

Renewal Theory

(1) Recurrent Times:

① Forward:

Def: For $\varphi = (t_n)$ renewal process. $t_{N(t)} \leq t \leq t_{N(t)+1}$
the forward recurrent time is $A(t) =$

$$t_{N(t)+1} - t. \quad \left(\begin{array}{c} t_n \\ \text{---} \\ t \end{array} \xrightarrow{\text{---}} \begin{array}{c} t_{n+1} \\ \text{---} \end{array} \right) A(t)$$

Ex: If $\varphi \sim \text{Pois}(\lambda)$. Then $A(t) \sim \text{Exp}(\lambda) : P(A(t) \leq x)$

$$= 1 - P(T_{N(t)+1} > x+t) = 1 - P(N(t, t+x] = 0) = 1 - e^{-\lambda x}$$

Denote: $X = (X_n)$ i.i.d interarrival times. $X \sim F$

prop. i) $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t A(s) ds = E(X^2) / 2E(X) \cdot n.s.$

ii) $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t I_{\{A(s) > a\}} ds = \lambda E(X - a)^+ \cdot n.s.$

iii) $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t P(A(s) > a) ds = \lambda E(X - a)^+.$

Pf: 1) Intuitively, $\int_0^t A(s) ds \approx \sum_{i=1}^{N(t)} \frac{X_i^2}{2}.$

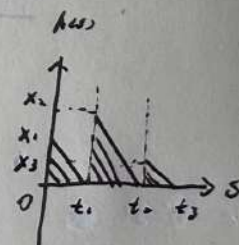
Actually, we have:

$$\frac{1}{t} \sum_{i=1}^{N(t)} \frac{X_i^2}{2} \leq \frac{1}{t} \int_0^t A(s) ds \leq \frac{1}{t} \sum_{i=1}^{N(t)+1} \frac{X_i^2}{2}.$$

Then apply ERT. as before

2) $\int_0^t I_{\{A(s) > a\}} ds \approx \sum_{i=1}^{N(t)} (X_i - a)^+.$

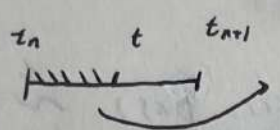
3) By DCT. in the expectation.



② Backward:

Def: For $\varphi = \{t_n\}$ renewal process. $t_{n-1} \leq t \leq t_n$.

$B(t) = t - t_{n-1}$ is backward recurrence time.

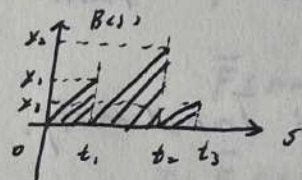
 $(B(t) \sim F(x) = \begin{cases} 1 & x > t \\ 1 - e^{-\lambda x} & 0 \leq x \leq t \end{cases} \text{ if } \varphi \sim \text{P.O.I.})$

prop, i) $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t B(s) \lambda ds = E(X) / 2 E(X)$. a.s.

ii) $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t I_{\{B(s) > a\}} \lambda ds = \lambda E(X - a)^+$. a.s.

iii) $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(B(s) > a) \lambda ds = \lambda E(X - a)^+$.

pf: Similar as before:



(2) Distribution:

① Equilibrium Dist.:

Denote: $\bar{F}(x) = 1 - F(x) = p(X > x)$. tail prob.

Rmk: $\lambda E(X - a)^+ = \lambda \int_a^\infty p(X > \eta) \lambda \eta$
 $= \lambda \int_a^\infty \bar{F}(\eta) \lambda \eta.$

Def: As $a \geq 0$ varies. Rmk defines a tail prob. of a dist. We denote cdf of it by F_e which is called equilibrium dist. of F .

$F_e(x) =: \lambda \int_0^x \bar{F}(\eta) \lambda \eta.$ Set $X_e \sim F_e$ r.v.

Rmk: X_e is anti. since density $= F_e' = \lambda \bar{F}$ exists.

