CS480/680: Introduction to Machine Learning

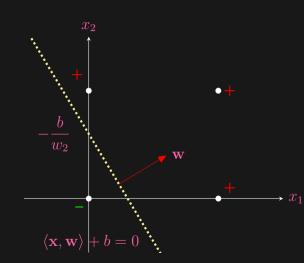
Lec 03: Logistic Regression

Yaoliang Yu



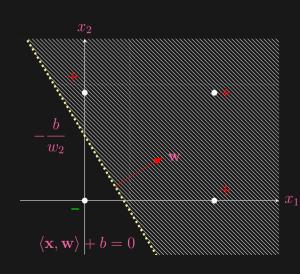
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- How confident we are about the prediction ŷ?
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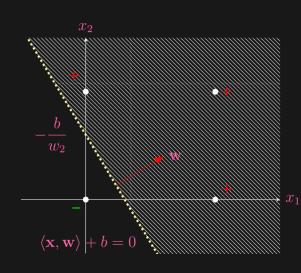
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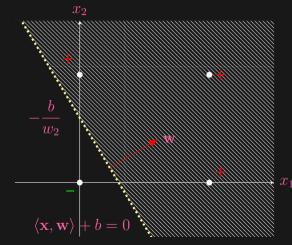
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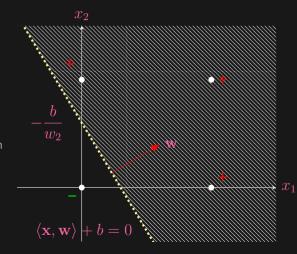


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 - in fact was used in multi-class preceptron
 - real-valued: hard to interpret
 - many ways to transform into $\left[0,1\right]$
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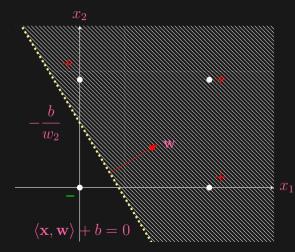


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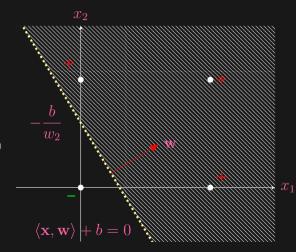


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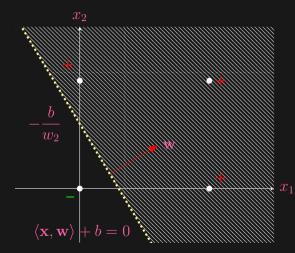


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- $Y_i \overset{i.i.d.}{\sim} \operatorname{Bernoulli}(q)$ for some $q \in [0, 1]$
- How to evaluate a probabilistic forecast \hat{p} ?
- Scoring function: $s: \mathcal{Y} \times [0,1] \to \mathbb{R}, \ s(\mathsf{y},p)$ scores the "fitness"
- Scoring rule: $\mathbb{S}:[0,1]\times[0,1]\to\mathbb{R}$, $\mathbb{S}(q,p):=\mathbb{E}_{\mathsf{Y}\sim\mathrm{Bernoulli}(q)}[s(\mathsf{Y},p)]$
- (Strict) properness (truthfulness): $q = \operatorname{argmin}_p \mathbb{S}(q, p)$
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- The resulting entropy is exactly Shannon's entropy
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- Model postulates $Y|X = x \sim Bernoulli(p(x; w))$, i.e. Pr(Y = 1|X = x) = p(x; w)
- Given $(\mathbf{X}_i, \mathbf{y}_i), i = 1, \dots, n$, assume independence

$$Pr(\mathsf{Y}_1 = \mathsf{y}_1, \dots, \mathsf{Y}_n = \mathsf{y}_n | \mathsf{X}_1 = \mathbf{x}_1, \dots, \mathsf{X}_n = \mathbf{x}_n) = \prod_{i=1}^n Pr(\mathsf{Y}_i = \mathsf{y}_i | \mathsf{X}_i = \mathbf{x}_i)$$
$$= \prod_{i=1}^n [p(\mathbf{x}_i; \mathbf{w})]^{\mathsf{y}_i} [1 - p(\mathbf{x}_i; \mathbf{w})]^{1-\mathsf{y}_i}$$

• Maximizing the conditional log-likelihood:

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The Logit Transform

• $p(\mathbf{x}; \mathbf{w}) : \mathcal{X} \to [0, 1]$, how to parameterize using \mathbf{w}

$$\log n(\mathbf{x}; \mathbf{w}) = (\mathbf{x}; \mathbf{w})^{\frac{1}{2}}$$

- Logit transform: $\log \frac{p(\mathbf{x}; \mathbf{w})}{1 p(\mathbf{x}; \mathbf{w})} = \langle \mathbf{x}, \mathbf{w} \rangle$
 - i.e., bile vada iato ia ali alifile idilictio
- ullet Equivalently, the sigmoid transformation: $p(\mathbf{x};\mathbf{w}) = rac{1}{1+\exp(-\langle \mathbf{x},\mathbf{w}
 angle})$

