
IME692A Assignment 2

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1 Question 1

This question discusses about a Multiple Linear Regression Model used for the `cadata.txt` dataset which consists of features like **Median House Value, Median Income, Housing Median Age, Total Rooms, Population, Total Bedrooms, Latitude, Longitude, Households**. The equation describing the MLR model is represented below:

$$\ln(MHV) = \beta_0 + \beta_1(MI) + \beta_2(MI^2) + \beta_3(MI^3) + \beta_4 \ln(MA) \\ + \beta_5 \ln\left(\frac{TR}{P}\right) + \beta_6 \ln\left(\frac{B}{P}\right) + \beta_7 \ln\left(\frac{P}{H}\right) + \beta_8 \ln(H) + e$$

The beta values are mentioned in the questions and the errors are calculated for the given MLR model to the data. These errors are further tested for conditions like **zero expectation** since the independent variables should have no information regarding the expected value of the errors. The errors follow a **Gaussian distribution** with a zero mean and the errors are also independent of each other.

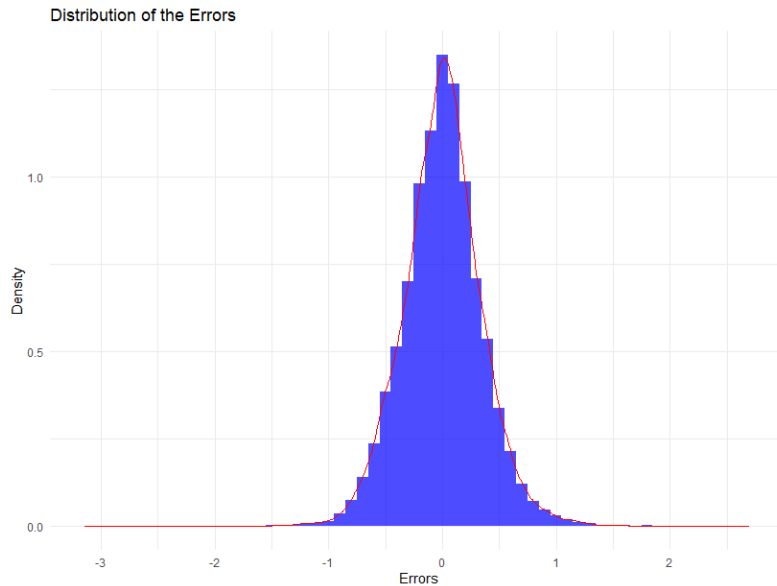


Figure 1: Distribution of Error values

Then we calculate the **Ordinary Least Squares Estimator** $\hat{\beta} = (X^T X)^{-1} X^T y$. As for the second portion of the question, we are required to choose **100** random samples with a **sample size** of **200**. We are then required to calculate the expected value of $\hat{\beta}$ which is $E[\hat{\beta}_i] \forall \beta_i, i = 0 \dots 8$. We then

β Index	Actual Value β_i	Expected value $E[\hat{\beta}_i]$	Density for $\hat{\beta}_i \sim \mathcal{N}(\beta_i, \sigma^2(X^T X)^{-1})$
0	11.4939	11.5111	0.5626
1	0.4790	0.5270	0.7425
2	-0.0166	-0.0173	0.8167
3	-0.0002	-0.0350	0.7794
4	0.1570	0.1306	0.7753
5	-0.8582	-0.8616	0.7756
6	0.8043	0.8203	0.7716
7	-0.4077	-0.4231	0.7746
8	0.0477	0.0440	0.8074

Table 1: Expected Values and pdfs for $\hat{\beta}_i$

calculate the probability density values for each of the $\hat{\beta}_i$ corresponding to the distribution $\hat{\beta}_i \sim \mathcal{N}(\beta_i, \sigma^2(X^T X)^{-1})$. The given table shows the expected values and the probability densities for β_i 's.

