

LOWER BOUNDS on SAMPLE Complexity.

LECTURE 12
28/3/2021

THEOREM.

let \mathcal{H} be any hypothesis class over a domain X with a finite VC-dim d .

For every $\epsilon, \delta \in (0, 1)$, for any sampling distribution D and any true labelling $f: X \rightarrow \{+1, -1\}$,

$$\Pr_{S \sim D^n} [\exists h \in \mathcal{H} : L_S(h, f) = 0 \wedge L_D(h, f) > \epsilon] \leq \delta$$

whenever,

$$n \geq \frac{8d}{\epsilon} \log\left(\frac{8e}{\delta}\right) + \frac{4}{\epsilon^2} \log\left(\frac{2}{\delta}\right).$$

DEFN.

A hypothesis class \mathcal{H} over a domain X is said to be PAC-learnable with a sample complexity $s: (0, 1)^2 \rightarrow N$ if

for any $\epsilon, \delta \in (0, 1)$,

for any prob. dist D on X , and

any true labelling $f: X \rightarrow \{+1, -1\}$,

$$\Pr_{S \sim D^n} [\exists h \in \mathcal{H} : L_S(h, f) = 0 \wedge L_D(h, f) > \epsilon] \leq \delta$$

whenever. $n \geq s(\epsilon, \delta)$.

Note: Some definition insist that $s(\epsilon, \delta)$ is polynomially bounded in $1/\epsilon$ and $1/\delta$.

THEOREM (equivalently)

A hypothesis class H with a finite VC-dimension d is PAC-learnable with sample complexity

$$s(\varepsilon, \delta) \leq \frac{8d}{\varepsilon} \log\left(\frac{8e}{\varepsilon}\right) + \frac{4}{\varepsilon} \log\left(\frac{2}{\delta}\right)$$
$$\in O\left(\frac{1}{\varepsilon^2}(d \log \frac{1}{\varepsilon} + \log \frac{1}{\delta})\right)$$

$\boxed{\begin{array}{l} \text{VC-dim}(H) < \infty \\ \xrightarrow{\quad} \\ H \text{ is PAC} \\ \text{learnable} \end{array}}$

CONVERSE

$$\text{VC-dim}(H) = \infty \stackrel{?}{\Rightarrow} \text{Not PAC-learnable}$$

TIGHTNESS

For VC-dim finite, can we reduce the sample complexity?

$VC\ Dim(\mathcal{H}) = \infty \rightarrow$ NOT PAC learnable

Given $\text{f} \in \mathcal{H}$ s.t. $\forall D, \forall \delta, \forall \epsilon, \exists S \subseteq D$

Ans: Yes
 $VC\ dim(\mathcal{H})$ is infinite
 $\Leftrightarrow \mathcal{H}$ is PAC-learnable.

CLAIM:

$$S_H(y_1, y_2) \geq \frac{1}{2} VC\ Dim(\mathcal{H})$$

Given \mathcal{H} : Any hypothesis class of \mathcal{H}

Given X : Domain of \mathcal{H}

$T = \{x_1, \dots, x_d\} \subseteq X$: A set shattered by \mathcal{H}

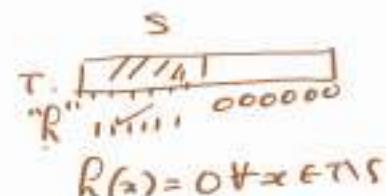
Clever choice:

$$\mathcal{D} : P(x_i) = \begin{cases} 1 & \text{if } x_i \in T \\ 0 & \text{if } x_i \notin T \end{cases}$$

$$\varepsilon, \delta = \frac{1}{2}$$

Let $n < d/2$ and $S \sim \mathcal{D}^n$

\Rightarrow At least half the points in T are not in S Training data



$$\Rightarrow \exists h : \bar{\epsilon}_S(h, f) = 0 \wedge \bar{\epsilon}_{\mathcal{D}}(h, f) > \frac{1}{2}$$

$$\Rightarrow P_{S \sim \mathcal{D}^n} [\exists h : \bar{\epsilon}_S(h, f) = 0 \wedge \bar{\epsilon}_{\mathcal{D}}(h, f) > \varepsilon] = 1$$

$$> \delta = \frac{1}{2}$$

Hence $S_H(y_1, y_2) \geq d/2$.

(Note: S_H is decreasing in ε and δ so it is no better for smaller ε and δ)

Can we have a smaller algorithm?

Let A' be the small alg.

$$R^* = A(S) \quad (\text{Deterministic})$$

"Best choice" among

$$\{R_{\text{best}} : \mathcal{L}_S(R, f) = 0\}$$

We will beat it with same D but
different f 's.

Let $f(x) = \begin{cases} +1, & x \in X \setminus T \\ \{-1, 0\}, & x \in T. \end{cases}$

("Probabilistic Method")
 $f \sim F$

$$n < d/2$$

For any set S of at most n elements
from T (At least $\frac{1}{2}q|T|$ in outside S)
 $|T \setminus S| \geq \frac{1}{2}|T|$

$$\begin{aligned} \mathbb{E}_{f \sim F} [\mathcal{L}_D(A(S), f)] &= \frac{1}{2} |T \setminus S| \cdot \underbrace{|x \in T \setminus S|}_{\text{at least } \frac{1}{2}|T|} \cdot \underbrace{|x \in S|}_{0} \\ &> \frac{1}{4} \end{aligned}$$

Hence $\exists f$ s.t.

$$\forall S \quad \mathcal{L}_D(A(S), f) > \frac{1}{4} - \epsilon$$

$$\Rightarrow P_{S \sim D^n} [\mathcal{L}_D(A(S), f) > \frac{1}{4}] = 1$$

Can a randomised algorithm work?

Ans: No:

Finite VC-dimension

- Smaller Sample Complexity?

Final Answer:

$$S_H(\varepsilon, \delta) = \Omega\left(\frac{d}{\varepsilon} \log \frac{1}{\varepsilon} + \frac{1}{\varepsilon} \log \frac{1}{\delta}\right)$$

so light up to constants.

$\log \frac{1}{\varepsilon}$ term can be removed if we make
"smart algorithms"

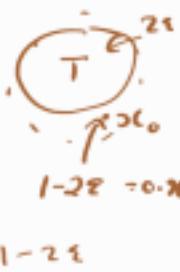
Idea: $S_H(\varepsilon, \delta) \geq \Omega(d/\varepsilon)$.

$$S_H(h, h) \geq d/2$$

$T = \{x_1, \dots, x_d\}$ shattered by H

$$x_0 \in X \setminus T$$

$$\mathcal{D} : P(x) = \begin{cases} 0, & x \in X \setminus (T \cup \{x_0\}) \\ 1-2\varepsilon, & x = x_0 \\ \frac{2\varepsilon}{d}, & x \in T. \end{cases}$$



$$\text{Let } n < \frac{d}{8\varepsilon}$$

$$\mathbb{E}_{S \sim \mathcal{D}^n} [|S \cap T|] = n P_{x \sim \mathcal{D}} \{x \in T\} \\ = n \cdot 2\varepsilon \\ \leq \frac{d}{8\varepsilon} \cdot 2\varepsilon \\ = d/4.$$

