# Fourth Home Assignment Data Analysis

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#### Question 1

From probability theory we have the following definitions and properties:

$$(a) p_X(x) = \sum_{y \in \mathcal{Y}} p_{XY}(x, y)$$

(b) If X and Y are independent, then  $P_{XY}(x,y) = p_X(x)p_Y(y)$ 

$$(c) \ \mathbb{E}[X] = \sum_{x \in \mathcal{X}} x p_X(x)$$

X and Y are discrete random variables that take values from in  $\mathcal{X}$  and  $\mathcal{Y}$ .  $p_X$  is the distribution of X,  $p_Y$  the distribution of Y and  $p_{XY}$  the distribution of X and Y.

1)

We prove the following identity:

$$\mathbb{E}[X+Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

By (c), the expected value of X is given by:

$$\mathbb{E}[X] = \sum_{x \in \mathcal{X}} x p_X(x)$$

Therefore the expected value of X + Y would be

$$\begin{split} \mathbb{E}[X+Y] &= \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} (x+y) p_{XY}(x,y) \\ &= \sum_{x \in \mathcal{X}} x \sum_{y \in \mathcal{Y}} p_{XY}(x,y) + \sum_{y \in \mathcal{Y}} y \sum_{\in \mathcal{X}} p_{XY}(x,y) \\ &= \sum_{x \in \mathcal{X}} x p(x) + \sum_{y \in \mathcal{Y}} y p(y) \\ &= \mathbb{E}[X] + \mathbb{E}[Y] \end{split}$$

In the last step we use the definition for the expected value of a random variable. We have now shown that  $\mathbb{E}[X+Y] = \mathbb{E}[X] + \mathbb{E}[Y]$ .

2)

To prove the following identity, we use that the random variables X and Y are independent.

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$$

We can write  $\mathbb{E}[XY]$  as

$$\mathbb{E}[XY] = \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} xyp_{XY}(x, y)$$

This is where we use that X and Y are independent - using property (b):

$$\sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} xy p_X(x) p_Y(y)$$

This can be reduced to prove our identity

$$\sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} xy p_X(x) p_Y(y) = \sum_{x \in \mathcal{X}} x p_X(x) \sum_{y \in \mathcal{Y}} y p_Y(y)$$
$$= \mathbb{E}[X] \mathbb{E}[Y]$$

This proves the identity  $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$ .

3)

A bag has 2 red apples and 2 green apples. There is taken 2 apples from the bag without putting them back into the bag. Let X be the first apple and let Y be the second apple. The joint distribution table of X and Y is seen below:

X / Y	Red	Green
Red	$\frac{1}{6}$	$\frac{2}{6}$
Green	$\frac{2}{6}$	$\frac{1}{6}$

The probability of apple X being red is:

$$\mathbb{E}[X = \text{Red}] = \frac{1}{2}$$

Which is the same probability for apple Y being red. We have that

$$\mathbb{E}[X = \operatorname{Red} \wedge Y = \operatorname{Red}] = \frac{1}{6}$$

Since  $\frac{1}{2}\frac{1}{2} = \frac{1}{4} \neq \frac{1}{6}$  then

$$\mathbb{E}[XY] \neq \mathbb{E}[X]\mathbb{E}[Y]$$

in this example.

4)

The identity to be proved:

$$\mathbb{E}[\mathbb{E}[X]] = \mathbb{E}[X]$$

We know that  $\mathbb{E}[X] = k$  and that  $\mathbb{E}[k] = k$ . That means taking the expected value of an expected value will just return the constant you already found. This can be done more than 2 times and it will always be the constant k that is your result.

### Question 2

1)

N balls are drawn from a bin with 2N balls uniformly and without replacement. The fraction of red balls is  $\varepsilon$  and  $0 < \varepsilon \le \frac{1}{2}$ . The other fraction are green balls  $(1 - \varepsilon)$ . We are asked to show that

$$\mathbb{P}\{N \text{ green balls are pulled in a row}\} \leq e^{-N\varepsilon}$$

The probability of getting N green balls in a row can be defined as

$$\prod_{i=0}^{N-1} \frac{2N(1-\varepsilon)-i}{2N-i}$$

We subtract i since it is without replacement, so the number of balls in the bin decreases. This can be rewritten:

$$\begin{split} \prod_{i=0}^{N-1} \frac{2N(1-\varepsilon)-i}{2N-i} &= \prod_{i=0}^{N-1} \frac{2N-2N\varepsilon-i}{2N-i} \\ &= \prod_{i=0}^{N-1} 1 - \frac{2N\varepsilon}{2N-i} \end{split}$$

From the assignment text, we have:

$$1 + x < e^x$$

This can be used on the probability of drawing N green balls in a row we found, so we get

$$\begin{split} \prod_{i=0}^{N-1} 1 - \frac{2N\varepsilon}{2N-i} &\leq \prod_{i=0}^{N-1} e^{-2N\varepsilon/(2N-i)} \\ &= e^{\sum_{i=0}^{N-1} -2N\varepsilon/(2N-i)} \end{split}$$

