

## TUTORIAL 8

### TEAM MEMBERS:

Devansh Singh Rathore (111701011)

Pellakuru Vamsi(111801033)

ANISH - 111801003

### Q1.

#### Given -

G is strongly connected digraph

P is transition probability matrix of a random walk on G with n vertices

#### To show -

$\pi$  is strictly positive

#### Proof by contradiction -

Let's assume  $\pi$  is not strictly positive.

i.e. there exists  $i \in V(G)$ , s.t.  $\pi_i = 0$ . ----- Statement 1

Since  $\pi$  is a stationary distribution

$$\pi P = \pi$$

$$\text{So, for all } x \in V(G), \sum_{k=1}^n P_{kx} \pi_k = \pi_x$$

$$\text{So, for } x = i, \pi_i = \sum_{k=1}^n P_{ki} \pi_k \text{ -----Equation 1}$$

Since LHS = 0, for RHS to be zero to satisfy the equation,

$\pi_k = 0$  for all  $k$  where  $P_{ki}$  is non-zero. (There exists non-zero  $P_{ki}$  values since the graph is strongly connected) ----- Statement 2

For statement 1 to be true statement 2 needs to be true as well.

Now we use the equation 1 on all the 'k' s in statement 2. -----let us name this process Recursive step 1

We will get the corresponding  $\pi_{k'} = 0$  where  $k'$  are the points which are connected to the points  $k$ . ----- Observation 1

If we repeat the recursive step 1 (at most n times) we will eventually visit all the vertices in the graph G (since G is strongly connected).

And by using observation 1 repeatedly we can say that  $\pi_a = 0$  for all a

But this false because  $\sum_{a=1}^n \pi_a = 1$  since  $\pi$  is a probability vector.

Which means our assumption that  $\pi$  is not strictly positive is wrong.

**Hence the  $\pi$  is strictly positive. Hence Proved.**

## Q2.

**Given -**

X is geometric random variable with mean p i.e.

$$\forall k \in N^+, P[X = k] = (1 - p)^{k-1} p.$$

**To show -**

$$E[X] = 1/p$$

**Proof -**

$$\begin{aligned} E[X] &= \sum_k k \cdot E[X = k] \\ &= 1 \times E[X = 1] + 2 \times E[X = 2] + \dots \\ &= 1 \times (1 - p)^0 p + 2 \times (1 - p)^1 p + \dots \end{aligned} \quad --1$$

$$(1-p) E[X] = 1 \times (1 - p)^1 p + 2 \times (1 - p)^2 p + \dots \quad --2$$

From (1) - (2):

$$E[X] - E[X] + p E[X] = p + p(1 - p) + p(1 - p)^2 + \dots$$

$$\begin{aligned} \text{Therefore, } p E[X] &= p(1 / (1 - (1 - p))) \\ &= p(1/p) \\ &= 1 \end{aligned}$$

Therefore  $E[X] = 1 / p$ .

Hence Proved.

### Q3.

1)

Let suppose it has a stationary distribution  $R$ .

So

$$\begin{aligned} R_j &= R(j-1) \frac{1}{2} + R(j+1) \frac{1}{2} \\ &= (\frac{1}{2})^n (\text{sigma of } n \text{th neighbours of } j) \end{aligned}$$

Since  $\sum R_j$  for  $j$  belong to  $Z = 1$  (prob. vector)

Therefore  $R_j \leq (\frac{1}{2})^n$  and  $n \rightarrow \infty$

So  $R_j \rightarrow 0$  as  $n \rightarrow \infty$ , for every  $j$  in  $Z$ .

Which is not possible as  $\sum R_j = 1$ .

For every  $j$  in  $Z$

### Q3

1) alternate solution

If we expand  $P_{ij}$  we will get an equation like  $\pi_i = \pi_{i+1}$  for all  $i$  from which we can say that  $\pi$  is uniform.

The from that we can say that  $\pi_i = 1/n$  but here  $n$  is infinity.

So  $\pi_i = 0$ . For all  $i$ .

This does not satisfy the condition  $\sum \pi_i = 1$

From this we can say that this random walk has no stationary distribution.