3+2x = 1$$

$$\text{So, } f_x(x) = \begin{cases} \int_0^{1-x} 2dy = 1 & 0 \leq x \leq \frac{1}{2} \\ \int_0^{1-x} 2dy + \int_{\frac{3}{2}-x}^1 2dy = 1 & \frac{1}{2} < x \leq 1 \end{cases}$$

Similarly marginal PDF of y = Integration of $f_{xy}(x,y)$ wrt to x

When $0 \leq y < \frac{1}{2}$ Area of shaded region is given by the area b/w the curve $xy=\frac{1}{2}$ & $xy=1$

$$\text{So, } \left| \int_{\frac{1}{2}-y}^{1-y} 2dx \right| \quad \text{when } 0 \leq y < \frac{1}{2}$$

$$= 2[1-y - \frac{1}{2}+y] = 1$$

& When $\frac{1}{2} \leq y < 1$ Area of shaded region is given by the sum of area b/w curve $xy=1$ & $x=0$ and $xy=\frac{3}{2}$

and $x=1$

$$\text{So, } \left| \int_0^{1-y} 2dx + \int_{\frac{3}{2}-y}^1 2dx \right| = 1 \quad \text{when } \frac{1}{2} \leq y < 1$$

$$= 2(1-y) + 2(1-\frac{3}{2}+y) = 1$$

$$\underline{\text{So,}} \quad f(x, y) = \begin{cases} \int_{1/2-y}^{1-y} 2dx = 1 & 0 \leq y < 1/2 \\ \int_{1-y}^1 2dx + \int_{1/2-y}^1 2dx = 1 & 1/2 \leq y \leq 1 \end{cases}$$

$= 1$

$$\boxed{3} \quad E(X/Y = 1/4) = \int_{-\infty}^{\infty} x f_{X/Y}(x/y) dx$$

$$\underline{\text{Now}} \quad f_{X/Y}(X/Y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = 1.$$

$\left[\begin{array}{l} \text{when } y = \frac{1}{4} \\ x+y = 1/2 \\ \Rightarrow x = 1/4 \\ x+y = 1 \\ \Rightarrow x = 3/4 \end{array} \right]$

(from Graph) $= \begin{cases} 2 & 1/4 \leq x \leq 3/4 \\ 0 & \text{else} \end{cases}$

$$\underline{\text{So,}} \quad E(X/Y = 1/4) = \int_{1/4}^{3/4} x \cdot 2dx = \frac{9}{16} - \frac{1}{16} = \frac{8}{16} = \frac{1}{2}$$

$$\therefore E(X/Y = 1/4) = \frac{1}{2}$$

$$\underline{\text{Now}} \quad \text{var}(X/Y = 1/4) = \int_{-\infty}^{\infty} x^2 f_{X/Y}(x/y) dx$$

$$= \int_{1/4}^{3/4} x^2 \cdot 2dx = \frac{2}{3} \left[\frac{27}{64} - \frac{1}{64} \right] = \frac{13}{96}$$

~~4~~

$$\therefore \boxed{\text{var}(X/Y = 1/4) = \frac{13}{96}}$$

$$[4] \quad f_{x|y}(x|y) = \frac{f_{x,y}(x,y)}{f_y(y)}$$

$$f_{x|y}(x|y=\frac{3}{4}) = \frac{f_{x,y}(x,y)}{1}$$

When $y = \frac{3}{4}$ $f_{x,y}(x,y)$ is in shaded region

When $x+y=1 \Rightarrow x=1-\frac{3}{4}=\frac{1}{4}$

$x+y=\frac{3}{2} \Rightarrow x=\frac{3}{2}-\frac{3}{4}=\frac{3}{4}$

So, $f_{x,y}(x|y=\frac{3}{4}) = \begin{cases} 2, & 0 \leq x \leq \frac{1}{4} \\ & \frac{3}{4} \leq x \leq 1 \\ 0, & \text{otherwise} \end{cases}$