Note that in a long term:

$$\lim_{t \rightarrow \infty} \int_0^t \frac{p(A(s) < x)}{t} ds = \lim_{t \rightarrow \infty} \int_0^t \frac{p(B(s) < x)}{t} ds = 1 - \lambda \overline{E}(X - x)^+ \\ = \lambda \int_0^x \overline{F}(\eta) d\eta = F_e(x).$$

\Rightarrow The stationary dist of $A(s), B(s) \stackrel{d}{\sim} F_e$

② Spread:

Def: Spread as length of interarrival time is

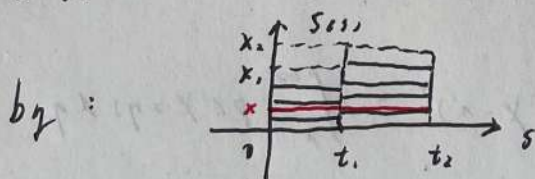
$$S(t) =: t_{\mu_{n+1}} - t_{\mu_n} = A(t) + B(t) = X_{\mu_{n+1}}$$

Remark: Immediately, $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t S(s) ds = \overline{E}(X) / \overline{E}(X)$.

Prop. i) $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t I_{\{S(s) > x\}} ds = \lambda \overline{E}(X I_{\{X > x\}})$ a.s.

$$ii) \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(S(s) > x) ds = \lambda \overline{E}(X I_{\{X > x\}})$$

Pfc: $\frac{1}{t} \int_0^t I_{\{S(s) > x\}} ds \approx \frac{1}{t} \sum_{i=1}^{N(t)} X_i I_{\{X_i > x\}}$



Remark: In a long term:

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(S(s) < x) ds = \lambda \overline{E}(X I_{\{X < x\}})$$

$$\text{Denote: } \overline{F}_s(x) = \lambda \overline{E}(X I_{\{X > x\}})$$

$$= \lambda \int_x^\infty (t-x)^+ + x dF$$

$$= \lambda \overline{E}(X - x)^+ + \lambda x \overline{F}(x).$$

$$= \overline{F}_e(x) + \lambda x \overline{F}(x).$$

by prop above. the stationary dist. of
 $S(x) \sim 1 - \bar{F}_S(x) = F_S(x)$. Assume r.v. X_s .

$\Rightarrow X_s$ may not be conti. generally. (F' doesn't exist)
 $\bar{E}(X_s) = \bar{E}(X) / \bar{E}(x)$.

(3) Inspection Paradox:

prop. $S(t) \geq_{\text{stoch}} X$. i.e. $p(X > x) \leq p(S(t) > x)$. $\forall t, x \geq 0$.

Moreover. $p(X_s \geq x) \geq p(X \geq x)$. $X_s \geq_{\text{stoch}} X$.

Pf: 1') $p(S(t) > x \mid N(t) = n, t_n = s) = p(X_{n+1} > x \mid X_{n+1} > t-s)$
 $= \bar{F}(\max\{x, t-s\}) / \bar{F}(t-s)$
 $\geq \bar{F}(x)$.

$\Rightarrow \bar{E}_{N, t_N}(p(S(t) > x \mid \square)) = p(S(t) > x) \geq \bar{F}(x)$

2') By $p(X_s \geq x) = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(S(s) > x) ds \geq p(X \geq x)$

Rmk. Note that: $\bar{E}(S(t)) \geq \bar{E}(X)$

$\bar{E}(X_s) \geq \bar{E}(X)$

\Rightarrow The prop. above implies a paradox: Observing makes the expectation of lifetime be longer than usual.

That's because if we observe at time t , then the part of lifetime $< t$ will be ignored.

e.g. A extreme case:

X is r.v. of lifetime of bulbs. $\begin{cases} p(X=0) = 0.9 \\ p(X=1) = 0.1 \end{cases}$

Then. all bulbs we observed has lifetime 1.

The right way to estimate lifetime of them

is SLLN: $\frac{1}{n} \sum_{i=1}^n X_k \approx E(X)$. for large n .

(4) Renewal Reward Thm:

Suppose R_i is r.v. of reward between t_i, t_{i+1} .

$\Rightarrow R(t) = \sum_{i=1}^{N(t)} R_j$ Assume $(X_i, R_i)_{i \in \mathbb{Z}^+}$ i.i.d. but R_i

can depend on X_i . (X, R) denote a cycle.

Thm. (RRT)

For a positive recurrent renewal process

where a reward R_i is earned during a

cycle with length X_i . If $E(|R_i|) < \infty$.

