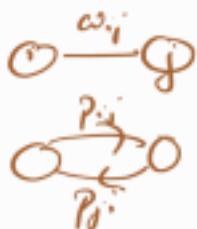
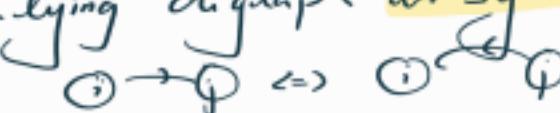


NORMALIZED CONDUCTANCE

Defn. A Markov chain with TPM P and a stationary distribution π is called time reversible if

$$\forall i, j \quad \pi_i P_{ij} = \pi_j P_{ji}.$$

\Rightarrow Underlying digraph is symmetric



$\rightarrow G$ can be obtained from an undirected graph H by doubling each edge. (H may have self loops)

$\rightarrow P$ can be encoded as weights on the edges of H .

$$w_{ij} = \pi_i P_{ij} (= \pi_j P_{ji})$$

Claim: If I know all w_{ij} values then I can compute all P_{ij} values.

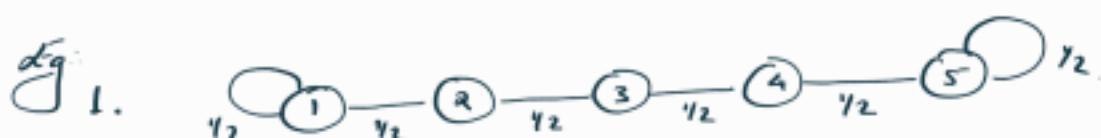
Qn. Can we find P_{ij} from w_{ij} 's?

$$\begin{aligned} \hat{v}_{ii} &= \sum_j w_{ij} & (w_{ij} = \hat{\pi}_i P_{ij}) \\ \checkmark P_{ij} &= \frac{w_{ij}}{\hat{v}_{ii}} & \sum_j w_{ij} \\ && = \sum_j \hat{\pi}_i P_{ij} \\ \text{Some time } w_{ij}'s &\text{ may be scaled by an} \\ \text{unknown constant } &(w_{ij} = c \hat{\pi}_i P_{ij}). & = \hat{\pi}_i \sum_j P_{ij} \\ && = \hat{\pi}_i \end{aligned}$$

In that case,

$$\begin{aligned} \checkmark c &= \sum_i \sum_j w_{ij} & \sum_i \sum_j w_{ij} \\ w_i &= \sum_j w_{ij} & = \sum_i \sum_j c \hat{\pi}_i P_{ij} \\ P_{ij} &= w_{ij}/w_i & = c \sum_i \hat{\pi}_i \sum_j P_{ij} \\ && = c \end{aligned}$$

Hence the weighted undirected graph \mathcal{W}
(with possibly self loops) defines the random walk.



(No $w_{ij} \Rightarrow w_{ij} = 1$).



"RANDOM WALKS ON
WEIGHTED UNDIRECTED GRAPHS"

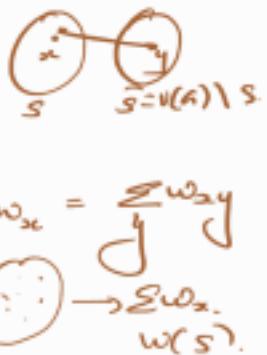


(P need not be
symmetric)

Defn. The normalised conductance $\phi(a)$ of a weighted undirected graph G is defined as

$$\phi(a) = \min_{\substack{S \subseteq V(G) \\ 0 < w(S) \leq w(\bar{S})}} \left\{ \frac{\sum_{x \in S, y \in \bar{S}} w_{xy}}{w(S)} \right\}$$

$$\text{where } w(S) = \sum_{x \in S} w_{xx} \quad \left| = \min_{S \subseteq V(G)} \left\{ \frac{\sum_{x \in S, y \in \bar{S}} w_{xy}}{\min(w(S), w(\bar{S}))} \right\} \right.$$



Intuition. (Assume $\sum_{x,y} w_{xy} = 1$, so that $(c=1)$)

$$\sum_{x,y} w_{xy} = \hat{v}_x \text{ and } w_{xy} = \hat{v}_x P_{xy}$$

$$P_{xy} = P_n \left[\text{next state} = y / \text{curr. st.} = x \right]$$

$$= P_n \left[\text{next move in along the arc } xy / \text{curr. st.} = x \right]$$

$$w_{xy} = \hat{v}_x P_{xy} = P_n \left[\text{next move in along the arc } xy \right] \text{ in steady state}$$



$$\sum_{x \in S, y \in \bar{S}} w_{xy} = P_n \left[\text{next move in an escape from } S \right] \text{ in SS.}$$

$$\frac{1}{w(S)} \sum_{x \in S, y \in \bar{S}} w_{xy} = P_n \left[\text{next state in outside } S / \text{curr. st. in } S \right] \text{ in SS.}$$

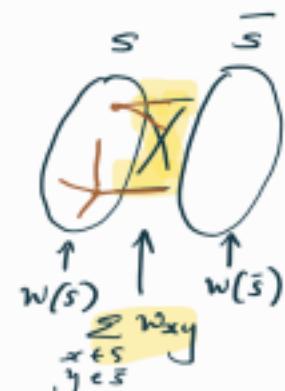
= "Escape probability from S ".

ϕ = minimum escape prob among
all sets $S \subseteq V(\mathcal{A})$ with
 $0 < w(S) \leq \frac{1}{2}$

Defn. (Norm. Cond of ϕ cut).

for $S \subseteq V(\mathcal{A})$ with $w(S) > 0$,

$$\phi(S, \bar{S}) := \frac{\sum_{x \in S, y \in \bar{S}} w_{xy}}{\min\{w(S), w(\bar{S})\}} \leq 1$$



$\phi(\mathcal{A}) = \min \phi(S, \bar{S})$, where the
minimum is over all non-trivial
cuts of \mathcal{A} .

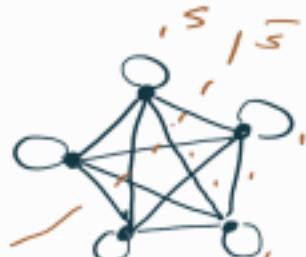
EXAMPLES

1. K_n with self loops

$$w_{ij} = 1 \quad \# \text{edges}$$

$$w_i = n \quad \# \text{vertices}$$

$$w(S) = n/|S|$$



$$P_{ij} = \frac{w_{ij}}{w_i} = \frac{1}{n}$$

Let (S, \bar{S}) be a cut with

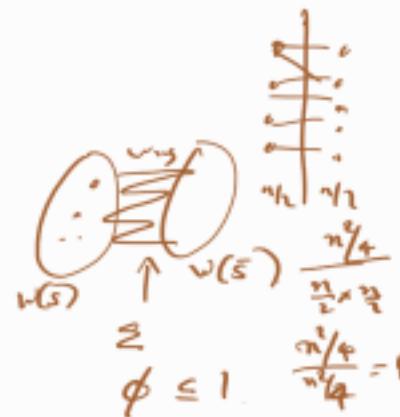
$$k = |S| \leq |\bar{S}|$$

$$\phi(S, \bar{S}) = \frac{k \times (n-k)}{\min\{nk, n(n-k)\}}$$

$$= \frac{k \times (n-k)}{nk} = \frac{n-k}{n}$$

$$\phi(\mathcal{A}) = \min_{k \leq \frac{n}{2}} \left(\frac{n-k}{n} \right) = \frac{1}{2}$$

