# Meta-analysis of the Interoceptive Accuracy Scale (IAS) Structure and its Subjective

# **Correlates**

3 Ana Neves

School of Psychology, University of Sussex

5 Author Note

Ana Neves (D) https://orcid.org/0009-0006-0020-7599

This preprint is a non-peer-reviewed work from the **Reality Bending Lab**.



- Author roles were classified using the Contributor Role Taxonomy (CRediT;
- https://credit.niso.org/) as follows: Ana Neves: Project administration, Data curation, Formal
- 11 Analysis, Investigation, Visualization, Writing original draft, Writing review & editing
- 12 Correspondence concerning this article should be addressed to

13 Abstract

- Blabla the abstract blabla.
- 15 Keywords: keyword1, keyword2, keyword3

# Meta-analysis of the Interoceptive Accuracy Scale (IAS) Structure and its Subjective Correlates

16

17

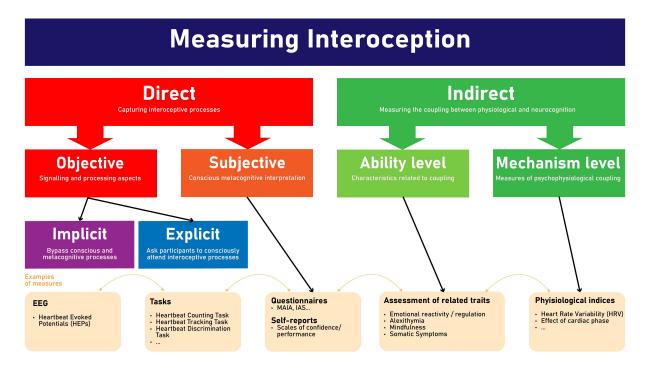
Interoception is referred to the process of sensing, interpreting and integrating information pertaining to internal organs, such as the heart, the lungs or the gut (Khalsa, Adolphs, Cameron, Critchley, Davenport, Feinstein, Feusner, Garfinkel, Lane, Mehling, Meuret, et al., 2018). While recent research emphasizes a key role of interoception in a variety of processes (e.g., emotion regulation, decision making) and of outcomes (physical and psychological well being), the field remains clouded by concerns about how interoception is assessed.

Various measures of interoception have been developed (see Figure 1), forming a combination of "objective" and "subjective" assessments (i.e., physiological tasks such as the heart beat counting or tracking vs. questionnaires and subjective scales involving a metacognitive reflection), "explicit" and "implicit" paradigms (i.e., directing participants' awareness and 10 attention to interoceptive processes vs. measuring interoception unbeknownst to them), various 11 interoceptive modalities (e.g., cardioception, respiroception, gastroception) and theoretical 12 dimensions (e.g., accuracy, sensitivity, awareness). While there is no consensus as to which 13 particular approach provides the most accurate and "pure" measure of interoception and 14 interoceptive abilities (assuming it is a unidimensional construct), it is instead plausible that each 15 measure has strengths and limitations, and a utility dependent on the context and goal at hand 16 (Jahedi & Méndez, 2014). 17

Although the use of subjective self-report questionnaires to measure deeply embodied functions might seem paradoxical at first, recent redefinitions of interoception emphasize the role of high-level and metacognitive elaboration of interoceptive information. These redefinitions provide theoretical grounding to support the idea that some facets of interoception, including participants' metacognitive beliefs, can be assessed subjectively (Khalsa, Adolphs, Cameron, Critchley, Davenport, Feinstein, Feusner, Garfinkel, Lane, Mehling, & others, 2018; Suksasilp & Garfinkel, 2022). Moreover, the notion that self-reports might not reflect the same processes as other interoception tasks might be important to contextualize the apparent lack of convergence

Figure 1

Different ways in which interoception can be measured.



- between measures in the field (Desmedt et al., 2022). For instance, existing findings typically
- show weak or no correlations between questionnaires and objective measures (Brand et al., 2023;
- including for measures of the same theorethical dimensions, e.g., task-based accuracy
- vs. self-reported accuracy, Murphy et al., 2019), such as the Heartbeat Counting Task (HCT,
- Schandry, 1981) and the Heartbeat Detection Task (HDT, Kleckner et al., 2015). Perhaps more
- surprisingly, low correlations have been observed even among questionnaires, suggesting the
- targeting of different facets related to interoception.
- One striking example concerns the assessment of "interoceptive sensibility," defined as the self-reported tendency to focus on and detect internal sensations (Garfinkel et al., 2015), or more
- narrowly as the subjective ability to focus on interoceptive signals (Khalsa, Adolphs, Cameron,
- <sup>36</sup> Critchley, Davenport, Feinstein, Feusner, Garfinkel, Lane, Mehling, & others, 2018). A recent
- 37 systematic review suggested that various questionnaires designed to assess this dimension actually
- measure distinct constructs, leading researchers to mistakenly treat them as equivalent measures

(Desmedt et al., 2022). Notably, sensibility in this review also included self-report scales from Khalsa, Adolphs, Cameron, Critchley, Davenport, Feinstein, Feusner, Garfinkel, Lane, Mehling, and others (2018)'s 8-facet model of interoception, which divides self-report measures into 41 sensibility and self-report scales as the capacity to evaluate one's personal experiences of internal bodily states, assess their consequences, and express them through speech or movement. Some of 43 the studies included the Multidimensional Assessment of Interoceptive Awareness scale (MAIA, Mehling et al., 2012; MAIA-2, Mehling et al., 2018), Body Perception Questionnaire [BPO; Porges (1993)], Private subscale of the Body Consciousness Questionnaire (PBCS, Miller et al., 1981), Body Awareness Questionnaire (BAQ, Shields et al., 1989), and Eating Disorder Inventory 47 (Garner et al., 1983; EDI, Garner, 1991). The inexistent to moderate correlations between these questionnaires underlines the need for a more comprehensive understanding of what is being measured with different questionnaires and dimensions, as well as their potential overlaps with other constructs (e.g., alexithymia, body awareness).

A recently developed scale with a rapidly growing popularity is the Interoceptive 52 Accuracy Scale (IAS, Murphy et al., 2019). The IAS consists of 21 Likert-scale items that query how accurately one can perceive different bodily signals, with one item per physiological modality such as respiration ("I can always accurately perceive when I am breathing fast"), heart ("I can always accurately perceive when my heart is beating fast"), skin ("I can always accurately perceive when something is going to be ticklish"), arousal or bodily functions like coughing ("I can always accurately perceive when I am going to cough") or urinating ("I can always accurately perceive when I need to urinate"). Notably, the IAS' statements are about specific 59 interoceptive behaviours, which is a notable difference with other popular interoception questionnaires, such as the MAIA-2, which contains more general and metacognitive items (e.g., 61 "I trust my body sensations", "I can notice an unpleasant body sensation without worrying about 62 it"), as well as dimensions related to attention regulation (e.g., Not-distracting) or emotion regulation (e.g., Not-worrying).

**note to dom**: review this. 1. yes I think they are general interoceptive signals (I added the

65

items that loaded to both factors for Murphy's and Koike's validation) 2. Do we want to add that lin created a short verison of the scale? 3. Yes kioke did named their factors

Although the original validation study suggested a two-factor structure for the IAS, 68 reflecting the perception of general interoceptive signals (i.e., urinate, hungry, defecate, thirsty, pain, heart, taste, breathing, temperature, mulscles, affective touch, vomit, sexual arousal), and that of signals that may be difficult to perceive solely through interoceptive information (i.e., itch, tickle, cough, burp, bruise, blood sugar, sneeze, wind). The authors however underlined its acceptable but imperfect fit (Murphy et al., 2019, p. 127), and several follow-up studies have indeed identified different optimal solutions. For instance, Brand et al. (2023) and Golino and Epskamp (2017) (the latter using a new approach to structure analysis, Exploratory Graph 75 Analysis), reported a 1-factor solution, while Lin et al. (2023) and Campos et al. (2021) found 76 bifactor solutions - one general factor above a set of lower-level factors (Rodriguez et al., 2016) -77 to be the best fit. Discussions have also been focused on specific items. For instance, Murphy et 78 al. (2019) notes that some items might measure direct interoceptive signals such as cardioception, while others might capture phenomena not perceivable through interoceptive signals alone (e.g., 80 "bruising"; p. 119). Lin et al. (2023) also highlights their correlation analysis, showing five 81 locally dependent pairs and three items (touch, blood sugar, bruise) with exceptionally high difficulty and low discrimination, and Campos et al. (2021) reported "tickle" to be the only item that reflected more specific factors than the general factor. Localization issues also arose in Lin et al. (2023) validation of the IAS, where both "itch" and "tickle" correspond to the same Chinese character, leading to their collapse into a single item. Notably, the only other study to report a 2-factor solution was conducted by Koike and Nomura (2023), who performed an Exploratory Factor Analysis (EFA) assuming 2 factors to align with the findings from the original validation paper. According to the authors, the first factor appears to reflect cutaneous sensations (e.g., itching, tickling, coughing, burping, affective touch, bruising, passing gas, sneezing, muscle sensations, sexual arousal, and taste), while the second factor seems to represent visceral 91 sensations (e.g., urination, defecation, hunger, thirst, pain, breathing, fatigue/blood sugar,

temperature, vomiting, and heartbeat).