Then: $\lim_{t \rightarrow \infty} \frac{R(t)}{t} = E(R)/E(X)$ a.s.

$$\lim_{t \rightarrow \infty} \frac{E(R(t))}{t} = E(R)/E(X).$$

p.f. 1°) Suppose $R_i \geq 0$. or set $R_i = R_i^+ - R_i^-$.

$$N(t): \frac{1}{t} \sum_{i=1}^{N(t)} R_i \leq \frac{R(t)}{t} \leq \frac{1}{t} \sum_{i=1}^{N(t)} R_i.$$

Apply ERT.

$$2^\circ) |R(t)|/t \leq Y(t) = \frac{1}{t} \sum_{i=1}^{N(t)} |R_i| \xrightarrow{\text{a.s.}} E(|R|)/E(X)$$

$\Rightarrow |R(t)|/t$ is bdd a.s. By DCT.

Application: We can see some byproduct as "reward" to calculate its long time rate by RRT.

i) Let $R(t) = \int_0^t A(s) ds$. Conti reward.

$$R_j = \int_{t_{j-1}}^{t_j} A(s) ds = \int_{t_{j-1}}^{t_j} (t_j - s) \lambda ds$$

$$\stackrel{s=t_{j-1}+u}{=} \int_0^{X_j} (X_j - u) \lambda du = \frac{X_j^2}{2}$$

$$\Rightarrow \text{By RRT: } \frac{1}{t} \int_0^t A(s) ds \xrightarrow{t \rightarrow \infty} E\left(\frac{X^2}{2}\right) / E(X).$$

ii) Similarly, $R(t) = \int_0^t B(s) ds$.

iii) $R(t) = \int_0^t I_{\{A(s) > x\}} \lambda ds$. $R_i = \int_0^{X_i} I_{\{X_i - s > x\}} \lambda ds$

(5) Central Limit Thm:

Consider renewal process $\{t_n\}$ with interarrival times

$$X_n = t_n - t_{n-1}, \text{ s.t. } E(X) = \frac{1}{\lambda}, \text{ Var}(X) = \sigma^2.$$

Thm. CCLT for counting process

$$Z(t) = \frac{N(t) - \lambda t}{\sigma \sqrt{\lambda^3 t}} \xrightarrow[t \rightarrow \infty]{} Z \sim N(0,1).$$

RMk: It implies: $E(N(t)) \sim \lambda t$. $\text{Var}(N(t)) \sim \sigma^2 \lambda^3 t$

Pf: The key is: $P(N(t) < n) = P(t_n > t)$.

$$\text{Let } n(t, x) = \lfloor \lambda t + x \sqrt{\sigma^2 \lambda^3 t} \rfloor.$$

$$\Rightarrow P(Z(t) < x) = P(N(t) < n(t, x)) \\ = P(t_{n(t, x)} > t)$$

$$= P\left(\frac{t_{n(t,x)} - \mu(t,x)/\lambda}{\sigma \sqrt{\mu(t,x)}} > \frac{t - \mu(t,x)/\lambda}{\sigma \sqrt{\mu(t,x)}} \right)$$

Note by CLT, $\frac{t_n - \mu/\lambda}{\sigma \sqrt{\mu}} \xrightarrow{d} Z \sim N(0,1)$.

$$\text{So: } P(Z(t) < x) \xrightarrow{t \rightarrow \infty} P(Z > -x) = P(Z < x)$$

Remark: We don't need "Renewal" actually,
but require: $\varphi = \{t_n\}$ satisfies CLT.

(6) Delayed Renewal Process:

Note that the counting process $N(t)$ of renewal process $\varphi = \{t_n\}$ generally doesn't have stationary increment like poisson process.

Def: φ_s is a shift by time s version of point process $\varphi = \{t_n\}$, which is $\{t_n + s\}$, moving the origin to be $t=s$. with counting process;
 $N_s(t) = N(s+t) - N(s)$.

Remark: $\varphi_s \stackrel{d}{\sim} \varphi$. $\forall s \geq 0 \Leftrightarrow N(t)$ has stationary increment.

