

# 2024-ST-'53-65'

EE24BTECH11023

- 1) Let  $\{X_n\}_{n \geq 1}$  be a time homogeneous discrete-time Markov chain with state space  $\{0, 1, 2\}$  and transition probability matrix

$$P = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \end{pmatrix}$$

Which of the following statements is/are true?

- 0 and 1 are recurrent states.
  - 2 is a transient state.
  - The Markov chain has a unique stationary distribution.
  - The Markov chain is irreducible.
- 2) Let  $X_1, X_2, \dots, X_n$  be a random sample of size  $n (\geq 2)$  from a population having Poisson distribution with mean  $\lambda$ , where  $\lambda > 0$  is an unknown parameter. If  $T_1 = \bar{X}$  and  $T_2 = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2}$ , where  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ , then which of the following statements is/are true?
- $T_1$  is an unbiased estimator of  $\lambda$ .
  - $T_2$  is an unbiased estimator of  $\sqrt{\lambda}$ .
  - $T_2^2$  is an unbiased estimator of  $\lambda$ .
  - Both  $T_1$  and  $T_2$  are estimators of  $\lambda$  as well as  $\lambda^2$ .
- 3) Let  $X_1, X_2, X_3$  be a random sample of size 3 from a population having Bernoulli distribution with parameter  $p$ , where  $p \in (0, 1)$  is unknown. Define

$$T_1(X_i, X_j, X_k) = X_i - X_j(1 - X_k), \quad T_2(X_i, X_j, X_k) = \frac{1}{2}(X_i X_j + X_j X_k),$$

for  $i, j, k = 1, 2, 3$  with  $i \neq j \neq k$ . Let  $x_1, x_2, x_3$  denote realizations from the random sample. Then which of the following statements is/are true?

- $T_1(X_1, X_2, X_3)$  has the same distribution as  $T_1(X_2, X_3, X_1)$ , but  $T_1(x_1, x_2, x_3) \neq T_1(x_2, x_3, x_1)$  for some realization  $x_1, x_2, x_3$ .
  - $T_2(X_1, X_2, X_3)$  and  $T_2(X_3, X_2, X_1)$  are both unbiased estimators for  $p^2$ .
  - $T_1(X_1, X_2, X_3)$  and  $T_1(X_2, X_3, X_1)$  are both unbiased estimators for  $p^2$ , and  $T_1(x_1, x_2, x_3) = T_1(x_2, x_3, x_1)$  for all realizations  $x_1, x_2, x_3$ .
  - $T_2(x_1, x_2, x_3) = T_2(x_2, x_3, x_1)$  for all realizations  $x_1, x_2, x_3$ .
- 4) Let  $X_1, X_2, \dots, X_n$  be a random sample of size  $n (\geq 2)$  from a population having probability density function

$$f(x; \lambda) = \begin{cases} \frac{1}{\lambda} e^{-x/\lambda}, & x > 0 \\ 0, & \text{otherwise} \end{cases}$$

where  $\lambda > 0$  is an unknown parameter. Let  $T_1 = \sum_{i=1}^n X_i$  and  $T_2 = (\sum_{i=1}^n X_i)^{-1}$ . For any positive integer  $v$  and any  $\alpha \in (0, 1)$ , let  $X_{v,\alpha}^2$  denote the  $(1 - \alpha)$ -th quantile of the central

chi-square distribution with  $\nu$  degrees of freedom. Consider testing  $H_0 : \lambda = \lambda_0$  against  $H_1 : \lambda > \lambda_0$ . Then which of the following tests is/are uniformly most powerful at level  $\alpha$ ?

- $H_0$  is rejected if  $\frac{2}{\lambda_0} T_1 > X_{2n, \alpha}^2$ .
  - $H_0$  is rejected if  $\frac{2}{\lambda_0} T_1 > X_{2n, 1-\alpha}^2$ .
  - $H_0$  is rejected if  $\frac{\lambda_0}{2} T_1 > X_{2n, \alpha}^2$ .
  - $H_0$  is rejected if  $\frac{\lambda_0}{2} T_1 > X_{2n, 1-\alpha}^2$ .
- 5) Let  $\{1, 6, 5, 3\}$  and  $\{11, 7, 15, 4\}$  be realizations of two independent random samples of size 4 from two separate populations having cumulative distribution functions  $F(\cdot)$  and  $G(\cdot)$ , respectively, and probability density functions  $f(\cdot)$  and  $g(\cdot)$ , respectively. To test  $H_0 : F(t) = G(t)$  for all  $t$  against  $H_1 : F(t) \geq G(t)$  with strict inequality for some  $t$ , let  $U_{MW}$  denote the Mann-Whitney U-test statistic. Let, under  $H_0$ ,  $P(U_{MW} > 12) \leq 0.10$ ,  $P(U_{MW} > 14) \leq 0.05$ ,  $P(U_{MW} > 15) \leq 0.025$ , and  $P(U_{MW} > 16) \leq 0.01$ . Then, based on the given data, which of the following statements is/are true?
- $H_0$  is rejected at level 0.10.
  - $H_0$  is rejected at level 0.05.
  - $H_0$  is rejected at level 0.025.
  - $H_0$  is rejected at level 0.01.

