# 106349 - Advanced probability

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#### Abstract

### 1 Introduction. Summary of course through an example. Branching process

We have an individual that gives a birth to a random number of offsprings – random variable X. X define a distribution, i.e.,  $P: \mathbb{Z}^+ \to [0,1]$ , i.e.,  $P(X=k) \in [0,1]$ , and  $\sum_{k=0}^{\infty} P(X=k) = 1$ .

**Definition 1.1.**  $f_X(\theta) = \sum_{k=0}^{\infty} \theta^k P(X=k)$  – moment-generating function.

The series is absolutely convergent for  $\theta \in [-1, 1]$  since k sums to 1. For  $\theta \in (-1, 1)$ ,  $f_x$  is analytic, thus we can differentiate it term-by-term:

$$f_X'(\theta) = \sum_{k>1} \theta^{k-1} P(X=k)$$

Since,  $f_X$  is analytic, knowing it means knowing P(X = k) and vice versa. Note that  $f_X(0) = P(X = 0)$  and  $f_X(1) = 1$ . Also

$$f_X'(1) = \sum_{k>0}^{\infty} kP(X=k) = \mathbb{E}X = \mu$$

$$\lim_{\theta \to 1} \frac{f_X(1) - f_X(\theta)}{1 - \theta} = \lim_{\theta \to 1} \frac{1 - f_X(\theta)}{1 - \theta}$$

Note also that  $f_X$  is convex, since second derivative is positive.

Size of  $n^{th}$  generation Let  $\left(X_r^{(n)}\right)_{n,r=1^{\infty}}$ , where n is generation and r is offspring number (index) in  $n^{th}$  generation.

Assume  $X_r^{(n)}$  are i.i.d. (independent, identically distributed) random variables. Identically distributed means

$$P(X_n^r = k) = P(X = k)$$

Independence means

$$P(\forall i < J X_{r_i}^{n_i} = k) = \prod_{i=1}^{J} P(X_{r_i}^{n_i} = k)$$

Define  $z_1 = X_1^1$ .  $z_2 = \sum_{r=1}^{z_1} X_r^2$  an so on:

$$z_{n+1} = \sum_{r=1}^{z_n} X_r^n$$

We want to study asymptotics of  $z_n$ . Given U and V taking values in  $\mathbb{Z}^+$ ,

$$\mathbb{E}[U|V=k] = \sum_{i=0}^{\infty} jP(U=j|V=k)$$

, where

$$P(U = j | V = k) = \frac{P(U = j, V = k)}{P(V = k)}$$

If U, V are independent, P(U=j|V=k)=P(U=j) and thus  $\mathbb{E}[U|V=k]=\mathbb{E}U$ .

**Definition 1.2.** Define random variable  $\mathbb{E}[U|V]$  such that

$$\mathbb{E}[U|V] = \mathbb{E}[U|V = k]$$

if V = k.

**Definition 1.3** (Tower property).

$$\mathbb{E}\big[\mathbb{E}[U|V]\big] = \mathbb{E}U$$

Define

$$f_n = \sum_{k=0}^{\infty} \sum_{k=0}^{\infty} \theta^k P(z_n = k) = \mathbb{E}\theta^{z_n}$$

Theorem 1.1.

$$f_{n+1}(\theta) = f_n(f_X(\theta))$$

or

$$f_n(\theta) = \underbrace{f \circ f \circ \cdots \circ f}_{n \text{ times}}(\theta))$$

*Proof.* Use tower property with  $U^{z_{n+1}}$  and  $V = \theta^{z_n}$ . By tower property

$$\mathbb{E}[\theta^{z_{n+1}}] = \mathbb{E}\big[\mathbb{E}[\theta^{z_{n+1}}|\theta^{z_n}]\big]$$

$$\mathbb{E}\big[\mathbb{E}[\theta^{z_{n+1}}|\theta^{z_n}]\big] = \sum_{k=0}^{\infty} P(z_n = k) \mathbb{E}[\theta^{z_{n+1}}|\theta^{z_n} = k]$$

What is  $\mathbb{E}[\theta^{z_{n+1}}|\theta^{z_n}=k]$ ?