The probability distributions are also represented in the following bar plot:

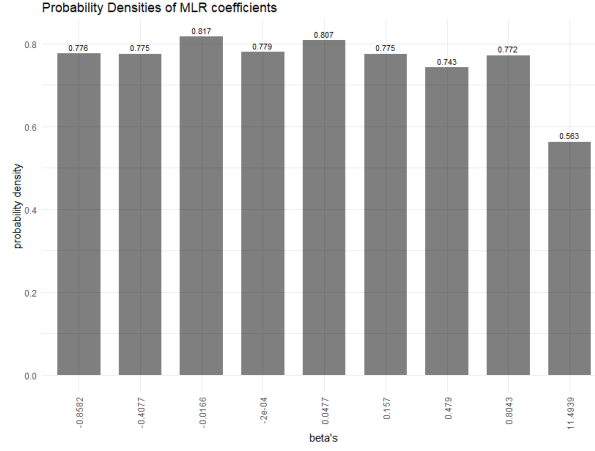


Figure 2: Caption

The second part involves the calculations of **90%**, **95%** and **99%** confidence intervals for cases corresponding to **known variance (Z distribution)** and **unknown variance (t distribution)**. The plots for the confidence intervals for each of the $\hat{\beta}_i$'s are represented below:

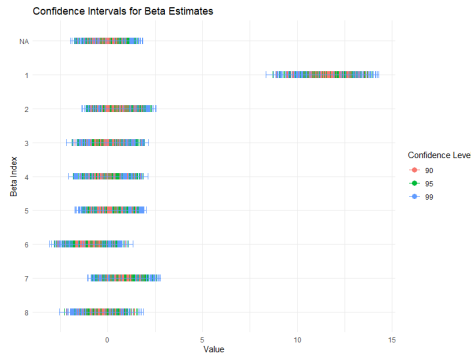


Figure 3: T distribution Confidence Intervals

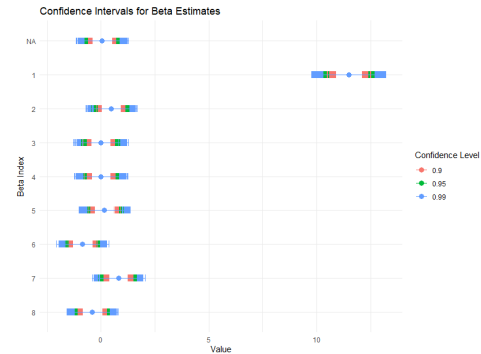


Figure 4: Z distribution Confidence Intervals

The data samples from **20600** to **20640** are calculated using the estimated values of $\hat{\beta}$ and compared to the original one. The total deviation is around **55.0203** . The table for true values vs estimated values is given below:

Index	Value	Index	Value
20600	-2.43861331	20621	1.33648232
20601	-0.55377642	20622	-1.18527660
20602	-1.80225954	20623	-1.02931905
20603	-2.51177635	20624	-0.84893814
20604	-2.89936993	20625	-0.48928029
20605	-1.88715946	20626	1.71713742
20606	-2.35839122	20627	-1.45428905
20607	-2.02958382	20628	0.43717245
20608	-1.98800848	20629	-1.20333717
20609	-1.94451654	20630	-2.57668380
20610	-1.61251297	20631	-0.08315018
20611	-2.63626494	20632	-0.19283215
20612	-2.68061651	20633	-0.54832525
20613	-2.27428666	20634	-1.34027409
20614	-2.50547289	20635	-0.06188863
20615	-1.54043292	20636	-2.28538561
20616	-1.23961149	20637	-0.68049152
20617	-1.74645184	20638	-2.15750645
20618	0.08541121	20639	-1.89336037
20619	-1.41760042	20640	-1.52173984
20620	-0.97845290		

