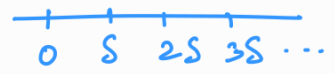


→ Suppose we're tossing a coin at time stamp - $\delta, 2\delta, 3\delta, 4\delta, \dots$.

Given p is the prob. of getting H in a single toss.



Let $X = k\delta$ where k : no. of tosses to get the first heads.

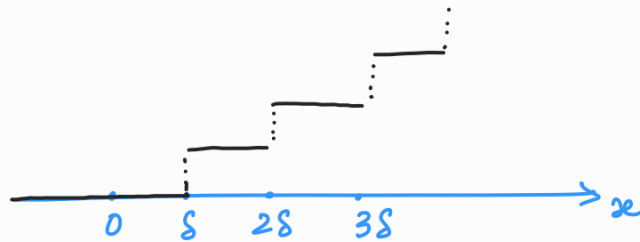
$$P(X = k\delta) = (1-p)^{k-1} p$$

$$F_X(n\delta) = 1 - (1-p)^n, \quad n \in \mathbb{N}.$$

$F_X(x)$ where $x \in \mathbb{R}$?

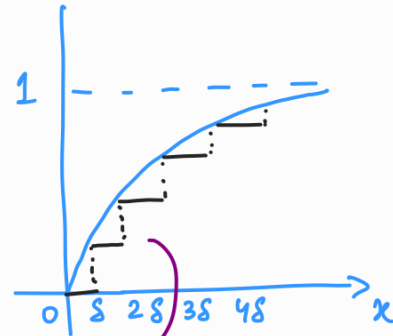
Write ^{this} expression formally $\forall x \in \mathbb{R}$

Right conti:
func



$$\rightarrow f_X(x) = \lambda e^{-\lambda x}, \quad x \geq 0.$$

$$F_X(x) = \int_{-\infty}^x f_X(u) du = 1 - e^{-\lambda x}, \quad x \geq 0.$$



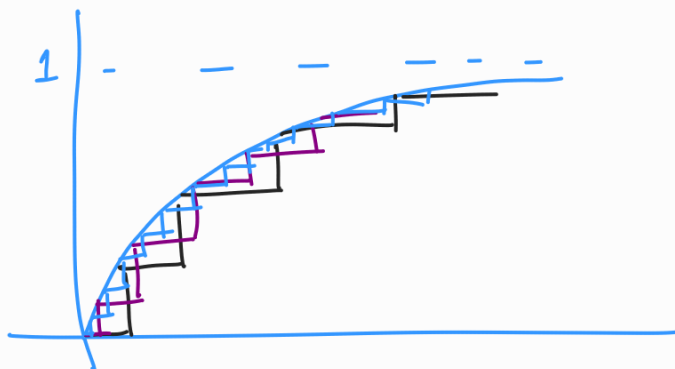
Graph similar to geo cdf

Intersection values: $e^{-\lambda n\delta} = (1-p)^n$

For $\delta > 0$, choose p as \Rightarrow We have p as a func of δ .

→ Starting with exponential RV, constructed geo RV for any δ .

→ If we decrease the value of δ , the plot of ^{CDF of} geo RV approaches exp. RV CDF plot.



\Rightarrow As δ decreases,
geo RV \rightarrow exp. RV

→ Exp. RV is essentially a geo RV where the time gap b/w 2 consecutive losses is infinitesimally small.

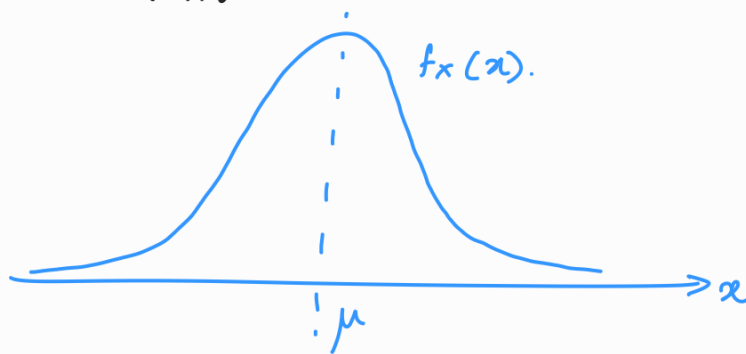
Formally, above can be stated as :

$$\lim_{\delta \rightarrow 0} \underbrace{F_x(x)}_{\text{Geo RV CDF}} = F_x^{\text{Exp}}(x)$$

Exercise.

• Gaussian RV (Normal RV).

$$f_x(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad x \in \mathbb{R}$$



$$\rightarrow \int_{-\infty}^{\infty} f_x(x) dx = 1 \quad \rightarrow \text{Prove it!}$$

$$I = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx.$$

$$t = \frac{x-\mu}{\sigma}.$$

(Change of variable).

$$\Rightarrow dt = \frac{dx}{\sigma}$$

↳ Whenever there is linear scaling of var. of interest

$$\Rightarrow I = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$$

↳ Can't do integration by parts of this. (Try it out).

Consider I^2 instead of I and if we show $I^2=1$ and I has to be ± 1

$$I^2 = \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt}_I \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-s^2/2} ds}_I$$

because $f_x(x)$ is a non-negative quantity.
So $\int f_x(x)$ would be positive.

