

Too beautiful to be fake: Attractive faces are less likely to be judged as artificially generated

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ABSTRACT

Technological advances render the distinction between artificial (e.g., computer-generated faces) and real stimuli increasingly difficult, yet the factors driving our beliefs regarding the nature of ambiguous stimuli remain largely unknown. In this study, 150 participants rated 109 pictures of faces on 4 characteristics (attractiveness, beauty, trustworthiness, familiarity). The stimuli were then presented again with the new information that some of them were AI-generated, and participants had to rate each image according to whether they believed them to be real or fake. Despite all images being pictures of real faces from the same database, most participants did indeed rate a large portion of them as ‘fake’ (often with high confidence), with strong intra- and inter-individual variability. Our results suggest a gender-dependent role of attractiveness on reality judgements, with faces rated as more attractive being classified as more real. We also report links between reality beliefs tendencies and dispositional traits such as narcissism and paranoid ideation.

1. Introduction

Advancements in technology have now made it possible to create near-perfect simulations that are indistinguishable from reality with an ease, affordability and accessibility that are unprecedented in Human history. These artificial, yet realistic constructs permeate all areas of life through immersive works of fiction, deep fakes (real-like images and videos generated by deep learning algorithms), virtual and augmented reality (VR and AR), artificial beings (artificial intelligence “bots” with or without a physical form), fake news and skewed narratives, of which ground truth is often hard to access (Nightingale & Farid, 2022). Such developments not only carry important consequences for the technological and entertainment sectors, but also for security and politics - for instance if used for propaganda and disinformation, recruitment into malevolent organizations, or religious indoctrination (Pantserev, 2020). This issue is central to what has been coined the “post-truth era” (Lewandowsky et al., 2017), in which the distinction (and lack thereof)

between authentic and simulated objects will play a critical role.

While not all simulations have achieved perfect realism, such as Computer Generated Images (CGI) in movies or via recent algorithms such as GANs or diffusion model, which often include distortions or lack certain key details distinguishing them from real images (Corvi et al., 2022; McDonnell & Breidt, 2010), it is fair to assume that these technical limitations will become negligible in the near future. This is particularly true in the field of face generation, where face-generation algorithms are already able to create stimuli that are virtually indistinguishable from real photos (Moshel et al., 2022; Nightingale & Farid, 2022; Tucciarelli et al., 2020). Such a technological feat, however, leads to a new question: if real and fake stimuli cannot be differentiated based on their objective “physical” characteristics, how can we form judgements regarding their nature?

Literature shows that the context surrounding a stimulus often plays an important role in the assessment of its reality (a process henceforth referred to as simulation monitoring, Makowski, 2018; Makowski,

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Sperduti, et al., 2019). With the extensive search and processing of cues within ambiguous stimuli being an increasingly complex and cognitively effortful strategy (Michael & Sanson, 2021; Susmann et al., 2021), people tend to draw on peripheral contextual cues (Fig. 1), such as the source of the stimulus (e.g., which journal was the information published in), and its credibility, authority and expertise, to help facilitate their evaluation (Michael & Sanson, 2021; Petty & Cacioppo, 1986; Susmann et al., 2021). However, the atomization and decontextualization of information allowed by online social media (where text snippets or video excerpts are often mass-shared with little context) makes this task progressively difficult (Berghel, 2018; Y. Chen et al., 2015). Thus, in the absence of clear contextual information, what drives our beliefs of reality?

Evidence suggests that inter-individual characteristics play a crucial role in simulation monitoring, with factors such as cognitive style, prior beliefs, and personality traits (Bryanov & Vziatysheva, 2021; Ecker et al., 2022; Sindermann et al., 2020). For instance, individuals with stronger analytical reasoning skills have been found to better discriminate real from fake stimuli (Pehlivanoglu et al., 2021; Pennycook & Rand, 2019), and prior knowledge or beliefs about the stimulus influences one's perception of it by biasing the attention deployment towards information that is in line with one's expectations (Britt et al., 2019). Furthermore, dispositional traits, such as high levels of narcissism and low levels of openness and conscientiousness, have been associated with greater susceptibility to fake news (Piksa et al., 2022; Sindermann et al., 2020). Interestingly, a recent review suggested that narcissism was related to a strong self-perceived ability at detecting lies (Turi et al., 2022), which could translate to participants scoring high on narcissism providing more clear cut and confidence responses. Conversely, those high in honesty-humility tend to be more conservative in their judgements to ensure fairness (Liu et al., 2020), likely resulting in lower confidence ratings.

Beyond stimulus- and individual-related characteristics, evidence suggests that the interaction between the two (i.e., the subjective reaction associated with the experience of a given stimulus), contributes to simulation monitoring decisions. For instance, the intensity of experienced emotions have been shown to increase one's sense of presence - the extent to which one feels like "being there", as if the object of experience was real - when engaged in a fictional movie or a VR environment (Makowski et al., 2017; Sanchez-Vives & Slater, 2005). Indeed, participants' self-reported emotional arousal were found to significantly predict the probability that they would perceive images as real (Azevedo et al., 2020). Conversely, beliefs that emotional stimuli were fake (e.g., that emotional scenes were not authentic but instead involved actors and movie makeup) were found to result in emotion down-regulation (Makowski, Sperduti, et al., 2019; Sperduti et al., 2017). In line with these findings, studies on susceptibility to fake news have also found heightened stimulus emotionality to be associated with greater belief (Bago et al., 2022; Martel et al., 2020), and higher neurophysiological arousal was predictive of judging realistic images as real (Azevedo et al., 2020). Additionally, other factors, such as the stimuli's perceived self-relevance (Goldstein, 2009; Sperduti et al., 2016), as well as familiarity (Begg et al., 1992), could also play a role in guiding our appraisal of a stimulus. For instance, Miller et al. (2023) reported that participants were more likely to mistakenly identify AI-generated faces as real because they perceived them as more familiar.