118

Note to dom: Ferentzi et al. (2021) suggests that all MAIA dimensions, except for
Not-Worrying and Not-Distracting, relate to a general interoceptive awareness factor. The low
correlations of NW and ND with this general factor indicate that these dimensions do not
contribute significantly to the measurement of general interoceptive awareness. Instead, NW
appears to measure low emotionality or neuroticism, while ND does not capture a coherent
underlying structure.

Regarding its validity, the IAS has naturally been compared to other interoception-related 100 measures, and shows a positive correlations with most facets of the MAIA (Mehling et al., 2018), 101 except for the Not-Distracting and Not-Worrying subscales (Brand et al., 2023) - which were 102 highlighted previously as related to non-interoceptive abilities (Ferentzi et al., 2021) Interestingly, 103 findings on the correlation between the IAS and the body awareness dimension of the Body 104 Perception Questionnaire (BPQ-A, Porges, 1993) have been mixed: some studies report small 105 positive correlations (Brand et al., 2023; Campos et al., 2021; Koike & Nomura, 2023), while 106 others find small negative correlations (Lin et al., 2023) or no correlation at all (Murphy et al., 107 2019). Small positive correlations have also been observed with the "observation" and 108 "description" subscales of the Five Facet Mindfulness Questionnaire (FFMQ, Baer et al., 2006; 109 Brand et al., 2023; Koike & Nomura, 2023), as well as with the "non-reactivity" and "acting with 110 awareness" subscales (Koike & Nomura, 2023). Additionally, the IAS has shown a positive correlation with the interoceptive awareness subscale of the Eating Disorder Inventory (Lin et al., 2023) and a negative correlation with the Interoceptive Confusion Questionnaire (Brand et al., 2023; ICQ, Brewer et al., 2016; Murphy et al., 2019). Lastly, small positive correlations have also 114 been reported with the Interoceptive Attention Scale (Koike & Nomura, 2023; IATS, Lin et al., 115 2023), though studies have also found no correlation between these measures (Gabriele et al., 2022). 117

While assessing the predictive validity of an interoception scale can be conceived as theoretically challenging, expected negative associations were observed between the IAS and

```
alexithymia (Brand et al., 2023; Campos et al., 2021; Koike & Nomura, 2023; Lin et al., 2023;
120
    Murphy et al., 2019), somatic symptoms (Brand et al., 2023; Koike & Nomura, 2023; Lin et al.,
121
    2023), depressive symptoms (Brand et al., 2023; Koike & Nomura, 2023; Lin et al., 2023), anxiety
122
    (Brand et al., 2023), neuroticism (Brand et al., 2023) and self-esteem (Murphy et al., 2019).
123
           TODO: can you update the fig. 1's box (assessmeent of related traits) to remove the
124
    question marks and put the the stronger interoception correlates - as in with what?
125
           Given The current study aims at 1) clarifying the structure of the IAS with a meta-analytic
126
    approach that leverages existing data and contrast the traditional CFA/SEM factor-based analyses
127
    with network-based ones such as EGA; 2) the second part will provide an overview of the
128
    dispositional correlates of the IAS, clarifying the pattern of associations which is key to better
129
    understand the nature, place and role of interoception questionnaires within a larger context.
130
    Study 1
131
           The goal of study 1 is to re-analyse and assess the factor structure of the IAS by taking
132
    advantage of the large number of open-access datasets (Arslanova et al., 2022; Brand et al., 2022;
133
    Brand et al., 2023; Campos et al., 2021; Gaggero et al., 2021; Lin et al., 2023; Murphy et al.,
134
    2019; Todd et al., 2022; Von Mohr et al., 2023). While combining these studies might provide a
135
    more robust and generalizable understanding of the IAS' factor structure, we also additionally
136
    provide an individual analysis (i.e., on all samples separately) to add nuance to the general picture,
137
    as all studies differ in their sample sizes, demographics, language, and procedure.
138
    Methods
139
           Datasets. Our search focused on studies citing the original IAS validation paper (Murphy
140
    et al., 2019), identifying 136 papers (as of 01/05/2024). To qualify for inclusion, papers needed to
141
    (1) provide accessible data in open-access, (2) employ the IAS as a measure, and (3) report
142
    individual IAS items scores. We also included the data of two unpublished (but already
    open-access) studies. A total of 10 studies was included (see Table 1).
           The total number of participants was 32,214 participants (Mean = 48.6 \pm 13.1, 71.6\%
```

Female).

Sample	Subsample	Language	Z	Difference	Age (Mean ± SD)	Range	Female %	Availability
Murphy et al., (2020)								osf.io/3m5nh
	Sample 1	English	451		$25.8 \pm 8.4$	18-69	69.4%	
	Sample 2	English	375		$35.3 \pm 16.9$	18-91	70.1%	
Gaggero et al., (2021)		English and Italian	814		$24.9 \pm 5.3$	18-58	60.3%	osf.io/5x9sg
Campos et al., (2022)		Portuguese	515		$30.7 \pm 10.5$	18-72	59.6%	osf.io/j6ef3
Todd et al., (2022)		English	802		$48.6.6 \pm 14.1$ *	18-92*	*%05	osfio/ms354
Arslanova et al., (2022)		English	143		$28.5 \pm 7.6$	18-73	46.8%	osf.io/mp3cy
Brand et al., (2022)		German	619		$43.9 \pm 14.5$	18-78	78.7%	osf.io/xwz6g
Brand et al., (2023)								osf.io/3f2h6
	Sample 1	German	522		$23.4 \pm 6.7$	18-79	79.5%	
	Sample 2	German	1993		$32.0 \pm 12.6$	16-81	77.7%	
	Sample 3	German	802		$27.3 \pm 9.3$	18-72	96.89	
Lin et al., (2023)								osf.io/3eztd
	Sample 1	Chinese	1166	Collapsed "Itch" and "Tingling"	$32.5 \pm 8.4$	16-60	57.0%	
	Sample 2	Chinese	200	Collapsed "Itch" and "Tingling"	$37.4 \pm 7.4$	20-60	56.2%	
VonMohr et al., (2023)		English	21843		$56.5 \pm 14.4$	18-93	73.2%	osf.io/7p9u5
Makowski et al., (2023a)		English	485	Analog scales	$30.1 \pm 10.1$	18-73	50.3%	github.com/RealityBending/IllusionGameReliability
Makowski et al., (2023b)		English	836	Analog scales	$25.1 \pm 11.3$	17-76	53.0%	github.com/Dominique Makowski/PHQ4R
Makowski et al., (2023c)		English	104	Analog scales	$21.6 \pm 5.0$	18-50	2/9/	github.com/RealityBending/InteroceptionPrimals
Poreiro et al., (2024)		English	107		$26.8 \pm 9.2$	18-57	74.8%	osf.io/49wbv
Poreiro et al., unpublished		English	131		$30.9 \pm 12.0$	18-60	75.9%	
Total			32214		$48.6 \pm 13.1$	17-93	71.6%	

rmation taken from the sample description of relevant paper rather than recomputed.

