

# Conti. Time Markov Chain

## (1) Definitions:

① Def: A stochastic process  $(X_t)_{t \geq 0}$  with discrete state space  $S$  ( $|S| = S$ ) is CTMC if  $P(X_{s+t} = j | X_s = i, (X_u)_{0 \leq u < s}) = P(X_{s+t} = j | X_s = i) = P_{ij}(t)$ . in dept with  $s$ .  $\forall i, j \in S$ .

Remark: i) Trans prob. matrix  $P(t) = (P_{ij}(t))_{S \times S}$ .  $P(0) = I$ .

ii)  $H_i$  is the holding time for  $i$  when it enter state  $i$ .

Remark:  $H_i$  has memoryless property. So it's exponential dist.  $\exists (\lambda_i)$  s.t.  $H_i \sim \text{Exp}(\lambda_i)$ .

A CTMC can be described by:

i) Transition Matrix  $P = (P_{ij})_{S \times S}$ . Describes how chain changes at transition epochs. ( $P_{ii} = 0$ )

ii) Set of transition rates  $(\lambda_i)_{i \in S}$

e.g. If  $X(t) = i$ . Then, "next time it changes" has prob.  $\lambda_i t$

Def:  $Q = P'(0)$ . infinitesimal generator of  $(X_t)_{t \geq 0}$ .



prop.  $Q = P'(0) = (P'_{ij}(0))_{S \times S}$  satisfies:  $\forall i \neq j \in S$ .

$$P'_{ij}(0) = a_i P_{ij}, \quad P'_{ii}(0) = -a_i, \quad \text{So: } \sum_{j \in S} P'_{ij}(0) = 0$$

pf:  $(N_i(t))_{t \geq 0}$  is Poisson counting process with rate  $a_i$ .  
recall it has "little occ" property.

$$\begin{aligned} P'_{ij}(0) &= \lim_{h \downarrow 0} P_{ij}(h)/h = \lim_h P_{ij}(P(N_i(h)=1))/h \\ &= P_{ij} \lim_{h \downarrow 0} (a_i h + o(h))/h = a_i P_{ij}. \end{aligned}$$

$$\begin{aligned} P'_{ii}(0) &= \lim_h (P_{ii}(h) - 1)/h = - \lim_h \frac{P_i(X(h) \neq i)}{h} \\ &= \lim_h - \frac{P(N_i(h)=1)}{h} = -a_i. \end{aligned}$$

② Def: Let  $(z_n)$  is seq of times that transition occurs.  
 $(X_n) = (X(z_n)) \hookrightarrow (X(t))$  is embedded DTMC  
from  $(X(t))_{t \geq 0}$  with transition matrix  $P = (P_{ij})$

prop. (Chapman-Kolmogorov Equation)

$$P_{ij}(t+s) = \sum_{k \in S} P_{ik}(t) P_{kj}(s), \quad \text{i.e. } P(t+s) = P(t)P(s)$$

Def: A CTMC is explosive if the transition in a finite time is infinite.

e.g.,  $P_{i,i+1} = 1, \forall i \geq 0, \quad a_i = 2^i, i \geq 0, \quad S = \mathbb{Z}^+ \cup \{0\}.$

$$E(\sum_i H_i) = \sum 1/a_i < \infty.$$

thm:  $\sum_i 1/a_i < \infty \Rightarrow$  It's nonexplosive.

### ③ Queue Models:

#### i) FIFO m/m/1:

"FIFO" means "first in queue first out of queue".

suppose Arrivals to the queue are Poisson at rate  $\lambda$ . ( $S_n$ ) service times  $\overset{i.i.d}{\sim} \text{Exp}(\mu)$ .

$X(t)$  is number of customers in the system.

##### 1) Holding time:

$$H_0 \sim \text{Exp}(\lambda), \quad H_i = \min \{S_i, X\} \sim \text{Exp}(\lambda + \mu).$$

$X$  is time until arrival. (Note:  $p(S_i = X) = 0$ )

##### 2) Transition Prob.:

$$p(0,1) = \lambda, \quad p(i, i+1) = p(X < S_i) = \lambda / (\lambda + \mu).$$

$$p(i, i-1) = \mu / (\lambda + \mu).$$

$$\text{From: } a_i = \begin{cases} \lambda, & i=0 \\ \lambda + \mu, & i \geq 1. \end{cases} \Rightarrow \text{Obtain } Q = P'(\infty).$$