Def: A delayed renewal process is a renewal process where the first arrival time $t_1 = X_1$ indeptly has a different dist. $\sim F$.
For $(X_n)_{n \geq 2}$ interarrival times. are i.i.d. $\sim F$

Remark: 1) ERT remains valid. F doesn't need

to have finite first moment.

ii) φ is stationary renewal process $\Leftrightarrow t, (s) = A(s)$
of φ_s has the same list. $\forall s \geq 0$.

As $s \rightarrow \infty$. φ_s has a limiting list. of the delayed
version φ . denoted by $\varphi^* = \{t_n^*\}_{n \geq 1}$. $t_i^* \sim F_e$.
With its counting process $N^*(t)$. forward recurrent
time process $\{A^*(t)\}_{t \geq 0}$.

Def. φ^* is stationary version of φ .

Prop. φ^* is stationary renewal process:

$$p(A^*(u) \leq x) = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(A(s+u) \leq x) ds \\ = \lim_{t \rightarrow \infty} \frac{1}{t} \int_u^{t+u} p(A(s) \leq x) ds = F_e(x)$$

$$\Rightarrow A^* \sim F_e. \text{ i.e. } t_i^* \sim F_e. \varphi_s^* = \varphi^*.$$

prop. Stationary version φ^* of renewal process with
rate $\lambda = E(X)^{-1}$ satisfies: $E(N^*(t)) = \lambda t$. and
 $\lambda = E(N^*(1))$.

Pf. Denote $\mu(t) = E(N^*(t)) \Rightarrow \mu(n) = n\mu(1). \forall n$.

So: $\mu(t) = t\mu(1)$. With $\mu(t)/t \rightarrow \lambda$. So $\mu(1) = \lambda$.

follows from ERT. $\lim_{t \rightarrow \infty} \lim_{s \rightarrow \infty} \square = \lim_{s \rightarrow \infty} \lim_{t \rightarrow \infty} \square$

(7) Renewal Equations:

For renewal process $\varphi = \{t_n\}$ with $p(X \leq y) = F(y)$.

$$\begin{aligned}
 \text{Note: } p(A(t) > x) &= p(A(t) > x | X_1 > t) p(X_1 > t) + \int_0^t p(\square | X_1 = s) p(X_1 = s) ds \\
 &= p(X_1 - t > x) + \int_0^t p(A(t-s) > x) dF(s) \\
 &= \bar{F}(t+x) + \int_0^t p(A(t-s) > x) dF(s)
 \end{aligned}$$

Denote: $\bar{F}(x+t) = Q(t)$. $H(t) = p(A(t) > x)$.

$\Rightarrow H(t) = Q(t) + H * F(t)$. renewal equation.

By Iteration: $H(t) = Q + (Q + H * F) * F$

$$= \dots = Q + \sum_{n=1}^{\infty} Q * F^{*n}. \quad H \text{ is unknown}$$

prop. $m(t) = E(N(t))$ satisfies: $H(t) = Q + Q * m(t)$. (*)

Pf: $E(N(t)) = E\left(\sum_i I_{[t_i, t_{i+1})}\right) = \sum F^{*n}$.

since $F^{*n}(t) = p\left(\sum_{i=1}^n X_i \leq t\right) = p(t_n \leq t) \xrightarrow{n \rightarrow \infty} 0$

Thm. (key renewal Thm)

If the renewal equation holds for given non-lattice F with mean $1/\lambda$ and Q is DRI (directly Riemann integrable, i.e. $\int_0^\infty Q(t) dt$ exists)

Then solution for (*) holds. $\lim_{t \rightarrow \infty} H(t) = \lambda \int_0^\infty Q(t) dt$.

Rmk: These remain valid for delayed case:

in sense, $F_0 = F$. F_1 is delay X_1 's dist. then.

$$H_1 = Q_1 + H_0 * F_1, \quad H_0 = Q_0 + Q_0 * m_0, \quad m_0 = E(N(t))$$

$$\text{With } H_1 = Q_1 + Q_0 * m_1, \quad m_1 = E(N(t))$$

$$\text{And } \lim_{t \rightarrow \infty} Q_0 * m_1 = \lim_{t \rightarrow \infty} Q_0 * m_0 = \lambda \int_0^\infty Q_0(t) dt.$$

Pf: It's equi. with Blackwell's Thm