$$\sum_{x \in S, y \in \bar{S}} w_{xy} = |E(S, \bar{S})|$$



2. P_n with end loops



$$w_{ij} = 1 \text{ if edges}$$

$$\text{let } S \subseteq [n]$$

$$w(S) = 2|S|$$

Hence $|S| < |\bar{S}| \Rightarrow w(S) \leq w(\bar{S})$
 $(1 \leq i \leq n/2)$

$$\phi(S, \bar{S}) = \sum_{x \in S, y \in \bar{S}} w_{xy} / w(S)$$

$$= \frac{|E(S, \bar{S})|}{2|S|} \leftarrow \text{Minimize}$$



$$|S| = k \leq n/2$$

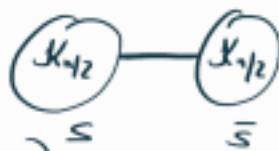
$$w(a) = \min_{|S| \leq n/2} \phi(S, \bar{S}) \approx \frac{1}{2 \cdot \frac{n}{2}} = \frac{1}{n}$$

3. A disconnected $\Rightarrow \phi(a) = 0$.



Pick S to be $V(C_i)$ with $w(V(C_i))$ smallest.

4. Dumbbell



$$(w_{ij} = 1 \text{ if edges})$$

$$w(S) = \frac{n}{2} \cdot \frac{n}{2} + 1 \approx n^2/4$$

$$|E(S, \bar{S})| = 1.$$

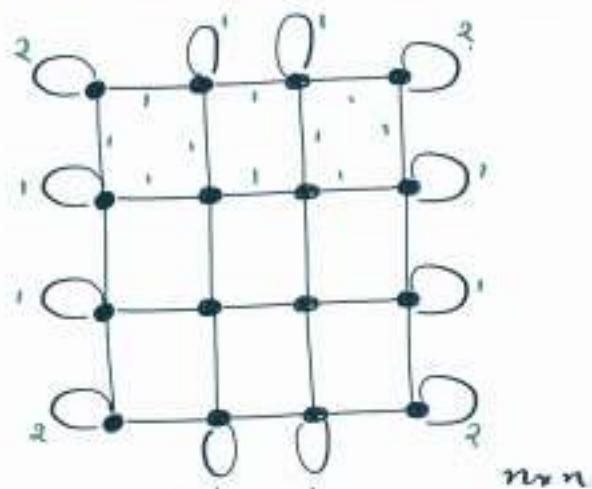
$$\phi(S, \bar{S}) = \frac{4}{n^2}$$

$$\phi(a) \leq \frac{4}{n^2} \text{ "Worse than path"}$$



No monotonicity

2D-lattice with end loops



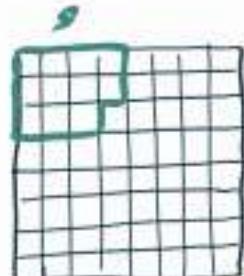
$$|s| = k = m^2$$

$w(s) = 4|s|$ hence
 $|s| < |\bar{s}| \Rightarrow w(s) \leq w(\bar{s})$.

Obs 1.

If $|s| \leq \frac{1}{4}n^2$ then a corner square has least no. of escape edges relative to $|s|$

$$\mathcal{E}(s, \bar{s}) \approx 2\sqrt{|s|}$$



Hence $\phi(s, \bar{s}) = \frac{\mathcal{E}(s, \bar{s})}{w(s)}$
 $\approx \frac{2\sqrt{|s|}}{\pi|s|}$
 $= \frac{1}{2\sqrt{|s|}} \leq \frac{1}{2\sqrt{\frac{n^2}{4}}} = \underline{\underline{\frac{1}{n}}}$

Obs 2.

If $\frac{n^2}{4} < |s| \leq \frac{n^2}{2}$, then a margin strip has least no. of escape edges.

$$\mathcal{E}(s, \bar{s}) \approx n$$



Hence $\phi(s, \bar{s}) = \frac{n}{4|s|}$
 $\leq \frac{n}{4 \cdot \frac{n^2}{2}} = \underline{\underline{\frac{1}{2n}}}$

$$\phi(a) = \gamma_{2n} \quad P_n - \gamma_n$$

$$= \gamma_{2\sqrt{N(\omega)}}$$

6. d -dimensional lattice with end loops

$$|V(a)| = n^d$$

$$\phi(a) = \frac{1}{d} n \quad (\text{Exercise})$$

$$= \frac{1}{d} |V(a)|^{1/d}$$

Why is $\phi(a)$ important?

$$\Omega\left(\frac{1}{\phi(a)}\right) \leq \underset{\text{time}}{\text{Mixing}} \leq O\left(\frac{1}{\phi(a)}\right).$$

Mixing Time

"No. of steps by which any starting distribution gets trapped in an ϵ -neighbourhood of the stationary distribution."

Formal defn depends on

1. Distance measure used
2. We use $p(t)$ or $a(t)$.

Defn (BHK Defn 4.1)

Fix any $\epsilon > 0$. The ϵ -mixing time of a random walk is the smallest integer t such that for any starting distribution P , $\|a(t) - \bar{v}\|_1 < \epsilon$, where \bar{v} is the stationary distribution of the random walk.

Mixing Time & Normalized Conductance

LECTURE 19
25/Apr/2021

$$\cdot \phi(\varepsilon) = \phi(s, \bar{s}) = \frac{\sum_{x \in s, y \in \bar{s}} w_{xy}}{\min\{w(s), w(\bar{s})\}}$$



$$\cdot \phi = \min_{\forall t, s \subseteq V(G)} \phi(s).$$

$$\cdot \tau_\varepsilon = \max_s \min_t \|a_s(t) - \hat{v}\|_1 < \varepsilon$$

Goal: $\Omega\left(\frac{\tau_s}{\phi}\right) \leq \tau_\varepsilon \leq O\left(\frac{\log(\hat{v}_{\min})}{\varepsilon^2 \phi^2}\right)$

$$\hat{v}_{\min} = \min_i \hat{v}_i$$

Tool: Probability flow across a cut (s, \bar{s}) .

Let $q = pP$ (current distn p
Next distn q).

for a set $s \subseteq V(G)$,

net prob. out-flow = $\Delta_p(s) = P(s) - q(s)$.

$$(Net \Delta_p(s)) = \sum_{i \in s, j \in \bar{s}} (p_i p_{ij} - p_j p_{ji})$$

$$P(s) = \sum_{i \in s} p_i$$

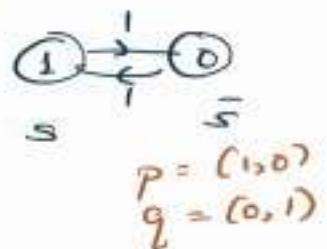
$$= \sum_{i \in s, j \in \bar{s}} \left(\frac{p_i}{v_i} w_{ij} - \frac{p_j}{v_j} w_{ij} \right)$$

$$= \sum_{i \in s, j \in \bar{s}} \left(\frac{p_i}{v_i} - \frac{p_j}{v_j} \right) w_{ij}$$

$$= \sum_{i \in s, j \in \bar{s}} (v_i - v_j) w_{ij} \quad (v_i = \underline{p_i/v_i})$$

Obs: 1. $\Delta_p(s)$ can be as large as 1.

and as small as -1



2. $\Delta_n(s) = 0 \nabla s \subseteq V(\mathcal{L})$

3. $\Delta_{\alpha(t)}(s) \leq \frac{2}{t} \nabla s \subseteq V(\mathcal{L})$
(+ starting distn o.)