Statistical Analysis. To examine the factor structure of the IAS, a two-steps approach was employed. First, Exploratory Graph Analysis (EGA), was used to estimate the dimensions via network estimation and community detection, alongside assessing the stability of dimensions and items using the bootstrapping techniques (Golino & Epskamp, 2017). The selection of EGA was motivated by its capability to handle complex, multidimensional data and provide robust dimension estimates. A novel network psychometrics - Unique variable analysis (UVA, Christensen et al., 2023) - approach based on the weighted topological overlap will be computed to evaluate which items have substantial local dependence (> 0.25). Subsequently, exploratory factor analysis (EFA) was employed followed by confirmatory factor analysis (CFA).

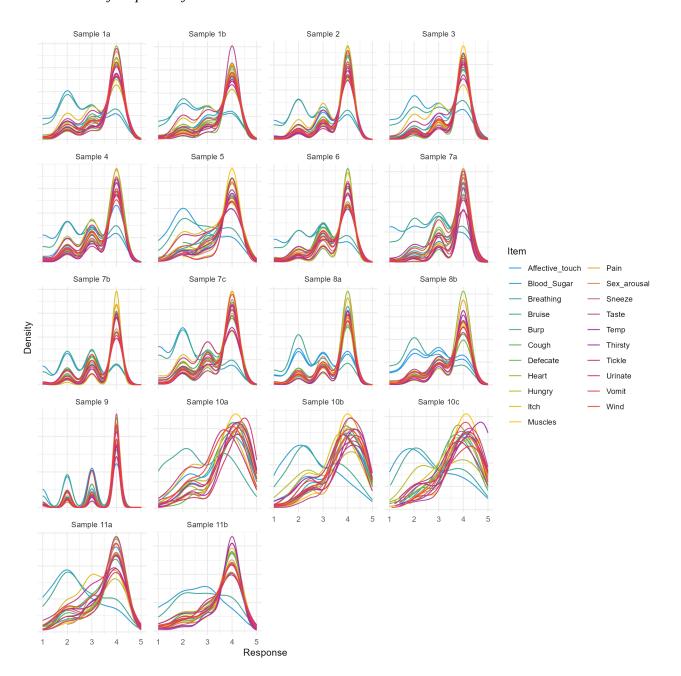
#### Results

Visualizing the distribution of the items for all samples suggests the presence of a consistent modal value (Figure 2). In other words, participants are most likely to answer 4/5 (i.e., agree) on most items (but "affective touch", "blood sugar", and "bruise" that exhibit a different distributional pattern). Additionally, one can note the low density on extreme values (1 and 5), meaning that the bulk of answers (i.e., 99%) varies between 3 values. The interindividual variability seems improved in the samples using an analogue scale, displaying a more continuous and progressive spread of answers.

**Correlations.** The correlation analysis revealed that the items overall have positive intercorrelation patterns with no clear structure emerging. This remains the same across all samples. However, there are possibly some higher-order groupings emerging for the 2 analog-scale samples.

EGA.. The UVA revealed that there are two large to very large redundant variables when taking all samples into account. Namely, "itch" and "tickle", where "tickle" should be removed, and "itch" should be kept. There are several more items that are moderately to largely redundant, namely, "wind" and "burp", and "urinate" and "defecate". On top of that, "sneeze" and "cough", "heart" and "breathing", and "hungry" and "thirsty" seem to have small to moderate redundancy. These findings are rather consistent across the samples with minor differences, such as that when

**Figure 2**Distribution of responses for all items across various datasets.



the questionnaire had an analog scale, there seems to be no large to very large redundant items but 174 "itch" and "tickle" remain moderately to largely redundant and "heart" and "breathing" small to 175 moderately redundant in one sample. 176

According to the network analysis, using the Walktrap and Louvain algorithms applied to 177 Glasso networks, a 4-factor structure fits the questionnaire best across all data sets. This is rather 178 consistent within the data sets, where some samples indicate 3-factor structure, and some a 179 5-factor structure would fit well too. The 4-factor structure model with the best fit entails the 180 following items per group: 1) itch, tickle, bruise, blood sugar; 2) burp, wind, cough, sneeze, 181 vomit; 3) affective touch, sexual arousal, muscles, temperature, pain, and taste; 4) Heart, 182 breathing, hungry, thirsty, urinate, and defecate. 183

Stability analysis, employing 500 bootstrap iterations, also favoured the 4-factor solution for its greater stability. Most items, except for 'affective touch,' demonstrated stability levels exceeding 0.90, indicating structural consistency and reliability (Christensen & Golino, 2021). These findings underscore the robustness of the identified 4-dimensional structure.

184

185

186

194

195

197

When accounting for all samples, the factor analysis reveals that a 4-factor structure fits 188 best. The exploratory factor analysis revealed that 4 latent factors (oblimin rotation) accounted for 189 41.58% of the total variance of the original data (MR1 = 14.45%, MR3 = 11.76%, MR2 = 8.09%, 190 MR4 = 7.28%). Since UVA identified "tickle" as the item to be removed—and it also had the lowest uniqueness value in factor analysis with a similar loading to "itch"—" it was excluded from 192 subsequent analyses.

CFA was computed with the removal of the item "tickle" as it was constantly flagged as redundant. This analysis compared 5 models: a single-factor solution, a 4-factor solution, a 5-factor solution, a 6-factor solution and a 7-factor solution. The latter was preferred in most 196 datasets, including with indices that penalise increased number of parameters (such as BIC). There was no evidence for higher order factors.

## 99 Discussion

214

215

218

219

223

224

In this study, several datasets were analyzed for a meta analysis of the structure of the IAS.

The findings reveal that a 4-factor model fits the IAS best. Additionally, the lowest level structure

(pairs of items) seem to be the most robust, especially for samples using Likert scales (some

higher-order groupings might emerge for the 2 analog-scale samples). There was no clear

evidence for higher-order factors.

These findings contrast with previous research, which all found that 2-factor model (Koike & Nomura, 2023; Murphy et al., 2019), 1-factor model (Brand et al., 2023) and bifactor model (Campos et al., 2021; Lin et al., 2023) fits the data best. While this analysis also revealed an okay fit for the 1-factor model, the 4-factor model was superior. The 4-factor structure reveals different 'hubs' of items that are related, not only in this structure analysis, but also in underlying mechanisms. The 'wind-burp-cough-sneeze-vomit' category, for example, only entails items that are linked to excretion through the mouth. The other categories are organized similarly. This organization and structure is useful for further analysis, as the data can be analyzed and interpreted according to a grouping that is coherent in result, as well as underlying mechanisms.

**note to dom**: what stats do you mean here? the UVA one?

There are several items that show redundancy suggesting that adapting the IAS would be beneficial for validity. Based on the given results, we suggest removing the tickle, while keeping the itch item [todo: stats?]. Other items with slight redundancy were "hungry" and "thirsty", "urinate" and "defecate", and "sneeze" and "cough".

Interestingly, Lin et al. (2023) also found that "tickle" and "itch" were redundant, which led them to exclude one of them. Although, the reason being that the character for both words is the same in the Chinese language. On top of that, they came up with a shortened version of the IAS, excluding further items, resulting in a 12-item IAS, which aligns with our findings, suggesting that further items are ambiguous as to whether they should be removed. Their 12-item IAS included the "hunger", "breath", "urinate", "taste", "vomit", "cough", "temperature", "sexual arousa", "wind", "muscle", "pain", "itch" items.

In contrast, other findings also found "tickle" to be redundant but did not suggest excluding items (Campos et al., 2021).

The findings indicate a high proportion of answers at 4 (see Figure 1), especially when using a 5-step scale. The analogue scale shows a more dispersed distribution, with some answers indicating the highest 5/5, which was not the case in Likert-scales. Therefore, we recommend using an analog scale for the IAS.

Before this paper, the IAS had not yet been used or analyzed with an analog scale, rather than a five-step scale. Therefore, this study provides a novel approach to improving the IAS in a simple manner.

#### Limitations and Future Directions

There are several limitations to the IAS. There are some redundant items, the 5-point scale does not provide great variability, and the structure could be improved. Therefore, improving the IAS, or creating a new questionnaire investigating interoception could be useful to achieving reliable and accurate indication of interoceptive awareness.