Table 2: Data Values

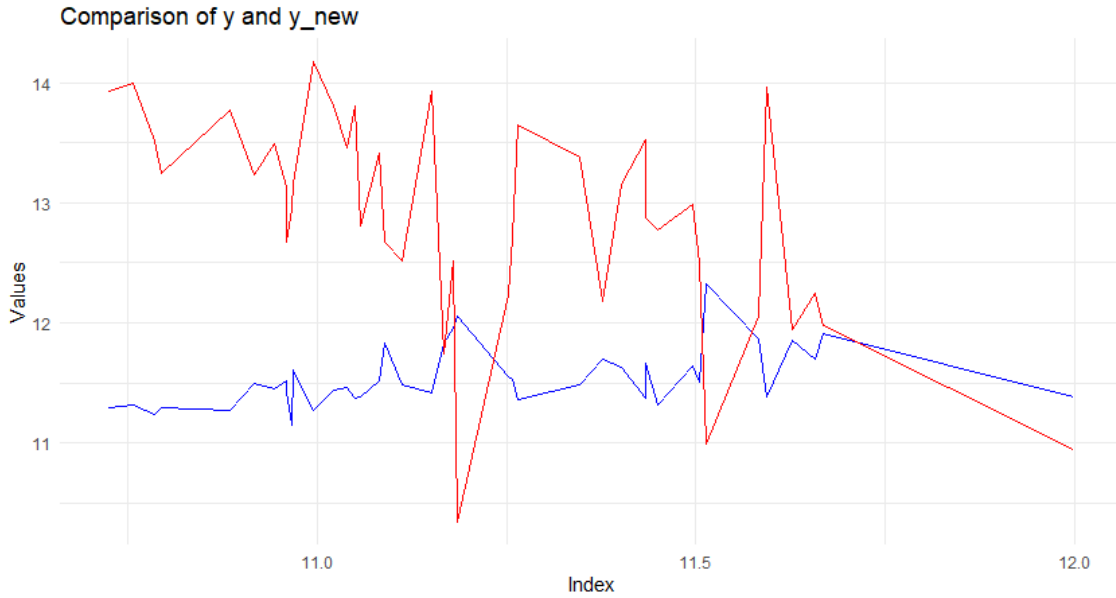


Figure 5: Variation of y and $y_{\text{estimated}}$

As for the **b** part, the tables for the loss values across different models:

- for **Model 1**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.6723**.

- for **Model 2**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.6723**.
- for **Model 3**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.67422**.
- for **Model 4**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.67422**.
- for **Model 5**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.67423**.
- for **Model 6**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.67423**.
- for **Model 7**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.67422**.
- for **Model 8**, the value of $\hat{\beta}$ is $[11.39946, 0.28464, -0.36161, 0.04353, 0.4154, -0.9716, 0.78819, -0.42018, 0.4959]^T$ and the value of $\hat{\theta}$ is **11.67423**.

Model no.	$\lambda 1$	$\lambda 2$	$\lambda 3$	$\lambda 4$
1	23.3076	47.5032	71.6987	95.8943
2	-0.7091	-0.5303	-0.3516	-0.1729
3	23.2900	47.5801	71.8702	96.1603
4	-0.7997	-0.5994	-0.3992	-0.1989
5	-0.3949	-0.2958	-0.1968	-0.0978
6	-0.3997	-0.2994	-0.1992	-0.0989
7	0.0048	0.0036	0.0024	0.0012
8	0.00026	0.00005	0.0007	0.0010

Table 3: Risk Values for $\hat{\lambda}_i$

Model no.	$\lambda 1$	$\lambda 2$	$\lambda 3$	$\lambda 4$
1	0.82236	1.64473	2.46709	3.28940
2	0.0896	0.1793	0.2690	0.3587
3	0.8223	1.6447	2.4670	3.2894
4	0.0896	0.1793	0.2690	0.3586
5	0.00081	0.00076	0.00071	0.00067
6	0.00013	0.00025	0.00037	0.00050
7	0.00081	0.00076	0.00071	0.00067
8	0.00013	0.00025	0.00034	0.00050

