

# MACS 201 : Hilbert spaces and probability

## 1 Hilbert spaces

**Def.** Let  $\mathcal{H}$  be a complex linear space. An **inner-product** on  $\mathcal{H}$  is a function  $\langle \cdot | \cdot \rangle : \mathcal{H} \times \mathcal{H} \rightarrow \mathbf{C}$  which satisfies the following properties :

- (i)  $\forall (x, y) \in \mathcal{H} \times \mathcal{H}, \langle x | y \rangle = \overline{\langle y | x \rangle},$
- (ii)  $\forall x, y, z \in \mathcal{H} \forall (\alpha, \beta) \in \mathbf{C} \times \mathbf{C}, \langle \alpha x + \beta y | z \rangle = \alpha \langle x | z \rangle + \beta \langle y | z \rangle,$
- (iii)  $\forall x \in \mathcal{H}, (\langle x | x \rangle = 0) \iff (x = 0)$

Then  $\|\cdot\| : x \mapsto \sqrt{\langle x | x \rangle} \geq 0$  defines a norm on  $\mathcal{H}$ . Both are continuous.

**Th.** For all  $x, y \in \mathcal{H}$ , we have :

- a) *Cauchy-Schwarz inequality* :  $|\langle x | y \rangle| \leq \|x\| \cdot \|y\|,$
- b) *triangular inequality* :  $|\|x\| - \|y\|| \leq \|x - y\| \leq \|x\| + \|y\|,$
- c) *Parallelogram inequality* :  $\|x + y\|^2 + \|x - y\|^2 = 2\|x\|^2 + 2\|y\|^2.$

**Def.** An inner-product space  $\mathcal{H}$  is called an Hilbert space if it is complete.

**Prop.** For all measured space  $(\Omega, \mathcal{F}, \mu)$ , the space  $L^2(\Omega, \mathcal{F}, \mu)$  endowed with  $\langle f | g \rangle = \int f \bar{g} d\mu$  is a Hilbert space.

**Def.** Two vectors  $x, y \in \mathcal{H}$  are orthogonal if  $\langle x | y \rangle = 0$  which we denoted by  $x \perp y$ . If  $\mathcal{S}$  is a subspace of  $\mathcal{H}$ , we write  $x \perp \mathcal{S}$  if  $\forall s \in \mathcal{S}, x \perp s$ . Also we write  $\mathcal{S} \perp \mathcal{T}$  if all vectors in  $\mathcal{S}$  are orthogonal to  $\mathcal{T}$ .

**Not.** If  $\mathcal{H} = \mathcal{A} + \mathcal{B}$  and  $\mathcal{A} \perp \mathcal{B}$  we will denote  $\mathcal{H} = \mathcal{A} \oplus \mathcal{B}$ .

**Def.** Let  $\mathcal{E}$  be a subset of an Hilbert space  $\mathcal{H}$ . The orthogonal set of  $\mathcal{E}$  is defined as  $\mathcal{E}^\perp = \{x \in \mathcal{H} \mid \forall y \in \mathcal{E}, \langle x | y \rangle = 0\}$ .

**Th.** If  $\mathcal{E}$  is a subset of an Hilbert space  $\mathcal{H}$ , then  $\mathcal{E}^\perp$  is closed.

### Orthogonal and orthonormal bases

**Def.** Let  $E$  be a subset of  $\mathcal{H}$ . It is an orthogonal set if for all  $(x, y) \in E \times E, x \neq y, x \perp y$ . If moreover  $\forall x \in E, \|x\| = 1$ , we say that  $E$  is orthonormal.

**Th.** Let  $(e_i)_{i \geq 1}$  be an orthonormal sequence of an Hilbert space  $\mathcal{H}$  and let  $(\alpha_i)_{i \geq 1} \in \mathbf{C}^{\mathbf{N}}$ . The series  $\sum_{i=1}^{\infty} \alpha_i e_i$  converges in  $\mathcal{H}$  if and only if  $\sum_i |\alpha_i|^2 < \infty$ , in which case  $\|\sum_{i=1}^{\infty} \alpha_i e_i\|^2 = \sum_{i=1}^{\infty} |\alpha_i|^2$ .

**Prop.** Let  $x \in \mathcal{H}$  (Hilbert space) and  $E = \{e_1, \dots, e_n\}$  a finite orthonormal set of vectors. Then  $\|x - \sum_{k=1}^n \langle x | e_k \rangle e_k\|^2 = \|x\|^2 - \sum_{k=1}^n |\langle x | e_k \rangle|^2 = \inf\{\|x - y\|^2, y \in \text{Span}(e_1, \dots, e_n)\}$ .

**Cor** (Bessel inequality). Let  $(e_i)_{i \geq 1}$  be an orthonormal sequence of a Hilbert space  $\mathcal{H}$ . Then  $\forall x \in \mathcal{H}, \sum_{i=1}^{\infty} |\langle x | e_i \rangle|^2 \leq \|x\|^2$ .

**Def.** A subset  $E$  of a Hilbert space  $\mathcal{H}$  is said dense if  $\overline{\text{Span}(E)} = \mathcal{H}$ . An orthonormal dense sequence is called a Hilbert basis.