Due to their popularity as a target of CGI technology and the prospect offered with facial features that can be experimentally manipulated, AI-generated images of faces are increasingly used to study face processing (Dawel et al., 2021), in particular in relationship with saliency or emotions, as well as to other important components of face evaluation, such as trustworthiness or attractiveness (Balas & Pacella, 2017; Calbi et al., 2017; Sobieraj & Krämer, 2014; Tsikandilakis et al., 2019). Interestingly, artificially created faces rated as more attractive (by an independent group of raters) were perceived as less real (Tucciarelli et al., 2020). Conversely, Lefooghe et al. (2022) reports that

attractiveness ratings were significantly lower when participants who were told that the faces were AI-generated were compared to those who had no prior knowledge. Similarly, when participants are informed that faces are AI-generated, the perceived artificiality leads to lower trust ratings (Wang & Nishida, 2024), even when they are real faces (Lefooghe et al., 2022). In contrast, when participants are unaware that the faces are AI-generated, trust ratings for these synthetic faces tend to increase (Nightingale & Farid, 2022). Whereas this line of evidence suggests that reality beliefs have an effect on face attractiveness and trustworthiness ratings, the opposite question - whether attractiveness and trustworthiness contribute to the formation of reality beliefs - has received little attention to date.

AI-generated content, in particular realistic images, is becoming commonplace and carries important risks for misinformation and blackmailing (Viola & Voto, 2023), emphasizing the need to understand the different components that come into play in the formation of reality beliefs. This exploratory study primarily aims at investigating the effect of facial attractiveness on simulation monitoring, i.e., on the beliefs that an image is real or artificially generated. The affective reality theory (Makowski, 2018, 2023) posits that the default tendency is to believe that experiences are real, with emotional and bodily reactions playing a pivotal role in reinforcing or challenging this belief. According to the theory, there is a quadratic (inverse U-shaped) relationship between affect and reality judgements: stimuli that elicit mild to moderate emotional and/or bodily reactions tend to enhance the perception of realness, increasing confidence in reality. However, when emotional or bodily responses become too intense or overwhelming, the default belief shifts towards non-reality beliefs as an emotion regulation mechanism. In other words, extreme emotional reactions can potentially trigger beliefs of reality denial ("it cannot be real") as a protective mechanism to help individuals cope with distress. The present experiment can be put in relation with the first part of this proposal - that of a positive relationship between embodied or emotional reactions and appraisals of reality. Following this hypothesis, faces rated as either highly attractive or unattractive - and eliciting stronger reactions - would likely be judged as real. We expect a similar pattern with trustworthiness, where faces judged as highly trustworthy or untrustworthy will be more likely to be perceived as real. Finally, we anticipate a positive relationship between familiarity and perceived realness, as familiar faces tend to be judged as more real (Miller et al., 2023). Additionally, we will further explore the link shared by dispositional traits, such as personality and attitude towards AI, with simulation monitoring tendencies. Importantly, this study does investigate the discriminative accuracy between "true" photos and "true" artificially-generated images (which we consider more a technological issue than a psychological one), focusing on the beliefs that a stimulus is real or fake, independently of its true. In other words, the present study investigates the psychological process that leads to different beliefs of reality, rather than the discrimination between real faces and actual AI-generated ones, which largely depends on the technological quality of the AI-generation process.

2. Methods

All the material (preregistration,¹ experiment demo, experiment code, raw data, analysis script with complementary figures and analyses, etc.) is available at <https://github.com/RealityBending/FakeFace>.

¹ This approach diverges from the preregistration in several key ways. First, the phrasing of items was modified from "Assuming the face you saw was of a real individual, how..." to "I find this person..." A new attractiveness scale (i.e., Beauty) was introduced to capture a more objective measure of attraction. Finally, the data analysis method was altered from Bayesian Mixed Models due to computational limitations, as we were unable to run these models on a high-performance cluster.

Determinants of Simulation Monitoring

« Is this information **real** or **fake**? »

