

The co-variance matrix E is :

4.0000	-2.7500
-2.7500	2.0000

K_cap is :

-2.0000	1.0000
1.5000	-0.5000

The Best Linear Unbiased Estimate (BLUE) for x is:

0.6194

0.4591

The co-variance matrix E is :

0.0679	-0.0260
-0.0260	0.1129

K_cap is :

-0.1050	0.0525	0.1895
0.1872	0.1564	-0.1313

The Best Linear Unbiased Estimate (BLUE) for x is:

-1.4303
1.8791

The co-variance matrix E is :

0.0487	0.0054
0.0054	0.0618

K_cap is :

-0.0540	0.1046	0.1480	-0.0518
0.1041	0.0715	-0.0637	0.0843

The Best Linear Unbiased Estimate (BLUE) for x is:

-1.2201
1.5368

$$2(a) \quad M = \begin{bmatrix} -1 \\ 0 \\ 1 \end{bmatrix}$$

$$\Sigma = \begin{bmatrix} 2 & 2 & 1 \\ 2 & 4 & 2 \\ 1 & 2 & 2 \end{bmatrix}$$

$$\Sigma_{11} = \begin{bmatrix} 2 & 2 \\ 2 & 4 \end{bmatrix} ; \quad \Sigma_{12} = \begin{bmatrix} 1 \\ 2 \end{bmatrix}$$

$$\Sigma_{21} = \begin{bmatrix} 1 & 2 \end{bmatrix} ; \quad \Sigma_{22} = \begin{bmatrix} 2 \end{bmatrix}$$

$$M_{xy} = \begin{bmatrix} -1 \\ 0 \end{bmatrix} ; \quad M_z = \begin{bmatrix} 1 \end{bmatrix}$$

$$M_{12} = M_1 + \sum_{i2} \Sigma^{-1} (x_2 - M_2)$$

$$\Rightarrow M_{xy/z} = M_{xy} + \sum_{i2} \Sigma^{-1} (x_2 - M_z)$$

$$M_{xy/z} = \begin{bmatrix} -1 \\ 0 \end{bmatrix} + \begin{bmatrix} 1 \\ 2 \end{bmatrix} \left(\frac{1}{2} \right) (z - 1)$$

$$= \begin{bmatrix} -1 \\ 0 \end{bmatrix} + \begin{bmatrix} 1 \\ 2 \end{bmatrix} (0.5z - 0.5)$$

$$= \begin{bmatrix} -1 \\ 0 \end{bmatrix} + \begin{bmatrix} 0.5z - 0.5 \\ 2(0.5z - 0.5) \end{bmatrix}$$

$$= \begin{bmatrix} -1 \\ 0 \end{bmatrix} + \begin{bmatrix} 0.5z - 0.5 \\ z - 1 \end{bmatrix}$$

$$M_{xy/z} = \begin{bmatrix} 0.5z - 1.5 \\ z - 1 \end{bmatrix}$$

$$\Sigma_{y|z} = \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21}$$

$$= \begin{bmatrix} 2 & 2 \\ 2 & 4 \end{bmatrix} - \begin{bmatrix} 1 \\ 2 \end{bmatrix} \begin{pmatrix} 1 \\ 2 \end{pmatrix}^{-1} \begin{bmatrix} 1 & 2 \end{bmatrix}$$

$$= \begin{bmatrix} 2 & 2 \\ 2 & 4 \end{bmatrix} - \begin{bmatrix} 1 & 2 \\ 2 & 4 \end{bmatrix} \begin{pmatrix} 1 \\ 2 \end{pmatrix}^{-1}$$

$$= \begin{bmatrix} 2 & 2 \\ 2 & 4 \end{bmatrix} - \begin{bmatrix} 0.5 & 1 \\ 1 & 2 \end{bmatrix} = \begin{bmatrix} 1.5 & 1 \\ 1 & 2 \end{bmatrix}$$

$$\Sigma_{xy|z} = \begin{bmatrix} 1.5 & 1 \\ 1 & 2 \end{bmatrix}$$

b) $\Sigma_{11} = 1.5$; $\Sigma_{22} = 1$; $\Sigma_{11} = 1$; $\Sigma_{22} = 2$

$$N_y = \mu_y + \Sigma_{12} \Sigma_{22}^{-1} (y - \mu_y) =$$

$$\Sigma_y = \Sigma_{12} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21}$$