"Linearity
of Expectation"

$$x > 0$$

$$P\{X > k \cdot \frac{x}{\varepsilon}\} \leq$$

$$\mathbb{P}_{S \sim \mathcal{D}^n} \left[|S \cap T| \geq \frac{d}{2} \right] \leq \frac{1}{2} \quad (\text{Markov Inequality})$$

$$\mathbb{P}_{S \sim \mathcal{D}^n} \left[|S \cap T| < \frac{d}{2} \right] > \frac{1}{2}$$

$$\left(|S \cap T| < \frac{d}{2} \right) \Rightarrow \exists h : \mathcal{L}_S(h, f) = 0 \wedge \mathcal{L}_D(h, f) > \varepsilon$$

$$\text{Hence } \mathbb{P}_{S \sim \mathcal{D}^n} \left[\exists h : \mathcal{L}_S(h, f) = 0 \wedge \mathcal{L}_D(h, f) > \varepsilon \right] > \frac{1}{2}$$

$$\text{i.e. } S_n(\varepsilon, \gamma_2) \geq d/8\varepsilon \\ = \omega(d/\varepsilon)$$

$$\frac{d}{8\varepsilon} \leq S_n(\varepsilon, \gamma_2) \leq \frac{8d\log(\gamma_2)}{\varepsilon} + \frac{4}{\varepsilon}\log\left(\frac{2}{n}\right) \\ = \frac{8d\log(\gamma_2)}{\varepsilon} + \frac{8}{\varepsilon}$$

\hookrightarrow ≈ 64 -factor gap ")

Rule of thumb : d/ε

Making full use of PAC learning

We proved that

H is PAC-learnable

if and only if

VC-dimension of H is finite

What more can one ask for?

1. Can we tolerate a non-zero in-sample error?

$$\exists L \in H : (\hat{e}_S(L, f) = 0) \wedge (\hat{e}_D(L, f) > \varepsilon)$$

2. Can we use a combination of simple (small VC-dim) classifiers?

3. Anything b/n d and d+1?

4. Boosting

Non-Zero In-Sample Error

Goal

$$\mathbb{P}_{S \sim D^n} \left[\exists h \in H : \underbrace{\left| \hat{E}_D(h, f) - \hat{E}_S(h, f) \right|}_{\text{err}} > \varepsilon \right] \leq \delta$$

$$= \mathbb{P}_{S \sim D^n} \left[\bigcup_{h \in H} C_h \right] \leq \delta, \text{ where}$$

$$C_h = \left| \frac{1}{n} \sum_{x \in S} \mathbb{I}_{h(x) \neq f(x)} \right| > \varepsilon.$$

Step 1. Fix an arbitrary $h \in H$.

Let $\mu = \hat{E}_D[h, f]$ (a deterministic quantity)

$X = \hat{E}_S[h, f]$ (random variable)

$$= \frac{1}{n} \left| \left\{ x \in S : h(x) \neq f(x) \right\} \right|$$

$$= \frac{1}{n} (X_1 + X_2 + \dots + X_n),$$

where $X_i = \begin{cases} 1 & \text{if } h(x_i) \neq f(x_i) \\ 0, o/w \end{cases}$

($S = \{x_1, \dots, x_n\}$)

independent

So X is the average of n Bernoulli random variables.

Now our $\mathbb{E}X = \frac{1}{n} \sum \mathbb{E}X_i$ (linearity of expectation)

$$= \frac{1}{n} \sum_{x_i \sim D} \mathbb{P}[h(x_i) \neq f(x_i)]$$

$$\boxed{\mathbb{E}X = \hat{E}_D(h, f)}$$

$$= \frac{1}{n} \sum \hat{E}_D(h, f)$$

$$= \frac{1}{n} \sum \mu$$

$$= \mu //$$

H : hypo class
X : training set
F : true labelling
D : sampling dist
 $\exists h \in H : \hat{E}_S(h, f) = 0 \wedge \hat{E}_D(h, f) > \varepsilon$.
"mis-leading"

More training data
(S) which permit even one $h \in H$
which is a very different in-sample and out-of-sample case.



If X is a $\{0, 1\}$ -RV,
then

$$\begin{aligned} \mathbb{E}X &= P(X_i = 1) \cdot 1 \\ &\quad + P(X_i = 0) \cdot 0 \\ &= P(X_i = 1) \\ &= P \end{aligned}$$

Hence

$$C_h = |x - \mu| > \varepsilon.$$

C_h : Event that X deviates from its mean by more than ε .

Hoeffding Bounds

Let X_1, X_2, \dots, X_n be independent $\{0, 1\}$ -random variables with $P(X_i=1) = p$.

Let $X = \frac{1}{n}(X_1 + \dots + X_n)$. Then

$$\begin{aligned} EX &= \frac{1}{n} \sum_{i=1}^n E[X_i] \\ &= p. \end{aligned}$$

$$(a) P(X > p + \varepsilon) \leq e^{-2\varepsilon^2 n}$$

$$(b) P(X < p - \varepsilon) \leq e^{-2\varepsilon^2 n}$$

$$\boxed{P\{ |X - p| > \varepsilon \} \leq 2e^{-2\varepsilon^2 n}}$$

Hence $P_{S \sim D^n}[C_h] \leq 2e^{-2\varepsilon^2 n}$ (Exercise: Finite H)

$$\text{Recall: } P_{S \sim D^n}[A_h] \leq (1 - \varepsilon)^n$$

$$\begin{aligned} A_h &: (\bar{\epsilon}_S(h, f) = 0) \wedge \\ &\quad (\bar{\epsilon}_D(h, f) \geq \varepsilon), \\ 1+x &\leq e^x \end{aligned}$$

$$A_h = (\bar{\epsilon}_S(h, f) = 0) \wedge (\bar{\epsilon}_D(h, f) \geq \varepsilon) : B = \bigcup_{h \in H} B_h$$

$$B_h = (\bar{\epsilon}_S(h, f) = 0) \wedge (\bar{\epsilon}_{S'}(h, f) > \varepsilon/2) : P(B_h / A_h) \geq \frac{1}{2}$$

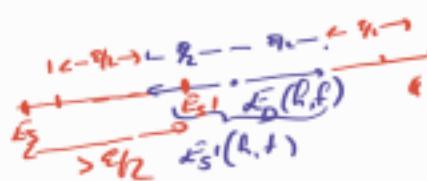
$$C_h = |\bar{\epsilon}_D(h, f) - \bar{\epsilon}_S(h, f)| > \varepsilon : P(B) \leq \delta_h$$

$$D_h = |\bar{\epsilon}_{S'}(h, f) - \bar{\epsilon}_S(h, f)| > \varepsilon/2 : \Rightarrow P(A) \leq \delta.$$

CLAIM: $P(D_h / C_h) \geq \frac{1}{2}$

PROOF: Let $D'_h = |\bar{\epsilon}_{S'}(h, f) - \bar{\epsilon}_D(h, f)| \leq \varepsilon/2$.