$$\mathbb{E}[\theta^{z_{n+1}}|\theta^{z_n}=k] = \mathbb{E}\left[\theta^{\sum_{j=1}^k X_j^{n+1}}|\theta^{z_n}=k\right] \stackrel{\text{independence}}{=} \mathbb{E}\left[\theta^{\sum_{j=1}^z X_j^{n+1}}\right] \stackrel{\text{independence}}{=} \prod_{j=1}^k \mathbb{E}\left[\theta^{X_j^{n+1}}\right] \stackrel{\text{i.d.}}{=} (f_X(\theta))^k$$

Thus

$$\mathbb{E}\big[\mathbb{E}[\theta^{z_{n+1}}|\theta^{z_n}]\big] = \sum_{k=0}^{\infty} P(z_n = k)(f_X(\theta))^k = f_n(f(\theta))$$

Also we can say

$$\mathbb{E}[\theta^{z_{n+1}}|z_n] = (f_X(\theta))^{z_n}$$

Study of  $z_n$  What is  $\pi_n = P(z_n = 0) = f_n(0) = f(\pi_{n-1})$ , probability that population is extinguished. Since  $z_{n-1} = 0 \Rightarrow z_n = 0$ , i.e.  $\pi_n$  is non-decreasing.

Let  $P(z_n = 0 \text{ for some n}) = \pi$ .

We hope that  $\{z_n = 0\}$  such that

$$\bigcup_{n} \{z_n = 0\} = \{z_n = 0 \text{ for some n}\}\$$

i.e.,  $\pi = \lim_{n \to \infty} \pi_n$ . We call  $\pi$  the extinction probability.

**Theorem 1.2.** If  $\mu = \mathbb{E} > 1$  then  $\pi$  is a unique root of  $\pi = f(\pi)$  and  $\pi \in [0,1)$ . If  $\mu \leq 1, \pi = 1$ .

If we look at  $f(\pi)$  and  $\pi$ , they intersect in 1, and they can intersect in two points since f(x) is convex. There is second intersection iff  $f'(1) = \mu > 1$ .

Construction of  $X_n^r$  Construct set  $\Omega$ ,  $f_{n,r}: \Omega \to \mathbb{Z}^+$  and  $\mathcal{F}$  a collection of subsets of  $\Omega$  with  $P: \mathcal{F} \to [0,1]$ . Let  $\Omega = \mathbb{Z}^+ \times \mathbb{Z}^+$ ,  $\mathcal{F} = \{0,1\}^{\Omega}$ .

The problem is when we have infinitely number of variables.

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**Example** Example of not well-behaved triple  $(\Omega, \mathcal{F}, P)$ .  $\Omega = \mathbb{N}$ . Now  $\mathcal{F} = \{C \subset \mathbb{N} : C \text{ has density}\}$ .

C has density means

$$\frac{|C \cap \mathbb{N}|}{n} \stackrel{n \to \infty}{\to} \rho(C)$$

However, for  $C(m) = \{1, 2, \dots, m\}, \forall m \quad \rho(c_m), \text{ and }$ 

$$\rho\Big(\bigcup C_m\Big) = 1$$

Thus  $(\mathbb{N}, \mathcal{F}, \rho)$  is not a good probability space, since it doesn't fulfills this  $\pi_n \to \pi$  property. Note we can define other probabilities on naturals, for example

$$P(\{i\}) = 2^{-i}$$

Asymptotics of  $z_i$  Assuming  $\pi \in (0,1)$ , what is behavior of  $z_n$ ?

**Definition 1.4.**  $z_n$  is a Markov chain if

$$P(z_{n+1} = j | z_i = k_i \quad \forall i \le n) = P(z_{n+1} = j | z_n = k_n)$$

We can use to compute expectation:

$$\mathbb{E}[z_{n+1}|z_i = k_i \quad \forall i < n] = E[z_{n+1}|z_n = k_n]$$

Then, since  $E\left[\sum_{i=1}^{J} X_i^n\right] = J\mu$ 

$$E[z_{n+1}|z_n] = \mu z_n$$

Let  $M_n = \frac{z_n}{u^n}$  then  $\mathbb{E}[M_n] = 1$ . Also

$$\mathbb{E}[M_{n+1}|z_0,\ldots,z_n]=M_n$$

This is a definition of martingale with respect to  $z_0, \ldots, z_n$ .