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• $p(\mathbf{x}; \mathbf{w}) : \mathcal{X} \to [0, 1]$, how to parameterize using **w**?

$$- p(\mathbf{x}; \mathbf{w}) = \langle \mathbf{x}, \mathbf{w} \rangle?$$
$$- \log p(\mathbf{x}; \mathbf{w}) = \langle \mathbf{x}, \mathbf{w} \rangle?$$

• Logit transform: $\log \frac{p(\mathbf{x}; \mathbf{w})}{1 - p(\mathbf{x}; \mathbf{w})} = \langle \mathbf{x}, \mathbf{w} \rangle$

• Equivalently, the sigmoid transformation: $p(\mathbf{x};\mathbf{w}) = \frac{1}{1+\exp(-\langle \mathbf{x},\mathbf{w} \rangle)}$

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Logistic Regression

$$\min_{\mathbf{w}} \sum_{i=1}^{n} -\mathsf{y}_i \log[p(\mathbf{x}_i; \mathbf{w})] - (1 - \mathsf{y}_i) \log[1 - p(\mathbf{x}_i; \mathbf{w})]$$

ullet Plug in the parameterization $p(\mathbf{x}; \mathbf{w}) = rac{1}{1 + \exp(-\langle \mathbf{x}, \mathbf{w}
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$$\min_{\mathbf{w}} \sum_{i=1}^{n} \left[\log[1 + \exp(-\langle \mathbf{x}_i, \mathbf{w} \rangle)] + (1 - \mathsf{y}_i) \langle \mathbf{x}_i, \mathbf{w} \rangle \right]$$

ullet Note the label encoding $\mathsf{y}_i \in \{0,1\}$; if instead, $\mathsf{y}_i \in \{\pm 1\}$, then

$$\min_{\mathbf{w}} \sum_{i=1}^{\infty} \left[\log[1 + \exp(-\mathsf{y}_i \langle \mathbf{x}_i, \mathbf{w} \rangle)] \right]$$
logistic loss

Logistic Regression

$$\min_{\mathbf{w}} \sum_{i=1}^{m} -\mathsf{y}_i \log[p(\mathbf{x}_i; \mathbf{w})] - (1 - \mathsf{y}_i) \log[1 - p(\mathbf{x}_i; \mathbf{w})]$$

• Plug in the parameterization $p(\mathbf{x}; \mathbf{w}) = \frac{1}{1 + \exp(-\langle \mathbf{x}, \mathbf{w} \rangle)}$

$$\min_{\mathbf{w}} \sum_{i=1}^{n} \left[\log[1 + \exp(-\langle \mathbf{x}_i, \mathbf{w} \rangle)] + (1 - \mathsf{y}_i) \langle \mathbf{x}_i, \mathbf{w} \rangle \right]$$

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Logistic Regression

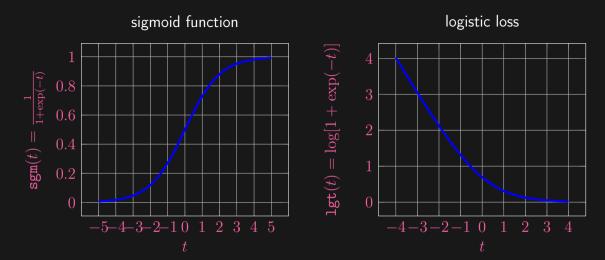
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D. R. Cox. "The Regression Analysis of Binary Sequences". Journal of the Royal Statistical Society. Series B (Methodological), vol. 20, no. 2 (1958), pp. 215–242.

$$p(\mathbf{x}; \mathbf{w}) = \operatorname{sgm}(\langle \mathbf{x}, \mathbf{w} \rangle) = \frac{1}{1 + \exp(-\langle \mathbf{x}, \mathbf{w} \rangle)}$$

- ullet $\hat{f y}=1$ iff $p({f x};{f w})=\Pr({f Y}=1|{f X}={f x})>rac{1}{2}$ iff $\langle {f x},{f w}
 angle>0$
- Decision boundary remains to be $H := \{\mathbf{x} : \langle \mathbf{x}, \mathbf{w} \rangle = 0\}$
- Can predict $\hat{\mathbf{y}} = \operatorname{sign}(\langle \mathbf{x}, \mathbf{w} \rangle)$ as before, but now with confidence $p(\mathbf{x}; \mathbf{w})$

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• Logistic regression estimates the posterior probability $\eta(\mathbf{x}) := \Pr(\mathsf{Y} = 1 | \mathsf{X} = \mathbf{x})$ under the linear odds ratio assumption