Given C_h , $D'_h \Rightarrow D_h / C_h$



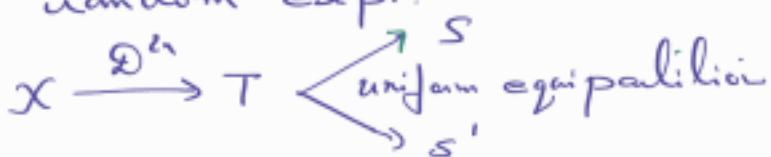
$$\left| \begin{array}{l} \text{if } A \rightarrow B, \\ P(A) = P(B). \end{array} \right. \quad \begin{array}{c} \xrightarrow{\varepsilon} \\ \text{if } s_i > r_i \end{array}$$

$$\begin{aligned}
 P(D_R/C_n) &\geq P(D'_R/C_n) & P\left\{|X - \mu_0| > \frac{\epsilon}{2}\right\} &\leq 2e^{-\frac{\epsilon^2 n}{4}} \\
 &\geq |x - \mu_0| \leq \frac{\epsilon}{2} \\
 &\quad \text{Average of } n \text{ Bernoulli's} \\
 &\geq 1 - 2e^{-\frac{\epsilon^2 n}{2}} \\
 &\geq \frac{1}{2} \quad \therefore \underline{n > 4/\epsilon^2}
 \end{aligned}$$

Hence $P(D/c) \geq \frac{1}{2}$ when $c = 1$ — (1)

$$\mathcal{D} = \bigcup_{L_{FH}} \mathcal{D}_L \text{ and } C = \bigcup_{L_{FH}} C_L$$

New random except



For any h , $P_{S,S \cup D^+} [D_h]$

$$= P_{T \sim D^{2n}_{\text{vec}}(S, S')} [D_L]$$

$$= \mathbb{P}_{S,S'} \left[|E_S(k,f) - E_{S'}(k',f)| > \varepsilon_h \right]$$

$$\leq P_{S,S'} \left[(|\bar{d}_S(h,f) - \mu| > \frac{\epsilon}{4}) \vee (|\bar{d}_{S'}(h,f) - \mu| > \frac{\epsilon}{4}) \right]$$

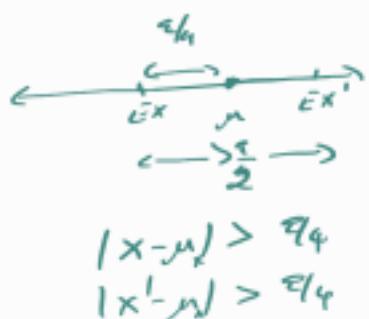
$$\leq 2.2 e^{-2\frac{\varepsilon^2}{16}n} \quad (\text{Hoeffding}).$$

$$= -A e^{-\frac{\varepsilon^2}{2n}} \quad -(*)$$

$$|x - x'| > \epsilon/2$$

$$\mathcal{E}^x = \mathcal{E}^{x'} = \mu$$

$$= \mathcal{E}_p$$



Putting it all together.

$$P_{S \cup D^n} \left[\bigcup_{h \in H} C_h \right] \leq \delta$$

$\uparrow P(D/C) \geq 1/2$

$$P_{S \cup D^n} \left[\bigcup_{h \in H} D_h \right] \leq \delta/2$$

\uparrow Smartest Idea

$$P_{S \cup D^n}^{\text{shattered}} \left[\bigcup_{\substack{[h] \in H \\ T}} D_h \right] \leq \delta/2$$

\uparrow Shattering

$$g_H(2n) P_{S, S'} [D_h] \leq \delta/2$$

$\uparrow (*)$

$$g_H(2n) 4e^{-\frac{\varepsilon^2 n}{8}} \leq \delta/2$$

\uparrow Sauer's Lemma.

$$\binom{2n}{\leq d} e^{-\frac{\varepsilon^2 n}{8}} \leq \delta/8$$

(VCdim(H) = d)

$$n \geq \underbrace{\frac{c}{\varepsilon^2} \left(d \ln(\delta/\varepsilon) + \ln(1/\delta) \right)}_{\text{---}}$$

THEOREM

VC Dimension(H) finite

\Leftrightarrow H is UNIFORM PAC LEARNABLE

(H has Uniform Convergence Property).

Sample Complexity: $O\left(\frac{1}{\varepsilon^2} (d \log \delta/\varepsilon + \log 1/\delta)\right)$

$$C_e : |E_S(h, t) - E_D(h, t)| > \varepsilon$$

error

$$D_h : |E_S(h, t) - E_{S'}(h, t)| > \varepsilon$$

$$R = h \text{ if } h|_T = h'|_T$$

(T = S \cup S')

Concave?

$$\text{VC dim } (H) = \infty$$

$\Rightarrow H$ is not PAC-learnable

$\Rightarrow H$ is not uniform PAC-learnable.

$$(P(A) \leq P(C)).$$

1. Approximation vs. Generalisation Tradeoff

Lecture 14
4/3/2021

We know

$$P_{S \sim D^n} [\exists h : |E_S(h, f) - E_D(h, f)| > \varepsilon]$$

$$= P_{S \sim D^n} [C]$$

$$\leq 2 P_{S, S' \sim D^n} [D]$$

$$\leq 2 g_H(2n) P_{S, S' \sim D^n} [D]$$

$$\leq 2 g_H(2n) 4 e^{-\frac{\varepsilon^2}{8} n}$$

$$\leq 8(2n)^d e^{-\frac{\varepsilon^2}{8} n}$$

$$\begin{aligned} E_S &= 0 \\ |E_D - E_S| &< \varepsilon \\ E_D - E_S &< \varepsilon \\ E_D &< E_S + \varepsilon \end{aligned}$$

Generalization error

— (1)

Earlier we solved for n so that (1) $\leq \delta$.

We can also solve for ε assuming n is given.

$$8(2n)^d e^{-\frac{\varepsilon^2}{8} n} = \delta$$

$$e^{\frac{\varepsilon^2}{8} n} \geq \frac{8(2n)^d}{\delta}$$

$$\varepsilon^2 \geq \frac{8}{n} \left[\ln \left(\frac{8}{\delta} \right) + d \ln 2n \right]$$

$$\varepsilon \geq \sqrt{\frac{8}{n} [d \ln 2n + \ln \frac{8}{\delta}]}$$

$$= \underline{\varepsilon} \approx \tilde{O}(\sqrt{d/n})$$

Hence

$$|E_m - E_D| \leq \underline{\varepsilon} \approx \rho^{1-\delta}$$

(Generalization error)

$$E_D \leq E_{in} + \underbrace{R}_{\uparrow}$$

Approximation Generalisation
 E_{in}

Occam's
Razor"

R increases with increase in d/n ✓

E_{in} is likely to decrease with d/n . ✓

Hence choice of d (i.e. choice of N)
is a **trade off**.

"Amateur ML engineers" tend to choose
a larger d/n so that E_{in} (which is
staring at their face) is low.
"Overfitting"

"Professional ML engineers" are more
careful about R

(Prof. Abu-Mostafa, CALTECH)

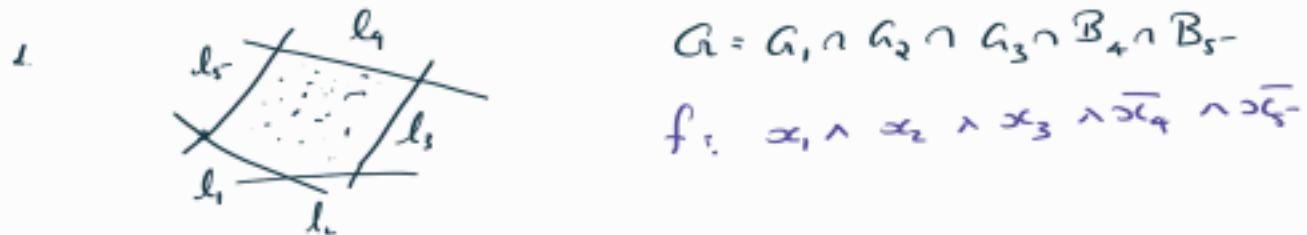
Finding the sweet spot?

- Bias-Variance trade off
- Regularisation
- Validation

(Not in this course).

COMBINING HYPOTHESIS CLASSES

Scenarios



2



"Perception w/o"

Let H be a hypothesis class that is obtained as a combination of k hypothesis classes H_1, \dots, H_k .

$$\text{i.e. } h(x) = f(h_1(x), \dots, h_k(x))$$

where $h_i \in H_i$ and

$f: \{0,1\}^k \rightarrow \{0,1\}$ is any Boolean function.