Let  $(\Omega, \mathcal{F}, P)$  we say S happens almost surely (a.s.) if

$$P(\{w \in \Omega : S \text{ is true for w}\}) = 1$$

**Theorem 1.3** (Martingale convergence theorem). If  $M_n$  is a positive martingale then  $\lim_{n\to\infty} M_N = M_\infty$  exists a.s. and

•  $\mu \leq 1$ .  $M_{\infty} = 0$  a.s. That means  $\mathbb{E}M_{\infty} = 0$  but  $\mathbb{E}M_{=}1$ , i.e.,

$$\mathbb{E}\Big[\liminf_{n\to\infty} M_n\Big] < [\liminf_{n\to\infty} \mathbb{E}[M_n]$$

•  $\mu > 1$ . If  $M_{\infty} > 0$  with positive probability then  $z_n \sim \mu^n M_{\infty}$ .

Lemma 1.1 (Fatou's lemma).

$$\mathbb{E} \Big[ \liminf_{n \to \infty} M_n \Big] \leq \liminf_{n \to \infty} \mathbb{E} [M_n]$$

Theorem 1.4.

$$\mathbb{E}[M_{\infty}] = 1 \iff \mu > 1 \quad \text{ and } \mathbb{E}[X \log(X)] < \infty$$

# 2 Overview of measure theory

Notation

- $\bullet$  S is a set.
- $\mathcal{A}$  is algebra of subsets of S
  - 1.  $S \in \mathcal{A}$
  - 2.

$$E \in \mathcal{A} \Rightarrow E^C \in \mathcal{A}$$

, where  $E^C=S\setminus E$ 

3.

$$E_1, E_2 \in \mathcal{A} \Rightarrow E_1 \cup E_2 \in \mathcal{A}$$

meaning

$$E_1, E_2 \in \mathcal{A} \Rightarrow E_1 \cap E_2 \in \mathcal{A}$$

- $\mathcal{F}$  is a  $\sigma$ -algebra if lest item works for countable union.
- $E\Delta F = E \setminus F \cup F \setminus E$

**Definition 2.1.** A measurable space is a pair  $\{S, \mathcal{F}\}$ .

**Proposition 2.1.** If we have  $(\mathcal{F}_i)_{i\in I}$ , then  $\bigcap_{i\in I} \mathcal{F}$  is also a  $\sigma$ -algebra.

**Definition 2.2.** Let C be a collection of subsets of S.  $\sigma(C)$  is a smallest  $\sigma$ -algebra containing C ( $\sigma$ -algebra generated by C). It is easy to construct one

$$I = \{ \mathcal{F} : \mathcal{F} \supset C \}$$

and then

$$\sigma(C) = \bigcap_{\mathcal{F} \in I} \mathcal{F}$$

**Definition 2.3.** Let  $\{S, \mathcal{F}\}$  be a topological space.  $\mathcal{B}(X)$  (Borel  $\sigma$ -algebra) is defined as  $\sigma$ -algebra generated by open sets. We denote  $\mathcal{B} = \mathcal{B}(\mathbb{R})$ .

Exercise

$$\pi(\mathbb{R}) = \{(-\infty, x], x \in \mathbb{R}\}\$$

Show that  $\sigma(\pi(\mathbb{R})) = B$ 

**Definition 2.4.** Additive set function on a collection of sets  $\mathcal{F}$  is

$$\mu: \mathcal{F} \to [0, \infty)$$

$$\forall E, F \in \mathcal{F} \ E \cap F = \emptyset \quad \mu(E \cup F) = \mu(E) + \mu(F)$$

We say  $\mu$  is  $\sigma$ -additive if same holds of countable infinite sets

$$\forall \{E_i\}_{i=1}^{\infty} E_i \cap E_j = \emptyset \quad \mu(E \cup F) = \sum_{i=1}^{\infty} \mu(E_i)$$

**Definition 2.5.** A triple  $(S, \mathcal{F}, \mu)$  is a measure space if  $\mathcal{F}$  is a  $\sigma$ -algebra on S and  $\mu$  is  $\sigma$ -additive on  $\mathcal{F}$ .