## 240 Study 2

232

233

236

237

238

241

242

245

248

Study 2 aims to investigate correlates of the IAS. Correlations of the IAS will be computed to assess the relationship between subjective interoceptive accuracy and other subjective measures of interoception, mood, psychopathology, personality, and beliefs. Investigating correlates will help validate the IAS, as well as other interoceptive measures in the future.

#### Methods

Materials. The questionnaires used for the IAS correlates are listed in Table 2 (TODO:

add the rest of the questionnaires, sample items and references).

**Statistical analysis.** Correlations will be computed using the correlation package under a Bayesian framework (Ben-Shachar et al., 2020).

Questionnaire	Number of Dimensions	Assessment	Number of Items	Scoring
Interoceptive Related				
MAIA-2	8	Interoception	37	6-point Likert scale
BPQ	2	Body awareness and autonomic reactivity	49	5-point Likert scale
TAS-20	3	Alexithymia	20	5-point Likert scale
BVAQ	2	Alexithymia	40	5-point Likert scale
Mood				
BDI-II	1	Severity of depressive symptoms	21	0-to-3-point values
PHQ-4	2	Anxiety and depressive symptoms	4	5-point Likert scale
STAI-T	1	Trait anxiety	20	4-point Likert scale
GAD-2	1	General Anxiety	2	4-point Likert scale
Personality				
NEO-FFI	1	Neuroticism	12	5-point Likert scale
Mini IPIP6	6	Personality	24	Analogue scales
BFI	5	Personality	44	5-point Likert scale
PID-5-SF	5	Dysfunctional personality traits	25	4-point Likert scale
Psychopathology				
SPQ-BRU	4	Schizotypy	32	5-point Likert scale
MSI-BPD	1	Borderline personality disorder	10	Analogue scales
ASQ - Short	5	Autistic Traits	28	4-point Likert scale
Beliefs and Misbeliefs				
GCB	5	Conspiracy beliefs	15	5-point Likert scale
PI-18	1	Beliefs about the world	99	6-point Likert scale
LIE scale	4	Lying tendencies	16	Visual analogue scales

## Results

# Average correlations

The EGA components captured grouping of items such as 'wind' and 'burp, 'cough' and 'sneeze', 'muscle' and 'pain'. These groupings were used in correlational analysis to analyse how much each pairing is associated with other factor such as Alexithymia and with Mood disorders (see figure 2).

notecompute internal consistency

Correlations with body measures. Alexithymia was assessed in the samples with the Bermond–Vorst Alexithymia Questionnaire [BVAQ; Vorst and Bermond (2001)] and the Toronto Alexithymia Scale[TAS; Bagby et al. (1994)].

The BVAQ consists of 5 subscales - fantasising, idenitying, analysing; verbalising and emotionalising - assessed with 40 items on a 5-point Likert scale, from 'defnitely applies to me' to

'in no way applies to me'. Additionally, the BVAQ reduces these subscales into two high order factors, an affective component and a cognitive one, with high scores being indicative of high proneness to alexithymia.

On average, the cognitive component of the BVAQ was weakly and negatively correlated with all IAS pairs of items with the biggest correlation being with the Itch/Bruise pair (r = -0.112) and the lowest correlation beeing with the Muscle/Pain pair (r = -0.244). The affective component of the BVAQ was positively, but very weak, correlated with all pairs, with the biggest correlation being with the Itch/Bruise pair (r = 0.107). The only exception was a negative correlation with the Urinate/Defecate pair (r = -0.036).

The TAS contains 20-items rated on a 5-point forced scale, from 'strongly disagree' to 'strongly afree', divided into 3 dimensions - difficulty identitying feelings, difficulty describing feelings, and externally thinking. High scores on this scale also reflect higher alexithymia.

All the three dimensions assessed with the TAS were on average negatively correlated with all pairs of IAS items. The difficulty describing feelings had its strongest correlation with Hungry/Thirsty (r = -0.179) and weakest with the Wind/Burp (r = -0.117). while, the difficulty describing feelings had its strongest correlation with Muscle/Pain (r = -0.247) and weakest with Itch/Bruise (r = -0.157). Lastly, the external thinking dimension was more correlated with the Cough/Sneeze pair (r = -0.138) and less correlated with the Hungry/Thirsty (r = -0.018).

The studies within our sample used the Body Perception Questionnaire short-form (BPQ-SF) and the very-short form (BPQ-VSF) to assess interoception (Cabrera et al., 2018). The BPQ-SF comprised of 46 items on a 5-point Likert scale assessing body awareness (26 items) and autonomic reactivity (21 items). The BPQ-VSF comprises of 12 items from the body awareness subscale of the BPQ-SF. In this study, all scores assessing these two dimensions were grouped together, hence no disitintion is made between awareness measured with the BPQ-SF and the BPQ-VSF, or eith scores obtained only using the awareness subscale.

note to add: discuss later

In general, all pairs of the IAS were positively, and weakily, associated with the body

<sup>289</sup> awareness subscales, while negative and weakily correlated with the autonomic reactivity
<sup>290</sup> subscale. The strongest correlation identified between the IAS pairs and the body awereness
<sup>291</sup> subscale was with the

Heart/Breathing pair (r = 0.151) whilst the strongest correlation with the autonomic reactivity was with the Urinate/Defecate pair (r = -0.235). The weakest correlation between the body awareness and the IAS was with the Hungry/Thirsty pair (r = 0.055) and between the autonomic reactivity and the IAS was with the Heart/Breathing pair (r = -0.106).

The MAIA was one of the most commonly used measures of interoception in our study,
with nine samples reporting its use. This 37-item questionnaire assesses eight state-trait
dimensions of interoception: Noticing, Not-Distracting, Not-Worrying, Attention Regulation,
Emotional Awareness, Self-Regulation, Body Listening, and Trust. Responses are rated on a scale
from 0 (Never) to 5 (Always).

On average, all MAIA dimensions were positively and weakly to moderately correlated with IAS pairings. Notably, the strongest correlations were observed between the Noticing dimension and the Heart/Breathing pairing (r = 0.394), Trusting and Hungry/Thirsty (r = 0.347), and Attention Regulation and Heart/Breathing (r = 0.334). The Not-Distracting and Not-Worrying subscales were generally positively correlated with IAS pairings, with a few exceptions:

Not-Distracting showed minimal correlation with Cough/Sneeze (r = 0.0206) and Heart/Breathing (r = -0.007), while Not-Worrying had a low correlation with Itch/Bruise (r = 0.031).

The Interoceptive Confusion Questionnaire (ICQ; Brewer et al. (2016)) was used to assess individuals' difficulties in interpreting non-affective physiological states, such as pain and hunger.

The ICQ consists of 20 items rated on a scale from 1 ("Does not describe me") to 5 ("Describes me very well"), with higher scores indicating greater interoceptive confusion.

The ICQ showed weak to moderate negative correlations with all IAS pairings. The strongest correlation was observed with the Hunger/Thirsty pairing (r = 0.348), while the weakest was with the Itch/Bruise pairing (r = 0.207).