This neat trick can be used since  $e^x \cdot e^x = e^{x+x}$ . The constants can be moved outside the sum, so

$$\prod_{i=0}^{N-1} 1 - \frac{2N\varepsilon}{2N-i} \le e^{-2N\varepsilon \sum_{i=0}^{N-1} 1/(2N-i)}$$

The sum can never exceed 1 and it can never be below  $\frac{1}{2}$ , which means the factor  $-2N\varepsilon$  is at max  $-N\varepsilon$  and at minimum  $-2N\varepsilon$ , i.e. that

$$e^{-2N\varepsilon\sum_{i=0}^{N-1}1/(2N-i)} \le e^{-N\varepsilon}$$

If we put the two above inequalities together, we get:

$$\prod_{i=0}^{N-1} 1 - \frac{2N\varepsilon}{2N-i} \le e^{-2N\varepsilon \sum_{i=0}^{N-1} 1/(2N-i)} \le e^{-N\varepsilon}$$

Where the left hand side was the probability of drawing N green balls in a row, so we have shown the inequality  $\mathbb{P}\{N \text{ green balls are pulled in a row}\} \leq e^{-N\varepsilon}$ .

## Question 3

This question is about the growth function

1)

We are asked to show that

$$m_{\mathcal{H}}(2N) \le m_{\mathcal{H}}(N)^2$$

If we split the set of 2N into two sets with N points, the most dichotomies we can get is the cross product of the two sets. From "Learning From Data", page 45, we have the upper bound on the growth function

$$m_{\mathcal{H}}(N) \leq 2^N$$

So in each set we have at most  $2^N$  dichotomies. The cross product (and the max number of dichotomies in  $m_{\mathcal{H}}(2N)$  is  $2^N \cdot 2^N$ . So we can insert this upper bound of the left side of the inequality we wanted to show

$$2^N \cdot 2^N \le m_{\mathcal{H}}(N)^2$$

If the maximum dichotomies is reached each set of N that means we can insert that number on our right side as well:

$$2^N \cdot 2^N \le \left(2^N\right)^2 \text{ or }$$
$$2^{2N} \le 2^{2N}$$

Which shows the inequality.

2)

To write a generalization bound that only involves  $m_{\mathcal{H}}(N)$ , we use equation 2.12 from "Learning From Data":

$$E_{out}(g) \le E_{in}(g) + \sqrt{\frac{8}{N} \ln \frac{4m_{\mathcal{H}}(2N)}{\delta}}$$

The upper bound that was shown in previous question can easily be substituted into this equation at the cost of a looser bound:

$$E_{out}(g) \le E_{in}(g) + \sqrt{\frac{8}{N} \ln \frac{4m_{\mathcal{H}}(N)^2}{\delta}}$$

### Question 4

1)

We are asked to prove the following by induction

$$\sum_{i=0}^{d} \binom{N}{i} \le N^d + 1$$

**Basis**: The statement should hold for d = 0:

$$\sum_{i=0}^{0} \binom{N}{i} \le N^0 + 1$$

$$1 < 2$$

It does.

**Inductive step:** We assume our first equation holds for d and want to show it holds for d+1 and  $d \ge 1$  as well:

$$\sum_{i=0}^{d+1} \binom{N}{i} \le N^{d+1} + 1$$

We can take out the last element in the sum so we get

$$\binom{N}{d+1} + \sum_{i=0}^{d} \binom{N}{i} \le N^{d+1} + 1$$
$$= N^{d} + (N-1)N^{d} + 1$$

This is where we use the assumption so we insert the upper bound instead of the sum and get:

$$\binom{N}{d+1} + N^d + 1 \le N^d + (N-1)N^d + 1$$
$$\binom{N}{d+1} \le (N-1)N^d$$
$$= N^{d+1} - N^d$$

The right hand side grows faster for all N and  $d \le 1$ , but not for d = 0. However, we have already shown that it holds for d = 0.

2)

We can use the above result to derive a bound on  $m_{\mathcal{H}}(N)$  by using equation 2.9 from "Learning From Data":

$$m_{\mathcal{H}}(N) \le \sum_{i=0}^{d} \binom{N}{i}$$

Simply substitute the upper bound in the inequality with the one we found in (1):

$$m_{\mathcal{H}}(N) \le N^d + 1$$

Again, this bound is a bit looser than equation 2.9.

3)

The equation 2.12 in "Learning From Data" states:

$$E_{out}(g) \le E_{in}(g) + \sqrt{\frac{8}{N} \ln \frac{4m_{\mathcal{H}}(2N)}{\delta}}$$

Just like before we replace the factor  $m_{\mathcal{H}}(N)$  with the upper bound we found in (2):

$$E_{out}(g) \le E_{in}(g) + \sqrt{\frac{8}{N} \ln \frac{4 \cdot ((2N)^d + 1)}{\delta}}$$

For the bound to be meaningful, d should be (alot) less than N.