Table 4: Loss values for $\hat{\lambda}_i$

2 Question 2

2.1 Chi Square Distribution

For the first part of the problem, we are asked to find the **probability distribution function** (pdf) of X_{\min} and X_{\max} of a set of random variables X_i for all i up to n . Here $X_{\min} = \min(X_1, X_2, X_3, \dots, X_n)$ and $X_{\max} = \max(X_1, X_2, X_3, \dots, X_n)$ where the random variables follow a **Chi-square distribution** with pdf represented as

$$f(x) = \frac{1}{2^{n/2}\Gamma(n/2)} x^{n/2-1} e^{-x/2}$$

The calculations for the $f_{X_{\min}}(x)$ and $f_{X_{\max}}(x)$ are presented as follows:

$$F_{X_{\min}}(x) = 1 - P(X_{\min} > x) \quad (1)$$

$$F_{X_{\min}}(x) = 1 - P(\min(X_1, X_2, X_3, \dots, X_n) > x) \quad (2)$$

$$F_{X_{\min}}(x) = 1 - P(X_1 > x)P(X_2 > x)P(X_3 > x) \dots P(X_n > x) \quad (3)$$

$$F_{X_{\min}}(x) = 1 - \prod_{i=1}^n P(X_i > x) \quad (4)$$

$$F_{X_{\min}}(x) = 1 - \prod_{i=1}^n (1 - F(x)) \quad \text{where } F(x) = P(X \leq x) \quad (5)$$

Now, for the calculation of the **Cumulative probability distribution** $F(x)$, we have:

$$F(x) = \int_{-\infty}^x f(t) dt \quad (6)$$

$$F(x) = \int_{-\infty}^0 f(t) dt + \int_0^x f(t) dt \quad (7)$$

$$F(x) = \int_0^x \frac{1}{2^{n/2}\Gamma(n/2)} t^{n/2-1} e^{-t/2} dt \quad (8)$$

Performing a variable substitution of $p = t/2$, we get the following integral which can be expressed in terms of the **lower incomplete gamma function** where $\gamma(n, x) = \int_0^x t^{n-1} e^{-t} dt$.

$$F(x) = \int_0^{x/2} \frac{1}{2^{n/2}\Gamma(n/2)} (2p)^{n/2-1} e^{-p} 2dp \quad (9)$$

$$= \int_0^{x/2} \frac{1}{2^{n/2}\Gamma(n/2)} 2^{n/2} p^{n/2-1} e^{-p} 2dp \quad (10)$$

$$= \frac{\int_0^{x/2} e^{-p} p^{n/2-1} dp}{\Gamma(n/2)} \quad (11)$$

$$= \frac{\gamma(n/2, x/2)}{\Gamma(n/2)} \quad (12)$$

$$F_{X_{\min}} = 1 - \left(1 - \frac{\gamma(n/2, x/2)}{\Gamma(n/2)}\right)^{n-1} \quad (13)$$

$$f_{X_{\min}}(x) = \frac{dF_{X_{\min}}}{dx} \quad (14)$$

$$f_{X_{\min}}(x) = \frac{nx^{n/2-1}e^{-x/2}}{2^{n/2}\Gamma(n/2)} \left(1 - \left(1 - \frac{\gamma(n/2, x/2)}{\Gamma(n/2)}\right)^{n-1}\right) \quad (15)$$

The steps for the calculation of $F_{X_{\max}}(x)$ are mentioned as follows:

$$F_{X_{\max}}(x) = P(X_{\max} \leq x) \quad (16)$$

$$F_{X_{\max}}(x) = P(\max(X_1, X_2, X_3, \dots, X_n) \leq x) \quad (17)$$

$$F_{X_{\max}}(x) = P(X_1 \leq x)P(X_2 \leq x)P(X_3 \leq x) \dots P(X_n \leq x) \quad (18)$$

$$F_{X_{\max}}(x) = \prod_{i=1}^n P(X_i \leq x) \quad (19)$$

$$F_{X_{\max}}(x) = \prod_{i=1}^n F(x) \quad \text{where } F(x) = P(X \leq x) \quad (20)$$