**Correlations with mood measures.** Mood disorders were assessed using several 315 standardized measures, including the General Anxiety Disorder-2 [GAD-2; Kroenke et al. 316 (2007)], the State-Trait Anxiety Inventory [STAI; Spielberger (1970)] and its shorter version, the 317 STAI-5 (Zsido et al., 2020), Beck's Depression Inventory [BDI; Beck et al. (1996)], and the 318 Mood and Feelings Questionnaire [MFQ; Messer et al. (1995)]. Additionally, the Patient Health 319 Questionnaire (PHQ) was administered in its 2-item [PHQ-2; Kroenke et al. (2003)], 9-item 320 [PHQ-9; Kroenke et al. (2001)], and 15-item [PHQ-15; Kroenke et al. (2002)] versions. Finally, 321 borderline personality traits were assessed using the McLean Screening Instrument for Borderline 322 Personality Disorder [MSI-BPD; Zanarini (2003)]. 323 The GAD-2, a brief screening tool for generalized anxiety disorder, consists of two items 324 rated on a scale from 0 (not at all) to 3 (nearly every day). The STAI, a 40-item questionnaire 325 rated on a 4-point Likert scale (0 to 3), measures both state and trait anxiety. However, in our 326 study, most participants primarily completed the trait anxiety subscale. In some samples, a shorter 327 5-item version (STAI-5) was used to assess both state and trait anxiety. 328 On average, anxiety measures showed weak negative correlations with all IAS pairs. 329 Notably, the strongest correlations between the IAS pairings and the GAD-2, STAI-T and STAIT-5 330 were observed with the Hungry/Thirsty pair (r = -0.168, r = -0.270 and r = -0.248, respectively). The BDI consists of 21 items measuring the severity of depressive symptoms on a scale 332 from 0 to 3. The total score is calculated by summing the highest responses, which are then 333 compared to six depression severity levels, ranging from 1–10 (normal fluctuations in mood) to 334 over 40 (extreme depression). The PHQ-2 includes two items assessing the frequency of 335 depressive symptoms and anhedonia. The PHQ-2 is derived from the PHQ-9, a nine-item 336 screening tool used to assess depression severity and monitor treatment response. Both 337 questionnaires are measured on on a scale from 0 (not at all) to 3 (nearly every day) 338 Depression measures showed weak to moderate negative correlations with IAS pairings. 339 The BDI (r = -0.372), PHQ-2 (r = -0.148), and PHQ-9 (r = -0.241) correlated most with the 340 Hungry/Thirsty pair, while the MFQ correlated most with Heart/Breathing (r = -0.345) pair.

The PHQ-15 is a 15-item questionnaire that assesses somatic symptoms on a 3-point scale 342 (e.g., back pain). It exhibited its strongest correlation with the Hungry/Thirsty pair (r = -0.241) 343 and, on average, showed weak negative correlations with all other IAS pairings. 344

345

346

347

348

349

350

351

352

353

354

355

356

357

358

361

364

365

366

367

368

Lastly, the MSI-BPD is a 10-item questionnaire used to assess personality disorder, where items are rated on a dichotomous scale of 1 (present) and 0 (absent). The MSI-BPD also showed its strongest negative correlation with the Hungry/Thirsty pair (r = -0.140) and was negatively correlated with all other pairings, except for Cough/Sneeze, which showed a slight positive correlation (r = 0.0219).

Correlations with psychopathology measures. Maladaptive personality traits were assessed using the Personality Inventory for DSM-5 Short Form [PID-5-SF; Thimm et al. (2016)], which measures five domains: disinhibition, antagonism, detachment, negative affect, and psychoticism. The scale consists of 25 items rated on a 4-point Likert scale, ranging from 0 (very false or often false) to 3 (very true or often true).

On average, all maladaptive personality traits assessed by the PID-5-SF were weakly and negatively correlated with IAS pairings. The strongest correlation was observed between the psychoticism dimension and the Muscle/Pain pairing (r = -0.173).

Schizotypy was assessed using the Schizotypal Personality Questionnaire – Brief Revised Updated (SPQ-BRU; Davidson et al. (2016)), which consists of 32 items rated on a 5-point Likert 359 scale ranging from strongly agree to strongly disagree. This questionnaire evaluates four primary 360 dimensions: cognitive-perceptual (positive), interpersonal (negative), disorganized, and social anxiety. These dimensions are further divided into nine secondary factors: constricted affect, 362 eccentricity, magical thinking, lack of close friends, odd speech, referential thinking, social 363 anxiety, suspiciousness, and unusual perceptions.

On average, all nine factors were weakly and negatively correlated with IAS pairings, with correlations ranging from r = -0.170 (between lack of close friends and Muscles/Pain) to r = 0.102(between magical thinking and Itch/Bruise).

The short version of the Autism-Spectrum Quotient (ASQ-Short; Hoekstra et al., 2011)

was used to assess five autistic traits: social skills, adherence to routines, cognitive flexibility
(switching), imagination, and patterns/numbers. The questionnaire consists of 28 items rated on a
4-point Likert scale, ranging from 1 (definitely agree) to 4 (definitely disagree).

Overall, all pairings were weakly and negatively correlated with the ASQ dimensions, except for the Itch/Bruise and Heart/Breathing pairings, which showed weak positive correlations with the patterns/numbers trait (r = 0.184 and r = 0.038, respectively). The strongest correlation was observed between the imagination trait and the Wind/Burp pairing (r = -0.218).

372

375

**Correlations with personality measures.** The Big Five Inventory-Short Form [BFI-S; 376 Lang et al. (2011)] and the Mini International Personality Item Pool [Mini-IPIP6; Sibley et al. 377 (2011)] were used to assess general personality traits. The BFI-S consists of 15 items rated on a 378 7-point Likert scale, ranging from 1 (strongly disagree) to 7 (strongly agree), measuring five 379 personality factors: Openness, Conscientiousness, Extraversion, Agreeableness, and Neuroticism. The Mini-IPIP6 assesses six personality traits—Extraversion, Agreeableness, Conscientiousness, Neuroticism, Openness, and Honesty-Humility—using 24 items. While this questionnaire is typically scored on a 7-point Likert scale from 1 (very inaccurate) to 7 (very accurate), an analogous scale was used in the respective sample. Lastly, the Neuroticism subscale of the NEO Five-Factor Inventory [Neo-FFI; Costa and McCrae (1992)] was used to assess Neuroticism, 385 consisting of 12 items rated on a 5-point Likert scale from 1 (strongly disagree) to 5 (strongly 386 agree). 387

To assess correlations with the IAS pairings, scores were grouped across personality dimensions due to the overlap among these traits (with the exception of Honesty-Humility). On average, IAS pairings were positively associated with most personality dimensions, though these correlations were generally weak. The strongest correlation was observed between Conscientiousness and the Hungry/Thirsty pairing (r = 0.164). Both Honesty-Humility and Neuroticism were weakly and negatively correlated with the IAS pairings, with the highest correlations observed for the Hungry/Thirsty pairing (r = -0.217 and r = -0.200, respectively).

**Correlations with other measures.** The IAS was also correlated with primal world 395 beliefs, as measured by the Primal Inventory [PI-18; J. D. Clifton and Yaden (2021)], which 396 assesses beliefs about the world being alive, good, safe, and enticing. Items that evaluate neutral 397 beliefs about the hierarchical order of importance in the world (i.e., hierarchical), as well as 398 beliefs about the comprehensibility of most things and situations (i.e., understandable), and the 399 belief that the world is characterized by flux (i.e., changing) were added as well. The scale 400 contains 18 items ranging from 5 (Strongly agree) to 0 (strongly disagree). 401 Overall, most primal beliefs show weak positive correlations with all pairings of the IAS. 402 The strongest correlation is between the hierarchical belief and the Hungry/Thirsty pairing (r =403 0.181). Some beliefs, however, exhibit negative correlations with certain pairings. These negative 404 correlations range from r = -0.0940 between the changing belief and the Hungry/Thirsty pairing, to r = -0.00490 between the Enticing belief and the Itch/Bruise pairing. The Generic Conspiracist Beliefs Scale [GCBS; Brotherton et al. (2013)] was used to 407 assess five facets of conspiracy beliefs: Extraterrestrial, Global Conspiracies, Government 408 Malfeasance, Information Control, and Personal Wellbeing. The scale comprises 15 items rated 409 on a 5-point Likert scale, ranging from definitely not true (1) to definitely true (5). 410 Overall, the GCBS showed a weak but positive correlation with all facets of the IAS, with 411 the strongest correlation observed between Global Conspiracies and Hungry/Thirsty (r = 0.140). 412 Negative correlations were found within the Global Conspiracies, Extraterrestrial, and 413 Information Control facets, though these were small, ranging from r = -0.0101 to r = -0.0167. 414 Lastly, the Lying Profile Questionnaire [LIE; Makowski, Pham, et al. (2023)] a 16 item 415 visual analog scale was used to assess 4 dispositional lying simensions: ability; negativity, 416 contextuality, and frequency. Overall, most lying profile dimensions show weak correlations with IAS pairings. Ability 418 exhibits primarily weak positive correlations, with the strongest observed for Wind/Burp (r = 419 0.082). In contrast, Frequency tends to show weak negative correlations, ranging from Wind/Burp 420