Recall $\alpha(t) = \frac{1}{t}(p^{(0)} + \dots + p^{(t)})$.

$$(p^{(0)} = \text{---}, \quad p^{(i+1)} = p^{(i)} p)$$

$$\begin{aligned} \alpha(t)p - \alpha(t) &= \frac{1}{t}(p^{(1)} + \dots + p^{(t+1)}) - \\ &\quad \frac{1}{t}(p^{(0)} + \dots + p^{(t)}) \\ &= \frac{1}{t}(p^{(t+1)} - p^{(0)}) \end{aligned}$$

$$\| \alpha(t)p - \alpha(t) \|_1 = \frac{1}{t} \| p^{(t+1)} - p^{(0)} \|_1 \quad \| p - q \|_1 \\ p = (0, 1, 0) \\ q = (1, 0, 0)$$

$$\Delta_{\alpha(t)}(s) = \frac{2}{t}$$

Use full observation

$$\begin{aligned} \| p - q \|_1 &= \sum_i |p_i - q_i| \\ &= \sum_{p_i > q_i} (p_i - q_i) + \sum_{p_i < q_i} (q_i - p_i) \\ &= 2 \sum_{p_i > q_i} (p_i - q_i) \end{aligned}$$

$$\underbrace{\sum_{i \in S} p_i - \sum_{i \in S} q_i}_S$$

$$\begin{aligned} \alpha(t) &= \sum_{i \in S} p_i - \sum_{i \in S} q_i \\ &= \sum_{i \in S} (p_i - q_i) \\ &\leq \sum_{i \in S} |p_i - q_i| \\ &= \sum_i |p_i - q_i| \\ &= \| p - q \|_1. \end{aligned}$$

Exercise

Direction 1. ($\frac{1}{\phi} \leq \bar{\tau}_{\eta_2}$)

Let $s: \hat{v}_i(s) \leq \hat{v}_i(\bar{s})$

Consider the starting distribution

to be $\sigma = \pi|_S$ "restricted to S "

$$\sigma_i = \begin{cases} \hat{v}_i / \hat{v}(s) & , i \in S \\ 0 & , i \in \bar{S} \end{cases} \quad (v_i = \hat{v}_i(s) = \sigma_i(s))$$

$$\begin{aligned} \Delta_\sigma(s) &= \sum_{\substack{i \in S \\ j \in \bar{S}}} (v_i - v_j) \sigma_{ij} \\ &= \sum_{i \in S} \left(\hat{v}_i(s) - 0 \right) \sigma_{ij} \\ &= \frac{1}{\hat{v}(s)} \sum_{i \in S, j \in \bar{S}} \sigma_{ij} \\ &= \phi(s, \bar{s}) \quad (\phi(s)) \end{aligned}$$

for the distribution $\sigma = \pi|_S$,

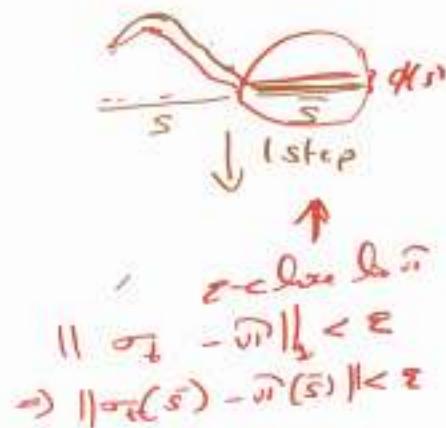
$$\Delta_\sigma(s) = \phi(s, \bar{s}).$$

$$\underline{\Delta_{\pi|_S}(s) = \phi(s, \bar{s})}$$

$$\sigma(s) = 0$$

In the subsequent moves the flow is only smaller (why?)

$$\begin{aligned} \tau_e &> \frac{\hat{v}_i(\bar{s}) - \eta_L}{\phi(s)} \\ &\geq \frac{\eta_2 - \eta_L}{\phi(s)} \\ &\approx \frac{1}{2\phi(s)}. \end{aligned}$$



$$\bar{\tau}_{\eta_2} \geq \frac{1}{2\phi(s)} \geq \frac{1}{2\phi}$$

Other Direction. ($\tau_c \leq O\left(\frac{\log(\sqrt{a_0})}{\epsilon^2 \phi^2}\right)$)

Let $a = a(t)$.

$$S = \{i : a_i \geq \bar{a}_i\}$$

$$= \{i : v_i \geq 1\}$$

Vertices in S are lucky.
" " \bar{S} " unlucky.

$$(v_i = a_i / \bar{a}_i)$$

$$\Delta_a(s) = \sum_{i \in S, j \in \bar{S}} (v_i - v_j) w_{ij}$$

$$> \sum_{i \in S, j \in \bar{S}} (v_i - 1) w_{ij} \quad (v_j < 1 \forall j \in \bar{S})$$

$$= \sum_{i \in S} (v_i - 1) \sum_{j \in \bar{S}} w_{ij} \quad (*)$$



$$= \sum_{i \in S} (v_i - 1) w_i|_{\bar{S}} \quad (w_i|_{\bar{S}} = \sum_{j \in \bar{S}} w_{ij})$$

Over simplified idea

$$\text{if } \forall i \in S \quad (v_i - 1) \approx c > 0$$

$$\begin{aligned} \text{Then } (*) &\Rightarrow \Delta_a(s) \geq c \sum_{i \in S, j \in \bar{S}} w_{ij} \\ &\geq c \phi(s) \min \{\hat{v}(s), \hat{v}(\bar{s})\} \\ &\geq c \phi(s) \min \{\hat{v}(s), \frac{\epsilon}{2}\} \\ &\quad (\hat{v}(\bar{s}) < \frac{\epsilon}{2} \Rightarrow \sum_{\bar{a}_i > a_i} (\bar{a}_i - a_i) < \frac{\epsilon}{2} \\ &\quad \Rightarrow \|\bar{a} - a\|_1 < \frac{\epsilon}{2}) \end{aligned}$$

$$\geq \underline{c \phi(s) \hat{v}(s) \frac{\epsilon}{2}}$$

$$\text{But } \Delta_a(s) = \frac{2t}{t}$$

$$\text{Hence } c \leq \frac{4}{t \phi(s) \hat{v}(s) \epsilon}$$

$$\left| \begin{array}{l} c \phi(s) \hat{v}(s) \frac{\epsilon}{2} \leq \frac{2}{t} \\ c \leq \frac{4}{t \phi(s) \hat{v}(s) \epsilon} \end{array} \right.$$

Now,

$$\begin{aligned}\|a - \hat{v}_i\|_1 &= \sum_i |a_i - \hat{v}_i| \\ &= 2 \sum_{i \in S} (a_i - \hat{v}_i) \quad (\text{ } a_i \geq \hat{v}_i) \\ &= 2 \sum_{i \in S} (v_i - 1) \hat{v}_i\end{aligned}$$

