Test of significance for Debberence at Means:

Let \overline{x}_1 be the mean of a sample of size n_1 brown a population with mean u_1 , and variance ε_1^2 and \overline{x}_2 be the mean of an independent random sample at size n_2 brown another population with mean u_2 and variance ε_2^2 . Then since sample sizes are large, $\overline{x}_1 \sim N(u_1, \frac{\varepsilon_1^2}{n_1})$ and $\overline{x}_2 \sim N(u_2, \frac{\varepsilon_2^2}{n_2})$

Also $\overline{\chi}_1 - \overline{\chi}_2$, being the debterence at two independent normal variates is also a normal variate. Then $\overline{\chi}_1$ (standard normal variate) corresponding to $\overline{\chi}_1 - \overline{\chi}_2$ is given by

$$Z = (\overline{X}_1 - \overline{X}_2) - E(\overline{X}_1 - \overline{X}_2)$$

$$S. E(\overline{$$

B) The means of two single large samples ab1000, and 2000 members are 67.5 inches and
68.0 inches respectively. Can the samples be
regarded as drawn from the same population of
standard denation 2.5 penches? Toss 1 at 5%
level of significance.

Sul: Green, n= 1000, n2 = 2000 71 = 67.5 inches, 50 = 68.0 inches Null hypothesis Hoily= 1/2 and 8 = 2.5 inches, 1e the samples have been drawn brown the same population of standard denation 2. suchy. Alternative heppothesis ' H, : U, Fly (Two failed) Test statistics, Under Ho, the test statistics is Z= x,-X 52(t, + t) (2.5)2(1000 200) $\frac{2.5 \times \sqrt{\frac{1}{1000} + \frac{1}{2000}}}{2.5 \times 0.0387} = -5.1$ [Z] 73· The calculated Z is highly significant and we reject the null hypothesis and conclude that camples are containly not brom the same population centh standard deviation 2.5 mehrs.

Remark O In the above foroblem 6,2= 62 = 62 . i.e. it the Samples have been drawn troom the populations certh Common S.D 6 wen under Ho: li, = 4, の一方十九 3) If 6,2 + 62 and 6, and of are unknown, then they are estimated . Groson cample values. This result ca some error, which is practically immakinal et comple are large.

Thuse estimates by large samples are given by $\hat{G}_{1} = S_{1}^{2} \approx 8_{1}^{2} \quad \text{and} \quad \hat{G}_{2}^{2} = S_{2}^{2} = 8_{2}^{2} \quad (for samples are large)$ $Z = \overline{X_{1} - \overline{X_{2}}}$ $\sqrt{\frac{8_{1}^{2}}{n_{1}} + \frac{8_{2}^{2}}{n_{2}}}$

Shoppers are chosen at random in super market A located ch a certain section at the city. Their overage weekly tood expendeture is Rs 250 certs a standard denoition ab hs 40 for 400 women sheppers chesen at random in sufer mosket 'B' in another section at the uty; the average weekly bood expenditure is RS220 with a Standard denation at Rs 55. Test at 1% level out signebicance ceelether the querage weekly bood expenditure ob two population shoppers are equal.

Sil n = 400, x1 = 250, 8, = 40 n2 = 400, X2 = 220 182 = 55 Null hypothesis Ho: 11=12 ce, the grerage weekly bood expendetures at the two populations of shoppers are equal. Alternative by pothers H1: 4, 742 (Two texted) 1 6,2 + 6,2 Since of 2 and of are not known, we can take 62= 5,2 and 62= 82

Z = 71-72 250-220 1- (40)2 + (ST)2 $= 30 \times 20$ = 8.82 Calculated 1217 2.58 So average & expendeture at two populations at shoppers in market A and market B debber signeticantly.

(3) The average hously wages at a sample at 150 workers on plant A' was Rs 2.56 with a standard deviation at Rs 1.08. The average hously wage at a sample at 200 workers in plant B' was Rs 2.87 centre a standard deviation at Rs 1.28 can an applicant sately assume that hously wages paid by plant B' are higher than those paid by plant A'?