(r = -0.062) to Muscles/Pain (r = -0.088). Contextuality displays mixed correlations, with

421

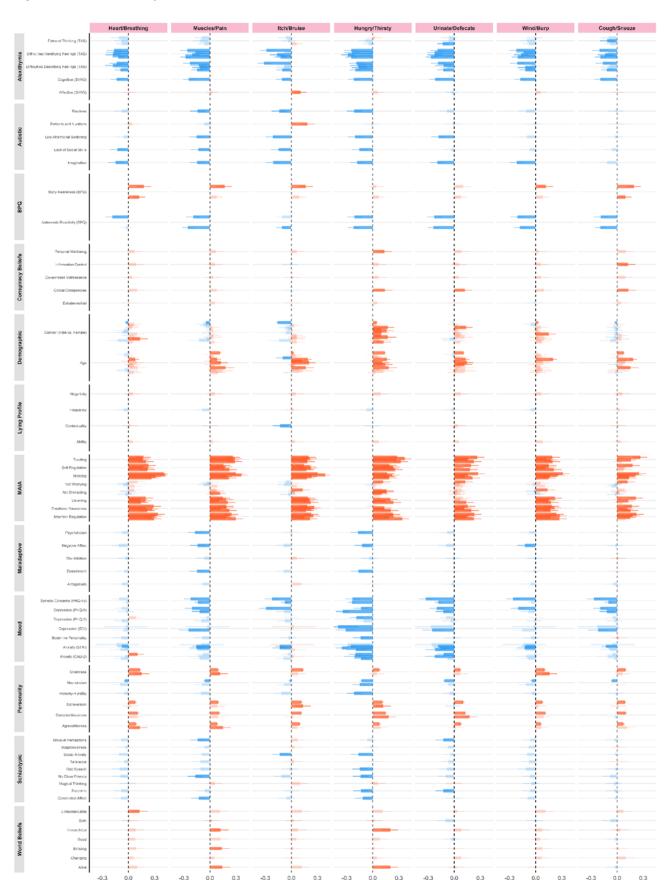
Itch/Bruise showing the strongest negative association (r = -0.127), while Urinate/Defecate has a

 $_{423}$  small positive correlation (r = 0.045). Finally, Negativity is consistently positively correlated with

all pairings, with the strongest relationship found for Hungry/Thirsty (r = 0.090).

Figure 3

Figure 2. Correlates of the IAS



## Discussion

Our findings confirm that interoception exists within a complex network of correlates.

Among these, alexithymia exhibits the strongest negative correlation with the IAS, whereas the

MAIA questionnaire shows the strongest positive correlation. These correlates not only help

explain different aspects of interoception but also serve as valuable tools for validating

interoceptive measures.

While our results reveal various correlations with the IAS, they are limited to the scope of the given questionnaire. Nonetheless, they provide valuable insights into how interoception may relate to different psychological and personality traits. The results show a consistent pattern of correlations with other measures and highlight interesting exploratory results, such as correlations between primal world beliefs with the IAS.

Our analysis found a strong negative correlation between alexithymia and IAS scores, aligning with previous research (Brand et al., 2023; Herbert et al., 2011; Murphy et al., 2019). Similarly, a negative correlation between autism and interoceptive awareness was observed in our sample, consistent with prior findings (DuBois et al., 2016).

Conspiracy beliefs did not strongly correlate with IAS scores, though a slight positive correlation was present. To our knowledge, this relationship has not been previously explored. However, prior studies have suggested connections between interoception and (political) beliefs, potentially pointing to shared underlying mechanisms (Ruisch et al., 2022a).

The relationship between interoception and lying profiles was also weak. This contrasts with previous research suggesting associations between interoception and deception (Makowski, Lau, et al., 2023), warranting further investigation.

Mood and IAS scores exhibited a strong negative correlation, consistent with prior studies that have documented similar findings (Solano López & Moore, 2018). Additionally, personality traits correlated with interoceptive accuracy scores, reinforcing existing research linking personality and interoception (Erle et al., 2021).

We also observed negative correlations between schizotypy and interoception, in line with

previous studies that identified a similar relationship with interoceptive awareness, particularly in individuals at risk for psychosis (Torregrossa et al., 2022).

Interestingly, world beliefs demonstrated significant positive correlations with interoception. While this relationship has not been previously documented, other forms of belief, such as political ideology, have been linked to interoception (Ruisch et al., 2022b). Further research is needed to determine whether world beliefs, which shape our perception of reality (J. D. W. Clifton, 2020), are meaningfully connected to interoception.

Overall, our findings highlight the broad relevance of interoception across various 459 cognitive and affective traits, underscoring its significance in both research and clinical contexts. 460 By identifying numerous correlates of the IAS, we contribute not only to a deeper understanding of interoception's role in daily life but also to the ongoing validation of the IAS and other 462 interoceptive measures. This analysis lays an important foundation for the development of new 463 interoceptive assessment tools, further advancing our comprehension of interoception and its impact on human experience.

# **General Discussion**

452

453

454

461

467

468

469

470

471

472

474

Our analyses revealed that the IAS follows a four-factor structure with an uneven distribution. While the findings indicate that the IAS measures interoception adequately, there is room for improvement. Additionally, different correlation measures with the IAS suggest opportunities for further exploration of how interoception is assessed. In the following section, we discuss the strengths and shortcomings of the IAS, followed by proposed steps to enhance interoception measurement.

Overall, the IAS is straightforward in its sensation-centered items. However, several areas for improvement emerge from this study. Firstly, redundant items should be removed, such as the "itch" item, as highlighted in our analysis. Previous research also suggests redundancy between itch and tickle items Campos et al. (2021). Interestingly, while Campos et al. (2021) does not recommend the removal of either, Lin et al. (2023) argues for removing the itch item due to their overlapping character representation.

Furthermore, this study recommends using analog scales instead of 5-point scales. The limited variability of the 5-point scale often results in most responses clustering around 3 or 4. As shown in Figure 2, adopting an analog scale significantly increases variability. However, even with an analog scale, IAS variability remains constrained. Greater variability allows for better differentiation among participants, making dispersion an essential factor for obtaining meaningful results. Enhancing variability would therefore be beneficial for the IAS.

Despite these improvements, certain limitations persist in the IAS that affect its accuracy. Notably, some modalities are underrepresented—for instance, heart perception is measured by only one item. Expanding modality coverage would enhance variability within each category, leading to more nuanced results. Moreover, the IAS lacks a clear theoretical or empirical structure, with only small item groupings. Ideally, a scale should allow for clear groupings that support meaningful data analysis. In this study, each group contained only two items, resulting in low scores and limited variability. Additionally, some IAS items are ambiguous, with their interpretation depending on context. For example, an item about perceiving heartbeats and another about vomiting could both relate to anxiety, leading to results that may differ from initial expectations. Thus, the grouping and structure of the IAS require refinement. 

Another concern is that all IAS items are phrased positively, which may influence participant responses. While positive phrasing has advantages, it can also introduce response bias, leading to unidimensional results. A more balanced phrasing approach, incorporating both positively and negatively framed items, could yield more accurate responses.

Given these considerations, it is clear that context-specific, cross-modal items—such as integrating cardioception and respiroception—are needed. Recognizing the necessity for a refined interoception scale, this study proposes the development of the Multidimensional Interoceptive Inventory (MInt). This new scale will be designed to align with recent findings on the IAS and interoception research while allowing for direct comparison with IAS correlates.

[TO DO: add - previous work suggests the importance of physiological contexts (Vlemincx et al., 2021)] I would rather put that in the discussion in the suggestions for better

## 506 scales

#### 507 Conclusion

The IAS is a valuable tool for measuring interoception compared to existing
questionnaires and methods. However, refining or even redesigning the questionnaire could lead
to a more precise and comprehensive assessment. This study highlights the need for a new
interoception scale to advance research in the field. By identifying various correlates of the IAS,
this work paves the way for future investigations into optimal interoceptive measures, ultimately
laying the foundation for the development of a more effective interoception survey.