(Over simplified)

$$\|a - \hat{v}_i\|_1 \approx 2c \hat{v}_i(s)$$

$$\leq \frac{8}{t\phi(s)\epsilon}$$

$$\leq \frac{8}{t\phi\epsilon} \quad (\phi = \phi(s))$$

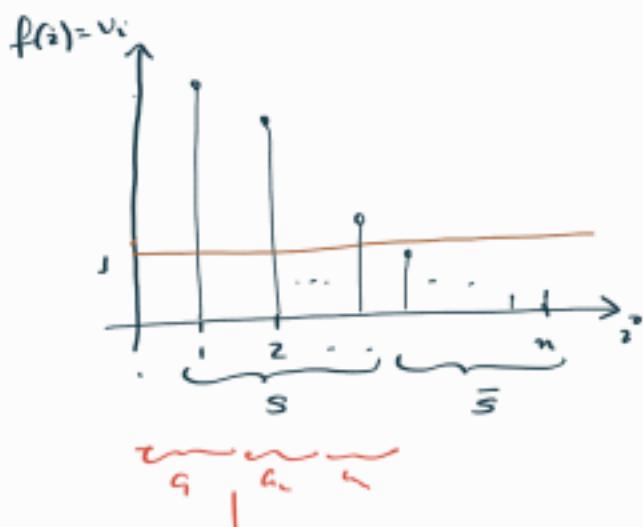
$$\text{If } t > \frac{8}{\phi\epsilon^2}, \text{ then } \frac{8}{t\phi\epsilon} < \epsilon. \quad \left| \frac{8}{t\phi\epsilon} < \epsilon \right.$$

The above over simplification is wrong.

v_i 's is never like a single step.

Relabel the vertices so that

$$v_1 \geq v_2 \geq \dots \geq v_n.$$



"A more careful analysis is required.
and you end up with

$$t_i \leq O\left(\frac{\log(\frac{1}{\epsilon_{\min}})}{\epsilon^3 \phi^2}\right)$$

$$= O\left(\frac{\log n}{\epsilon^3 \phi^3}\right) \text{ for uniform}$$

HITTING & COVERING TIMES

LECTURE 20
25/Apr/2021

Defn. For a random walk α on

$$V(\alpha) = [n],$$

1. Hitting time from i to j : $h(i,j)$

= expected time for the random walk to hit j for the first time if we start at i .

2. Hitting time of j

$$h(j) = \max_{i \in [n]} h(i,j).$$

(worst case)

3. Hitting time of α

$$h(\alpha) = \max_{i,j \in [n]} h(i,j).$$

4. Covering time of α : $c(\alpha)$

expected time for the random walk starting at i to reach every mode at least once.

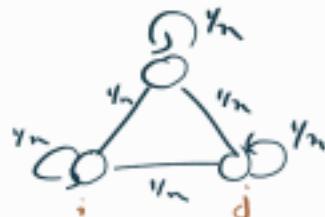
5. Covering time of α :

$$c(\alpha) = \max_i c(i).$$

MIXING TIME vs.
HITTING TIME vs.
COVERING TIME.

EXAMPLES

1. K_n with self loops



(a) Mixing time

$$\hat{P} = \begin{pmatrix} \frac{1}{n} & \frac{1}{n} & \dots & \frac{1}{n} \\ \vdots & & & \\ \frac{1}{n} & \dots & \frac{1}{n} \end{pmatrix}$$

$$\begin{aligned} P &= \frac{1}{n} - 1 \text{st step} \\ (1-p)P &= \frac{2}{n} - 2 \text{nd step} \\ P(\text{you will hit } j \text{ in } k \text{ steps}) &= (1-p)^{k-1} P \end{aligned}$$

For any starting distribution $P^{(0)}$

$$\hat{P}^{(t)} = (\frac{1}{n}, \dots, \frac{1}{n}) = \hat{\pi} \quad \leftarrow \text{Verify.}$$

$$\begin{aligned} \alpha(t) &= \frac{1}{t} (P^{(0)} + (t-1) \hat{\pi}) \\ &= \frac{1}{t} P^{(0)} + \left(1 - \frac{1}{t}\right) \hat{\pi} \end{aligned}$$

\downarrow
Geometric
RV
 $\text{Expected} = 1/p$

$$\| \alpha(t) - \hat{\pi} \|_1 = \| \frac{1}{t} (P^{(0)} - \hat{\pi}) \|_1 \leq \frac{2}{t}$$

$$\frac{2}{t} \leq \varepsilon \quad \uparrow \quad t \geq 2/\varepsilon$$

$$\Rightarrow \boxed{\bar{\tau}_\varepsilon < 2/\varepsilon} \quad \text{"constant"}$$

(If we define mixing in terms of $p(t)$ rather than $\alpha(t)$ then $\bar{\tau}_\varepsilon = 1$)

(b) Hitting time.

Let $i \neq j$

$$h(i, j) = E \left[\underbrace{\# \text{ steps to reach } j \text{ from } i}_{X} \right]$$

But X is geometric with $p = 1/n$.

$$P_n \{ X = k \} = (1-p)^{k-1} p$$

Hence $h(i,j) = \mathbb{E}[X] = \frac{1}{p}$ (Refer & Verify)

$$= n. \quad (\gg \underline{\bar{x}_e})$$

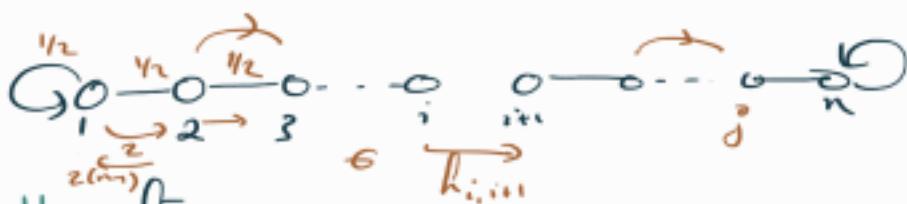
$\hookrightarrow h(a) = n$

(c) Covering time:

Equivalent to coupon collector problem.

Hence $C(K_n) \approx n \ln n$. \Rightarrow hitting time \gg mixing time.

2. Path with end loops



(a) Hitting time

$$h(1,2) = \mathbb{E}\{\text{Geom r.v. with mean } k_2\}$$

$$= 2.$$

$$h(2,3) = \frac{1}{2} + \frac{1}{2}(1 + h(1,2))$$

$$h(2,3) = \frac{1}{2} + \frac{1}{2}(1 + h(1,2))$$

$$\frac{1}{2}h(2,3) = \frac{1}{2} + \frac{1}{2}(1 + h(1,2))$$

$$h(i,i+1) = \frac{1}{2} + \frac{1}{2}[h(i-1,i+1) + 1]$$

$$R(2,3) = 2 + R(1,2)$$

$$= \frac{1}{2} + \frac{1}{2}[h(i-1,i) + h(i,i+1) + 1]$$

$$h(i,i+1) = 2 + h(i-1,i). \quad (h(2,3) = 2 + 2 = 4)$$

$$= 2i$$

$$h(i,j) = h(i,i+1) + h(i+1,i+2) + \dots + h(j-1,j)$$

$$(i < j) = 2i + 2(i+1) + \dots + 2(j-1)$$

$$= (i+j-1)(j-i)$$

$$\begin{aligned}
 \max_{i,j} h(i,j) &= h(1,n) \\
 &= 2(1+2+\dots+n-1) \\
 &= (n-1)n \\
 &\leq n^2 // P_n: \boxed{h(n) = n^2}
 \end{aligned}$$