SI: n= 150, x, = 2.56, 8, = 1.08 12=200, 72=2.87, B2=1.28 Neel hypothesis: Ho: 4, -42, there is no signebigan debberence between the mean wages at workers in plant A and plant B. Alternative hypothesis H,; le, < 42 $Z = \chi_1 - \chi_2$ 2.56 - 287 Here calculated Z < -1.645 (critical value ab So nell hypothesis rejected Zon lebt facted ats y level ob seg) ale conclude that average housely wage at plant B is certainly higher Dit-test bir single mean Suppose we want to kest: O If a random sample xi (i=1,2, - n) of size n has been drawn brown a normal population with specified mean, say 110, or (1) If the sample mean debbers significantly brown the hypothetical value les els the population mean. Under the nell hypothesis Ho? (1) The sample how been drawn troom the population usth mean lo or (11) There is no significant dibberence between the sample meen & and the population mean lo the statistic where $\overline{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$ and $s^2 = \frac{1}{(n-1)} \sum_{i=1}^{n} (x_i - \overline{x})^2$ bollow. student's t-distribution conth (n-1) degree of breedom.

Assumptions for Student's t-lest

The tollowing assumptions are made in the

Student's t-test

D The parent population brom which the sample is

drawn is normal:

D The sample observations are independent, i.e. the

sample is random

The population standard denation of is anknown.

A machinist is making engine ports with axle diameters of 0.700 inch. A random sample ab 10 parts shows a mean diameter at 0.742 inch certh a standard deviation of 0.040 inch. Compute the stalistic you would use to test whether the work is meeting the specification. Also state how would you proceed.

Solution. Here we are given:

 $\mu = 0.700$ inche, $\bar{x} = 0.742$ inche, s = 0.040 inche and n = 10 Null Hypothesis, $H_0: \mu = 0.700$, i.e., the product is conforming to specifications. Alternative Hypothesis, $H_1: \mu \neq 0.700$

Test Statistic. Under
$$H_0$$
, the test statistic is : $t = \frac{\bar{x} - \mu}{\sqrt{S^2/n}} = \frac{\bar{x} - \mu}{\sqrt{s^2/(n-1)}} \sim t_{(n-1)}$

$$t = \frac{\sqrt{9} \ (0.742 - 0.700)}{0.040} = 3.15$$

How to proceed further. Here the test statistic 't' follows Student's t-distribution with 10 - 1 = 9 d.f. We will now compare this calculated value with the tabulated value of t for 9 d.f. and at certain level of significance, say 5%. Let this tabulated value be denoted by t_0 .

- (i) If calculated 't', viz., $3.15 > t_0$, we say that the value of t is significant. This implies that \bar{x} differs significantly from μ and H_0 is rejected at this level of significance and we conclude that the product is not meeting the specifications.
- (ii) If calculated $t < t_0$, we say that the value of t is not significant, i.e., there is no significant difference between \overline{x} and μ . In other words, the deviation (\overline{x} - μ) is just due to fluctuations of sampling and null hypothesis H_0 may be retained at 5% level of significance, i.e., we may take the product conforming to specifications.

Example 16.6. The mean weekly sales of soap bars in departmental stores was 146.3 bars per store. After an advertising campaign the mean weekly sales in 22 stores for a typical week increased to 153.7 and showed a standard deviation of 17.2. Was the advertising campaign successful?

Solution. We are given : n = 22, $\bar{x} = 153.7$, s = 17.2.

Null Hypothesis. The advertising campaign is not successful, i.e., $H_0: \mu = 146.3$

Alternative Hypothesis, $H_1: \mu > 146.3$ (Right-tail).

Test Statistic. Under H_0 , the test statistic is: $t = \frac{\overline{x} - \mu}{\sqrt{s^2/(n-1)}} \sim t_{22-1} = t_{21}$

$$t = \frac{153.7 - 146.3}{\sqrt{(17.2)^2/21}} = \frac{7.4 \times \sqrt{21}}{17.2} = 9.03$$

Conclusion. Tabulated value of t for 21 d.f. at 5% level of significance for single-tailed test is 1-72. Since calculated value is much greater than the tabulated value, it is

highly significant. Hence we reject the null hypothesis and conclude that the advertising campaign was definitely successful in promoting sales.