Use Polar coordinate system

So we just need to show $I^2=1$ and hence we show $I=1$.

→ Sum of sq. in Cartesian → Convert to polar coord.

If we see scalar shift in var. of interest, change of var.

Use polar coord. change of var on $\int_{s=-\infty}^{\infty} \int_{t=-\infty}^{\infty} \frac{1}{2\pi} e^{-(s^2+t^2)/2} ds dt$.

$$s = r \cos \theta, \quad t = r \sin \theta$$



Exercise

$$I = \int_{r=0}^{\infty} \int_{\theta=0}^{2\pi} \frac{1}{2\pi} e^{-r^2/2} r dr d\theta$$

→ Solve this also. Show it's 1.

Mean of Gaussian : μ

Expectation of Gaussian : $E[X] = \int_{-\infty}^{\infty} x f_x(x) dx = \underline{\mu}$.

variance of Gaussian : σ^2

↪ Use int. by parts

$$E[X^2] = \int_{-\infty}^{\infty} x^2 \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{\sigma^2}} dx$$

Change of var (\because scalar shift of x).

$$t = \frac{x-\mu}{\sigma}$$

$$\int_{-\infty}^{\infty} t^2 e^{-t^2/2} dt. \quad \longrightarrow \text{Int. by parts.}$$

We know $\int t e^{-t^2/2} dt$

$$\text{So } t^2 e^{-t^2/2} = t \underbrace{\left(t e^{-t^2/2} \right)}_{\substack{\text{like} \\ \text{split this}}}$$

→ If X is a Gaussian/Normal RV, then we denote it as

$$X \sim N(\underbrace{\mu}_{\text{Mean}}, \underbrace{\sigma^2}_{\text{Variance}})$$

Properties of Gaussian RV

If $X \sim N(\mu, \sigma^2)$, then

(i) $Y = aX + b \sim N(a\mu + b, \sigma^2 a^2), \quad a \neq 0$

$$X \sim f_x$$

$$Y = aX + b \sim f_y.$$

f_y in terms of f_x ?

→ For conti. RV, PDF comes from its CDF. ←

$$F_x(x) = \int_{-\infty}^x f_x(u) du \quad \& \quad f_x(x) = F_x'(x).$$

For conti. RV,
always start
with CDF.
Can't talk
abt PDF
directly.

So first we find F_y and then taking derivative gives f_y .

$$F_y(y) = P(Y \leq y)$$

$$= P(aX + b \leq y) = P(aX \leq y - b)$$

$$= \begin{cases} P\left(X \leq \frac{y-b}{a}\right) & \longrightarrow \text{When } a > 0 \\ P\left(X \geq \frac{y-b}{a}\right) & \longrightarrow \text{When } a < 0. \end{cases}$$

$$= \begin{cases} = F_x\left(\frac{y-b}{a}\right), & a > 0 \\ = 1 - F_x\left(\frac{y-b}{a}\right), & a < 0. \end{cases}$$

$$f_Y(y) = F_Y'(y) = \frac{1}{|a|} f_x\left(\frac{y-b}{a}\right)$$

Notice this is X .

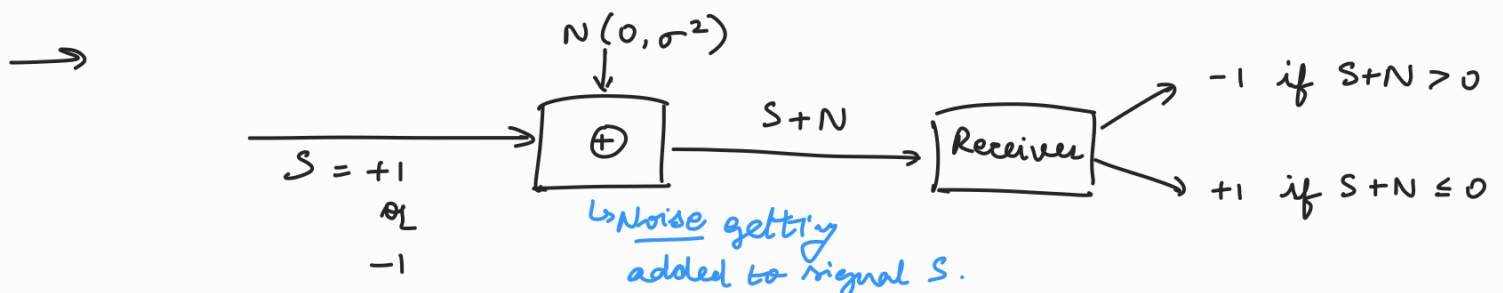
$$\text{Now } f_Y(y) = \frac{1}{\sqrt{2\pi\sigma^2 a^2}} e^{-\left(\frac{y-b}{a} - \mu\right)^2 / 2\sigma^2} \Rightarrow Y \text{ is also Normal.}$$

$$f_Y(y) = \frac{1}{\sqrt{2\pi\sigma^2 a^2}} e^{-\frac{(y - (a\mu + b))^2}{2a^2\sigma^2}}$$

So Mean = $a\mu + b$, Var = $a^2\sigma^2$

• Standard Normal RV : $N(0, 1) = Z$.

$$F_Z(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2/2} dt = \Phi(z)$$



Find the prob. of error when $S = -1$ is sent. : $P(N > 1)$