- 6) Let  $(X_1, X_2, X_3)$  have  $N_3(\mu, \Sigma)$  distribution with  $\mu = \begin{pmatrix} 2 \\ -3 \\ 1 \end{pmatrix}$  and  $\Sigma = \begin{pmatrix} 25 & -2 & 4 \\ -2 & 4 & 1 \\ 4 & 1 & 9 \end{pmatrix}$ . For which of the following vectors  $a$ ,  $X_2$  and  $X_2 - a^T \begin{pmatrix} X_1 \\ X_2 \end{pmatrix}$  are independent?

- $a = \begin{pmatrix} 0 \\ 1 \end{pmatrix}$
  - $a = \begin{pmatrix} -1 \\ -1 \end{pmatrix}$
  - $a = \begin{pmatrix} 0 \\ 2 \end{pmatrix}$
  - $a = \begin{pmatrix} 2 \\ 2 \end{pmatrix}$
- 7) Let  $A$  be a  $2 \times 2$  real matrix such that the trace of  $A$  is 5 and the determinant of  $A$  is 6. If the characteristic polynomial of  $(A + I_2)^{-1}$  is  $x^2 - bx + c$ , where  $I_2$  is the  $2 \times 2$  identity matrix, then  $\frac{b}{c}$  equals \_\_\_\_\_ (in integer).
- 8) Let  $\{X_n\}_{n \geq 1}$  be a time homogeneous discrete-time Markov chain with state space  $\{0, 1, 2\}$  and transition probability matrix

$$P = \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \end{pmatrix}$$

If  $P(X_0 = 0) = P(X_0 = 1) = \frac{1}{4}$ , then  $32E(X_2)$  equals \_\_\_\_\_ (integer).

- 9) Let  $(X, Y)$  be a random vector having joint moment generating function given by

$$M_{X,Y}(u, v) = \frac{e^{\frac{u^2}{2}}}{(1 - 2v)^3}, \quad -\infty < u < \infty, -\infty < v < \frac{1}{2}.$$

Then  $E(\frac{6X^2}{Y})$  equals \_\_\_\_\_ (rounded off to two decimal places).

- 10) Let  $\{N(t)\}_{t \geq 0}$  be a Poisson process with rate  $\lambda$ , where  $\lambda > 0$  is an unknown parameter. Starting from the origin, an intercity road has  $N(t)$  number of potholes up to a distance of  $t$  kilometers. Starting from the origin, potholes are found at the following distances (in kilometers):

0.9, 1.3, 1.8, 2.7, 3.4, 4.1, 4.7, 5.5, 6.2, 6.8, 7.4, 8.1, 8.9, 9.2, 9.7.

Based on the above data, the method of moment estimate of  $\lambda$  equals \_\_\_\_\_ (rounded off to two decimal places).

- 11) Let  $X_1, X_2, X_3, X_4$  be a random sample of size 4 from a population having uniform distribution over the interval  $(0, \theta)$ , where  $\theta > 0$  is an unknown parameter. Let  $X_{(4)} = \max\{X_1, X_2, X_3, X_4\}$ . To test  $H_0 : \theta = 1$  against  $H_1 : \theta = 0.1$ , consider the critical region that rejects  $H_0$  if  $X_{(4)} < 0.3$ . Let  $p$  be the probability of the Type-I error. Then  $100p$  equals \_\_\_\_\_ (rounded off to two decimal places).
- 12) Let a random sample of size 4 from a normal population with unknown mean  $\mu$  and variance 1 yield the sample mean of 0.16. Let  $\Phi(\cdot)$  be the cumulative distribution function of the standard normal random variable. It is given that  $\Phi(2.28) = 0.989$ ,  $\Phi(1.96) = 0.975$ , and  $\Phi(1.64) = 0.95$ . If the likelihood ratio test of size 0.05 is used to test  $H_0 : \mu = 0$  against  $H_1 : \mu \neq 0$ , then the power of the test at the sample mean equals \_\_\_\_\_ (rounded off to three decimal places).
- 13) Consider the multiple linear regression model

$$y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \epsilon_i, \quad i = 1, 2, \dots, 25$$

where  $\beta_0, \beta_1$ , and  $\beta_2$  are unknown parameters, and  $\epsilon_i$ 's are uncorrelated random errors with mean 0 and finite variance  $\sigma^2 > 0$ . Let  $F_{\alpha, m, n}$  be such that  $P(Y > F_{\alpha, m, n}) = \alpha$ , where  $Y$  is a random variable with an  $F$ -distribution with  $m$  and  $n$  degrees of freedom. Suppose that testing

$$H_0 : \beta_1 = \beta_2 = 0 \quad \text{against} \quad H_1 : \text{At least one of } \beta_1, \beta_2 \text{ is not } 0$$

involves computing  $F_0 = 11 \frac{R^2}{1-R^2}$  and rejecting  $H_0$  if the computed value  $F_0$  exceeds  $F_{\alpha, 2, 22}$ . Given that  $F_{0.025, 2, 22} = 4.38$  and  $F_{0.05, 2, 22} = 3.44$ , the smallest value of  $R^2$  that would lead to rejection of  $H_0$  for  $\alpha = 0.05$  equals \_\_\_\_\_ (rounded off to two decimal places).