« Real » = genuine, authentic

« Fake » = artificial, simulated, deceptive

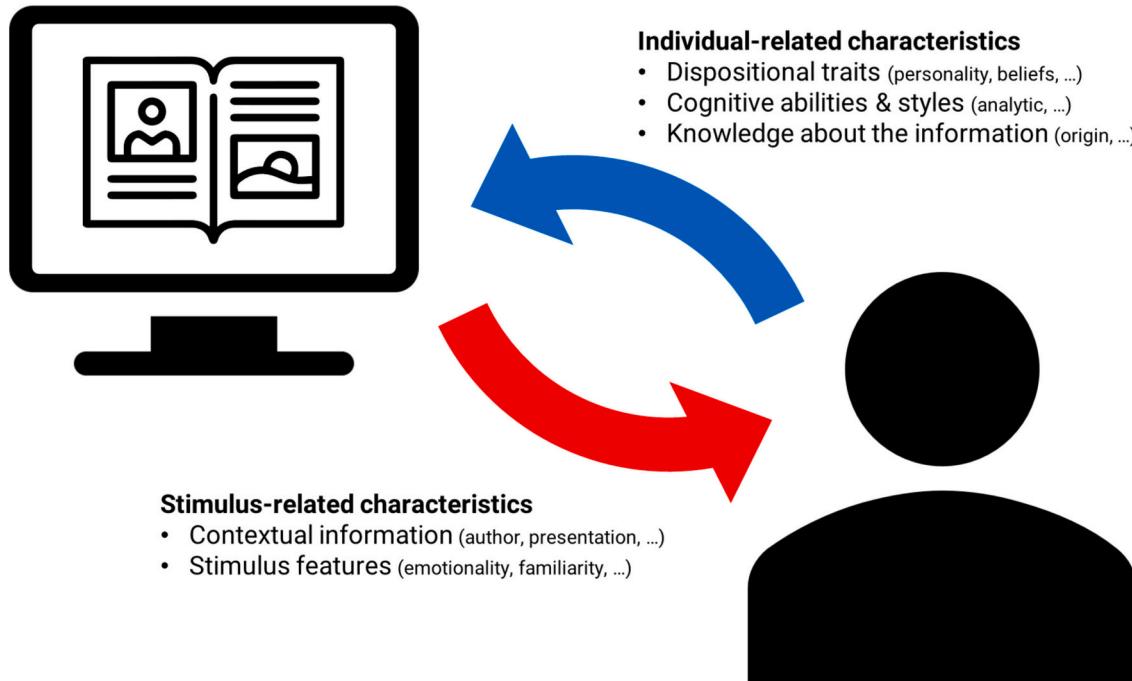


Fig. 1. The decision to believe that an ambiguous stimulus (of any form, e.g., images, text, videos, environments, ...) is real or fake depends of individual characteristics (e.g., personality and cognitive styles), stimulus-related features (context, emotionality), and their interaction, which can manifest for instance in our bodily reaction.

2.1. Ethics statement

This study was approved by the NTU Institutional Review Board (NTU IRB-2022-187) and all procedures performed were in accordance with the ethical standards of the institutional board and with the 1964 Helsinki Declaration. All participants provided their informed consent prior to participation and were incentivized after completing the study.

2.2. Procedure

In the first part of the study, participants answered a series of personality questionnaires presented in the order below. These include the *Mini-IPIP6* (24 items, Sibley et al., 2011) measuring 6 personality traits, the *SIAS-6* and the *SPS-6* (6 items each, Peters et al., 2012) assessing social anxiety levels, 5 items we devised pertaining to expectations about AI-generated image technology ("I think current Artificial Intelligence algorithms can generate very realistic images"), to potentially test and mitigate the potential effect of expectations/beliefs about AI. These items were mixed with 5 items from the general attitudes towards AI scale to lower the former's saliency and the possibility of it priming the subjects about the task, (GAAIS, Schepman & Rodway, 2020) the *FFNI-BF* (30 items, Jauk et al., 2023) measuring 9 facets of narcissism; the *R-GPTS* (18 items, Freeman et al., 2021) measuring 2 dimensions related to paranoid thinking; and the *IUS-12* (12 items, Carleton et al., 2007) measuring intolerance to uncertainty. Self-rated attractiveness was also assessed using 2 items - one measuring general attractiveness ("How attractive would you say you are?" Marcinkowska et al., 2021) and the other measuring physical attractiveness ("How would you rate

your own physical attractiveness relative to the average," Spielmann et al., 2020). 3 attention check questions were also embedded in the surveys. All Cronbach's alpha values were within the acceptable to excellent range, except for the neuroticism subscale of the Mini-IPIP6 and the negative subscale of the GAAI, which were poor, and the Expectations about AI scale, which was questionable (Gliem & Gliem, 2003; see supplementary material for the details of the reliability analysis).

In the second part of this study, images of neutral-expression faces from the validated American Multiracial Face Database (AMFD, J. M. Chen et al., 2021) were presented to the participants for 500 ms each, in a randomized order, following a fixation cross display (750 ms). The decision to present the faces for 500 ms was based on pilot studies, which demonstrated that this duration provides a sufficient perceptual window for decision-making and aligns with previous research indicating stable judgement levels and increased confidence beyond this exposure time (Willis & Todorov, 2006).

The AMFD is a recently validated database including a set of 110 pictures of homogeneous quality featuring diverse faces (particularly in terms of ethnicity), each (except one) posing with either a neutral or smiling expression. We selected all 109 neutral images (89 women and 20 men) to reduce the influence of confounding factors like affect. The AMFD primarily features racially ambiguous faces, representing multiple racial categories such as multiracial, Latinx, and white. The database includes 81 faces from individuals self-reporting two racial backgrounds and 29 from those with three or more racial backgrounds: 33 % Asian/White, 22 % Latinx/White, 11 % Asian/Latinx, 6 % White/Middle Eastern, 5 % Black/White, and 5 % Asian/Middle Eastern, with about 18

% identifying as other racial backgrounds.

After each stimulus presentation, ratings of *Trustworthiness* ("I find this person trustworthy") and *Familiarity* ("This person reminds me of someone I know") were collected using visual analog scales. Notably, as facial attractiveness is a multidimensional construct, encompassing evolutionary, sociocultural, biological as well as cognitive aspects (Han et al., 2018; Rhodes et al., 2006), we assessed attractiveness using 2 visual analog scales, measuring general *Attractiveness* ("I find this person attractive") and physical *Beauty* ("This face is good-looking"). This dual-scale approach aims to reflect two conceptually distinct dimensions: Attractiveness might capture personal, Self-relevant and subjective appeal, whereas Beauty might be related to a more "objective" decision based on aesthetic criteria that can be recognized independently of personal attraction. In other words, we wanted the experiment to be able to potentially capture scenarios where a face could be judged beautiful yet not attractive and vice versa.