(b) Covering line:
 $c(1) \leq h(1,n) \leq n^2$

What about $c(i)$ vs $h(i,1)$ & $h(i,n)$?

$$\begin{aligned}
 c(i) &\leq h(i,1) + h(1,n) \\
 &\leq 2h(1,n) \\
 &\leq 2n^2. \quad \text{Hence } c(n) \leq 2n^2
 \end{aligned}$$



(c) Mixing line:

$$\widehat{\gamma_1} = \min_{\gamma} \text{norm} \quad (\text{why?}) \\
 = (\gamma_n, \dots, \gamma_1)$$

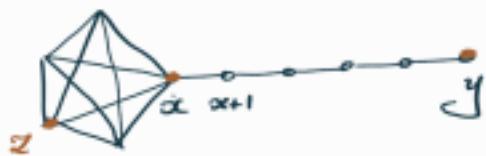
$$\widehat{\gamma}_{\min} = \gamma_n.$$

$$\begin{aligned}
 \phi &= \gamma_n \\
 \tau_e &= O\left(\frac{\log n}{\varepsilon^3 (\gamma_n)^2}\right)
 \end{aligned}$$

$$\tau_e = O\left(\frac{n^2 \log n}{\varepsilon^3}\right)$$

(Can be improved)

3. Lollipop graph



$$\text{Hitting time } h(y, x) \approx \frac{n^2}{4} = O(n^2).$$

$$\text{But } h(x, y) = \Theta(n^3).$$

$$\begin{aligned} h(x, x+1) &= \frac{1}{n/2} \cdot 1 + \left(1 - \frac{1}{n/2}\right) \times \left(\frac{n}{2} + h(x, x+1)\right) \\ &= \frac{2}{n} + \frac{n}{2} - 1 + \left(1 - \frac{2}{n}\right) h(x, x+1) \end{aligned}$$

$$h(x, x+1) \cdot \frac{2}{n} = \frac{2}{n} + \frac{n}{2} - 1$$

$$\begin{aligned} h(x, x+1) &= 1 + \frac{n^2}{4} - \frac{n}{2} \\ &\approx \frac{n^2}{4} \end{aligned}$$

$$\begin{aligned} h(x+i, x+i+1) &= \frac{1}{2} \cdot 1 + \frac{1}{2} (h(x+i-1, x+i+1) + 1) \\ &= 1 + \frac{1}{2} (h(x+i-1, x+i) + h(x+i, x+i+1)) \\ &= 2 + h(x+i-1, x+i). \end{aligned}$$

$$\approx 2i + \frac{n^2}{4}$$

$$\begin{aligned} h(x, y) &= \sum_{i=0}^{n/2-1} (2i + \frac{n^2}{4}) \\ &= \frac{n^2}{4} \cdot \frac{n}{2} + O(n^2) \\ &= \underline{\Theta(n^3)} \end{aligned}$$

$$h(x) = \Theta(n^3).$$

Beyond examples.

Thm 1. (Mean first recurrence theorem) (weighted undirected graphs).

$$h(i,i) = \frac{1}{\nu_i} \quad \forall i \in [n].$$

Proof: $\hat{\nu}_i$ = steady state prob for a random walk to be in mode i
 = proportion of time the walk is in mode i "ergodic property"
 \Rightarrow Expected time b/n visit to $i = \frac{1}{\nu_i}$

$$N \hat{\nu}_i$$

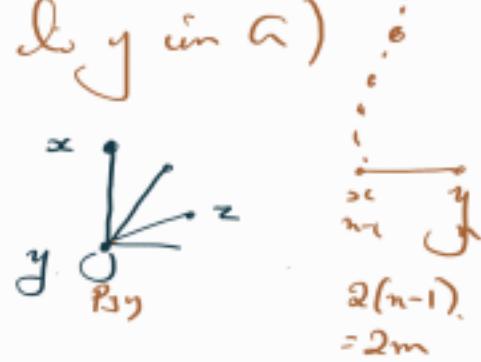
Thm 2. If
 ?? 1. All edge weights of G are 1, and

2. $x \sim y$ in G . (x in adj of y in G)

Then

$$h(x,y) \leq 2m,$$

when $m = |E(G)|$



Proof: $h(y,y) = p_{yy} + \sum_{z \in N(y) \setminus \{y\}} p_{yz} (1 + h(z,y))$

Hence $p_{yz} h(x,y) \leq h(y,y) = \frac{1}{\nu_y}$

$$\begin{aligned} h(x,y) &\leq \frac{1}{\nu_y} \frac{1}{p_{yz}} & (\hat{\nu}_y = \frac{d_y}{2m}) \\ &= \frac{1}{\frac{d_y}{2m} \cdot \frac{1}{p_{yz}}} = \underline{\underline{2m}} \end{aligned}$$

$$\left(\hat{\nu}_y = \frac{d_y}{2m} \right)$$

$$\text{Quiz 3}$$

Corollary: $h(a) \leq n^3$ for every graph
 $h(x, y) \leq 2m \cdot \text{dist}_a(x, y)$. $\forall x, y \in V(a)$

$$h(a) \leq 2m \cdot \text{diam}(a)$$



$$\leq 2mn$$

$$\leq \underline{n^3} // \quad (\text{Remember: Lollipop})$$

Thm 3.

$$c(a) \leq 4mn.$$

Proof. Let $z \in V(a)$ be arb.

Let T be a spanning tree of a rooted at z and let

$$z = x_1, x_2, \dots, x_{2n-1}$$

a DFS traversal of T .

(Each edge of T travelled twice).

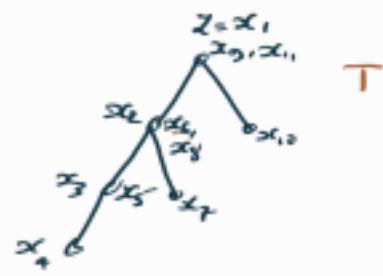
$c(z) \leq$ time to traverse T as

above

$$\leq \sum_{i=1}^{2n-2} d(x_i, x_{i+1}) \quad (\text{Linearity of expectation})$$

$$\leq 2n \cdot 2m$$

$$= \underline{4mn} //$$



Since z was arbitrary,

$$c(a) \leq 4mn.$$

This does not mean $c(a) \leq 2h(a)$.

(e.g.: K_n).

Thm 4.

$$c(a) \leq h(a) \times 2 \log_2 n.$$

Lemma: Let S be any subset of $V(a)$.
 $x \in V(a)$,

$c_{y_2}(x, S)$ = Expected line to hit
 half the vertices of S
 starting from x $\underbrace{h_x}_{R_x}$.

Then $c_{y_2}(x, S) = 2 \max_{y \in S} h(x, y) = 2h_x$

[Hence $c_{y_2}(x, S) \leq 2h(a) \nmid x \in V(a)$.]