**Prop.** Consider the measured space  $(\Omega, \mathcal{F}, \mu)$  and the Hilbert space  $\mathcal{H} = L^2(\Omega, \mathcal{F}, \mu)$ ,  $\overline{\text{Span}(\mathbf{1}_A, A \in \mathcal{F})} = \mathcal{H}$ .

**Th.** Let  $(e_i)_{i \geq 1}$  be a Hilbert basis of the Hilbert space  $\mathcal{H}$ . Then  $\forall x \in \mathcal{H}, x = \sum_{i=1}^{\infty} \langle x | e_i \rangle e_i$ .

**Th.** Let  $(e_i)_{i \geq 1}$  be an orthonormal sequence of the Hilbert space  $\mathcal{H}$ . The following assertions are equivalent :

- (i)  $(e_i)_{i \geq 1}$  is a Hilbert basis,
- (ii) if some  $x \in \mathcal{H}$  satisfies  $\forall i \geq 1, \langle x | e_i \rangle = 0$  then  $x = 0$ ,
- (iii)  $\forall x \in \mathcal{H}, \|x\|^2 = \sum_{i=1}^{\infty} |\langle x | e_i \rangle|^2$ .

**Th.** A Hilbert space  $\mathcal{H}$  is separable (i.e. contains a countable dense subset) if and only if it admits a Hilbert basis.

### Fourier series

Let  $\psi_n : x \mapsto \frac{1}{\sqrt{2\pi}} e^{inx}, n \in \mathbf{Z}$ . Let  $L^1(\mathbf{T})$  denote the set of  $2\pi$ -periodic locally integrable functions. For  $f \in L^1(\mathbf{T})$ , set  $\forall n \in \mathbf{N}, f_n = \sum_{k=-n}^n (\int_{\mathbf{T}} f \bar{\psi}_k) \psi_k$ .

**Th.** Suppose that  $f$  is a continuous  $2\pi$ -periodic function. Then the Cesaro sequence  $\frac{1}{n} \sum_{k=0}^{n-1} f_k$  converges uniformly to  $f$ .

**Cor.** Let  $\mu$  be a finite measure on the Borel sets of  $\mathbf{T} = \mathbf{R}/(2\pi\mathbf{Z})$ . The sequence  $(\phi_n)_{n \in \mathbf{Z}}$  is dense in the Hilbert space  $L^2(\mathbf{T}, \mathcal{B}(\mathbf{T}), \mu)$ .

**Cor.** The sequence  $(\phi_n)_{n \in \mathbf{Z}}$  is a Hilbert basis in  $L^2(\mathbf{T})$ . In particular,  $\forall f \in L^2(\mathbf{T}), f = \sum_{k=-\infty}^{\infty} \alpha_k \phi_k$  with  $\alpha_k = \frac{1}{\sqrt{2\pi}} \int_{\mathbf{T}} f(x) e^{-ikx} dx$  when the infinite sum converges in  $L^2(\mathbf{T})$ . The Parseval identity then reads  $\int_{\mathbf{T}} |f(x)|^2 dx = \sum_{k=-\infty}^{\infty} |\alpha_k|^2$ .

## Projection and orthogonality principle

**Th** (Projection theorem). Let  $\mathcal{E}$  be a closed convex subset of a Hilbert space  $\mathcal{H}$  and  $x \in \mathcal{H}$ . Then the following holds :

- (i) There exists a unique vector  $\text{proj}(x | \mathcal{E}) \in \mathcal{E}$  such that  $\|x - \text{proj}(x | \mathcal{E})\| = \inf_{w \in \mathcal{E}} \|x - w\|$ .
- (ii) If moreover  $\mathcal{E}$  is a linear subspace,  $\text{proj}(x | \mathcal{E})$  is the unique  $\hat{x} \in \mathcal{E}$  such that  $x - \hat{x} \in \mathcal{E}^\perp$ . It is called the orthogonal projection of  $x$  onto  $\mathcal{E}$ .

## 2 Probability

**Th** ( $\pi$  -  $\lambda$  theorem). If  $\mathcal{A} \subset \mathcal{C}$  with  $\mathcal{A}$  a  $\pi$ -system and  $\mathcal{C}$  a  $\lambda$ -system, then  $\sigma(\mathcal{A}) = \mathcal{C}$ .

**Th.** Let  $\mathcal{C}$  be a  $\pi$ -system on  $\Omega$  and  $\mathcal{F} = \sigma(\mathcal{C})$  the smallest  $\sigma$ -field containing  $\mathcal{C}$ . Then a probability measure  $\mu$  on  $(\Omega, \mathcal{F})$  is uniquely characterized by  $\mu(A)$  on  $A \in \mathcal{C}$ .

**Def.** Let  $X \in \mathcal{L}^1(\Omega, \mathcal{F}, \mathbf{P})$  and  $\mathcal{G}$  a sub- $\sigma$ -field of  $\mathcal{F}$ .