#### 514 References

- Arslanova, I., Galvez-Pol, A., Kilner, J., Finotti, G., & Tsakiris, M. (2022). Seeing through each other's hearts: Inferring others' heart rate as a function of own heart rate perception and perceived social intelligence. *Affective Science*, *3*(4), 862–877.
- Baer, R. A., Smith, G. T., Hopkins, J., Krietemeyer, J., & Toney, L. (2006). Using self-report assessment methods to explore facets of mindfulness. *Assessment*, *13*(1), 27–45.
- Bagby, R. M., Parker, J. D., & Taylor, G. J. (1994). The twenty-item toronto alexithymia scale—i.

  Item selection and cross-validation of the factor structure. *Journal of Psychosomatic*
- see Research, 38(1), 23–32.
- Beck, A. T., Steer, R. A., & Brown, G. (1996). Beck depression inventory–II. *Psychological*Assessment.
- Ben-Shachar, M. S., Lüdecke, D., & Makowski, D. (2020). Effectsize: Estimation of effect size indices and standardized parameters. *Journal of Open Source Software*, *5*(56), 2815.
- Brand, S., Meis, A. C., Tünte, M. R., Murphy, J., Woller, J. P., Jungmann, S. M., Witthöft, M.,
- Hoehl, S., Weymar, M., Hermann, C., & Ventura-Bort, C. (2023). A multi-site German
- validation of the Interoceptive Accuracy Scale and its relation to psychopathological symptom
- burden. Communications Psychology, 1(1). https://doi.org/10.1038/s44271-023-00016-x
- Brand, S., Petzke, T. M., & Witthöft, M. (2022). The differential relationship between self-reported interoceptive accuracy and attention with psychopathology. *Zeitschrift für*

- Klinische Psychologie Und Psychotherapie.
- Brewer, R., Cook, R., & Bird, G. (2016). Alexithymia: A general deficit of interoception. *Royal*Society Open Science, 3(10), 150664.
- Brotherton, R., French, C. C., & Pickering, A. D. (2013). Measuring belief in conspiracy theories:
- The generic conspiracist beliefs scale. *Frontiers in Psychology*, *4*, 279.
- Cabrera, A., Kolacz, J., Pailhez, G., Bulbena-Cabre, A., Bulbena, A., & Porges, S. W. (2018).
- Assessing body awareness and autonomic reactivity: Factor structure and psychometric
- properties of the body perception questionnaire-short form (BPQ-SF). *International Journal*
- of Methods in Psychiatric Research, 27(2), e1596.
- <sup>542</sup> Campos, C., Rocha, N. B., & Barbosa, F. (2021). *Untangling self-reported interoceptive attention*
- and accuracy: Evidence from the european portuguese validation of the body perception
- questionnaire and the interoceptive accuracy scale. http://dx.doi.org/10.31234/osf.io/a7wdj
- <sup>545</sup> Christensen, A. P., Garrido, L. E., & Golino, H. (2023). Unique variable analysis: A network
- psychometrics method to detect local dependence. *Multivariate Behavioral Research*, 58(6),
- 1165–1182.
- <sup>548</sup> Christensen, A. P., & Golino, H. (2021). Estimating the Stability of Psychological Dimensions via
- Bootstrap Exploratory Graph Analysis: A Monte Carlo Simulation and Tutorial. *Psych*, 3(3),
- 479–500. https://doi.org/10.3390/psych3030032
- <sup>551</sup> Clifton, J. D. W. (2020). Testing if primal world beliefs reflect experiences—or at least some
- experiences identified ad hoc. Frontiers in Psychology, 11.
- https://doi.org/10.3389/fpsyg.2020.01145
- <sup>554</sup> Clifton, J. D., & Yaden, D. B. (2021). Brief measures of the four highest-order primal world
- beliefs. *Psychological Assessment*, *33*(12), 1267.
- Costa, P. T., & McCrae, R. R. (1992). Normal personality assessment in clinical practice: The
- NEO personality inventory. *Psychological Assessment*, 4(1), 5.
- Davidson, C. A., Hoffman, L., & Spaulding, W. D. (2016). Schizotypal personality
- questionnaire—brief revised (updated): An update of norms, factor structure, and item content

- in a large non-clinical young adult sample. *Psychiatry Research*, 238, 345–355.
- Desmedt, O., Heeren, A., Corneille, O., & Luminet, O. (2022). What do measures of self-report
- interoception measure? Insights from a systematic review, latent factor analysis, and network
- approach. *Biological Psychology*, 169, 108289.
- DuBois, D., Ameis, S. H., Lai, M.-C., Casanova, M. F., & Desarkar, P. (2016). Interoception in
- Autism Spectrum Disorder: A review. *International Journal of Developmental Neuroscience*,
- 52(1), 104–111. https://doi.org/10.1016/j.ijdevneu.2016.05.001
- Erle, T. M., Mitschke, V., & Schultchen, D. (2021). Did my heart just leap or sink? The role of
- personality for the relation between cardiac interoception and well-being. *Personality and*
- Individual Differences, 170, 110493. https://doi.org/10.1016/j.paid.2020.110493
- Ferentzi, E., Olaru, G., Geiger, M., Vig, L., Köteles, F., & Wilhelm, O. (2021). Examining the
- factor structure and validity of the multidimensional assessment of interoceptive awareness.
- Journal of Personality Assessment, 103(5), 675–684.
- Gabriele, E., Spooner, R., Brewer, R., & Murphy, J. (2022). Dissociations between self-reported
- interoceptive accuracy and attention: Evidence from the interoceptive attention scale.
- Biological Psychology, 168, 108243.
- Gaggero, G., Bizzego, A., Dellantonio, S., Pastore, L., Lim, M., & Esposito, G. (2021).
- <sup>577</sup> Clarifying the relationship between alexithymia and subjective interoception. *PLoS One*,
- 578 *16*(12), e0261126.
- Garfinkel, S. N., Seth, A. K., Barrett, A. B., Suzuki, K., & Critchley, H. D. (2015). Knowing your
- own heart: Distinguishing interoceptive accuracy from interoceptive awareness. *Biological*
- Psychology, 104, 65–74.
- Garner, D. M. (1991). Eating disorder inventory-2; pro-fessional manual. *Psychological*
- Assessment Resources.
- Garner, D. M., Olmstead, M. P., & Polivy, J. (1983). Development and validation of a
- multidimensional eating disorder inventory for anorexia nervosa and bulimia. *International*
- Journal of Eating Disorders, 2(2), 15–34.

Golino, H. F., & Epskamp, S. (2017). Exploratory graph analysis: A new approach for estimating the number of dimensions in psychological research. *PloS One*, *12*(6), e0174035.

- Herbert, B. M., Herbert, C., & Pollatos, O. (2011). On the Relationship Between Interoceptive
- Awareness and Alexithymia: Is Interoceptive Awareness Related to Emotional Awareness?
- Journal of Personality, 79(5), 1149–1175. https://doi.org/10.1111/j.1467-6494.2011.00717.x
- Jahedi, S., & Méndez, F. (2014). On the advantages and disadvantages of subjective measures.
- Journal of Economic Behavior & Organization, 98, 97–114.
- https://doi.org/10.1016/j.jebo.2013.12.016
- Khalsa, S. S., Adolphs, R., Cameron, O. G., Critchley, H. D., Davenport, P. W., Feinstein, J. S.,
- Feusner, J. D., Garfinkel, S. N., Lane, R. D., Mehling, W. E., et al. (2018). Interoception and
- mental health: A roadmap. Biological Psychiatry: Cognitive Neuroscience and
- *Neuroimaging*, *3*(6), 501–513.
- Khalsa, S. S., Adolphs, R., Cameron, O. G., Critchley, H. D., Davenport, P. W., Feinstein, J. S.,
- Feusner, J. D., Garfinkel, S. N., Lane, R. D., Mehling, W. E., Meuret, A. E., Nemeroff, C. B.,
- Oppenheimer, S., Petzschner, F. H., Pollatos, O., Rhudy, J. L., Schramm, L. P., Simmons, W.
- K., Stein, M. B., ... Zucker, N. (2018). Interoception and Mental Health: A Roadmap.
- Biological Psychiatry: Cognitive Neuroscience and Neuroimaging, 3(6), 501–513.
- https://doi.org/10.1016/j.bpsc.2017.12.004
- Kleckner, I. R., Wormwood, J. B., Simmons, W. K., Barrett, L. F., & Quigley, K. S. (2015).
- Methodological recommendations for a heartbeat detection-based measure of interoceptive
- sensitivity. *Psychophysiology*, 52(11), 1432–1440.
- Koike, H., & Nomura, M. (2023). Development and validation of japanese versions of the
- interoceptive accuracy scale and interoceptive attention scale. SAGE Open, 13(4),
- 21582440231214639.
- Kroenke, K., Spitzer, R. L., & Williams, J. B. (2001). The PHQ-9: Validity of a brief depression
- severity measure. Journal of General Internal Medicine, 16(9), 606–613.
- Kroenke, K., Spitzer, R. L., & Williams, J. B. (2002). The PHQ-15: Validity of a new measure for