**Definition 2.6.**  $(S, \mathcal{F}, \mu)$  is finite if  $\mu(S) < \infty$  $(S, \mathcal{F}, \mu)$  is  $\sigma$ -finite if

$$\exists \{E_i, \, \mu(E_i) < \infty\}_{i=1}^{\infty} \quad S = \bigcup_{i=1}^{\infty} E_i$$

**Definition 2.7.** If  $\mu(S) = 1$ ,  $(S, \mathcal{F}, \mu)$  is probability space.

**Definition 2.8.** E is null if  $\mu(E) = 0$ .

**Definition 2.9.**  $\phi$  is said to be true almost everywhere with respect of  $\mu$  if

$$\mu(\lbrace X : \phi(X) = \text{False} \rbrace) = 0$$

#### Results from measure theory

**Definition 2.10.** A collection of sets  $\mathcal{D}$  is called a  $\pi$ -system if  $E, F \in \mathcal{D} \Rightarrow E \cap F \in \mathcal{D}$ 

**Theorem 2.2** (Uniquess). Let  $\mathcal{D}$  be a  $\pi$ -system generating a  $\sigma$ -algebra  $\mathcal{F}$ . Let  $\mu_1$  and  $\mu_2$  be two finite measures on  $\mathcal{F}$  which agree on  $\mathcal{D}$ . Then  $\mu_1 = \mu_2$ .

Collary 2.1.  $(S, \mathcal{F}, P_1), (S, \mathcal{F}, P_2)$  probability spaces, P1 = P2 on  $\pi$ -system  $\mathcal{D}$ , then  $P_1 = P_2$ .

**Theorem 2.3** (Carathéodory's extension theorem). Let  $\mathcal{A}$  be an algebra of sets.  $\mu_0: \mathcal{A} \to \mathbb{R}^+$   $\sigma$ -additive set function on  $\mathcal{A}$ . Then exists unique extension  $\bar{\mu}: \sigma(\mathcal{A}) \to \mathbb{R}^+$  such that  $m\bar{u} = \mu_0$ .

**Homework** Lebesgue on  $\mathbb{R}$ .  $\mathcal{A} = \{\text{open set}\}\$ . If we have

$$O = \bigcup_{i=1}^{\infty} (a_i, b_i)$$

then

$$\mu_0(O) = \sum_{i=1}^{\infty} b_i - a_i$$

Check that  $\mu_0$  is well defined and  $\sigma$ -additive.

**Lemma 2.1.**  $(S, \mathcal{F}, \mu)$  measure space.  $A, B \in \mathcal{F}$ , then

$$\mu(A \cup B) \le \mu(A) + \mu(B)$$

$$\mu\left(\bigcup_{i=1}^{\infty} F_i\right) \le \sum_{i=1}^{\infty} \mu(F_i)$$

If  $\mu(S) < \infty$ 

$$\mu(A \cup B) = \mu(A) + \mu(B) - \mu(A \cap B)$$

From that we get inclusion-exclusion:

$$\mu\left(\bigcup_{i=1}^{n} A_{i}\right) = \sum_{i=1}^{n} \mu(A)_{i} - \sum_{i \neq j} \mu(A_{i} \cap A_{j}) + \dots + (-1)^{n-1} \mu\left(\bigcap_{i=1}^{n} A_{i}\right)$$

Exercise Proof the lemma

**Lemma 2.2.** If  $F_n \subseteq F_{n+1}$  then

$$\mu\left(\bigcup_{i=1}^{\infty} F_i\right) = \lim_{n \to \infty} \mu(F_n)$$

If  $\mu(S) < \infty$  and  $F_n \supseteq F_{n+1}$  then

$$\mu\left(\bigcap_{i=1}^{\infty} F_i\right) = \lim_{n \to \infty} \mu(F_n)$$

*Proof.* Assume  $\mu(S) < \infty$ . Define  $F_{\infty} = \bigcup_{i=1}^{\infty} F_i$ . Let  $G_n = F_n \setminus F_{n+1}$ . Then

$$F_{\infty} = \bigcup_{i=1}^{\infty} G_i$$

Meaning

$$\mu(F_{\infty}) = \sum_{i=1}^{\infty} G_i$$

$$\mu(F_n) = \sum_{k=1}^n G_k$$

Since measure is finite, the tail of series tends to 0, thus

$$\mu(F_{\infty}) - \mu(F_n) = \sum_{k=n}^{\infty} G_k \to 0$$

Then we can take complements and get the second statement.