Example 16-7. A random sample of 10 boys had the following L.Q.'s: 70, 120, 110, 101, 88, 83, 95, 98, 107, 100. Do these data support the assumption of a population mean LQ. 100? Find a reasonable range in which most of the mean LQ. values of samples of 10 boys lie.

Solution. Null hypothesis, H_0 : The data are consistent with the assumption of a mean I.Q. of 100 in the population, i.e., $\mu = 100$.

Alternative hypothesis, H_1 : $\mu \neq 100$.

Test Statistic. Under H_0 , the test statistic is: $t = \frac{(\bar{x} - \mu)}{\sqrt{S^2/n}} \sim t_{(n-1)}$,

where \bar{x} and S^2 are to be computed from the sample values of I.Q.'s.

TABLE 16-1: CALCULATIONS FOR SAMPLE MEAN AND S.D.

$(x-\overline{x})^2$	$(x-\overline{x})$	x
739-84	- 27-2	70
519-84	22.8	120
163-84	12.8	110
14-44	3-8	101
84-64	-9-2	88
201-64	- 14-2	83
4-84	-2.2	95
0-64	0.8	98
96-04	9.8	107
7-84	2.8	100
1833-60		Total 972

Here
$$n = 10$$
, $\bar{x} = \frac{972}{10} = 97.2$ and $S^2 = \frac{1833.60}{9} = 203.73$

Tabulated $t_{0.05}$ for (10-1), i.e., 9 d.f. for two-tailed test is 2.262.

Conclusion. Since calculated t is less than tabulated $t_{0.05}$ for 9 d.f., H_0 may be ented at 5% level of since it. with the assumption of a significance and we may conclude that the data are consistent with the assumption of mean I.Q. of 100 in the population.

The 95% confidence limits within which the mean I.Q. values of samples of 10 s will lie are given by boys will lie are given by:

$$\bar{x} \pm t_{0.05} S / \sqrt{n} = 97.2 \pm 2.262 \times 4.514 = 97.2 \pm 10.21 = 107.41$$
 and 86.99
Hence the required 95% confidence:

Hence the required 95% confidence interval is [86-99, 107-41].

Example 16-8. The heights of 10 males of a given locality are found to be 70, 67, 62, 68, 61, 68, 70, 64, 64, 66 inches. Is it reasonable to believe that the average height is greater than 64 inches? Test at 5% significance level assuming that for 9 degrees of freedom P(t > 1.83) = 0.05.

Solution. Null Hypothesis, $H_0: \mu = 64$ inches Alternative Hypothesis, $H_1: \mu > 64$ inches

TABLE 16-2: CALCULATIONS FOR SAMPLE MEAN AND S.D.

x	70	67	62	68	61	68	70	64	64	66	Total 660
$x-\overline{x}$	4	1	-4	2	-5	2	4	-2	-2	0	0
$(x-\overline{x})^2$	16	1	16	4	25	4	16	4	4	0	90

$$\overline{x} = \frac{\sum x}{n} = \frac{660}{10} = 66;$$
 $S^2 = \frac{1}{n-1} \sum (x - \overline{x})^2 = \frac{90}{9} = 10$

Test Statistic. Under H_0 , the test statistic is:

$$t = \frac{\overline{x} - \mu}{\sqrt{S^2/n}} = \frac{66 - 64}{\sqrt{10/10}} = 2,$$

which follows Student's *t*-distribution with 10 - 1 = 9 d.f.

Tabulated value of t for 9 d.f. at 5% level of significance for single (right) tail-test is 1-833. (This is the value $t_{0.10}$ for 9 d.f. in the two-tailed tables given at the end of the chapter.)

Conclusion. Since calculated value of t is greater than the tabulated value, it is significant. Hence H_0 is rejected at 5% level of significance and we conclude that the average height is greater than 60 inches.

16.2.7. Critical Values of t. The critical (or significant) values of t at level of significance α and d.f. v for two-tailed test are given by the equation:

$$P[\mid t \mid > t_v(\alpha)] = \alpha$$

$$P[\mid t \mid < t_v(\alpha)] = 1 - \alpha \qquad ...(16.5)$$
...(16.5a)

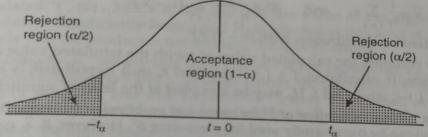


Fig. 16.2: Critical values of t-distribution

The values $t_v(\alpha)$ have been tabulated in Fisher and Yates' Tables, for different values of α and ν and are given in Table I at the end of the chapter.

