

Lecture 4

(12 August 2024)

Recap,

Continuity of Probability

$$P\left(\bigcup_{i=1}^{\infty} A_i\right) = \lim_{n \rightarrow \infty} P\left(\bigcup_{i=1}^n A_i\right)$$

Corollaries

$$(i) A_1 \subseteq A_2 \subseteq \dots \Rightarrow$$

$$P\left(\bigcup_{i=1}^{\infty} A_i\right) = \lim_{n \rightarrow \infty} P(A_n)$$

$$(ii) B_1 \supseteq B_2 \supseteq \dots \Rightarrow$$

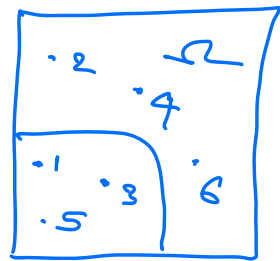
$$P\left(\bigcap_{i=1}^{\infty} B_i\right) = \lim_{n \rightarrow \infty} P(B_n)$$

$$(iii) P\left(\bigcup_{i=1}^{\infty} A_i\right) \leq \sum_{i=1}^{\infty} P(A_i)$$

[An elaborate proof for n sets just to illustrate explicit connections with inclusion-exclusion]

Conditional Probability

Roll a die.



$$P(\text{outcome is 1} \mid \text{outcome is even})$$

$$= 0$$

$$P(A|B) = 0 \text{ if } A \cap B = \emptyset$$

$$P(\text{Outcome} \in \{1, 2\} \mid \text{outcome is even})$$

$$= \frac{1}{3} = \frac{1/6}{2/6} = \frac{P(A \cap B)}{P(B)}$$

$P(A|B)$ is proportional to $P(A \cap B)$.
(for a fixed B)

$$P(A|B) = \alpha P(A \cap B)$$

$$P(B|B) = 1 \text{ intuitively}$$

$$\Rightarrow \alpha = \frac{1}{P(B)}.$$

Definition, The conditional probability of an event A given an event B is defined as

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \quad \text{if } P(B) > 0,$$

Exercise, Show that

$$P_B(A) \triangleq P(A|B) \quad \text{is}$$

a probability law w.r.t. Ω & \mathcal{F} ,
i.e., it satisfies all the three axioms.

Example, A family has two children. Assume that every birth results in a boy with probability $\frac{1}{2}$, independent of other births.

(i) What is the probability that both are boys, given that at least one is a boy?

(i) What is the probability that both are boys given that the younger one is a boy?

$$\Omega = \{BB \underline{B}G \underline{G}B \underline{G}G\}$$

$$P(BB | \text{at one is a boy})$$

$$= P(\{BB\} | \{B\underline{G} \underline{G}\underline{B} BB\})$$

$$= \frac{P(\{BB\} \cap \{B\underline{G} \underline{G}\underline{B} BB\})}{P(\{B\underline{G} \underline{G}\underline{B} BB\})}$$

$$= \frac{P(\{BB\})}{P(\{B\underline{G} \underline{G}\underline{B} BB\})} = \frac{\frac{1}{4}}{\frac{3}{4}} = \frac{1}{3}.$$

$$P(\{BB\} | \text{younger is a boy})$$

$$= \frac{P(\{BB\} \cap \{BB \underline{B}G\})}{P(\{BB \underline{B}G\})} = \frac{1}{2}.$$

Independence

Two events A and B are called independent events if

$$P(A \cap B) = P(A)P(B).$$

Interpretation: $P(A|B) = P(A) > 0$.