## 3 Mathematical statistics

### 3.1 Statistical modeling

**Def.** Let  $(\Omega, \mathcal{F})$  be a measurable space and  $\mathcal{P}$  a collection of probabilities on this space. Let  $X$  be a measurable function from  $(\Omega, \mathcal{F})$  to the observation space  $(\mathbf{X}, \mathcal{X})$ . We say that  $\mathcal{P}$  is a **statistical model** for the observation variable  $X$  and denote  $\mathcal{P}^X = (P^X)_{P \in \mathcal{P}}$  the corresponding collection of probability distributions.

It is usual in statistics to consider  $\Omega = \mathbf{X}, \mathcal{F} = \mathcal{X}$  and  $X(\omega) = \omega$ , in which case  $\forall P \in \mathcal{P}, P = P^X$ .

**Def.** Let  $\nu \in \mathbf{M}_+(\mathbf{X}, \mathcal{X})$  and  $\mathcal{P}$  be a statistical model for  $X$ . We say that  $\mathcal{P}$  is a  $\nu$ -dominated model for  $X$ , or that  $\mathcal{P}^X$  is  $\nu$ -dominated, if  $\forall P \in \mathcal{P}, P^X \ll \nu$ .

**Lem.** Let  $\nu \in \mathbf{M}_+(\mathbf{X}, \mathcal{X})$ . Consider a  $\nu$ -dominated model  $\mathcal{P}$  for the variable  $X$ . Then there exists a countable collection  $(P_n)_{n \geq 1}$  in  $\mathcal{P}$  such that  $\mathcal{P}^X$  is also dominated by  $\mu = \sum_{n \geq 1} 2^{-n} P_n^X$ .

**Def.** Let  $\mathcal{P}$  be a statistical model for the observation variable  $X$ . We say that  $\mathcal{P}$  is a **parametric model** for  $X$  if there exists a finite dimensional set  $\Theta$  such that  $\mathcal{P} = (P_\theta)_{\theta \in \Theta}$ .

**Def.** Let  $\mathcal{P}$  be a statistical model for  $X$ . Any finite dimensional quantity  $t(P^X)$  only depending on  $P^X$  as  $P \in \mathcal{P}$  is called an **identifiable parameter**.

**Def.** Let  $\mathcal{P}$  be a statistical model for  $X$ . A **statistic** in this context is any random variable  $T$  valued in  $(\mathbf{R}^d, \mathcal{B}(\mathbf{R}^d))$  with  $d \geq 1$ , defined by  $T = g(X)$  where  $g$  is a Borel function not depending on  $P \in \mathcal{P}$ .

If a statistic is used as a guess for a parameter  $t(P) \in \mathbf{R}^d$ , it is called an **estimator** of  $t(P)$ . In this case, the **bias** of  $T$  for estimating  $t(P)$  is defined as  $\text{Bias}(T, P) = \int T dP - t(P)$  whenever  $\int |T| dP < \infty$ . We say that  $T$  is an **unbiased estimator** of  $t(P)$  if  $\forall P \in \mathcal{P}, \int T dP = t(P)$ . The **quadratic risk** or **mean squared error** (in the case  $d = 1$ ) is defined by  $\text{MSE}(T, P) = \int (T - t(P))^2 dP = \text{Var}(T) + \text{Bias}(T, P)^2$ .

**Def.** Let  $T$  be a statistic valued in  $(\mathbf{R}^d, \mathcal{B}(\mathbf{R}^d))$  with  $d \geq 1$ . We say that  $T$  is a **sufficient statistic** for the model  $\mathcal{P}$  if, for all  $P \in \mathcal{P}$ , the conditional distribution of  $X$  given  $T$  does not depend on  $P$ , that is, there exists a probability kernel  $Q \subset \mathbf{R}^d \times \mathcal{X}$  such that, for all  $P \in \mathcal{P}, Q$  is a regular version of  $P^{X|T}$ .

**Lem.** Let  $S$  be a sufficient statistic associated to the Markov kernel  $Q$  and let  $T = g(X)$  be an unbiased estimator of the parameter  $t(P)$  (both real valued). Define  $T^R = \int g(x) Q(S, dx)$ . Then  $T^R$  is an unbiased estimator of the parameter  $t$  and its variance is smaller than that of  $T$ . As a consequence we have,  $\forall P \in \mathcal{P}, \text{MSE}(T^R, P) \leq \text{MSE}(T, P)$ .

**Th** (Fisher Factorization theorem). Let  $\nu \in \mathbf{M}_+(\mathbf{X}, \mathcal{X})$ . Consider a  $\nu$ -dominated model  $\mathcal{P}$  for  $X$  and let  $S = g(X)$  be a  $d$ -dimensional statistic. Then  $S$  is a sufficient statistic for the model  $\mathcal{P}$  if and only if there exists a non-negative Borel function  $h$  on  $\mathbf{X}$  such that  $\forall P \in \mathcal{P}$ , there exists a Borel function  $f_P : \mathbf{R}^d \rightarrow \mathbf{R}_+$  such that  $\frac{dP^X}{d\nu} = h \cdot f_P \circ g$ .