Proof:

$$\text{(1)} \quad \sum_{y \in S} h(x, y) \leq |S| h_x. \quad \begin{array}{c} \text{---} \\ x \\ \vdots \\ t = c_{y_2}(x, S) \end{array}$$

$$\text{(2)} \quad \sum_{y \in S} h(x, y) \geq \frac{|S|}{2} c_{y_2}(x, S)$$

$$\text{(1)} + \text{(2)} \Rightarrow c_{y_2}(x, S) \leq 2h_x \quad \blacksquare$$



Recurisvely applying the lemma $\log_2 n$ times to the set of un-hit vertices gives Thm 4.

$$c(a) = h(a) \left(1 + \frac{1}{2} + \frac{1}{4} + \dots + \frac{1}{n} \right). \quad \blacksquare$$

SINGULAR VALUE DECOMPOSITION

LECTURE 21
3/5/2021

SVD

$A_{m \times n}$

Consider m points x_1, \dots, x_n in \mathbb{R}^d ($d \gg 1$)

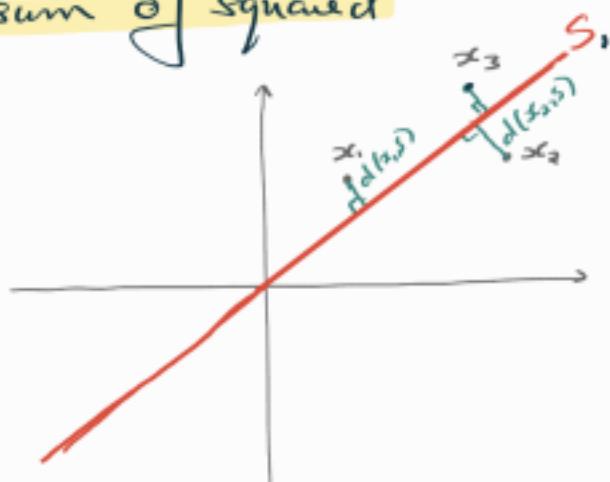
Can we find a lower-dimensional subspace S_k of \mathbb{R}^d s.t. x_1, \dots, x_n can be "approximated well" by points y_1, \dots, y_n in S_k ?

"DIMINISHABILITY REDUCTION"

"Approximated well" can have various interpretations.

SVD - interpretation

Choose the k -dimensional subspace which minimizes the sum of squared Euclidean distances



That is,

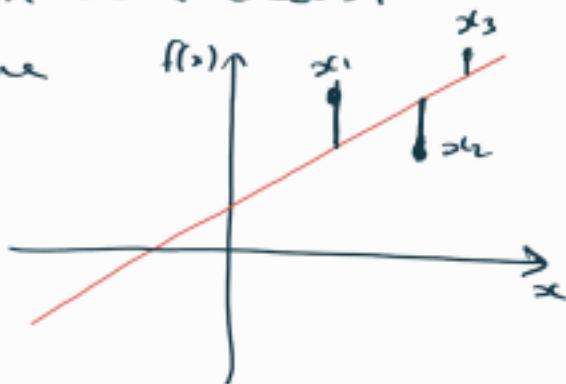
$$S_k = \arg \min_{S \in \mathcal{S}} \sum_{i=1}^n \text{dist}^2(x_i, S),$$

where the minimization is over all k -dimensional subspaces of \mathbb{R}^d

Notes

1. It is different from the least squares regression line

- vertical distance
- not necessarily through origin.



2. Once you find S_k , $\forall i \in [n]$ let

$y_i = \text{Proj}_{S_k}(x_i)$ be the **orthogonal** projection of x_i on S_k .

Then the matrix $B_k = \begin{bmatrix} -y_1 - \\ \vdots \\ -y_n - \end{bmatrix}_{n \times d}^{\leftarrow \text{rank } B_k}$

is the **best** k -rank approximation

for the data matrix $A = \begin{bmatrix} -x_1 - \\ \vdots \\ -x_n - \end{bmatrix}_{n \times d}$

That is

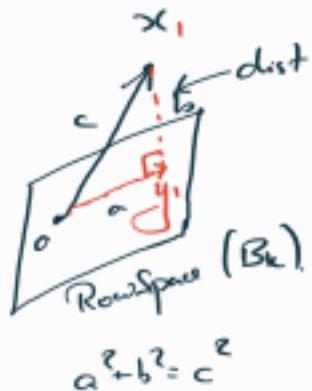
$$B_k = \underset{B \in \mathbb{R}^{n \times d}}{\arg \min} \|A - B\|_F,$$

where the minimization is over all $n \times d$ real matrices with rank $\leq k$.

$$\Gamma \|A\|_F^2 = \sum_{i=1}^m \sum_{j=1}^n a_{ij}^2 = \sum_{i=1}^m \|\text{row}_i(A)\|^2 = \sum_{j=1}^d \|\text{col}_j(A)\|^2$$

Proof Idea

$$\begin{aligned} \|A - B\|_F^2 &= \sum_{i=1}^n \sum_{j=1}^d (x_{ij} - y_{ij})^2 \\ &= \sum_{i=1}^n \|x_i - y_i\|^2 \\ &= \sum_{i=1}^n \text{dist}^2(x, s_i) \end{aligned}$$



But how do we find s_k ?

- Equivalently, we find an **orthonormal basis** $\{v_1, \dots, v_k\}$ for S_k .
- In fact we do more. We find an orthonormal basis $\{v_1, \dots, v_d\}$ of \mathbb{R}^d such that for any $k \in [d]$, $\{v_1, \dots, v_k\}$ gives the best-fit k -dim subspace for the data.
- $\text{dist}^2(x, s_k) + \|y_i\|^2 = \|x_i\|^2$ (Pythagorean)

↑	↑	↑
Minimise	Maximise	Independent of s_k

$$\sum_{i=1}^n \text{dist}^2(x_i, \varphi) + \sum_{i=1}^n \|y_i\|^2 = \sum_{i=1}^n \|\text{proj}_{S_k}(x_i)\|^2$$

Minimise
Maximise
Ind of S_k

$$S_k = \arg \max_S \sum_{i=1}^n \|P_{\text{proj}}(x_i)\|^2$$

$$- \parallel \text{Proj}_{\mathcal{S}}(x_i) \parallel^2 = \parallel \text{Proj}_{V_1}(x_i) \parallel^2 + \dots + \parallel \text{Proj}_{V_k}(x_i) \parallel^2$$

$$\begin{aligned}
 - \sum_{i=1}^n \langle x_i, v_i \rangle^2 &= \| \begin{bmatrix} \langle x_1, v_1 \rangle \\ \vdots \\ \langle x_n, v_n \rangle \end{bmatrix} \|^2 \\
 &= \| \begin{bmatrix} -x_1 \\ \vdots \\ -x_n \end{bmatrix} \begin{bmatrix} v_1 \\ \vdots \\ v_n \end{bmatrix} \|^2 \\
 &= \| A v_i \|^2
 \end{aligned}$$

$$\sum_{i=1}^n \|P_{\perp} j_a(x_i)\|^2 = \underbrace{\|A_{v_1}\|^2 + \dots + \|A_{v_k}\|^2}_{\text{Maximise}} - \textcircled{1}$$

- Suppose $k = 1$

- That is to find the best fit line (through origin) for the data.

$$\text{Then } v_1 = \arg \max_{v \in R^d} \|Av\|, \\ \|v\|=1.$$

(ties can be broken arbitrarily).

and $S_1 = \text{Span}\{f(0)\}$

Further $\max_{\|v\|=1} \|Av\| = \sigma_1(A)$ is called the First Singular Value of A , denoted by $\sigma_1(A)$

Finding v_1 and σ_1 are often from
classic numerical analysis problems.