Then $\text{VC-dim}(H) = ?$

Will depend on f (2^{2^k} possibilities)

An upper bound independent of f ? Yes.

Simplification: $H_i, \text{VC-dim}(H_i) \leq d$.

Claim: $\text{VC-dim}(H) \leq 2kd \ln kd$

PROOF:

Let S be a largest shattered set by H

$$\text{let } |S| = n.$$

$$\Rightarrow |\mathcal{H}|_S = 2^n \quad \text{--- (1)}$$

Consider $H^* = H_1 \times H_2 \times \dots \times H_k$

$$= \{(h_1, \dots, h_k) : h_i \in H_i\}$$

Obst 1. $|\mathcal{H}|_S \leq |\mathcal{H}^*|_S$

Obst 2. $|\mathcal{H}^*|_S = |\mathcal{H}_1|_S \times |\mathcal{H}_2|_S \times \dots \times |\mathcal{H}_k|_S$

$$\begin{aligned} |\mathcal{H}^*|_S &= |\mathcal{H}_1|_S \times \dots \times |\mathcal{H}_k|_S \\ &\leq \binom{n}{d_1} \times \dots \times \binom{n}{d_k} \\ &\leq n^{d_1} \times \dots \times n^{d_k} \\ &\leq n^{kd} \end{aligned}$$

Hence $2^n \leq n^{kd}$

$$n \stackrel{\uparrow}{\leq} 2^{kd \log kd} \quad \square$$

Fact: Perceptron metaclasses (neural networks with hard thresholds) have $\text{VC-dim } O(m \log m)$, where m is the no. of edges.

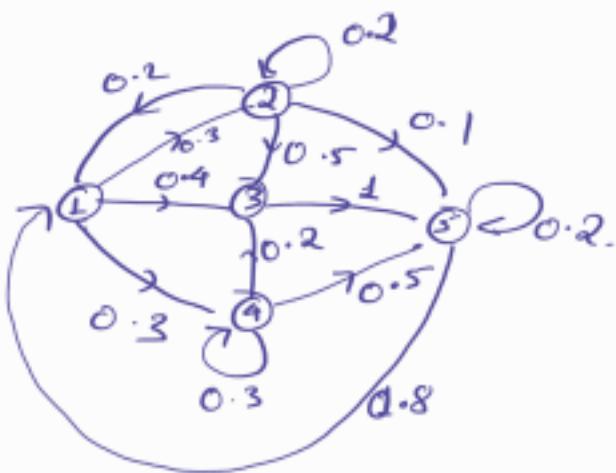
$$(\text{VC-dim}(H) = |S|)$$

$$\begin{aligned} h, h' &\in H \\ h|_S &\neq h'|_S \end{aligned}$$

$$h_{i_1}, f(h_1(x), \dots, h_n(x))$$

$$h'_i(x) = f(h'_1(x), \dots, h'_n(x))$$

$h'_1(x) \neq h_2(x)$ need
at least one $h'_i(x) \neq h_i(x)$



Edge weight = **transition probability**



$$P_{ij} = P[\text{Next State } j / \text{Current State } i]$$

Hence $\forall i$, $\sum_{j \in V(G) \text{ (containing from } i)} P_{ij} = 1$.

P1

$$P = \begin{bmatrix} P_{11} & \dots & P_{1n} \\ P_{21} & \dots & P_{2n} \\ \vdots & & \vdots \\ P_{n1} & \dots & P_{nn} \end{bmatrix}_{n \times n}$$

Transition Probability Matrix (TPM).

"Stochastic Matrix"

$$P_{ij} \geq 0$$

$$\sum_{j=1}^n P_{ij} = 1$$

(row sum)

P1 = Every row of P sums to 1.

$$\begin{bmatrix} P_{11} & \dots & P_{1n} \\ P_{21} & \dots & P_{2n} \\ \vdots & & \vdots \\ P_{n1} & \dots & P_{nn} \end{bmatrix}_{n \times n} \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}_{n \times 1} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}_{n \times 1}$$

Hence $\mathbb{1} = [1, 1, \dots, 1]^T$ is a (right) eigen vector
of P with eigen-value 1.

In fact, If P is a real non-negative
matrix then

$$P \text{ is a stochastic matrix} \iff P\mathbb{1} = \mathbb{1}$$

(P_2)

Corollary: If P_1 & P_2 are non S.M.s,
so is $P_1 P_2$. $P = P_1 P_2$

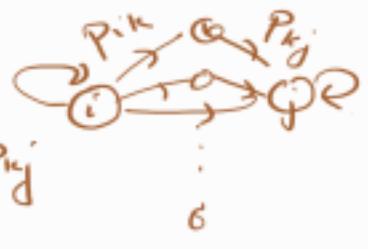
$$\begin{aligned} P_1 P_2 \mathbb{1} &= P_1 (P_2 \mathbb{1}) \\ &= P_1 \mathbb{1} \\ &= \mathbb{1}. \end{aligned} \quad P_{[i,j]} \geq 0$$

Corollary: If P is a S.M.
 $\forall k \in \mathbb{N}$, P^k is a S.M.

$$P^k_{[i,j]} = P[\text{State after } k \text{ steps} = j / \text{Current state} = i]$$

(Proof: Exercise.)

$$\begin{aligned} P[\text{Next state} = j / \text{Current state} = i] &= \sum_{k=1}^n P_{ik} P_{kj} \\ &= P^2_{[i,j]} \end{aligned}$$



Prop 3.

All eigen values of P have magnitude ≤ 1
i.e. $Px = \lambda x \Rightarrow |\lambda| \leq 1$.

Proof: Let $x = (x_1, \dots, x_n)$ be an eigen vector corresponding to λ s.t some $x_i = 1$ and $|P_{ij}| \leq 1 \forall j$ (Scaling).

We have $Px = \lambda x$

$$\langle (\text{eigen value of } P), x \rangle = \lambda x_i$$

$$\sum_{j=1}^n P_{ij} x_j = \lambda 1.$$

$$\lambda = \sum_{j=1}^n P_{ij} x_j$$

$$|\lambda| \leq \sum_{j=1}^n P_{ij} |x_j|$$

$$\leq \sum_{j=1}^n P_{ij} 1.$$

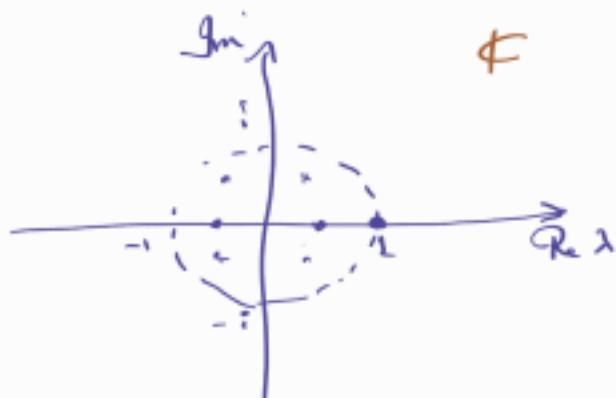
$$= 1.$$

$$x = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix}$$

$$x^T = [x_1 \dots x_n]$$

$$i: \begin{bmatrix} \rightarrow \\ \rightarrow \\ \vdots \\ \rightarrow \end{bmatrix} \begin{bmatrix} \uparrow \\ \uparrow \\ \vdots \\ \uparrow \end{bmatrix} \in \boxed{\mathbb{R}^n}$$

$$\begin{aligned} & \text{--- (i) } |z_1 + z_2| \leq |z_1| + |z_2| \\ & \text{--- (ii) } \begin{array}{l} z_1, z_2 \\ \text{--- } \end{array} \quad \begin{array}{l} z_1 \\ \text{--- } \end{array} \\ & (t_{ij} \text{ if } \Rightarrow |x_j| = 1 \forall j) \\ & \Rightarrow x_j = 1. \\ & \forall j: P_{ij} \geq 0 \end{aligned}$$

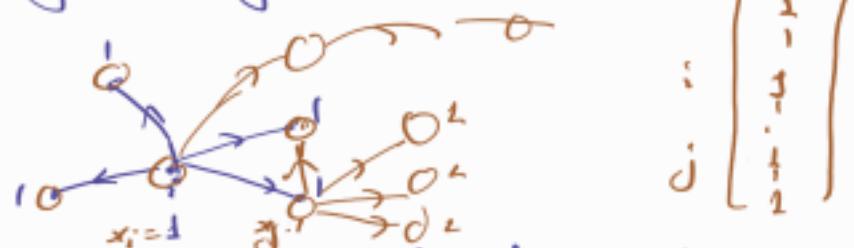


Next Question:

What is the multiplicity of $\lambda=1$?