#### ii) m/m/c:

##### 1) Holding times:

$$H_0 \sim \text{Exp}(\lambda), \quad H_1 = \min \{X, S_1\} \sim \text{Exp}(\lambda + \mu)$$

$$H_2 = \min \{X, S_1, S_2\} \sim \text{Exp}(\lambda + 2\mu), \dots$$

$$\Rightarrow H_i \sim \text{Exp}(\lambda + i\mu), \quad 0 \leq i \leq c, \quad H_i \sim \text{Exp}(\lambda + c\mu), \quad i \geq c.$$

##### 2) Transition prob.:

$$p(i, i+1) = p(X < \min_{1 \leq k \leq i} S_k) = \lambda / (\lambda + i\mu), \quad 0 \leq i \leq c.$$

$$p(i, i+1) = \lambda / (\lambda + c\mu), \quad i \geq c.$$



iii) M/M/∞:

$$H_i \sim \text{Exp}(\lambda + i\mu) \quad P_{i,i+1} = \lambda / (\lambda + i\mu) \quad \forall i \geq 0.$$

iv) Birth and Death Process:

Its  $P = (P_{ij})$  satisfies:  $P_{i,i+1} + P_{i,i-1} = 1$ .  $S = \mathbb{Z}^+ \cup \{0\}$ .

$$P_{i,i+1} = P(B_i < D_i) = \lambda_i / (\lambda_i + \mu_i). \quad B_i, D_i \text{ are time until}$$

Birth or death when there is population.  $B_i \sim \text{Exp}(\lambda_i)$ .

$D_i \sim \text{Exp}(\mu_i)$ . It's like M/M/1 model.

(2) Limit Theory:

Def: i)  $i \in S$  is recurrent / commutative with  $j \in S$  / accessible from  $j \in S$  if it holds in embedded DTMC.

ii)  $i \in S$  is positive recurrent if  $E(T_{ii}) < \infty$ .

Rmk: i) CTMC is positive recurrent  $\Leftrightarrow$  Embedded DTMC is positive recurrent.

ii) Positive recurrent is still a class property.

iii)  $P_j = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t P_{ij}(s) ds \in \mathbb{R}^+$  exists by RRT.

If it's finite. Then, let  $\vec{P} = (P_1, \dots, P_n)$  stationary list. for CTMC.  $P^* = \begin{pmatrix} \vec{P} \\ \vec{P} \end{pmatrix} = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(s) ds$ .

is called limiting prob. list.

Prop. If  $(X(t))_{t \geq 0}$  is positive recurrent CTMC.

Then  $p^*$  exists, unique.  $P_j = 1/\alpha_j E(T_{jj})$

Pf: Apply RRT:

$$L(Z_n) = L(T_{jj}^n). \quad R_n = \int_{Z_n}^{Z_{n+1}} I\{X(s) = j\} ds.$$

$N_j(t)$  is counting process for  $L(Z_n)$ .

$$N_j(t)/t \rightarrow 1/E(T_{jj}). \text{ a.s. by ERT.}$$

Combined with  $R_n \stackrel{a.s.}{\sim} M_j$ .

Cor. Under the condition above:  $P_j = \lim_{t \rightarrow \infty} P_{ij}(t)$ .

Cor. If  $X$  is null recurrent. Then  $p^*$  doesn't exist.  
and  $p_i = 0, \forall i \in S$

Rmk: If  $X_0 \stackrel{a.s.}{\sim} \vec{V}$ , initial dist. Then:

$X(t) \stackrel{a.s.}{\sim} \vec{V}(t)$ . From above. We have:

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t \vec{V}(s) ds = \vec{V} p^* = \vec{P}. \text{ So it}$$

means  $\vec{P}$  (or say  $p^*$ ) indept with  $\vec{V}$ .

### ① Stationary Version:

prop. For a positive recurrent CTMC with limiting

dist.  $P$ . If  $X_0 \sim \vec{P}$ . Then  $X(t) \sim \vec{P}, \forall t$ .

i.e.  $\vec{P} \cdot P(t) = \vec{P}, (\sum_i P_i P_{ij}(t) = P_j, \forall j \in S.)$

Rmk:  $\vec{P}$  is unique: if  $\vec{V} \cdot P(t) = \vec{V}$ , then:

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t \vec{V} \cdot P(s) ds = \vec{P} = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t \vec{V} ds = \vec{V}.$$



$$\underline{\text{Pf:}} \quad p^*(p(t)) = \lim_{n \rightarrow \infty} \int_0^u \frac{1}{n} p(s) p(t+s) ds$$

$$= \lim_{n \rightarrow \infty} \int_0^u \frac{1}{n} p(s+t) ds \quad (\text{by C-K equation})$$

$$= p^*$$

Define: Set  $X^* = \{X^*(t)\}_{t \geq 0}$  is the CTMC with  $X^*(0) \sim P$ , the limiting dist. stationary version.