$$f_{x,y}(x|y=\frac{3}{4}) = \begin{cases} 2, & 0 \leq x \leq \frac{1}{4} \text{ and } \frac{3}{4} \leq x \leq 1 \\ 0, & \text{else} \end{cases}$$

$$\therefore E(X|Y=\frac{3}{4}) = \frac{1}{2}$$

$$\text{Now } \text{Var}(X|Y=\frac{3}{4}) = \left\{ E(X^2|Y=\frac{3}{4}) - [E(X|Y=\frac{3}{4})]^2 \right\}$$

$$= \left[\frac{1}{12} - \left(\frac{1}{2} \right)^2 \right] = \frac{1}{12} - \frac{1}{4} = -\frac{1}{6}$$

$$\therefore \text{Var}(X|Y=\frac{3}{4}) = \frac{1}{12}$$

5.2

① $T_i - T_{i-1}$ = Random variable which is the number of shuffle that was done after $i-1$ cards went below n for next card to go below n .

Then, $P(T_i - T_{i-1} = 7)$

When exactly $i-1$ cards are under the original bottom card n . The chance that the current top card is inserted below

$$n \text{ is } \frac{(i-1)+1}{n} = i/n$$

For, current top card to go below n in the j th step, $(j-1)$ failure (if i goes to above bottom card n). Should occur and at j th step success (it goes below bottom card n)

$$P(T_i - T_{i-1} = j) = \frac{i}{n} \left(\frac{n-i}{n} \right)^{j-1}$$

$$= \frac{i}{n} \left(1 - \frac{i}{n} \right)^{j-1}$$

So, it is geometric distribution with $p = i/n$

Here, T_{i-1} depends on it. This is becoz. at a time the distribution of T_i depends on the immediate past. Where the next card would go depends only on the current card position. not on where other cards have been placed.

[2] Here card no. n is fixed. At beginning there is no card below n . Now, we observe that whenever we put a card below n , we put in a completely random position with respect to the rest of the card below n . So, we can say card below n are in uniformly random permutation.

Here T_{n-1} = first time that n becomes the top card of the deck so $(n-1)$ cards get below it.

At this point, we have a uniformly random permutation of all other cards $(n-1)$.

This is because bottom $(n-1)$ cards can take any 1 of possible $(n-1)!$ arrangement.

So, after $T_{n-1}+1$ shuffles all possible $n!$ arrangements are equally likely.

[3] Let T be the first time that ' n ' becomes the top card of the deck.

Then let $T = T_{n-1} + 1$

$$\Rightarrow T = (T_{n-1} - T_{n-2}) + (T_{n-2} - T_{n-3}) + \dots + (T_2 - T_1) + (T_1 - T_0)$$

$$E(T) = E(T_{n-1} - T_{n-2}) + \dots + E(T_1 - T_0)$$

$$= \frac{n}{n-1} + \frac{n}{n-2} + \dots + 1$$

($\because T_i - T_{i-1}$ has geometric distribution with $p = \frac{1}{n}$ so expected value in geometric $= \frac{1}{p} = n/i$ in case)

$$E(T) = n \left(\frac{1}{1} + \frac{1}{2} + \dots + \frac{1}{n-2} + \frac{1}{n-1} + \frac{1}{n} \right)$$

$$E(T) \approx n \cdot H_n \quad [H_n = n\text{th harmonic no}]$$

[4] Let T be the first time that bottom card 'n' becomes the top card of the deck.
Then the deck will be completely random. (proved in [2])

Also, $T = T_{n-1} + 1$

$$\& E(T) = n \cdot H_n \approx n \cdot \log n \quad [3]$$

Then if we shuffle K -times (say) so that with 99% chance the deck will be completely random or there is on $\frac{1}{100}$ chance for it not to be completely shuffled

$$\text{So, } P(T \geq K) \leq \frac{1}{100} \quad \text{--- (1)}$$

It means we need to shuffle more than K times with only 1% chance to not to be randomized shuffle