In the last part of the study, participants were informed that "about half" of the images previously seen were AI-generated (the instructions used a cover story explaining that the aim of the research was to validate a new face generation algorithm). The same set of stimuli was displayed again for 500 ms in a new randomized order. This time, after each display, participants were asked to express their belief regarding the nature of the stimulus using a visual analog scale (with *Fake* and *Real* as the two extremes). The study was implemented using *jsPsych* (De Leeuw, 2015), and the exact instructions are available in the experiment code.

2.3. Participants

Although the main part of the study relied on within-subject design (with 109 trials per participant), we also planned to do between-participants analyses, thus aiming at collecting a larger sample than traditionally used in experimental psychology (with budget availability as the main constraint). One hundred and fifty participants were recruited via *Prolific*, a crowd-sourcing platform recognized for providing high quality data [Peer et al., 2022; douglas2023data]. The only inclusion criterion was a fluent proficiency in English to ensure that the experiment instructions would be well-understood. Participants were incentivized with a reward of about £ 7.5 for completing the study, which took about 45 min to finish. Demographic variables (age, gender, sexual orientation, education and ethnicity) were self-reported on a voluntary basis.

We excluded 5 participants that either failed 2 ($\geq 66.6\%$) or more attention check questions, took an implausibly short time to finish the questionnaires or had incomplete responses. Out of the 5 participants excluded, 2 participants were excluded because they failed 2 out of 3 attention checks, 1 because they did not answer the sexual orientation question, which made further analysis impossible, and 2 had an abnormal low agreement ($r < 0.1$) between the beauty and the attractiveness ratings (possibly indicating random responses as these two scales exhibited a higher correlation for the other participants). The final sample included 145 participants (Mean age = 28.3, SD = 9.0, range: [19, 66]; Sex: 48.3 % females, 51.0 % males, 0.7 % others).

2.4. Data analysis

The real-fake ratings (measured originally with a $[-1, 1]$ analog scale) were converted into two scores, corresponding to two conceptually distinct mechanisms: the dichotomous *belief* (real or fake, based on the sign of the rating) and the *confidence* (the rating's absolute value) associated with that belief. The former was analyzed using logistic mixed models, which modelled the probability of assigning a face to the real (≥ 0) as opposed to fake (< 0). The latter, as well as the other face ratings (attractiveness, beauty, trustworthiness and familiarity), was modelled using mixed beta regressions (suited for outcome variables expressed in percentages). The models included the participants and stimuli as random **intercepts with no nested variables.

We started by investigating the effect of the procedure and instructions to check whether the stimuli (which were all images of real faces) were judged as fake in sufficient proportion to warrant their analysis. Additionally, we assessed the effect of the re-exposure delay, i.e., the time between the first presentation of the image (corresponding to the face ratings) and the second presentation (for the real-fake rating), as well as that of the presentation order to check whether for habituation or learning effects.

The determinants of reality beliefs were modelled separately for attractiveness, beauty, trustworthiness, and familiarity, using second order raw polynomials coefficients to allow for possible quadratic relationships (Fig. 2). Aside from attractiveness (conceptualized as a general construct), models for beauty, trustworthiness and familiarity were adjusted for the two remaining variables mutatis mutandis. The analysis focused on sexual-orientation relevant stimuli, i.e., on faces that were aligned with respect to the participants' sexual orientation (i.e., female faces for heterosexual males, male faces for homosexual males, etc.), and the models included the interaction with the participants' gender (as a sexual dimorphism has been reported in face appraisal processes). For the attractiveness and beauty models, we then added the interaction with the reported self-attractiveness (the average of the two questions pertaining to it) to investigate its potential modulatory effect. Finally, we investigated the inter-individual correlates of simulation monitoring with similar models (but this time, for all items regardless of the participant's gender or sexual orientation) for each questionnaire, with all of the subscales as orthogonal predictors.

The analysis was carried out using *R* 4.2 (R Core Team, 2022), the *tidyverse* (Wickham et al., 2019), and the *easystats* collection of packages (Lüdecke et al., 2019, 2020, 2021; Makowski et al., 2020; Makowski, Ben-Shachar, & Lüdecke, 2019). As all the details, scripts and complementary analyses are open-access, we will focus in the manuscript on findings that are highly statistically significant ($p < .01$).

3. Results

On average, across participants, 44 % of images (95 %~CI [0.12, 0.64]) were judged as fake and 56 % of images (95 %~CI [0.36, 0.88]) as real. An intercept-only model with the participants and images as random factors showed that the Intraclass Correlation Coefficient (ICC), which can be interpreted as the proportion of variance explained by the random factors, was of 9.0 % for the participants and 9.6 % for the stimuli.

While the delay of stimulus re-exposure stimulus did not have a significant effect on participants' beliefs of reality ($OR = 1.00$, 95%CI = [0.99, 1.00]), judgement confidence was found to be negatively associated with re-exposure delay when the faces were judged as real ($\beta = -0.006$, 95%CI = [-0.1, 0.002], $p = .004$). The presentation order also did not have an effect on the belief ($OR = 1.00$, 95%CI = [1.00, 1.00]) but was related to a decrease of confidence ($\beta_{real} = -0.003$, 95%CI = [-0.004, -0.002], $p < .001$; $\beta_{fake} = -0.002$, 95%CI = [-0.004, -0.0003], $p = .021$): items presented at the end of the session were judged with a similar bias but a decreased overall confidence.