If $\lambda=1$, then eqn(1) above has to be an equality.

i.e. $x_j = 1$ for all j s.t. $P_{ij} \neq 0$



If every vertex is reachable from i ,
then $\sum x_j = 1 \forall j \Rightarrow x = \begin{pmatrix} 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \end{pmatrix}$

++

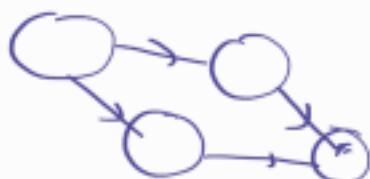
Prop 4.



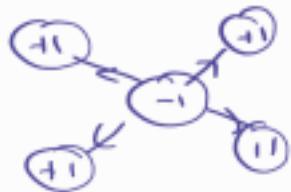
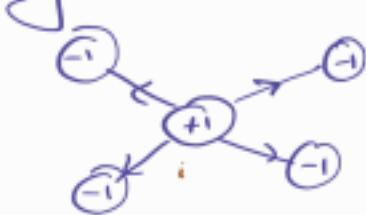
If the underlying graph G is strongly connected then the eigen value 1 has multiplicity 1.

= If 1 is the unique eigen vector (upto scaling) for $\lambda=1$.

Converse? Exercise:
(Hint. Sink component).



When do we get $\lambda = -1$.



iff G is "bipartite".

Strongly connected $\Rightarrow -1$ also has multiplicity 1.

Qn. Is P full rank?

Not necessarily

$$\begin{bmatrix} Y_3 & Y_3 & Y_3 \\ Y_3 & Y_3 & Y_3 \\ Y_3 & Y_3 & Y_3 \end{bmatrix}$$

Interestingly, the $n \times (n+1)$ matrix
 $P - I : 11$ has rank n .
(Exercise.)



Summary:

Let P be the TPM of a random walk on a finite directed graph α .

(1) $\sum_{j=1}^n P_{ij} = 1$ (TPM is a SM)

(2) $P \mathbf{1} = \mathbf{1}$, and hence
 $\lambda = 1$ is an eigen value.

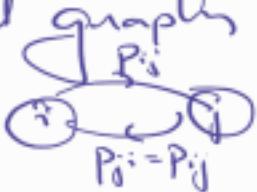
(3) All eigen values of P lie
in the closed unit disc
around 0 .

(4) $\lambda = 1$ has multiplicity 1 iff
 α is strongly connected.

(5) -1 is an eigen value of P
iff α is bipartite.

Note:

All these apply to undirected graphs
as well.



If G is undirected P is symmetric.

\Rightarrow Eigen values are real

$\Rightarrow \lambda \in [-1, 1]$

Now that we have learned so much
about P , let's move on and
study

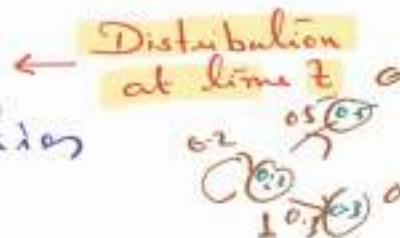
P^T

P^T is more important than P .

(x_1, \dots, x_n) is called a prob. vec if $x_i \geq 0$ & $\sum x_i = 1$

Why?

Let $p = (p_1, \dots, p_n)$ be the node probabilities at current step.

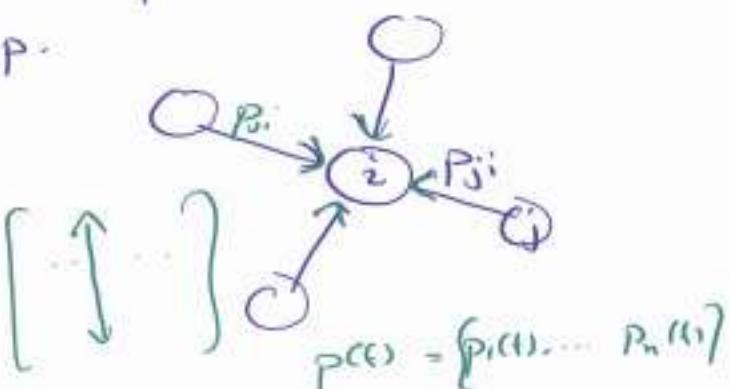


Then

$P^T P$ gives the node probability at the next step.

Distribution at time $t+1$

$$P'_i = \sum_{j=1}^n P_{ji} p_j \\ = [P^T P]_i$$



$$P(t+1) = P^T p(t)$$

(if $p(t)$ is written as a col vector)

$$p(t+1) = p(t) P$$

(if $p(t)$ is written as a row vector)

P^T captures

"Evolution of the random walk".

$$p(0) = [0, 0, 1, 0, 0, 0]$$

$$P^T p(0) = p(1) = \text{stabil. val}$$

$$P^T p(1) = p(2) = \dots$$

* Distribution at time t prob. vec $p(t)$

* Evolution

$$P(t+1) = P^T p(t)$$

Properties of P^T

(1) Eigen values (and multiplicities)
of P and P^T are the same.

Proof: Determinant

$$\det(P - \lambda I) = \det((P - \lambda I)^T)$$

$$= \det(P - \lambda I).$$

$x = (x_1, \dots, x_n) \in \mathbb{R}^n$ is called a probability vector if

$$(i) x_i \geq 0 \quad \forall i$$

$$(ii) \sum_{i=1}^n x_i = 1.$$

(2) x is a prob. vec $\Rightarrow P^T x$ is a prob. vec.

Proof: Let $y = P^T x$.
 $y_i \geq 0$, obvious.

$$\begin{aligned} \sum_{i=1}^n y_i &= \mathbf{1}^T y \\ &= \mathbf{1}^T P^T x \\ &= (\mathbf{P}\mathbf{1})^T x \\ &= \mathbf{1}^T x \\ &= 1. \end{aligned}$$

$$(\mathbf{1}, \dots, \mathbf{1}) \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = 1.$$

If G is strongly connected, then

(3) P^T also has a unique eigen vector
(up to scaling) for $\lambda=1$.

But is it a prob. vec?

$$P^T P = P$$

"stationary distribution"

STATIONARY DISTRIBUTIONS of Random Walks

LECTURE 16
11/4/2021

Defn. $\pi \in \mathbb{R}^n$ is called a stationary distribution of a random walk with transition prob. matrix P

i.e.