Using Markov inequality,

$$P(X \geq a) \leq \frac{E(X)}{a}$$

$$\Rightarrow P(T \geq K) \leq \frac{E(X)}{K}$$

$$\Rightarrow P(T \geq K) \leq \frac{n \log n}{K} \quad \text{--- (2)}$$

Comparing (1) & (2)

$$\frac{n \log n}{K} \leq \frac{1}{100}$$

$$\Rightarrow K \geq 100 n \log n$$

So, if we shuffle $100 n \log n$ times, with $\frac{99}{100}$ chance, the deck will be completely random

(5.3)

[1] The transition matrix P for this Markov chain is as follows.

$$P = \begin{bmatrix} 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ \frac{1}{4} & \frac{1}{4} & 0 & \frac{1}{4} & \frac{1}{4} & 0 \\ 0 & 0 & \frac{1}{2} & 0 & 0 & \frac{1}{2} \\ 0 & 0 & \frac{1}{2} & 0 & 0 & \frac{1}{2} \\ 0 & 0 & 0 & \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix}$$

[2] The chain is irreducible, because it is possible to go from any state to any other state. However, it is not aperiodic, because for any n even $p_{6,1}^{(n)}$ will be '0' and for any n odd $p_{6,5}^{(n)}$ will also be zero. This means that there is no power of P that would have all its entries strictly positive.

[3] for finding out the stationary distribution, let's consider the stationary state $S = [a_1 \ a_2 \ a_3 \ a_4 \ a_5 \ a_6]$

Now, we know that $S \cdot P = S$

\swarrow Stationary state \downarrow transition matrix \searrow stationary state

$$\Rightarrow [a_1 \ a_2 \ a_3 \ a_4 \ a_5 \ a_6] \begin{bmatrix} 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ \frac{1}{4} & \frac{1}{4} & 0 & \frac{1}{4} & \frac{1}{4} & 0 \\ 0 & 0 & \frac{1}{2} & 0 & 0 & \frac{1}{2} \\ 0 & 0 & \frac{1}{2} & 0 & 0 & \frac{1}{2} \\ 0 & 0 & 0 & \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix}$$

$$= [a_1 \ a_2 \ a_3 \ a_4 \ a_5 \ a_6]$$

Also, $a_1 + a_2 + a_3 + a_4 + a_5 + a_6 = 1$

$$\Rightarrow \left[\left(\frac{1}{4} a_3 \right) \left(\frac{1}{4} a_3 \right) \left(a_1 + a_2 + \frac{1}{2} a_4 + \frac{1}{2} a_5 \right) \left(\frac{1}{4} a_3 + \frac{1}{2} a_6 \right) \left(\frac{1}{4} a_3 + \frac{1}{2} a_6 \right) \right. \\ \left. \left(\frac{1}{2} a_4 + \frac{1}{2} a_5 \right) \right]$$

$$= [a_1 \ a_2 \ a_3 \ a_4 \ a_5 \ a_6]$$

$$\therefore \frac{1}{4} a_3 = a_1$$

$$\therefore \frac{1}{4} a_3 = a_2$$

$$a_1 + a_2 + \frac{1}{4} a_3 + \frac{1}{2} a_5 = a_3$$

③

$$\Rightarrow a_3 = 4a_1$$

$$\therefore a_3 = 4a_2$$

②

$$a_5 = \frac{1}{4} a_3 + \frac{1}{2} a_6$$

$$a_6 = \frac{1}{2} a_4 + \frac{1}{2} a_5$$

$$\Rightarrow 4a_5 = a_3 + 2a_6$$

$$2a_6 = a_4 + a_5$$

⑤

$$a_1 + a_2 + a_3 + a_4 + a_5 + a_6 = 1$$

⑥

Solving ①, ②, ③, ④, ⑤, ⑥ we get.

$$a_1 = \frac{1}{12}$$

$$a_2 = \frac{1}{12}$$

$$a_3 = \frac{4}{12}$$

$$a_4 = \frac{2}{12}$$

$$a_5 = \frac{2}{12}$$

$$a_6 = \frac{2}{12}$$

$$\therefore \text{Stationary dist } (S) = \left[\frac{1}{12} \ \frac{1}{12} \ \frac{4}{12} \ \frac{2}{12} \ \frac{2}{12} \ \frac{2}{12} \right]$$

[4]

Let t_1 be the expected number of steps until the chain hits state 5 for the first time given that $X_0 = 1$.