3.1. Determinants of simulation monitoring

Attractiveness had a significant positive and linear relationship ($R^2_{marginal} = 2.0\%$) with the belief that a stimulus was real ($\beta_{poly1} = 16.57$, 95%CI = [7.33, 25.82], $z = 3.51$, $p < .001$) for males, and a quadratic relationship for females ($\beta_{poly2} = 7.82$, 95%CI = [1.81, 13.84], $z = 2.55$, $p = .011$), with both non-attractive and attractive faces being judged as more real. Attractiveness was also found to have a significant positive and quadratic relationship with confidence in judging faces both as real ($\beta_{poly2} = 4.30$, 95%CI = [0.97, 7.64], $z = 2.53$, $p = .011$) and as fake ($\beta_{poly2} = 5.23$, 95%CI = [0.86, 9.60], $z = 2.35$, $p = .019$) for

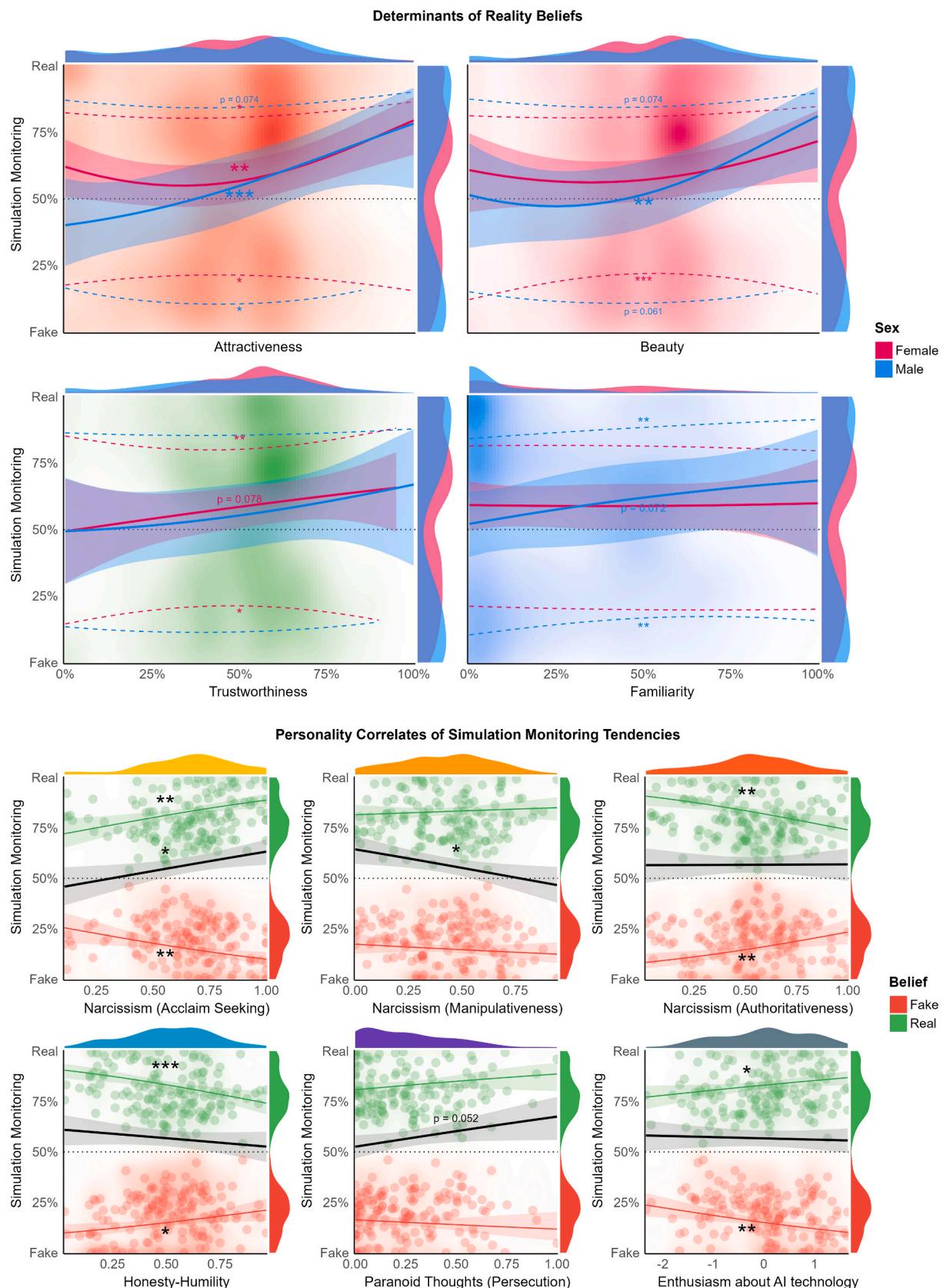


Fig. 2. Top part shows the effect of face ratings on 1) the probability of judging a face as real vs. fake (solid line) and 2) on the confidence associated with that judgement (dashed lines) depending on the sex. Bottom part shows the effect of personality traits on the belief (black line) and the confidence associated with it (colored lines). The points are the average per participant confidence for both types of judgements. Stars indicate significance ($p < .001^{***}$, $p < .01^{**}$, $p < .05^*$).

females. For males, however, a significant negative and quadratic relationship was found between attractiveness ratings and belief confidence only for faces judged as fake ($\beta_{poly2} = -9.92$, 95%CI = [-18.99, -0.86], $z = -2.15$, $p = .032$). There was no interaction with reported self-attractiveness.

Beauty, adjusted for trustworthiness and familiarity, had a significant positive and linear relationship ($R^2_{marginal} = 2.0\%$) with the belief that a stimulus was real ($\beta_{poly1} = 11.82$, 95%CI = [4.28, 20.21], $z = 2.76$, $p = .006$) for males only. No effect on confidence was found, aside from a quadratic relationship in females for faces judged as fake, suggesting that non-beautiful and highly beautiful faces were rated as fake with more confidence than average faces ($\beta_{poly2} = 7.84$, 95%CI = [3.39, 12.29], $z = 3.46$, $p < .001$). There was no interaction with reported self-attractiveness.