Example, Two fair dice are rolled.

$$A = \{\text{Sum is } 7\}, \quad B = \{1^{\text{st}} \text{ roll is } 1\}.$$

A and B are independent

$$C = \{\text{Sum is } 8\}$$

C and B are not independent

Three events A_1, A_2, A_3 are (mutually) independent if

$$P(A_i \cap A_j) = P(A_i)P(A_j) \quad i \neq j$$

$$P(A_1 \cap A_2 \cap A_3) = P(A_1)P(A_2)P(A_3).$$

Events A_1, A_2, \dots, A_n are (mutually) independent if

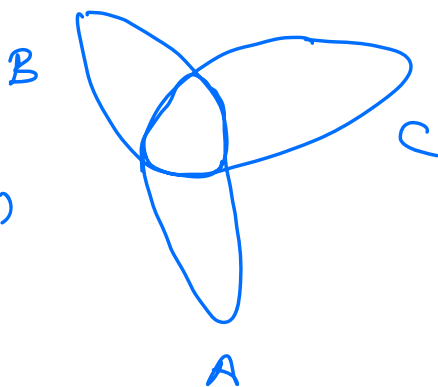
$$P\left(\bigcap_{i \in I} A_i\right) = \prod_{i \in I} P(A_i) \quad I \subseteq [1:n],$$

A_1, A_2, \dots, A_n are pairwise independent if

$$P(A_i \cap A_j) = P(A_i)P(A_j), \quad i \neq j.$$

Pairwise independence does not imply (mutual) independence.

Example.



$$\begin{aligned} P(A) &= P(B) = P(C) \\ &= \frac{1}{5} \end{aligned}$$

$$P(A \cap B) = P(B \cap C) = P(C \cap A) = P(A \cap B \cap C) = p$$

$p = \frac{1}{25}$ pairwise indep. but not mutually independent.

A collection of events A_1, A_2, \dots, A_n is called a partition of Ω if $\bigcup_{i=1}^n A_i = \Omega$ and $A_i \cap A_j = \emptyset$ $i \neq j$.

Total Probability Theorem

Let $\{A_1, A_2, \dots, A_n\}$ be a partition of Ω such that $P(A_i) > 0$ $\forall i \in [1:n]$. Then, for any arbitrary event B ,

$$P(B) = \sum_{i=1}^n P(B|A_i) P(A_i).$$

Proof, $P(B) = P(B \cap \Omega)$

$$= P(B \cap \bigcup_{i=1}^n A_i)$$

$$= P\left(\bigcap_{i=1}^n (B \cup A_i)\right)$$

$$\begin{aligned}
 &= \sum_{i=1}^n P(B \cap A_i) \\
 &\quad \text{(by additivity)} \\
 &= \sum_{i=1}^n P(B|A_i) P(A_i).
 \end{aligned}$$

Bayes' Theorem

Let $\{A_1, A_2, \dots, A_n\}$ be a partition of Ω such that $P(A_i) > 0 \ \forall i \in [1:n]$.
 Then, for any arbitrary event B with $P(B) > 0$,

$$P(A_i|B) = \frac{P(B|A_i) P(A_i)}{\sum_{j=1}^n P(B|A_j) P(A_j)}$$

Proof. $P(A_i|B) = P(A_i \cap B) / P(B)$

$$= P(B|A_i) P(A_i) / \sum_{j=1}^n P(B|A_j) P(A_j)$$

by total probability theorem,

Example, In answering a multiple-choice question, a student either knows the answer or guesses. Let P be the probability that the student knows the answer and $1-P$ be the probability that the student guesses. Assume that the student who guesses at the answer will be correct with probability $1/m$, where $m = \text{no. of multiple-choice options}$. Find

$$P(\text{student knew the answer} \mid \text{they answered correctly}).$$

Proof, $A_1 = \{\text{student knew the answer}\}$

$$A_2 = \{\text{student guesses}\}$$

$$B = \{\text{student answered question correctly}\}$$

$$P(A_1|B) = \frac{P(B|A_1)P(A_1)}{P(B)}$$

$$= \frac{P(B|A_1)P(A_1)}{P(B|A_1)P(A_1) + P(B|A_2)P(A_2)}$$

$$= \frac{1 \cdot p}{1 \cdot p + \frac{1}{m} \cdot (1-p)} = \frac{mp}{mp + 1 - p} \cdot$$