- evaluating the severity of somatic symptoms. *Psychosomatic Medicine*, 64(2), 258–266.
- Kroenke, K., Spitzer, R. L., & Williams, J. B. (2003). The patient health questionnaire-2: Validity
- of a two-item depression screener. *Medical Care*, 41(11), 1284–1292.
- Kroenke, K., Spitzer, R. L., Williams, J. B., Monahan, P. O., & Löwe, B. (2007). Anxiety
- disorders in primary care: Prevalence, impairment, comorbidity, and detection. Annals of
- Internal Medicine, 146(5), 317–325.
- Lang, F. R., John, D., Lüdtke, O., Schupp, J., & Wagner, G. G. (2011). Short assessment of the
- big five: Robust across survey methods except telephone interviewing. Behavior Research
- *Methods*, 43, 548–567.
- 623 Lin, X.-X., Shen, H.-R., Lin, J.-X., Zhang, Y.-H., Murphy, J., Wang, Y.-Z., Sun, Y.-B., Wang, N.,
- Wang, J.-Y., Wei, G.-X., & Luo, F. (2023). Psychometric validation and refinement of the
- Chinese Interoceptive Accuracy Scale (IAS) in general population and patients with chronic
- pain. Journal of Psychosomatic Research, 175, 111541.
- 627 https://doi.org/10.1016/j.jpsychores.2023.111541
- Makowski, D., Lau, Z. J., Pham, T., Te, A., Kirk, S., & Liauw Yong Tong, C. (2023). The heart
- can lie: The role of interoception and theory of mind in deception.
- http://dx.doi.org/10.31234/osf.io/p342w
- Makowski, D., Pham, T., Lau, Z. J., Raine, A., & Chen, S. A. (2023). The structure of deception:
- Validation of the lying profile questionnaire. Current Psychology, 42(5), 4001–4016.
- 633 Mehling, W. E., Acree, M., Stewart, A., Silas, J., & Jones, A. (2018). The multidimensional
- assessment of interoceptive awareness, version 2 (MAIA-2). *PloS One*, 13(12), e0208034.
- Mehling, W. E., Price, C., Daubenmier, J. J., Acree, M., Bartmess, E., & Stewart, A. (2012). The
- Multidimensional Assessment of Interoceptive Awareness (MAIA). *PLoS ONE*, 7(11),
- e48230. https://doi.org/10.1371/journal.pone.0048230
- Messer, S. C., Angold, A., Costello, E. J., Loeber, R., Van Kammen, W., & Stouthamer-Loeber,
- M. (1995). Development of a short questionnaire for use in epidemiological studies of
- depression in children and adolescents: Factor composition and structure across development.

- International Journal of Methods in Psychiatric Research, 5, 251–262.
- 642 Miller, L. C., Murphy, R., & Buss, A. H. (1981). Consciousness of body: Private and public.
- Journal of Personality and Social Psychology, 41(2), 397.
- Murphy, J., Brewer, R., Plans, D., Khalsa, S. S., Catmur, C., & Bird, G. (2019). Testing the
- independence of self-reported interoceptive accuracy and attention. Quarterly Journal of
- Experimental Psychology, 73(1), 115–133. https://doi.org/10.1177/1747021819879826
- Porges, S. (1993). Body perception questionnaire. Laboratory of Developmental Assessment,
- 648 *University of Maryland*, 10, s15327752jpa5304\_1.
- Rodriguez, A., Reise, S. P., & Haviland, M. G. (2016). Evaluating bifactor models: Calculating
- and interpreting statistical indices. *Psychological Methods*, 21(2), 137.
- Ruisch, B. C., Von Mohr, M., Naber, M., Tsakiris, M., Fazio, R. H., & Scheepers, D. T. (2022a).
- Sensitive liberals and unfeeling conservatives? *Interoceptive sensitivity predicts political*
- liberalism. Politics and the Life Sciences, 41(2), 256–275. https://doi.org/10.1017/pls.2022.18
- Ruisch, B. C., Von Mohr, M., Naber, M., Tsakiris, M., Fazio, R. H., & Scheepers, D. T. (2022b).
- Sensitive liberals and unfeeling conservatives? *Interoceptive sensitivity predicts political*
- liberalism. Politics and the Life Sciences, 41(2), 256–275. https://doi.org/10.1017/pls.2022.18
- 657 Schandry, R. (1981). Heart beat perception and emotional experience. *Psychophysiology*, 18(4),
- 483–488.
- <sup>659</sup> Shields, S. A., Mallory, M. E., & Simon, A. (1989). The body awareness questionnaire:
- Reliability and validity. *Journal of Personality Assessment*, 53(4), 802–815.
- 661 Sibley, C. G., Luyten, N., Purnomo, M., Mobberley, A., Wootton, L. W., Hammond, M. D.,
- Sengupta, N., Perry, R., West-Newman, T., Wilson, M. S., et al. (2011). The mini-IPIP6:
- Validation and extension of a short measure of the big-six factors of personality in new
- zealand. New Zealand Journal of Psychology, 40(3).
- 665 Solano López, A. L., & Moore, S. (2018). Dimensions of Body-Awareness and Depressed Mood
- and Anxiety. Western Journal of Nursing Research, 41(6), 834–853.
- https://doi.org/10.1177/0193945918798374

- Spielberger, C. D. (1970). Manual for the state-trait anxiety inventory (self-evaluation questionnaire). (*No Title*).
- Suksasilp, C., & Garfinkel, S. N. (2022). Towards a comprehensive assessment of interoception in a multi-dimensional framework. *Biological Psychology*, *168*, 108262.
- Thimm, J. C., Jordan, S., & Bach, B. (2016). The personality inventory for DSM-5 short form

  (PID-5-SF): Psychometric properties and association with big five traits and pathological

  beliefs in a norwegian population. *BMC Psychology*, 4, 1–11.
- Todd, J., Swami, V., Aspell, J. E., Furnham, A., Horne, G., & Stieger, S. (2022). Are some interoceptive sensibility components more central than others? Using item pool visualisation to understand the psychometric representation of interoception. *Plos One*, *17*(12), e0277894.
- Torregrossa, L. J., Amedy, A., Roig, J., Prada, A., & Park, S. (2022). Interoceptive functioning in schizophrenia and schizotypy. *Schizophrenia Research*, 239, 151–159.
- https://doi.org/10.1016/j.schres.2021.11.046
- Von Mohr, M., Silva, P. C., Vagnoni, E., Bracher, A., Bertoni, T., Serino, A., Banissy, M. J.,
- Jenkinson, P. M., & Fotopoulou, A. (2023). My social comfort zone: Attachment anxiety shapes peripersonal and interpersonal space. *Iscience*, 26(2).
- Vorst, H. C., & Bermond, B. (2001). Validity and reliability of the bermond–vorst alexithymia questionnaire. *Personality and Individual Differences*, *30*(3), 413–434.
- Zanarini, M. C. (2003). Zanarini rating scale for borderline personality disorder (ZAN-BPD): A
   continuous measure of DSM-IV borderline psychopathology. *Journal of Personality* Disorders, 17(3), 233–242.
- Zsido, A. N., Teleki, S. A., Csokasi, K., Rozsa, S., & Bandi, S. A. (2020). Development of the
   short version of the spielberger state—trait anxiety inventory. *Psychiatry Research*, 291,
   113223.