Trustworthiness, adjusted for beauty and familiarity, had a predominantly positive and linear relationship ($R^2_{marginal} = 2.0\%$) with the belief that a stimulus was real ($\beta_{poly1} = 6.44$, 95%CI = [-0.11, 13.00], $z = 1.93$, $p = .0054$) for females only. No effect on confidence was found for males, whereas a quadratic relationship was found for females for both faces judged as real ($\beta_{poly2} = 6.14$, 95%CI = [2.13, 10.14], $z = 3.00$, $p = .003$) as well as fake ($\beta_{poly2} = 6.12$, 95%CI = [1.49, 10.75], $z = 2.59$, $p = .001$), suggesting that non-trustworthy and highly trustworthy faces were rated with more confidence than average faces.

We did not find any significant relationships for familiarity adjusted for beauty and trustworthiness ($R^2_{marginal} = 2.0\%$). However, a significant positive and linear relationship was found between familiarity and the confidence judgements of rating faces as real ($\beta_{poly1} = 9.98$, 95%CI = [3.83, 16.13], $z = 3.18$, $p = .001$) whereas a negative linear relationship was found with those judged as fake ($\beta_{poly1} = -12.41$, 95%CI = [-20.27, -4.54], $z = -3.09$, $p = .002$) for males only. This hence suggests that males more confidently judge faces as real with when they are familiar, and as fake when they are unfamiliar.

Note that we also tested as predictors the normative attractiveness and trustworthiness scores (i.e., the average values from the stimuli database validation), which showed a significant positive linear relationship between beliefs of reality and attractiveness, as well as trustworthiness, only for males (see Supplementary Analysis for details).

3.2. Inter-individual correlates of simulation monitoring

The models including the personality traits suggested that *Honesty-Humility* had a significant negative relationship with the confidence associated with real as well as fake judgements ($\beta_{real} = -1.62$, 95%CI = [-2.55, -0.70], $z = -3.43$, $p < .001$; $\beta_{fake} = -1.16$, 95%CI = [-2.09, -0.23], $z = -2.45$, $p = 0.014$).

Significant positive associations were found between the probability of judging faces as real and dimensions of narcissism such as *Acclaim Seeking* ($\beta = 2.24$, 95%CI = [1.17, 4.27], $z = 2.44$, $p = .015$), and *Manipulativeness* ($\beta = 0.47$, 95%CI = [0.25, 0.87], $z = -2.4$, $p = 0.017$). Confidence judgements also shared significant links with narcissism through various facets, such as a positive relationship between the confidence for both real and fake judgements with *Acclaim Seeking* ($\beta_{real} = 1.65$, 95%CI = [0.59, 2.70], $z = 3.07$, $p = .002$; $\beta_{fake} = 1.62$, 95%CI = [0.56, 2.68], $z = 3.00$, $p = .003$), and a negative relationship with *Authoritativeness* ($\beta_{real} = -1.57$, 95%CI = [-2.58, -0.57], $z = -3.08$, $p = .002$; $\beta_{fake} = -1.49$, 95%CI = [-2.50, -0.48], $z = -2.89$, $p = .004$).

A positive trend was found in the relationship between the *Persecutory Ideation* dimension of paranoid thinking and the belief that the faces were real ($\beta = 1.87$, 95%CI = [0.99, 3.54], $z = 1.93$, $p = .054$).

The *Prospective Anxiety* aspect of intolerance to uncertainty shared a negative trend in its association with confidence ratings ($\beta_{real} = 1.43$, 95%CI = [0.10, 2.76], $z = 2.10$, $p = .036$; $\beta_{fake} = -0.91$, 95%CI =

[-1.93, 0.11], $z = -1.75$, $p = .081$). No significant effect was found for social anxiety.

Questions pertaining to the attitude towards AI were reduced to 3 dimensions through factor analysis, labelled AI-Enthusiasm (loaded by items expressing interest and excitement in AI development and applications), AI-Realness (loaded by items expressing positive opinions on the ability of AI to create realistic material), and AI-Danger (loaded by items expressing concerns on the unethical misuse of AI technology). However, only AI-Enthusiasm displayed a significant positive relationship with the confidence in both real and fake judgements ($\beta_{real} = 0.21$, 95%CI = [0.02, 0.40], $z = 2.20$, $p = .028$; $\beta_{fake} = 0.31$, 95%CI = [0.12, 0.50], $z = -8.90$, $p < 0.001$).

4. Discussion

This study aimed at investigating the effect of facial ratings (attractiveness, beauty, trustworthiness and familiarity) on simulation monitoring, i.e., on the belief that a stimulus was artificially generated. Most strikingly, despite all the stimuli being real faces from the same database, all participants believed (to high degrees of confidence) that a significant proportion of them were fake. This finding not only attests to the effectiveness of our instructions, but highlights the current levels of expectation regarding CGI technology. The strong impact of prior expectations and information on reality beliefs underlines the volatility of our sense of reality. In fact, stimuli-related and participant-related characteristics accounted together for <20 % of the beliefs variance, suggesting a large contribution of other subjective processes.