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- Possible to do the comparison without estimating $\eta(\mathbf{x})$ explicitly

sufficient but not necessary, be lazy

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L03 11/10

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L03 11/16

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L03 11/1₁

$$p(\mathbf{x}; \mathbf{w}) = F(\langle \mathbf{x}, \mathbf{w} \rangle)$$

- ullet $F:\mathbb{R}
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L03 12/16

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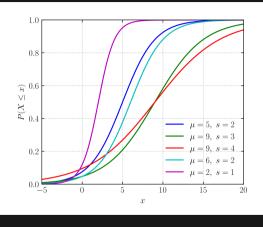
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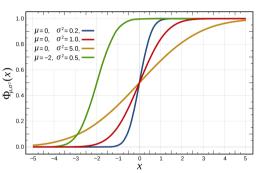
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L03 13/16

$$\min_{\mathbf{w}} \sum_{i=1}^{n} \left[\log[1 + \exp(-\mathsf{y}_{i} \langle \mathsf{x}_{i}, \mathsf{w} \rangle)] \right]$$

• Newton's algorithm

$$\mathbf{w} \leftarrow \mathbf{w} - \eta \cdot [\nabla^2 f(\mathbf{w})]^{-1} \cdot \nabla f(\mathbf{w})$$

- The gradient $\nabla f(\mathbf{w}) = \mathbf{X}(\hat{\mathbf{p}} \frac{\mathbf{y}+1}{2})$: changing target
- ullet The Hessian $abla^2 f(\mathbf{w}) = \sum_i \hat{p}_i (1-\hat{p}_i) \mathbf{x}_i \mathbf{x}_i^ op$: weighted by confidence
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Linear Regression vs. Logistic Regression

- least-squares: $\sum_{i=1}^{n} (y_i \hat{y}_i)^2$
- prediction: $\hat{y}_i = \langle \mathbf{x}_i, \mathbf{w} \rangle$
- objective: $\|\mathbf{y} \hat{\mathbf{y}}\|_2^2$
- grad: $\mathbf{w} \leftarrow \mathbf{w} \eta \mathbf{X}(\hat{\mathbf{y}} \mathbf{y})$
- Newton: $\mathbf{w} \leftarrow \mathbf{w} \eta (\mathbf{X} \mathbf{X}^{\top})^{-1} \mathbf{X} (\hat{\mathbf{y}} \mathbf{y})$

- cross-entropy: $\sum_{i=1}^n -\frac{1+y_i}{2}\log \hat{p}_i \frac{1-y_i}{2}\log (1-\hat{p}_i)$
- prediction: $\hat{y}_i = \text{sign}(\langle \mathbf{x}_i, \mathbf{w} \rangle), \ \hat{p}_i = \text{sgm}(\langle \mathbf{x}_i, \mathbf{w} \rangle)$
- objective: $\mathsf{KL}(\frac{1+\mathbf{y}}{2}\|\hat{\mathbf{p}})$
- grad: $\mathbf{w} \leftarrow \mathbf{w} \eta \mathbf{X} (\hat{\mathbf{p}} \frac{1+\mathbf{y}}{2})$
- Newton: $\mathbf{w} \leftarrow \mathbf{w} \eta (\mathbf{X} \hat{S} \mathbf{X}^{\top})^{-1} \mathbf{X} (\hat{\mathbf{p}} \frac{1+\mathbf{y}}{2})$

- Diagonal weight matrix $\hat{S} = \operatorname{diag} (\hat{\mathbf{p}} \odot (1 \hat{\mathbf{p}}))$
- Logistic regression = iteratively weighted linear regression

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• objective:
$$\|\mathbf{y} - \hat{\mathbf{y}}\|_2^2$$

• grad:
$$\mathbf{w} \leftarrow \mathbf{w} - \eta \mathbf{X}(\hat{\mathbf{y}} - \mathbf{y})$$

• Newton:
$$\mathbf{w} \leftarrow \mathbf{w} - \eta (\mathbf{X} \mathbf{X}^{\top})^{-1} \mathbf{X} (\hat{\mathbf{y}} - \mathbf{y})$$

• cross-entropy:
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• Softmax parameterization

$$\Pr(\mathsf{Y} = k | \mathsf{X} = \mathbf{x}; \mathsf{W} = [\mathsf{w}_1, \dots, \mathsf{w}_c]) = \frac{\exp(\langle \mathsf{x}, \mathsf{w}_k \rangle)}{\sum_{l=1}^c \exp(\langle \mathsf{x}, \mathsf{w}_l \rangle)}$$

- F I = G)
- Minimizing again the logarithmic loss:

$$\min_{\mathbf{W}} \hat{\mathbb{E}} \left[-\log \frac{\exp(\langle \mathbf{X}, \mathbf{w_Y} \rangle)}{\sum_{l=1}^{c} \exp(\langle \mathbf{X}, \mathbf{w_l} \rangle)} \right]$$

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- nonnegative and sum to 1
- Encode $y \in \{1, \dots, c\}$
- Minimizing again the logarithmic loss:

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