- Suppose $k=2$.

$$\text{let } v_1 = \arg \max_{\|v\|=1} \|Av\|$$

$$v_2 = \arg \max_{\substack{\|v\|=1, \\ v \perp v_1}} \|Av\|$$

$$V_2 = \text{span}\{Av_1, Av_2\}$$

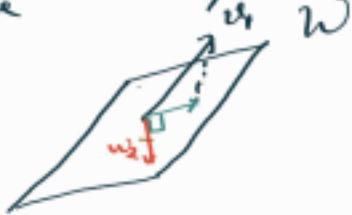
Claim: $S_2 = V_2$. (v_2 is the best-fit)
2D subspace

Proof: W be any 2D subspace.

- Pick $w_2 \in W$ s.t

$$w_2 \perp v_1 \text{ and } \|w_2\| = 1.$$

- Pick $w_1 \in W$ s.t
 $w_1 \perp w_2$ and $\|w_1\| = 1$.



So $\{w_1, w_2\}$ is an ONB for W

$$\|Aw_2\| \leq \max_{\substack{\|v\|=1 \\ v \perp v_1}} \|Av\| = \|Av_2\|.$$

$$\|Aw_1\| \leq \max_{\|v\|=1} \|Av\| = \|Av_1\|.$$

$$\text{So } \|Aw_1\|^2 + \|Aw_2\|^2 \leq \|Av_1\|^2 + \|Av_2\|^2$$

Explain $\{v_1, \dots, v_s\}$

SVD Algorithm.

Input : $A = \begin{bmatrix} \vdots & \vdots \\ -\alpha_1 & - \\ \vdots & \vdots \\ -\alpha_n & - \end{bmatrix}_{n \times d} \in \mathbb{R}^{n \times d}$

Output : An ONB $\{\mathbf{v}_1, \dots, \mathbf{v}_d\}$ of \mathbb{R}^d s.t
 $\forall k, \sum_{i=1}^k \|\mathbf{A}\mathbf{v}_i\|^2 = \max \sum_{i=1}^k \|\mathbf{A}\mathbf{w}_i\|^2,$
 where the maximisation is over all sets
 of k orthonormal vectors $\{\mathbf{w}_1, \dots, \mathbf{w}_k\}$.

$$\mathbf{v}_1 = \underset{\substack{\text{arg max} \\ \|\mathbf{v}\|=1}}{\mathbf{v}} \|\mathbf{A}\mathbf{v}\|, \quad \sigma_1 = \|\mathbf{A}\mathbf{v}_1\|$$

$$\mathbf{v}_2 = \underset{\substack{\text{arg max} \\ \|\mathbf{v}\|=1 \\ \mathbf{v} \perp \mathbf{v}_1}}{\mathbf{v}} \|\mathbf{A}\mathbf{v}\|, \quad \sigma_2 = \|\mathbf{A}\mathbf{v}_2\|$$

$$\mathbf{v}_3 = \underset{\substack{\text{arg max} \\ \|\mathbf{v}\|=1 \\ \mathbf{v} \perp \mathbf{v}_1, \mathbf{v}_2}}{\mathbf{v}} \|\mathbf{A}\mathbf{v}\|, \quad \sigma_3 = \|\mathbf{A}\mathbf{v}_3\|$$

:

$$\mathbf{v}_d = \underset{\substack{\text{arg max} \\ \|\mathbf{v}\|=1 \\ \mathbf{v} \perp \mathbf{v}_1, \dots, \mathbf{v}_{d-1}}}{\mathbf{v}} \|\mathbf{A}\mathbf{v}\|. \quad \sigma_d = \|\mathbf{A}\mathbf{v}_d\|$$

Return $(\{\sigma_1, \dots, \sigma_d\}, \{\mathbf{v}_1, \dots, \mathbf{v}_d\})$.

$\sigma_1, \sigma_2, \dots, \sigma_d$ ↑
 Singular values Singular vectors.

Proof of correctness:

Exercise. Hint : Induction on k .

Easy Observation:

$$\sigma_1^2 + \dots + \sigma_d^2 = \|A\|_F^2$$

Proof: If $k=d$, "Best-fit" = "perfect-fit"

$$\sum_{i=1}^n \text{dist}^2(x_i, \hat{s}) + \sum_{i=1}^n \|y_i\|^2 = \sum_{i=1}^n \|x_i\|^2$$

↓
Minimize Maximize Ind of S

\circ $\underbrace{\sum_{i=1}^n \|A \cdot \hat{w}_i\|^2}_{\text{by } 1} = \sum_{i=1}^d \sigma_i^2$ $\underbrace{\|A\|_F^2}$

SVD (contd...)

MATRIX MULTIPLICATION - Different pictures

$$\text{I} \quad A_{m \times n} \mathbf{x}_{n \times 1} = \mathbf{b}_{m \times 1}$$

$$c_{ij} = \sum_k a_{ik} b_{jk}$$

View 1.

$$\begin{bmatrix} 1 & 1 & \dots & 1 \\ a_1 & a_2 & \dots & a_n \\ 1 & 1 & \dots & 1 \end{bmatrix}_{m \times n} \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix}_{n \times 1}$$



$$\begin{bmatrix} x_1 & x_2 & \dots & x_n \\ \times & \times & \dots & \times \\ 1 & 1 & \dots & 1 \\ a_1 & a_2 & \dots & a_n \\ 1 & 1 & \dots & 1 \end{bmatrix} = \sum_{i=1}^n x_i \text{col}_i(A)$$

↗
 const | vector
 ↙
 lin-comb

View 2

$$\begin{bmatrix} a_1 & \dots \\ \vdots & \ddots \\ a_m & \dots \end{bmatrix}_{m \times n} \mathbf{x}_{n \times 1} = \begin{bmatrix} \langle a_1, \mathbf{x} \rangle \\ \langle a_2, \mathbf{x} \rangle \\ \vdots \\ \langle a_m, \mathbf{x} \rangle \end{bmatrix}$$

$$\text{II} \quad \mathbf{x}_{1 \times m} A_{m \times n} = \mathbf{b}_{1 \times n} \quad [b_1 \dots b_n]$$

- Two views : 1. Linear combination
of rows of A
2. New vector of
inner product of \mathbf{x}
with columns of A .

$$\begin{array}{r} x_1 = a_1 \\ + \\ \vdots \\ + \\ x_n = a_n \end{array}$$

$$\text{III} \quad A_{m \times n} B_{n \times p} = C_{m \times p}$$

View 1.

$$A \begin{bmatrix} 1 & & & \\ b_1 & \dots & b_p \\ 1 & & & \end{bmatrix} \cdot \begin{bmatrix} 1 & & & \\ c_1 & \dots & c_p \\ 1 & & & \end{bmatrix} = \begin{bmatrix} 1 & 1 & 1 \\ Ab_1 & Ab_2 & \dots & Ab_p \\ 1 & 1 & 1 \end{bmatrix}$$

where i-th col of C = $A \times$ i-th col of B
 (2 views)

View 2.

$$i\text{-th row of } C = (i\text{-th row of } A) \times B$$

$$\begin{bmatrix} -a_1- \\ \vdots \\ -a_m- \end{bmatrix} B = \begin{bmatrix} -a_1- \\ \vdots \\ -a_m- \end{bmatrix} = \begin{bmatrix} -a_1 B \\ \vdots \\ -a_m B \end{bmatrix}$$