- (i) π is a prob. vec., and
- (ii) $P^\top \pi = \pi$.

$$\begin{aligned} P^\top P \\ P^\top P \\ P^\top \pi = \pi \end{aligned}$$

Can we
have π to
be prob. vec?

$$P^\top P = P$$

Theorem 1. (Fundamental Theorem of Finite Markov Chains)

** A random walk on every finite graph has a stationary distbn.

Obs. finiteness is necessary.



(Many important infinite cases can be handled - but not in this course)

$$\begin{aligned} V(a) &= IN \\ L(a) &= \{i, i+1\} : a \in L \end{aligned}$$

$$P^\top P$$

Standard proofs:

Key idea "fixed point theorems"

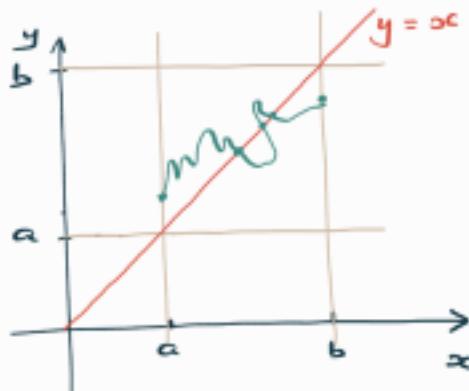
Let $f: X \rightarrow X$, then a point $x \in X$ s.t. $f(x) = x$ is called a fixed point of f .

$$\begin{aligned} \text{Domain } X \\ f: X \rightarrow X \\ X &= \{1, 2, 3\} \\ f(x) &= 4-x \\ \begin{array}{c} 1 \mapsto 3 \\ 2 \mapsto 2 \\ 3 \mapsto 1 \end{array} \end{aligned}$$

Example: Any continuous fn from a closed interval $[a, b]$ to itself has a fixed point.

$X = [a, b]$
 $f: X \rightarrow X$,
 continuous

Let $f: [a, b] \rightarrow [a, b]$, cont



Non-example

$f: [n] \rightarrow [n]$, derangement
 (cyclic shift) -

$1 \mapsto 2$
 $2 \mapsto 3$
 $3 \mapsto 1$

Fact 1. If X is a compact convex set and f is continuous then f has a fixed point (Brouwer's FP theorem, 1909).

$$X \subseteq \mathbb{R}^n$$

Compact =
 Closed & Bound

$[a, b]$ closed
 (a, b) not closed

Fact 2. Closed and bounded sets in \mathbb{R}^n are compact.

What's our X ?

$X = \text{set of all prob. vecs in } \mathbb{R}^n$
 $(n = |V(\mathcal{A})|)$

$$= \{(p_1, \dots, p_n) : p_i \geq 0, \sum p_i = 1\}$$

$f: X \rightarrow X$
 $p \mapsto P^T p$
 $p \in X \Rightarrow P^T p \in X$

Verify : X is ✓ (i), convex
✓ (ii), bounded
✓ (iii), closed.

$$\begin{aligned} \exists x \in X \text{ s.t } \\ f(x) = x, \\ \text{(i)} \quad p, q \in X \\ \lambda p + (1-\lambda)q \in X \\ (\text{Exercise}) \\ \text{(ii)} \quad X \text{ is bounded} \end{aligned}$$



Then $f: X \rightarrow X$
 $x \mapsto P^T x$

Verify - f is continuous

Done! f has a fixed point in X .

That is your stationary distn.

$$\begin{aligned} \|x\|_\infty &:= \max |x_i| \\ \{\|x\|_\infty : x \in X\} &\leq B \\ \|x\|_1 &:= \sum_{i=1}^n |x_i| \\ \|x\|_2 &:= \sqrt{\sum_{i=1}^n x_i^2} \end{aligned}$$

$$\begin{aligned} x, x+\delta x \\ f(x+\delta x) - f(x) &= f(\delta x) \end{aligned}$$

DIRECT PROOF.

Let x be any prob. vec.

Consider the sequence of prob. vecs

$$x = [p_1, p_2, \dots, p_n] \quad x, xP, xP^2, xP^3, \dots$$

$$(x, P^T x, (P^T)^2 x, \dots)$$

(Evolution of the random walk)

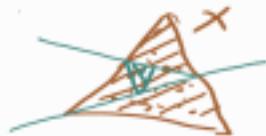
$x(0), x(1), \dots$, where

$$x(t) = xP^t$$

$$\begin{aligned} \text{Let } a(t) &= \frac{1}{t} (x(0) + x(1) + \dots + x(t-1)) \\ &= \frac{1}{t} (x + xP + xP^2 + \dots + xP^{t-1}) \end{aligned}$$

"long-term average"

Claim 1. $\forall t$, $a(t)$ is a prob. vec.



$$\therefore a(1), a(2), \dots \in X$$

X bounded $\Rightarrow a(1), \dots$ contains a convergent subseq
 $\underbrace{a(t_1)}_{3}, \underbrace{a(t_2)}_{7}, \dots$

(Proof: Pigeonhole principle).

$$X \text{ closed} \Rightarrow \lim_{n \rightarrow \infty} a(t_n) = a \in X,$$

(i.e. a is a prob. vec.)

$\forall t$,

$$\begin{aligned} a(t)P - a(t) &= \frac{1}{t} \left(\cancel{\alpha P + \cancel{\alpha P^2} + \dots} - \cancel{\alpha P^{t-1}} \right) \\ &= \frac{1}{t} \left(\cancel{\alpha} - \cancel{\alpha P} \right) \quad (\|P\|_\infty) \\ \|a(t)P - a(t)\|_\infty &\leq \frac{1}{t} \quad (\|\alpha\|_\infty = \max_i |x_i|) \end{aligned}$$

$$\|a(t_n)P - a(t_n)\|_\infty \leq \frac{1}{t_n} \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\Rightarrow \|aP - a\|_\infty = 0$$

$$aP = a$$

Hence a is a stationary distribution.

Summary.

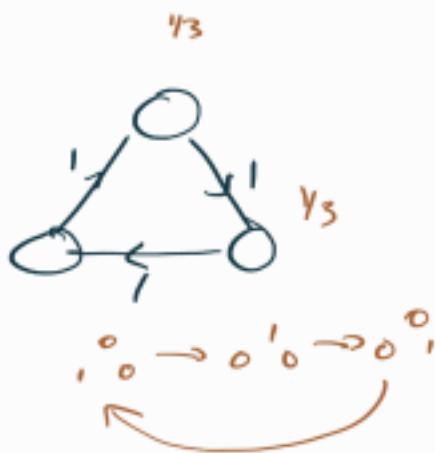
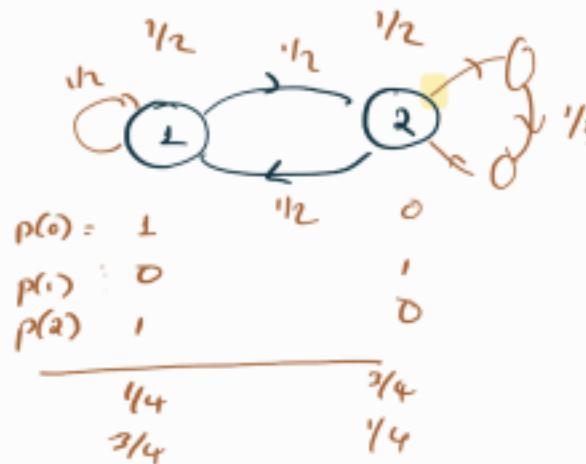
1. Every random walk
on a finite graph } has a stationary distribution π^* ($\pi^*P = \pi^*$)

2. If the graph is strongly connected } the stationary distribution is unique, and $a(t) \rightarrow \pi^*$ as $t \rightarrow \infty$

Qn. If $P(0)$ is an arb. starting distbn does $P(n) = P(0)P^n \rightarrow \pi^*$ as $n \rightarrow \infty$?