$$N_y = 0.5z - 1.5 + 1 \cdot \begin{pmatrix} 1 \\ 2 \end{pmatrix} (y - 2 + 1) = 0.5y - 1$$

$$\Sigma_{xy|z} = 1.5 - 1 \cdot 1 \cdot 1 = 1$$

$$\Rightarrow N_{x|y} = 0.5y - 1 ; \Sigma_{x|y} = 1$$

(c) $x|y, z = z$

$$\Sigma_{11} = [2] \quad \Sigma_{12} = [2 \ 1]$$

$$\Sigma_{22} = \begin{bmatrix} 4 & 2 \\ 2 & 2 \end{bmatrix} \quad \Sigma_{21} = \begin{bmatrix} 2 \\ 1 \end{bmatrix}$$

~~$$M_{x|y,z} = M_x + \Sigma_{12} \Sigma_{22}^{-1} \left(\begin{bmatrix} y \\ z \end{bmatrix} - M_{y,z} \right)$$~~

~~$$= -1 + [2 \ 1] \frac{1}{2} \left\{ \begin{bmatrix} y \\ z \end{bmatrix} - \begin{bmatrix} 0 \\ 1 \end{bmatrix} \right\}$$~~

~~$$= -1 + [2 \ 1] \frac{1}{2} \left\{ \begin{bmatrix} y \\ z-1 \end{bmatrix} \right\}$$~~

~~$$= -1 + [1 \ 0.5] \begin{bmatrix} y \\ z-1 \end{bmatrix}$$~~

$$\Sigma_{x|y,z} = 2 - [2 \ 1] \frac{1}{4} \begin{bmatrix} 2 & -2 \\ -2 & 4 \end{bmatrix} \begin{bmatrix} 2 \\ 1 \end{bmatrix} = 2 - [2 \ 1] \frac{1}{4} \begin{bmatrix} 2 \\ 0 \end{bmatrix} = 1.$$

$$M_{y|y,z} = M_x + \Sigma_{12} \Sigma_{22}^{-1} \left(\begin{bmatrix} y \\ z \end{bmatrix} - M_{y,z} \right)$$

$$= -1 + [2 \ 1] \frac{1}{4} \begin{bmatrix} 2y - 2z + 2 \\ -2y + 4z - 4 \end{bmatrix} = 0.5y - 1$$

(d) The results of b, c are identical

↳ It does not matter if the distribution conditioned first to z then to y or to both z, y .

problem (3). $S = \{y_1, \dots, y_n\} \rightarrow$ linearly independent, $M = \text{span}\{S\}$

$$\hat{x}_k = \underset{MEM}{\text{argmin}} \|x - m\|$$

a) $y_{k+1} \perp M_k$ $P_i: \exists \beta \text{ s.t. } \hat{x}_{k+1} = \hat{x}_k + \beta y_{k+1}$

$x - \hat{x}_k \perp M$ by pre-projection theorem.

$$x = \hat{x}_k + (x - \hat{x}_k) = \hat{x}_k + m \quad \text{where } m \in M_k^\perp$$

$M_{k+1} = \text{span}\{y_1, \dots, y_n, y_{k+1}\}$ { after we get the new measurement $y_{k+1} \in M_k^\perp$ }

$y_{k+1} \perp M_k$; $y_{k+1} \in M_k^\perp$ then using gram-Schmidt we can compute an orthogonal basis for M^\perp .

$$B := \{y_{k+1}, v_1, v_2, \dots, v_f\} \Rightarrow M^\perp = \beta y_{k+1} + \beta_1 v_1 + \dots + \beta_f v_f.$$

$$x = \hat{x}_k + \beta y_{k+1} \in M_{k+1}$$

$$\begin{cases} \{v_1, v_2, \dots, v_f\} \perp \{y_1, y_2, \dots, y_n\} \\ \{v_1, v_2, \dots, v_f\} \perp y_{k+1} \end{cases}$$