View 3.

$$\begin{bmatrix} -a_1- \\ \vdots \\ -a_m- \end{bmatrix} \begin{bmatrix} 1 & & & \\ b_1 & \dots & b_p \\ 1 & & & \end{bmatrix} = \begin{bmatrix} \langle ., . \rangle & \dots & \langle ., . \rangle \\ \langle ., . \rangle & \dots & \langle ., . \rangle \end{bmatrix}^C$$

$$c_{ij} = \langle a_i, b_j \rangle$$

$$= \langle i\text{-th row of } A, j\text{-th col of } B \rangle$$

$$c_{ij} = \sum a_{ik} b_{kj}$$

$$\begin{bmatrix} 1 & & & \\ Ab_1 & \dots & Ab_p \\ 1 & & & \end{bmatrix}$$

$$\sum_{i=1}^n \begin{bmatrix} 1 & & & \\ a_i & \dots & a_i \\ 1 & & & \end{bmatrix}$$

View 4.

$$\begin{bmatrix} 1 & & & \\ a_1 & \dots & a_n \\ 1 & & & \end{bmatrix} \begin{bmatrix} 1 & & & \\ -b_1 & \dots & -b_n \\ 1 & & & \end{bmatrix} = \sum_{i=1}^n \begin{bmatrix} 1 \\ a_i \\ 1 \end{bmatrix} [-b_i]_{n \times p}$$

$\underbrace{\qquad\qquad\qquad}_{m \times 1} \underbrace{\qquad\qquad\qquad}_{m \times p}$

PROJECTION MATRIX

Let $\cdot W$ be any k -dim subspace of \mathbb{R}^d
 $(k \leq d)$.

- $\{w_0, \dots, w_k\}$ be any ONB for W .

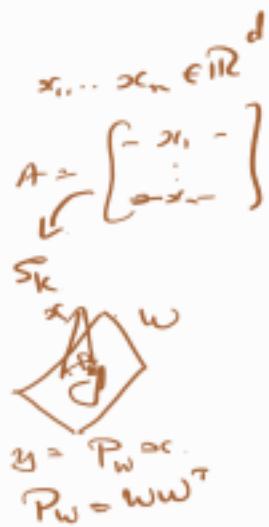
$$\cdot W = \left[\begin{array}{c|c|c|c} 1 & 1 & \dots & 1 \\ w_0 & w_1 & \dots & w_k \end{array} \right]_{d \times k}$$

Claim:

$$\forall x \in \mathbb{R}^d,$$

$$\text{Proj}_W(x) = \underbrace{WW^T}_{P_W, \text{ the projection matrix}} x$$

for W .



Proof: (manipulation)

$$\begin{aligned} \text{Proj}_W(x) &= \underbrace{\langle x, w_0 \rangle}_{\checkmark} w_0 + \dots + \underbrace{\langle x, w_k \rangle}_{\checkmark} w_k. \\ &= \left[\begin{array}{c|c|c|c} 1 & 1 & \dots & 1 \\ w_0 & w_1 & \dots & w_k \end{array} \right] \begin{bmatrix} \langle x, w_0 \rangle \\ \vdots \\ \langle x, w_k \rangle \end{bmatrix} \end{aligned}$$

$$= \left[\begin{array}{c|c|c|c} 1 & 1 & \dots & 1 \\ w_0 & w_1 & \dots & w_k \end{array} \right] \begin{bmatrix} -w_0 \\ \vdots \\ -w_k \end{bmatrix} \begin{bmatrix} 1 \\ x \end{bmatrix}$$

Check
 $W^T W = I$

$$= W W^T x.$$

Corollary: If $\{w_0, \dots, w_d\}$ is an ONB of \mathbb{R}^d
 then $W W^T = I$.

$$\text{Claim 2: } \mathbf{w}\mathbf{w}^T = \sum_{i=1}^k w_i w_i^T$$

Proof:

Standard property of matrix multiplication.

$$\boxed{\begin{aligned} C &= A B \\ p \times q &\quad p \times k \quad k \times q \\ C_{ij} &= \sum_{l=1}^k a_{il} b_{lj} \\ a_{il} b_{lj} &= (i,j)\text{-th entry} \\ &\text{of } \text{col}_l(A) \text{ row}_j(B). \end{aligned}}$$

Combining

Let $\sigma_1, \dots, \sigma_d$ and v_1, \dots, v_d be singular values & ^{sing.} vectors of a matrix $A_{n \times d}$.

Then, $VV^T = I$

$$VV^T = \sum_{i=1}^d v_i v_i^T$$

$$\begin{aligned} A &= AVV^T \\ &= A \sum_{i=1}^d v_i v_i^T \end{aligned}$$

$$= \sum_{i=1}^d (A v_i) v_i^T$$

$$= \sum_{i=1}^d \frac{u_i}{\sigma_i} u_i v_i^T,$$

(Notation
 $A v_i = \sum_{i=1}^d \frac{u_i}{\sigma_i} v_i$)

where $u_i = \frac{A v_i}{\|A v_i\|}$ are called
 the left singular vectors.

Also

$$\begin{aligned} A &= AVV^T \\ &= A \begin{bmatrix} 1 & & 1 \\ u_1 & \dots & u_d \\ 1 & & 1 \end{bmatrix} V^T \\ &= \underbrace{\begin{bmatrix} 1 & & 1 \\ \sigma_1 u_1 & \dots & \sigma_d u_d \\ 1 & & 1 \end{bmatrix}}_{A = \begin{bmatrix} a_1 & \dots & a_n \end{bmatrix}, C = \begin{bmatrix} c_1 a_1 & \dots & c_n a_n \end{bmatrix}} V^T \\ &= \begin{bmatrix} 1 & & 1 \\ u_1 & \dots & u_d \\ 1 & & 1 \end{bmatrix} \underbrace{\begin{bmatrix} \sigma_1 & & \\ & \ddots & \\ & & \sigma_d \end{bmatrix}}_{\Sigma, d \times d} V^T \\ &= U \Sigma V^T \end{aligned}$$

SVD (A)

$$\begin{aligned} A &= U \Sigma V^T \\ &= \sum_{i=1}^d \sigma_i u_i v_i^T \end{aligned}$$

Diagonal matrix
 V : cols form an o.n.b
 u :

L singular vector
left sing. vector
sing. values

$$U = \begin{bmatrix} 1 & & 1 \\ u_1 & \dots & u_d \\ 1 & & 1 \end{bmatrix}_{n \times d} \quad \Sigma = \begin{bmatrix} \sigma_1 & & 0 \\ 0 & \ddots & \\ & & \sigma_d \end{bmatrix}_{d \times d} \quad V = \begin{bmatrix} 1 & & 1 \\ v_1 & \dots & v_d \\ 1 & & 1 \end{bmatrix}_{n \times d}$$

Connection to eigen values ?

$$AA^T = U \Sigma V^T (U \Sigma V^T)^T$$

$$= U \Sigma \underbrace{V^T V}_{I} \Sigma U^T$$

$$= U \Sigma^2 U^T$$

$\|A\|$

"Eigen decomposition" of AA^T
real, sym.,
square.

$\sigma_1^2, \dots, \sigma_d^2$ are eigen values of AA^T