Juiz-1.

Q. F., F2 - CDFS.

Conti & sterethy inc.



+ xe1P

Show that 3 PV X, , X2 with CDF F, & F2

s.t X1 > X2.

 $(F_1 < F_2)$ 

We need to show that 3 x, . x 2. We need not show that

 $F_1(x) \leq F_2(x) \implies x_1 > x_2$ .

(In fact nec can 't show this lecrouse x, can le

less the/equal to x2).

Werong attempt

Contraporitnée of  $x_1 > x_2 : x_1 \leq x_2$ .

2×2 = 23 = 2 x1 = 2

Say Fi is CDF of X1, Fz is CDF of X2.

 $F_2(n) = P(x_2 \le n) \le P(x_1 \le n) = F_1(n)$ 

Contradiction.

This method won't work.

Becomse.

 $x_1 > x_2 \Rightarrow x_1(\omega) > x_2(\omega) + \omega$ .

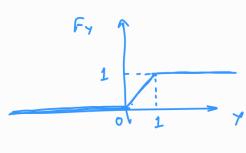
Nigotion of  $X_1(\omega) > X_2(\omega) : X_1(\omega) \leq X_2(\omega)$ 

Now me can't proceed by the above method.

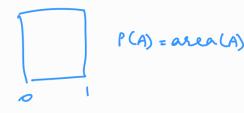
CDF: Montionic func.

CDF -> 1 AN N->00 CDF-> 0 AN M->-00.

Consider a RV 'Y' with CDF



Fx(n) = P(x < x)



Construct X1, X2 as

 $X_1 = F_1^{-1}(Y)$ ,  $X_2 = F_2^{-1}(Y)$ 

Claim: X, is RV whose CDF is F,,

X2 is EV whose CDF in F2 and

 $\times_1 > \times_2$ 

. Inverse is défined beconse

CDF is strictly inc. and continuon

!

One one

Onto

Proof:  $F_{x_i}(x) = P(x_i \le x)$ 

= P(F, -1(y) = x)

=  $P(Y \leq F_1(x))$ 

(: Fi is strictly one. we can write  $F_1^{-1}(y) \le x$  as  $Y \le F_1(x)$ .

 $= F_{\gamma} C F_{i} (\gamma i)$ 

= F<sub>1</sub>(x) when 0 ≤ y ≤ 1, then F<sub>y</sub>(y) = y (See above)

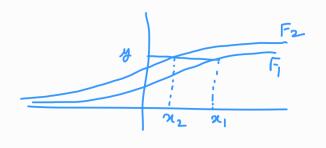
FILM) is CDF, So OFFILM) &1.

So now me've shown CDF of X, is Fi.

Illy we can show CDF of X2 is F2.

Now,  $F_1(x) < F_2(x) + \infty$ .

We need to show  $F_1^{-1}(y) > F_2^{-1}(y)$ ,  $y \in [0,1]$ 



Clearly 
$$x_1 > x_2$$

$$\Rightarrow F_1^{-1}(y) > F_2^{-1}(y)$$

$$\Rightarrow True \forall y \in [0,1]$$

$$\Rightarrow F_1^{-1}(y) > F_2^{-1}(y)$$

$$\Rightarrow x_1 > x_2$$

We know that  $E[X] := \sum_{x} x P_{X}(x)$ 

Here PMF of X is Px.

Now if PMF is Px14=y, then E[x14=y]= \frac{1}{2} n Px14(n/y)

When this conditioning is not there, the

E[(X-E[X|Y=y])2] -> This a valid expectation but this is not the X is not conditioned on Y. var(x|Y=y).

This is satisfied only when VOI (XIY=y) = E[(X-E[XIY=y])2 | Y=y].

NOT satisfied when vor (XIY=y) is taken as )

## Intuition behind law of total variance

We know that 
$$E[X|Y=y] = \begin{cases} xP_{X|Y}(x|y) \\ x \end{cases}$$
  
So  $\varphi(y) = E[X|Y=y]$ 

Here Y is RV.  $\phi(Y)$  is also RV.

$$\phi(Y) = \begin{cases} E[X|Y=y_1] \text{ neith perob. } P_Y(y_1) \\ E[X|Y=y_2] \text{ neith perob. } P_Y(y_2) \end{cases} \times g(X) = \begin{cases} g(x_1) \text{ with } \\ g(x_2) \text{ with } \\ \vdots \\ g(x_n) \text{ with } \end{cases}$$

$$P_{M_2}$$

$$\Psi(Y) = \begin{cases} & \text{val}(X|Y=Y_1) \text{ with perob. } P_Y(y_1) \\ & \vdots \\ & \vdots \end{cases}$$

var (x) = E [var (x/y)] + var (E[x/y])

We say vou (x) has 2 components.

Say we have 3 sets of stretents.

X: Quiz sore of a radon student. | Range (X): n1+n2+n3 volues

To compute var (x), we can compute treating all the students of  $S_1$ ,  $S_2$ ,  $S_3$  in one big set contains all students.

Y: Suction no. of random student. | Range (Y) = 21,2,33.

E[X|Y=1] - Avg. of scores in section 1.

$$E[\times \uparrow Y=2]$$
 ~ " 2

$$E[\times|Y=3]$$
  $\rightarrow$  11 3

Now var (X|Y=1) — By how much the quiz scores in  $S_1$  oberiate from the mean of quiz scores in  $S_1$ .

So Py (y1) var (XIY=1) + Py (y2) var (XIY=2) + Py (y3) var (XIY=3)

[ = E[vor(x 14)] = E[vor(x 14)] Expectation of variances within individual sections.

This does not capture the correlation

across sections, i.e., this doen't capture how quiz mores of S, compares

with scous of S2.

This is eastmed by

→ vac (E[x1Y])

5 Variance of averages between sections

 $E[g(x)|A] = \underset{x}{\text{sig}} g(x) P_{x|A}(x)$ 

E [g(x,y) |A] = Eng(x,y) PxylA (x,y)