Ans: Not always, but
Yes in most cases.

Some no cases.



Yes iff gcd of lengths of all directed cycles in G is 1!

Summary (Once again)

1. If you random walk }
on a finite graph } has a stationary
distribution π
2. If the graph is }
strongly connected } the stationary distribution
is unique, and
 $a(t) \rightarrow \pi$ as $t \rightarrow \infty$
3. If the graph is }
strongly connected and
 $\text{gcd}(\text{cycle lengths}) = 1$ } For any prob. vec $P^{(0)}$
 $P(t) \rightarrow \pi$ as $t \rightarrow \infty$

Applications

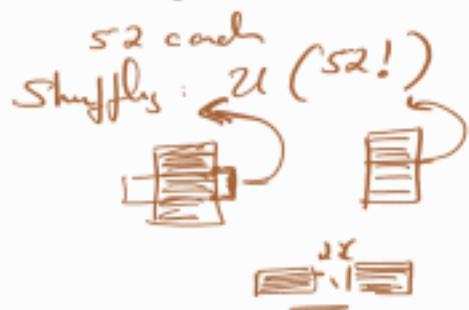
Analysing natural phenomena

- Brownian motion
- Financial market
- Text / Speech
- Genetics
- Web

Generating samples according to some target distribution (e.g.)

- Card shuffling
- Metropolis-Hastings alg.
- Gibbs sampling

$$(p_1, p_2, \dots, p_n) \uparrow \pi_n \quad n = \# \text{web pages}$$



MARKOV CHAIN MONTE CARLO METHODS

LECTURE 17
17/4/2021

Goal: Sample a point x
according to a distribution D on X .
 $x \sim D$

E.g.: σ is a permutation of $[n]$
chosen uniformly at random.

$x \in \mathbb{R}^n$ is sampled from an
n-dimensional gaussian

Key Idea: (π : dist/bn over X)

Let the domain X be finite ($|X| = n$)

Design a directed graph G and
transition probabilities P s.t.
 π is a stationary distribution of P

π will be
a n-length
prob. vec
 $\pi = (\pi_1, \dots, \pi_n)$
 $\Rightarrow \pi_i \geq 0,$
 $\sum_{i=1}^n \pi_i = 1$

Def 6: $P \rightarrow \pi$
Def 7: $\pi \rightarrow P$

Desirable properties:

- (1) G is strongly connected. : unique st. dist
- (2) $\gcd(\text{cycle length}) = 1$. : $\varphi^{(t)} \rightarrow \pi$
- (3) low degrees
- (4) Symmetric (regular for example).
- (5) Rapid mixing (convergence to
stationary distribution).

Easy Case: π is uniform on X . $\pi_i = \frac{1}{n}$

- G : connected non-bipartite undirected graph with $V(G) = X$
- : k -regular ($k > 2$)
- : $P_{ij} = \frac{1}{k} + \delta_{ij}$
- : Expander.



$$P^T = P$$

$\frac{1}{n}$ is eigen-vec of P
& P^T (since $P^T = P$)
 $\therefore (\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n}) = \pi$
is the s.t. dist of P

Non-uniform π

A simple sufficient condition:

If π is a prob. vec. & P is a st. mat. s.t.,

$$\forall i, j: \pi_i P_{ij} = \pi_j P_{ji} \quad (*)$$

$$P^T \pi = \pi$$



then π is a stationary dist'n of P



Proof:

$$\text{let } \omega = P^T \pi$$

$$\stackrel{P^T}{\Rightarrow} \left[\begin{array}{c} \xrightarrow{i} \\ \vdots \\ \xleftarrow{j} \end{array} \right] \left[\begin{array}{c} \downarrow \\ \vdots \\ \uparrow \end{array} \right]$$

$$\begin{aligned} \text{then } \omega_i &= \sum_{j=1}^n P_{ij} \pi_j \\ &= \sum_{j=1}^n P_{ij} \pi_j \quad (\text{by } *) \\ &= \pi_i \end{aligned}$$

□

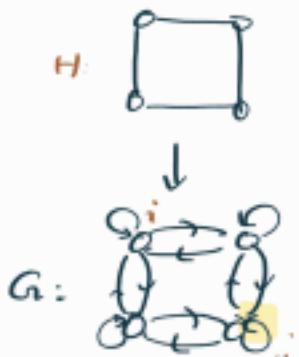
METROPOLIS - HASTING

Input: \hat{v}_i , a prob. dist on $[n]$
 (an n -length prob. vec)

Design of G :

Pick a "good" connected undirected graph H and replace each undirected edge with 2 opposite arcs and add a self loop at each node to get G .

Obs. G is strongly connected,
 $\text{gcd}(\text{cycle lengths}) = 1$.



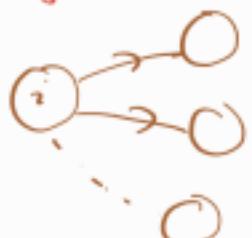
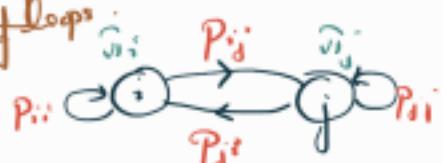
Design of P

$$(\text{dim: } \forall i, j \quad \hat{v}_i P_{ij} = \hat{v}_j P_{ji})$$

Attempt 1.

$$P_{ij} = \hat{v}_i \quad \forall j \in N^+(i)$$

✓ Automatically satisfied
 for missing edges of H
 and self loops



$$\text{Then, } \hat{v}_i P_{ij} = \hat{v}_j \hat{v}_j \quad \checkmark$$

$$\hat{v}_j P_{ji} = \hat{v}_j \hat{v}_i$$

$$\text{But is } \sum_{j=1}^n P_{ij} = 1?$$

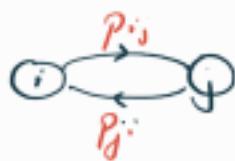
$$\sum_{j=1}^n P_{ij} = \sum_{j \in N^+(i)} \hat{v}_j \ll 1. \quad \text{if } H \text{ is sparse.}$$

We can fix this by changing P_{ii}

$$P_{ij} = \begin{cases} \hat{v}_i & , j \in N^+(i) \setminus \{i\} \\ 1 - \sum_{k \in N^+(i) \setminus \{i\}} \hat{v}_k & , j = i \end{cases}$$

Issue: too slow a walk.

Why not scale?



Subject to

1. equal scaling for P_{ij} and P_{ji}
2. $\sum_j P_{ij} \leq 1$.
(Deficit $\rightarrow P_{ii}$)

$\alpha =$ maximum out-degree of G .
not counting the self-loop.