Exercise Proof unconditionally

## 3 Recasting measure theory as probability

**Definition 3.1.** A probability space is a  $(\Omega, \mathcal{F}, P)$  is a measure space such that  $P(\Omega) = 1$ . We call  $\omega \in \Omega$  an outcome.  $E \in \mathcal{F}$  is an event. P(E) is probability of the event.

**Example** Tossing finite or infinite sequence of coins.

#### Tossing 4 coins

$$\Omega = \{HHHH, HHHT, HHTH, \dots, TTTH, TTTT\}$$

$$\mathcal{F} = 2^{\Omega}$$

$$P(\omega \in \Omega) = \frac{1}{|\Omega|}$$

### Tossing infinite number coins

$$\Omega = \{0, 1\}^{\mathbb{N}}$$

 $\Omega$  has a natural topology which is called a product topology. It is coarsest topology such that  $\pi_i: \Omega \to \{0,1\}$   $\pi_i(\omega) = \omega_i$  is continuous.

Let  $\mathcal{F} = \mathcal{B}(\Omega)$ .

Smallest  $\sigma$ -algebra such that

$$\pi_i^{-1}(0) \subset \Omega \in \mathcal{F}$$
 $\pi_i^{-1}(1) \subset \Omega \in \mathcal{F}$ 

$$\pi_i(\Omega, \mathcal{F}) \to \left( \{0, 1\}, \{0, 1\}^{\{0, 1\}} \right)$$

Natural  $\pi$ -system  $\mathcal{F}_n$  smallest  $\sigma$ -algebra making  $\pi_1, \ldots, \pi_n$  measurable. Note that

#### Proposition 3.1.

$$\bigcup_n \mathcal{F}_n 
eq \mathcal{F}$$

*Proof.* Define  $S_n(\omega) = \sum_{i=1}^n \omega_n$ .

$$X_n = \frac{S_n(\omega)}{n}$$

Define

$$Y(\omega) = \limsup X_n(\omega)$$

$$E = \left\{ \omega : Y(\omega) \ge \frac{1}{3} \right\}$$

$$E \in \mathcal{F} \setminus \bigcup_n \mathcal{F}_n$$

What  $\mathcal{F}_n$  looks like? For example,  $\mathcal{F}_2$  has 4 outcomes, deciding only first two tosses.

Note If we take  $(\Omega_4, \mathcal{F}^{(4)}, P_4)$ , restricting to  $(\Omega_3, \mathcal{F}^{(3)}, P_3)$ 

$$P_4(\{(0,0,0,\omega_4)\}) = P_3(\{(0,0,0)\})$$

Thus we want  $P_{fair}$  defined on  $\Omega$  to fulfill same property:

$$P_{fair}(E) = P_n(\tilde{E})$$

where  $E \in \mathcal{F}_n$  and  $\tilde{E} \in F^{(n)}$ .

**Definition 3.2.**  $E \subset \mathcal{F}$  occurs almost surely (a.s.) if P(E) = 1.

**Definition 3.3** ( $\limsup$  and  $\liminf$ ). Let  $\{E_n\}$  be a sequence of events.

$$\limsup E_n = \bigcap_{m} \bigcup_{n \geq m} E_n = \{E_n \text{ occurs infenetely often (i.o.)}\} = \{\omega \in \Omega: \forall m \ \exists n(\omega) > m \quad \omega \in E_n(\omega)\}$$