Thus $\{v_1, v_2, \dots, v_f\} \perp \text{span}\{y_1, y_2, \dots, y_{k+1}\} = M_{k+1}$

$$\Downarrow$$

$$m_{k+1}^\perp = \beta_1 v_1 + \dots + \beta_f v_f \in M_{k+1}^\perp$$

\Downarrow

$$x = m_{k+1} + m_{k+1}^\perp \Rightarrow m_{k+1}^\perp = x - m_{k+1} \Rightarrow (x - m_{k+1}) \perp M_{k+1}$$

$$x - \hat{x}_{k+1} \perp M_{k+1}$$

$$\langle x - \hat{x}_{k+1}, y_{k+1} \rangle = 0$$

$$\Rightarrow \langle \hat{x}_{k+1}, y_{k+1} \rangle = \langle x, y_{k+1} \rangle$$

$$\langle \hat{x}_k + \beta y_{k+1}, y_{k+1} \rangle = \langle x, y_{k+1} \rangle$$

$$\text{Since } y_{k+1} \perp M_k; \hat{x}_k \in M_k: \langle \hat{x}_k, y_{k+1} \rangle = 0$$

$$\Rightarrow \langle \hat{x}_k + \beta y_{k+1}, y_{k+1} \rangle = \beta \langle y_{k+1}, y_{k+1} \rangle = \langle x, y_{k+1} \rangle$$

$$\Rightarrow \beta = \frac{\langle x, y_{k+1} \rangle}{\langle y_{k+1}, y_{k+1} \rangle}$$

$$\hat{y}_{k+1}/k = \arg \min_{m \in M_k} \|y_{k+1} - m\|$$

$$y_{k+1} - \hat{y}_{k+1}/k \perp M_k \text{ by pre-projection theorem.}$$

$$v = y_{k+1} - \hat{y}_{k+1}/k, \text{ we get } v \perp \{y_1, y_2, \dots, y_k\}$$

we already showed that

$$\hat{x}_{k+1} = \hat{x}_k + \beta v$$

$$\beta = \frac{\langle x, v \rangle}{\langle v, v \rangle}$$

$$\Rightarrow \begin{cases} \hat{x}_{k+1} = \hat{x}_k + \beta (y_{k+1} - \hat{y}_{k+1}/k) \\ \beta = \frac{\langle x, y_{k+1} - \hat{y}_{k+1}/k \rangle}{\langle y_{k+1} - \hat{y}_{k+1}/k, y_{k+1} - \hat{y}_{k+1}/k \rangle} \end{cases}$$

$$M_{k+1} = \text{Span} \{y_1, y_2, \dots, y_k, y_{k+1} - \hat{y}_{k+1}/k\}$$

$$= M_{k+1}$$

$\hat{x}_{k+1} \in M_{k+1}$ must be the best estimator of x in M_{k+1}

$$\text{Since } \hat{y}_{k+1}/k \in M_k; \hat{y}_{k+1}/k \perp (y_{k+1} - \hat{y}_{k+1}/k)$$

$$\hat{x}_{k+1} = \hat{x}_k + \beta (y_{k+1} - \hat{y}_{k+1}/k)$$

$$\beta = \frac{\langle x, y_{k+1} \rangle - \langle x, \hat{y}_{k+1}/k \rangle}{\langle y_{k+1}, y_{k+1} \rangle - \langle y_{k+1}, \hat{y}_{k+1}/k \rangle}$$

$$\langle y_{k+1}, y_{k+1} \rangle - \langle y_{k+1}, \hat{y}_{k+1}/k \rangle$$

The co-variance matrix E is :

0.2778	-0.0278
-0.0278	0.1528

K_cap is :

0.2222
0.2778

The Minimum Variance Estimate for x is:

0.3417
0.4271

The co-variance matrix E is :

0.1937	-0.0813
-0.0813	0.1187

K_cap is :

-0.0375	0.1375
0.1125	0.0875

The Minimum Variance Estimate for x is:

0.4504
0.4963

The co-variance matrix E is :

0.0545	-0.0105
-0.0105	0.0828

K_{cap} is :

-0.0695	0.0599	0.1461
0.1287	0.1269	-0.0743

The Minimum Variance Estimate for x is:

-1.0134
1.2402

The co-variance matrix E is :

0.0437	0.0072
0.0072	0.0538

K_{cap} is :

-0.0441	0.0962	0.1297	-0.0429
0.0873	0.0674	-0.0474	0.0703

The Minimum Variance Estimate for x is:

-1.0296
1.2667

5(a) $\langle x, y \rangle = x^T y \rightarrow$ over determined system of eqⁿ

$$\hat{x} = (A^T A)^{-1} A^T b \Rightarrow A \hat{x} = b.$$

1118 for $y = Ax$ we get.