Since t-distribution is symmetric about t = 0, we get from (16.5)

$$P(t > t_v(\alpha)) + P[t < -t_v(\alpha)] = \alpha \implies 2P[t > t_v(\alpha)] = \alpha$$

$$\Rightarrow P[t > t_v(\alpha)] = \alpha/2 \qquad \therefore P[t > t_v(2\alpha)] = \alpha \qquad \dots (16.5b)$$

 t_v (2 α) (from the Tables at the end of the chapter) gives the significant value of t for a single-tail test [Right-tail or Left-tail-since the distribution is symmetrical], at level of significance α and v d.f.

Hence the significant values of t at level of significance 'a' for a single-tailed test can be obtained from those of two-tailed test by looking the values at level of significance 2a.

For example,

 t_8 (0.05) for single-tail test = t_8 (0.10) for two-tail test = 1.86 $t_{15}(0.01)$ for single-tail test = $t_{15}(0.02)$ for two-tail test = 2.60.

TABLE I.

SIGNIFICANT VALUES $t_{\nu}\left(\alpha\right)$ of t-Distribution (TWO-TAIL AREAS)

			$P[\mid t\mid > t_{v}(\alpha)]$] = α	de of	-	/	1	
d.f.		Pr	obability (Level					01	
(v)	0.50	0.10	0.05	0.02	0.01	0.1101		161	
1	1.00	6.31	12.71	31-82	63.66	636-62	U	1 18	
2	0.82	2.92	4.30	6.97	6.93	31.60	1	2 10	
3	0.77	2.35	3-18	4.54	5.84	12.94	1	3 4	
4	0.74	2.13	2.78	3.75 3.37	4·60 4·03	8-61	1	5	
5	0.73	2.02	2.57		Total Solder	6.86		6	
6	0.72	1.94	2.45	3.14	3.71	5.96	-	7 8	
7	0.71	1.90	2.37	3.00	3.50	5-41		9	
8	0.71	1.86	2.31	2.90	3.36	5.04	and the same of th		
9	0.70	1.83	2.26	2.82	3.25	4.78		11	
10	0.70	1.81	2.23	2.76	3-17	4.59	8 1	12 13	
11	0.70	1.80	2.20	2.72	3.11	4.44		14	
12	0.70	1.78	2.18	2.68	3.06	4.32	1	15	
13	0.69	1.77	2.16	2.65	3.01	4.22		16	
14	0.69	1.76	2.15	2.62	2.98	4.14		17	
15	0.69	1.75	2.13	2.60	2.95	4.07	7	18	
	personal reliable	1.75	2-12	2.58	2.92	4.0	12	20	
16	0.69	1.74	2.11	2.57	2.90	3.9	97	2	
17	0.69		2.10	2.55	2.88	3.	92	27	
18	0.69	1.73	2.09	2.54	2.86	3	3.88		
19	0.69	1.73	The second second	2.53	2.85	3	3.85		
20	0.69	1.73	2.09	of deposits to	d tree more and to	15	3.83		
01	0.69	1.72	2.08	2.52	2.83				
21	0.69	1.72	2.07	2.51	2.83				
22		1.71	2.07	2.50	2.8	1	3.77		
23	0.69	1000000	2.06	5.49	2.8	0	3.75		
24	0.69	1.71	The same of the same of	2.49	2 70		3.73		
25	0.68	1.71	2.06	S. J. C. P. S.	28	70	3.71	100	
06	0.68	1.71	2.06	2.48		260			
26		1.70	2.05	2.47	2.	2.77		3.67	
27	0.68		2.05	2.47	7 2	2.76		3.66	
28	0.68	1.70	The state of the s	1000000	1 2	2.76			
29	0.68	1.70	2.05		2.40		3.65		
30	0.68	1.70	2.04	2.4			2.79		
00	0.67	1.65	1.96	2.3	2.33 2.58			46	