Alternatively,  $(\Omega, \mathcal{F})$  and  $\{E_n\}$  there is a natural map

$$I:\Omega \to \left\{0,1\right\}^N$$

$$\omega \mapsto \{1_{E_n}(\omega)\}$$

where

$$1_E(\omega) = \begin{cases} 0 & \omega \notin E \\ 1 & \omega \in E \end{cases}$$

Now

$$\liminf E_n = \bigcup_m \bigcap_{n \geq m} E_n = \{E_n \text{ occurs eventually}\} = \{\omega \in \Omega: \exists m(\omega) \ \forall n \geq m(\omega) \quad \omega \in E_n(\omega)\}$$

**Remark** Since everything is countable, if  $E_n \in \mathcal{F}$ , then  $\limsup E_n$ ,  $\liminf E_n \in \mathcal{F}$ 

We can write

$$\left\{\frac{S_n}{n} \to \frac{1}{2}\right\} = \left\{\limsup \frac{S_n}{n} \le \frac{1}{2}\right\} \cap \left\{\liminf \frac{S_n}{n} \ge \frac{1}{2}\right\}$$

Choose  $q \in \mathbb{Q}^+$  and take a look at

$$\left\{ \liminf \frac{S_n}{n} > q \right\} = \liminf E_n(q)$$

where  $E_n = \left\{ \omega : \frac{S_n}{n} > q \right\}$ . In addition

$$\left\{\limsup \frac{S_n}{n} < q\right\} = \liminf F_n(q)$$

where  $F_n = \{\omega : \frac{S_n}{n} < q\}$ . Therefore  $\{\liminf \frac{S_n}{n} > q\} \in \mathcal{F}$ . Finally,

$$\left\{\liminf \frac{S_n}{n} \ge \alpha\right\} = \bigcap_{q \le \alpha} \left\{\liminf \frac{S_n}{n} > q\right\}$$

Lemma 3.1 (Fatou's lemma).

$$P\left[\liminf_{n\to\infty} E_n\right] \le \liminf_{n\to\infty} p(E_n)$$

Proof.

$$\liminf_{n \to \infty} E_n = \bigcup_m \bigcap_{n > m} E_n$$

Sets  $F_m = \bigcap_{n > m} E_n$  are increasing and  $F_n \subseteq E_n$ , thus

$$P\left[\liminf_{n\to\infty} E_n\right] = \lim_{n\to\infty} P(F_n) \le \liminf_{n\to\infty} P(E_n)$$

Lemma 3.2 (Fatou's lemma).

$$P\left[\limsup_{n\to\infty} E_n\right] \ge \limsup_{n\to\infty} p(E_n)$$

*Proof.* Note that  $(\limsup E_n)^C = \liminf E_n^C$ , thus this is straightforward form previous lemma.

**Lemma 3.3** (First Borel-Cantelli lemma). Let  $\{E_n\} \subseteq \mathcal{F}$  be a sequence of events s.t.  $\sum_n P(E_n) < \infty$ , then

$$P(E_n \text{ happens i.o.}) = 0$$

Proof.

$$P(E_n \text{ i.o.}) = P\left(\bigcap_{m} \bigcup_{n \ge m} E_n\right) \le P\left(\bigcup_{n \ge m} E_n\right) \le \sum_{n = m}^{\infty} P(E_n) \stackrel{m \to \infty}{\to} 0$$

Since  $P(E_n \text{ i.o.})$  is independent on m, it got to be 0.

Example Fix  $\epsilon > 0$ . Look at  $P\left(\left|\frac{S_n(\omega)}{n} - \frac{1}{2}\right| > \epsilon\right)$ .

Claim

$$P\left(\left|\frac{S_n(\omega)}{n} - \frac{1}{2}\right| > \epsilon\right) \le \frac{12}{\epsilon^4} \frac{1}{n^2}$$

By 3.3 
$$P\left(\left|\frac{S_n(\omega)}{n} - \frac{1}{2}\right| > \epsilon \text{ i.o.}\right) = 0 \text{ thus}$$

$$\left\{\frac{S_n}{n} \to \frac{1}{2}\right\} = \bigcap_{\epsilon > 0} \left\{ \left| \frac{S_n(\omega)}{n} - \frac{1}{2} \right| < \epsilon \text{ eventually} \right\} = 1$$