(2) is ensured if $\forall i, j \quad P_{ij} \leq \frac{1}{\alpha} \quad (i \neq j)$ — *

Scaling for $P_{ij} = \frac{\gamma_n}{\max\{\hat{\gamma}_j, \hat{\gamma}_i\}}$ $\alpha - \beta -$
scaled $\frac{\gamma_n}{\beta}$

Hence

$$\forall j \in \mathcal{N}(i) \setminus \{i\}, \quad P_{ij} = \frac{\frac{\gamma_n}{\max\{\hat{\gamma}_j, \hat{\gamma}_i\}} \times \hat{\gamma}_j}{\frac{1}{\alpha} \times \min\left\{\frac{1}{\hat{\gamma}_j}, \frac{1}{\hat{\gamma}_i}\right\} \times \hat{\gamma}_j}$$

$$= \frac{1}{\alpha} \min\left\{1, \frac{\hat{\gamma}_j}{\hat{\gamma}_i}\right\}$$

Metropolis-Hastings ($\mathcal{Q} \rightarrow H$)

$$\forall i \quad P_{ij} = \begin{cases} \frac{1}{\alpha} \min\left\{1, \frac{\hat{\gamma}_j}{\hat{\gamma}_i}\right\} & , j \in \mathcal{N}(i) \setminus \{i\} \\ 1 - \sum_{j \in \mathcal{N}(i)} P_{ij} & , j = i \end{cases}$$

Sanity check.

$$\begin{aligned}
 \text{Diagram: } & \text{A directed graph with two nodes, } i \text{ and } j. \text{ There is a self-loop on node } i \text{ and a directed edge from } i \text{ to } j. \\
 \widehat{\pi}_i \cdot P_{ij} &= \frac{\widehat{\pi}_j}{\alpha} \min \left\{ 1, \frac{\widehat{\pi}_i}{\alpha t_{ij}} \right\} \\
 &= \frac{1}{\alpha} \min \left\{ \widehat{\pi}_j, \frac{\widehat{\pi}_i}{t_{ij}} \right\} \\
 &= \frac{\widehat{\pi}_i}{\alpha} \min \left\{ 1, \frac{\widehat{\pi}_i}{\widehat{\pi}_j} \right\} \\
 &= \widehat{\pi}_j \cdot P_{di} \quad \checkmark
 \end{aligned}$$

$$\forall i, \sum_{j \in \mathcal{N}(i)} P_{ij} = 1. \quad \checkmark$$

Example: $\widehat{\pi} = (0.5, 0.3, 0.2)$

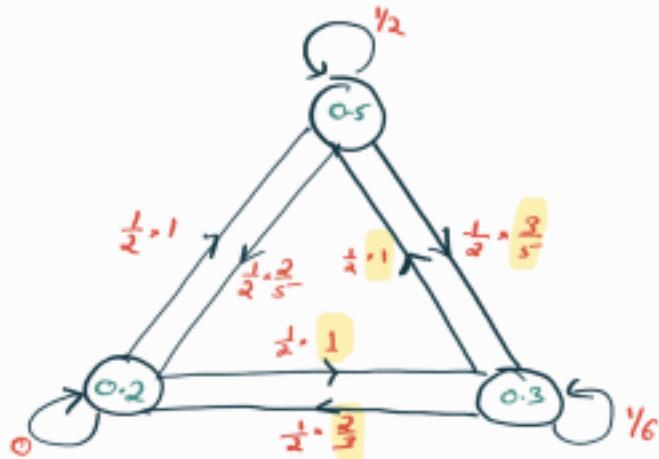
$H:$



$G:$



$n=2$.



Walker's rule

When at node i

(π_{ij})

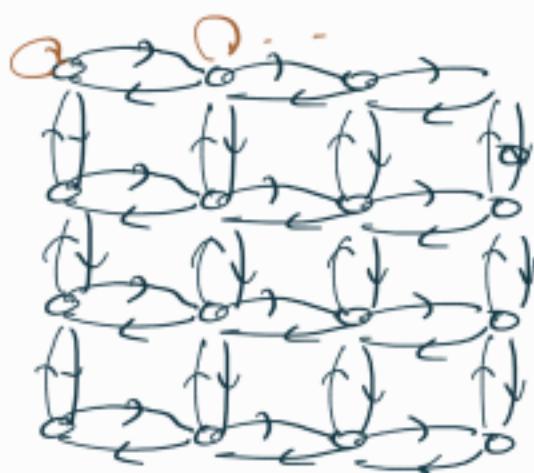
- Scaled each out edge (i, j) w.p. $1/n$

• If $(\widehat{\pi}_j \geq \widehat{\pi}_i)$
move to node j (w.p. 1)

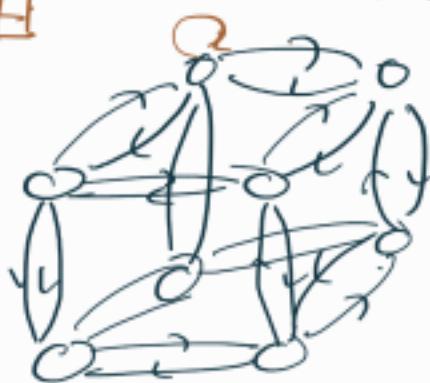
Else move to node j w.p. $\widehat{\pi}_j / \widehat{\pi}_i$

Usually,

Graph : d-dimensional lattice $[m]^d$



$$d=2, m=4$$



$$d=3, m=2$$

- $n = m^d \geq |X|$

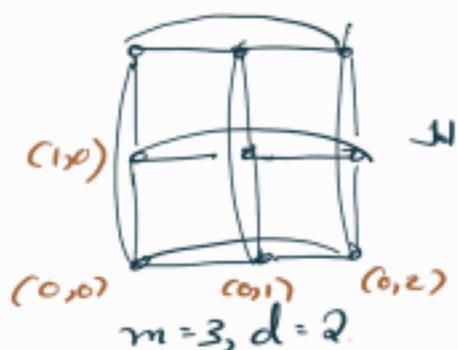
- Almost $2d$ -regular
(except boundary vertices),
 $\alpha = 2d$.

GIBBS SAMPLING

Domain: $X = [m]^d$
 - $\{(x_1, \dots, x_d) : x_i \in [m]\}$

Target: p , a distribution on X

$H = \underbrace{K_m \square K_m \square \dots \square K_m}_{d \text{ times}}$
 (Hamming graph)



$V(H) = X$
 $\{x, y\} \text{ is an edge of } H \iff$
 $x \text{ and } y \text{ differ in exactly one co-ordinate}$

Obs. Much denser than lattice
 $(M \cdot H)$.

Let $x = (x_1, x_2, \dots, x_d)$
 $y = (y_1, x_2, \dots, x_d) \quad , \quad y \neq x,$

(x and y are adjacent in H)

$$P_{xy} = \frac{1}{d} \pi(y/x_1, \dots, x_d)$$

$$\left[\pi(y/x_1, \dots, x_d) = \frac{\pi(y_1, x_2, \dots, x_d)}{\sum_{z=1}^m \pi(z, x_2, \dots, x_d)} \right]$$

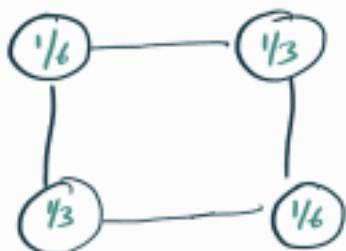
$$\begin{aligned} \text{Hence } \pi(x) P_{xy} &= \pi(y) P_{yx} \quad \forall x, y \\ &= \frac{\pi(x_1, x_2, \dots, x_d) \pi(y_1, x_2, \dots, x_d)}{\sum_{z=1}^m \pi(z, x_2, \dots, x_d)}. \end{aligned}$$

$$\sum_{y \in N^+(x)} p_{xy} = \underbrace{d + d + \dots + d}_{d \text{ times}} - 1.$$

Walker's Rule

- When at node (x_1, \dots, x_d)
 - Pick a dim $k \in [d]$ co.p. γ_d each
 - (Follow attempt 1 on dimension k)
Move to $(x_1, x_{k+1}, \dots, x_d)$
co.p. $\pi(y_k/x_1, \dots, x_{k-1}, x_{k+1}, \dots, x_d)$

Example:



$$y_3 \rightarrow y_6 \equiv y_2 \rightarrow y_3$$