Although attractiveness did not seem to be the primary drive underlying simulation monitoring of face images, we do nonetheless report significant associations, with different patterns observed depending on the participant's gender. The quadratic relationship found for female participants is aligned with our hypothesis that salient faces (i.e., rated as very attractive or very unattractive) are judged to be more real. The fact that this effect did not reach significance for beauty underlines that attractiveness judgement, and its role in simulation monitoring, is a multidimensional construct that cannot be reduced to physical facial attractiveness, in particular for women (Buunk et al., 2002; Qi & Ying, 2022). In fact, female participants were more confident in judging faces as fake only when they were rated very high or low on beauty, suggesting that physical beauty and attractiveness are not analogous in their effects on simulation monitoring decisions.

Interestingly, we found a significant positive linear relationship in male participants for both attractiveness and beauty on simulation monitoring that we could interpret under an evolutionary lens. Specifically, males purportedly place more emphasis on facial attractiveness as a sign of reproductive potential, as compared with females, who tend to value characteristics signaling resource acquisition capabilities (Buunk et al., 2002; Fink et al., 2006; Qi & Ying, 2022). It is thus possible that the evolutionary weight associated with attractiveness skewed the perceived saliency of attractive faces for men, rendering them significantly more salient than unattractive faces, and in turn distorting the relationship with simulation monitoring. However, future studies should test this saliency-based hypothesis by measuring constructs closer to salience and its effects, for instance using neuroimaging (Indovina & Macaluso, 2007; Lou et al., 2015) or physiological markers (e.g., heart rate deceleration, Skora et al., 2022).

Our findings do not support the existence of a strong link between perceived trustworthiness and reality judgements. Given prior evidence that faces presented as computer-generated were rated less trustworthy (Balas & Pacella, 2017; Hoogers, 2021; Liefoghe et al., 2022), we expected such a linear association to be more clearly present. However, our results suggest a relationship with confidence ratings, especially for women, whereby faces judged with low and high trustworthiness are judged as real and fake with higher confidence. One of the underlying mechanisms that possibly contributed to this dimorphism could be the

increased risk-taking aversion reported in females (explained evolutionarily as a compromise to their reproductive potential, [Van Den Akker et al., 2020](#)), to which perceived facial trustworthiness relates ([Hou & Liu, 2019](#)). Future studies should clarify the role of trustworthiness both as a predictor and outcome of reality beliefs.

Contrary to our hypothesis, we did not find familiarity to be significantly related to simulation monitoring decisions. Interestingly, there were significant linear relationships between familiarity and confidence judgements for males only, where familiarity increased the confidence of reality beliefs. Although the familiarity measure was not a “recognition” measure, evidence from studies pertaining to the latter could be linked, reporting better face memory for females ([Lewin & Herlitz, 2002](#); [Mishra et al., 2019](#); [Sommer et al., 2013](#)), as well as an overconfidence in face recall for males ([Bailey, 2021](#); [Herbst, 2020](#)). However, it should be noted that the distribution of familiarity ratings was strongly skewed, and only a low number of pictures was rated as highly familiar. As such, future studies should clarify this point by experimentally manipulating familiarity, for instance by modulating the amount of exposure to items before querying the simulation monitoring judgements.

Regarding the role of inter-individual characteristics in simulation monitoring tendencies, we found higher scores of honesty-humility - a trait related to an increased risk perception and aversion ([Levidi et al., 2022](#); [Weller & Thulin, 2012](#)) - to be related to a lower confidence in simulation monitoring judgements. Notably, greater narcissistic tendencies in dimensions such as acclaim seeking were associated with a higher number of faces judged as real. This is in line with recent research which found people with higher narcissism scores less likely to engage in analytical reasoning strategies such as reflective thinking ([Ahadzadeh et al., 2021](#); [Littrell et al., 2020](#)), and to be more vigilant and attentive to external stimuli ([Carolan, 2017](#); [Eddy, 2021](#); [Grapsas et al., 2020](#)).

Moreover, putting the significant positive links between narcissistic acclaim seeking and confidence judgements in perspective with the negative correlation between honesty-humility and narcissism ([Hodson et al., 2018](#)), we confirm previous evidence regarding the relationship between narcissistic grandiosity and over-confidence in decision-making ([Brunell & Buelow, 2017](#); [Campbell et al., 2004](#); [Chatterjee & Pollock, 2017](#); [O'Reilly & Hall, 2021](#)). Although an inverse effect was found for the narcissistic facet of authoritativeness, we interpret this relationship as related to a higher response assertiveness. Taken together, these results suggest that participants with low humility and high recognition desires are more confident in their judgement regarding the real or fake nature of ambiguous stimuli. Alternatively, participants with opposite traits might perceive a higher risk in the decision-making process and its potential consequences (e.g., being seen as bad at the task at hand), resulting in more conservative confidence ratings.

Our findings suggest - though with low certainty - a potential positive link between paranoid ideation and the tendency to believe that the stimuli were real. Given previous reports that people with higher levels of paranoia are more sensitive to cues of social threat ([Fornells-Ambrojo et al., 2015](#); [Freeman et al., 2003](#); [King & Dudley, 2017](#)), it is plausible that paranoid traits confer greater saliency and emotionality to observed faces, hence increasing perceptions of its realness. This hypothesis, if confirmed by future studies, would be in line with previous findings that persecutory delusions are predicted by a greater sense of presence in VR environments populated with virtual characters ([Freeman et al., 2005](#)).

Despite the ubiquity of AI, the literature pertaining to the influence of people's AI attitudes on simulation monitoring is scarce. Contrary to our expectations, we did not find evidence for the role of participants' expectations regarding the capabilities of AI technology (in terms of the realism of its productions). Instead, we found only one's enthusiasm about AI technology to be related to an increased confidence in simulation monitoring ratings. This could potentially be because participants with a highly positive attitude towards AI perceive themselves as having greater knowledge about AI and its capabilities ([Said et al., 2022](#)), hence permitting themselves to be more confident in their simulation

monitoring decisions. In fact, this result is in line with reports that AI attitudes interacts with people's perceived self-knowledge to influence their perception of the opportunities and risks accorded by AI applications ([Said et al., 2022](#)).