Performing a similar variable substitution of $p = t/2$ for the calculation of $F_{X_{\max}}(x)$ we get the following integral which can be expressed in terms of the **lower incomplete gamma function** where $\gamma(n, x) = \int_0^x t^{n-1} e^{-t} dt$.

$$F(x) = \int_0^{x/2} \frac{1}{2^{n/2} \Gamma(n/2)} (2p)^{n/2-1} e^{-p} 2dp \quad (21)$$

$$= \int_0^{x/2} \frac{1}{2^{n/2} \Gamma(n/2)} 2^{n/2} p^{n/2-1} e^{-p} 2dp \quad (22)$$

$$= \frac{\int_0^{x/2} e^{-p} p^{n/2-1} dp}{\Gamma(n/2)} \quad (23)$$

$$= \frac{\gamma(n/2, x/2)}{\Gamma(n/2)} \quad (24)$$

$$F_{X_{\max}} = \left(\frac{\gamma(n/2, x/2)}{\Gamma(n/2)} \right)^{n-1} \quad (25)$$

$$f_{X_{\max}}(x) = \frac{dF_{X_{\max}}}{dx} \quad (26)$$

$$f_{X_{\max}}(x) = \frac{nx^{n/2-1} e^{-x/2}}{2^{n/2} \Gamma(n/2)} \left(\frac{\gamma(n/2, x/2)}{\Gamma(n/2)} \right)^{n-1} \quad (27)$$

2.2 t distribution

For the first part of the problem, we are asked to find the **probability distribution function** (pdf) of X_{\min} and X_{\max} of a set of random variables X_i for all i up to n . Here $X_{\min} = \min(X_1, X_2, X_3, \dots, X_n)$ and $X_{\max} = \max(X_1, X_2, X_3, \dots, X_n)$ where the random variables follow a **t distribution** with pdf represented as

$$f(x) = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{\pi n} \Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{(n+1)}{2}}$$

Similar to the previous part, Now, for the calculation of the **Cumulative probability distribution** $F(x)$, we have:

$$F(x) = \int_{-\infty}^x f(u) du \quad (28)$$

$$F(x) = \int_{-\infty}^0 f(u) du + \int_0^x f(u) dt \quad (29)$$

$$F(x) = \frac{1}{2} + \left(\frac{1}{2} - \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{\pi n} \Gamma\left(\frac{n}{2}\right)} \int_x^\infty \left(1 + \frac{u^2}{n}\right)^{-\frac{n+1}{2}} dt \right) \quad (30)$$

$$F(x) = 1 - I_{x(u)}\left(\frac{n}{2}, \frac{1}{2}\right) \quad \text{where } x(u) = \frac{n}{u^2 + n} \quad (31)$$

$$(32)$$

Here the term $I_{t(u)}(t)(n/2, 1/2)$ refers to the **incomplete beta function**. The final steps for calculations of $f_{X_{\min}}(x)$ and $f_{X_{\max}}(x)$ are:

$$F_{X_{\max}} = \left(1 - I_{x(u)}\left(\frac{n}{2}, \frac{1}{2}\right)\right)^n \quad (33)$$

$$f_{X_{\max}}(x) = n \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi}\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}} \left(1 - I_{x(u)}\left(\frac{n}{2}, \frac{1}{2}\right)\right)^{n-1} \quad (34)$$

$$F_{X_{\min}} = \left(I_{x(u)}\left(\frac{n}{2}, \frac{1}{2}\right)\right)^n \quad (35)$$

$$f_{X_{\min}}(x) = n \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi}\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}} \left(I_{x(u)}\left(\frac{n}{2}, \frac{1}{2}\right)\right)^{n-1} \quad (36)$$