So, by total probability law,

$$t_1 = 1 + t_3$$

$$t_3 = 1 + \frac{t_1}{4} + \frac{t_2}{4} + \frac{t_4}{4} + \frac{t_5}{4}$$

$$t_2 = 1 + t_3$$

$$t_4 = 1 + \frac{t_3}{2} + \frac{t_5}{2}$$

$$t_5 = 1 + \frac{t_4}{2} + \frac{t_5}{2}$$

$$\Rightarrow t_4 = 1 + \frac{1}{2}(1 + \frac{t_3}{2}) + \frac{t_5}{2} = \frac{3}{2} + \frac{t_3}{4} + \frac{t_5}{2}$$

$$\therefore t_3 = 1 + \frac{1}{4}(1 + t_3) + \frac{1}{4}(1 + t_3) + \frac{t_4}{4}$$

$$\Rightarrow \frac{3t_3}{4} = \frac{3}{2} + \frac{t_3}{2}$$

$$\Rightarrow t_3 = \frac{3}{2} + \frac{t_3}{2} + \frac{t_4}{4}$$

$$\Rightarrow 3t_4 = 6 + 2t_3$$

$$t_3 = 3 + \frac{t_4}{2}$$

$$\Rightarrow 3t_4 = 6 + 6 + t_4$$

$$\Rightarrow 2t_4 = 12$$

$$\Rightarrow t_4 = 6$$

$$\Rightarrow t_3 = 3 + 6/2 = 6$$

$$\Rightarrow \boxed{t_1 = 1 + 6 = 7}$$

[5]

Let r_k be the mean return time to state k .

$$\text{Then } r_i = 1 + \sum_k t_k p_{ik}$$

When t_k is the expected time until the chain hits state k given $X_0 = k$

$$\begin{aligned} r_1 &= 1 + t_3 \cdot 1 \\ &= 1 + t_3 \end{aligned}$$

Here t_3 is the expected time until chain hits state 1 given $X_0 = 3$

Again by law of total probability

$$\text{we have, } t_1 = 0$$

$$t_2 = 1 + t_3$$

$$t_3 = 1 + t_3/4 + t_4/4 + t_5/4$$

$$t_4 = 1 + t_3/2 + t_5/2$$

$$t_5 = 1 + t_3/2 = t_4/2$$

$$t_6 = 1 + t_4/2 + t_5/2$$

$$\Rightarrow t_4 = t_5$$

$$\Rightarrow t_6 = 1 + t_4$$

$$\Rightarrow t_3 = 1 + t_3/4 + t_4/4$$

$$\Rightarrow t_3 = 1 + \frac{1+t_3}{4} + \frac{1}{2} + \frac{t_3}{4} + \frac{t_3}{4}$$

$$\Rightarrow 4t_3 = 4 + 1 + t_3 + 2 + t_3 + t_3$$

$$\Rightarrow 2t_3 = 7 + t_3$$

$$= 7 + (1 + t_3)$$

$$= 7 + (1 + 2t_3 - 2 - \frac{t_3}{2})$$

$$= 7 + (1 + 2t_3 - 2 - (\frac{1+t_3}{2}))$$

$$\Rightarrow \frac{t_3}{2} = 1\frac{1}{2}$$

$$\Rightarrow t_3 = 3$$

$$\underline{80} \quad r_1 = 1 + t_3 = 4$$

Ans

$$r_1 = \text{Mean time to return to state } l_m$$

$$r_1 = \frac{1}{\pi_1} = \frac{1}{1/4} = 4$$

from [3]

$$\therefore r_1 = 4$$