$$\hat{x} = (A^T A)^{-1} A^T y$$

$$\text{w.k.t } A = \begin{bmatrix} 1 & 2 \\ 3 & 4 \\ 5 & 0 \\ 0 & 6 \end{bmatrix} ; y = \begin{bmatrix} 1.5377 \\ 3.6948 \\ -7.7193 \\ 7.3621 \end{bmatrix}$$

$$A^T = \begin{bmatrix} 1 & 3 & 5 & 0 \\ 2 & 4 & 0 & 6 \end{bmatrix}$$

$$\left[\begin{bmatrix} 1 & 3 & 5 & 0 \\ 2 & 4 & 0 & 6 \end{bmatrix} \begin{bmatrix} 1 & 2 \\ 3 & 4 \\ 5 & 0 \\ 0 & 6 \end{bmatrix} \right]^{-1} \begin{bmatrix} 1 & 3 & 5 & 0 \\ 2 & 4 & 0 & 6 \end{bmatrix} \times \begin{bmatrix} 1.5377 \\ 3.6948 \\ -7.7193 \\ 7.3621 \end{bmatrix}$$

$$\hat{x} = \begin{bmatrix} -1.3169 \\ 1.4368 \end{bmatrix}$$

$$56) Q = E\{\varepsilon\varepsilon^T\} = I$$

$$A = \begin{bmatrix} 1 & 2 \\ 3 & 4 \\ 5 & 0 \\ 0 & 6 \end{bmatrix}$$

$$y = \begin{bmatrix} 1.5377 \\ 3.6948 \\ -7.7193 \\ 7.3621 \end{bmatrix}$$

$$Q = \begin{bmatrix} 1 & & 0 \\ & 1 & \\ 0 & & 1 \\ & & & 1 \end{bmatrix}$$

W.K.T

$$\hat{x} = (A^T Q^{-1} A)^{-1} A^T Q^{-1}$$

$$\hat{x} = \hat{K} y$$

$$\hat{x} = \begin{bmatrix} -1.3169 \\ 1.4368 \end{bmatrix}$$

$$57) E(\varepsilon\varepsilon^T) = Q = I_{4 \times 4} \text{ \& } P = E(x x^T) = 100 I_{2 \times 2}$$

$$A = \begin{bmatrix} 1 & 2 \\ 3 & 4 \\ 5 & 0 \\ 0 & 6 \end{bmatrix}$$

$$; y = \begin{bmatrix} 1.5377 \\ 3.6948 \\ -7.7193 \\ 7.3621 \end{bmatrix}$$

$$Q = \begin{bmatrix} 1 & & & \\ & 1 & & \\ & & 1 & \\ & & & 1 \end{bmatrix} = I_{4 \times 4}$$

$$P = 100 \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = 100 \times I_{2 \times 2}$$

$$\hat{K} = P C^T (C P C^T + Q)^{-1}$$

$$\hat{x} = \hat{K} y$$

$$\hat{x} = \begin{bmatrix} -1.3163 \\ 1.4365 \end{bmatrix}$$

con. Sol) Now $P = 10^6 \times I_{2 \times 2}$

$$\hat{x} = \begin{bmatrix} -1.3169 \\ 1.4368 \end{bmatrix}$$

Sol)	Least squares	BLUE	MVE $P = 10^6 I_{2 \times 2}$ $Q = I_{4 \times 4}$	MVE $P = 10^6 I_{2 \times 2}$ $Q = I_{4 \times 4}$
	$x = \begin{bmatrix} -1.3169 \\ 1.4368 \end{bmatrix}$	$x = \begin{bmatrix} -1.3169 \\ 1.4368 \end{bmatrix}$	$x = \begin{bmatrix} -1.3163 \\ 1.4365 \end{bmatrix}$	$x = \begin{bmatrix} -1.3169 \\ 1.4368 \end{bmatrix}$

Most of the results match.

i.e.; Value of \hat{x} from Least Squares = Value of \hat{x} from BLUE.

but at MVE; there is a small deviation but this deviation $(8) \rightarrow 0$ as $P \rightarrow \infty$ (i.e. $P = 10^6 I_{2 \times 2}$)

∴ even \hat{x} with MVE matches with \hat{x} from BLUE & least squares.

⇒ MVE reduces to BLUE when the co-variance of the unknown x becomes very large.

The co-variance matrix E is :

0.0437	0.0072
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0.0072	0.0538
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The Minimum Variance Estimate for x is:

-0.8836

1.0802