On a methodological level, although the order of presentation of the facial images was randomized to reduce effects of adaptation, participants were more confident in their judgements for faces perceived as real following a shorter re-exposure delay. Such shorter durations could be associated with the faces being better remembered and appearing more familiar, thereby triggering self-referential and autobiographical memory processing during the repeated display ([Abraham & Von Cramon, 2009](#); [Gobbini et al., 2013](#); [Taylor et al., 2009](#)). Indeed, this finding is consistent with studies in which fictional stimuli that were associated with familiarity up-regulated emotions, biasing its salience and perceived realness ([Makowski et al., 2017](#); [Sperduti et al., 2016](#)). However, if that was the case, we would expect shorter re-exposure delays to impact the decision bias as well towards reality, rather than simply the confidence. Future studies should further investigate the modulatory effects of types and degrees of familiarity on perceived realness judgements.

Several limitations have to be noted. The current experimental paradigm required participants to judge the realness of faces they had prior exposure to (which was done to prevent reality judgements from influencing the other ratings). Although the effect of re-exposure delay was negligible, the potential bias induced by face familiarity, that is by re-presenting the same face stimuli twice, as compared to judging completely new items, cannot be discarded. Future studies could examine that by incorporating novel face images or increasing the duration of the re-exposure delay.

Another issue is the impact on reality judgements of the prior explicit instruction that “about half of the faces were AI-generated and the other half real photos”. Given this prior information given to participants, it might seem like our enthusiasm pertaining to the finding that most people did indeed believe a high number of stimuli to be fake might be unwarranted, since it simply affirms participants followed the instructions. However, even if that was the case, the finding that our beliefs of reality can be so easily re-programmed with simple instructions and lead to high-confidence answers remains an interesting phenomenon. Moreover, it is to note that the paradigm did not explicitly instruct participants to balance their answers according to a certain distribution (e.g., 50–50) - merely providing them a description of the dataset (but participants could, and in some cases did, deviate substantially from the information provided). The fact that no presentation order effect was found on reality beliefs suggests that participants did not try to actively distribute their responses to match the instructions, in which case we would have expected a different pattern: for instance, the first few items judged as real (the initial “true” belief of the participants), and a bias would progressively appear towards responding “fake” (as participants realize that all stimuli are of similar nature and that they have to “make up” for the prevalence of their “real” answers to fulfill the expected proportion of responses given the instructions).

That said, the potential demand effect of the instructions still exists, and a control condition without the cover story with AI-generated images would in-principle be able to mitigate such confounds to some extent. However, the distinction real/fake is hard to operationalize and introduce to participants in a vacuum (simply instructing them to discriminate real from fake without providing some background information regarding the context and defining what is meant by “fake” seems hardly feasible). That being the case, future studies should study the impact of these higher-order expectations on ratings (for instance, [Tucciarelli et al., 2020](#) found that merely mentioning that some faces were AI generated decreased, on average, the trustworthiness ratings for all faces) as well as on the simulation monitoring process itself (i.e., the “criterion”: would people form and distribute judgements differently). This can be studied by modulating this expectation in a controlled fashion (e.g., “most of the images but a few are real” vs. “most of the

images but a few are fake') or inventing some implicit way of measuring reality belief that would not require the explicit introduction of the concept of fake vs. real to participants.

Finally, it is important to note that although consistent in their directions across models and variables, the magnitude of the effects found in the study was relatively small, suggesting that the facial appraisals measured in the study were not the key determinants of simulation monitoring. Hence, beyond exploring new potential mechanisms, future studies should include a more thorough debriefing to try to capture what conscious strategies (if any) the participants used (e.g., focusing on some features of the stimulus - like hair or eyes in the case of faces) to guide their reality beliefs. Additionally, the role of specific facial features, like perceived dominance, warmth or gender, would be an interesting avenue to explore in future studies, in particular with paradigms directly manipulating these dimensions (for instance using AI to generate faces of different characteristics).

In summary, the aim of the present study was to examine whether a subset of specific characteristics, in particular face attractiveness, significantly influences our simulation monitoring decisions. Notably, we found faces rated as attractive to be perceived as more real, with a possible sexual dimorphism affecting the shape of the relationship. We also found that inter-individual traits, such as narcissistic acclaim-seeking, honesty-humility, and paranoid ideation, were related to a systematic bias towards beliefs that the stimuli were real or fake. We believe that these findings provide the foundations to help us understand what drives reality beliefs in an increasingly reality-ambiguous world.

CRediT authorship contribution statement

Dominique Makowski: Writing – review & editing, Writing – original draft, Visualization, Validation, Supervision, Software, Resources, Methodology, Investigation, Funding acquisition, Formal analysis, Data curation, Conceptualization. **An Shu Te:** Writing – original draft, Validation, Project administration, Methodology, Formal analysis, Data curation. **Ana Neves:** Writing – review & editing, Formal analysis. **Stephanie Kirk:** Writing – original draft, Project administration, Data curation. **Ngoi Zi Liang:** Writing – original draft, Project administration, Data curation. **Panagiotis Mavros:** Writing – review & editing, Funding acquisition. **S.H. Annabel Chen:** Writing – review & editing, Writing – original draft, Supervision, Project administration, Funding acquisition.

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Declaration of competing interest

None.

Data availability

The datasets generated and/or analyzed during the current study are available in the GitHub repository <https://github.com/RealityBending/FakeFace>.

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