Rmk:  $X_s^* \sim X^*$  is only to see.  $\forall s \geq 0$  shift.

## ② Equations:

### i) Thm. (Kolmogorov Backward Equation)

For CTMC with infinitesimal generator  $Q = P'(0)$ .

We have:  $P'(t) = Q \cdot P(t)$  holds.  $\forall t \geq 0$ . Besides,

the unique solution is:  $P(t) = e^{Qt} =: \sum_{n=0}^{\infty} \frac{t^n Q^n}{n!}$

Rmk: i) It means  $P(t)$  is determined by  $Q$ .

ii) For forward equation:  $P'(t) = P(t)Q$ .

it will cause some problems on interchange the sum and limit.

But  $e^{Qt}$  is the common solution.

$$\underline{\text{Pf:}} \quad P(t+h) - P(t) = (P(h) - I)P(t) \\ = (P(h) - P(0))P(t).$$

### ii) Balance Equations:

Note that every time  $X(t)$  want to enter state  $i$ , then it must leave  $i$  state first.  $\Rightarrow$  The number of entering  $i$  diff the number of leaving  $i$  at most one.

Claim: the long run rates of these two will coincide.

Def: The balance equation for positive recurrent CTMC is  $\vec{P} \alpha = 0$ ,  $\alpha = P'(0)$ .

Rmk:  $\vec{P} \alpha = 0 \Rightarrow a_i p_i = \sum_{j \neq i} p_j a_j p_{ji}$ .

LHS is rate of leaving  $i$ .

RHS is rate of entering  $i$ .

Thm. A nonexplosive irreducible CTMC is positive recurrent  $\Leftrightarrow \exists$  unique probability solution  $\vec{P}$  for  $\vec{P} \cdot \alpha = 0$ , i.e.  $P_i > 0$ , and it's limiting dist. for CTMC.

Pf:  $(\Rightarrow) \exists \vec{P}$  is stationary dist.:

$$\vec{P} \cdot p(t) = \vec{P} \Leftrightarrow \vec{P} \cdot \frac{p(t) - I}{t} = 0, t \rightarrow 0$$

$$(\Leftarrow) \text{ Show } = \vec{P} \cdot p(t) = \vec{P}.$$

by backward equation:

$$\vec{P} \alpha p(t) = \vec{P} \cdot p'(t) = 0 = \frac{d}{dt} (\vec{P} \cdot p(t))$$

$$\therefore \vec{P} \cdot p(t) = \vec{P} \cdot p(0) = \vec{P}$$



Then it follows from a Lemma:

Lemma. For a CTMC. If it has stationary dist.

$\vec{p}$ . i.e.  $\vec{p} \cdot P(t) = \vec{p}$ . Then it's positive recurrent.

Pf:  $p_j = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p_{ij}(s) ds = 0 \Rightarrow \vec{p} = 0$ . Contradict!

if we assume it's null recurrent.

Thm. An irreducible CTMC with finite state space  $S$  is always positive recurrent.

Pf: Note that the embedded DTMC is irred.

So it's positive recurrent.

Denote:  $S = \{1, 2, \dots, b\}$ .  $z_n$  is return time to state 1 for DTMC.  $T_n$  is for CTMC.

Set  $Y_n \stackrel{\text{i.i.d.}}{\sim} \text{Exp}(\min\{a_1, \dots, a_b\})$ , where the holding time  $H_i \sim \text{Exp}(a_i)$ .

$$\Rightarrow E(T_n) \leq E\left(\sum_{i=1}^{z_n} Y_i\right) = E(z_n) / \min a_i < \infty$$

Rmk: When using Balance Equations to solve  $(P_i)$

stationary dist. if  $S$  is infinite. We need

use:  $a_i P_i = \sum_{j \neq i} P_j a_j P_{ji}$  rather than matrix form. (use it recursively)

③ General Case:



We treat  $P_{ii} = 0$  of transition matrix in the discussion above. But in general, we can set  $P_{ii} \in (0, 1)$ .

1) Set  $K$  is r.v. of total number of visiting state  $i$  before transitioning to state  $j \neq i$ .  $K \sim \text{Geocp}$ ,  $p = 1 - P_{ii}$ .

2) Renew holding time for  $i$ :  $\tilde{H}_i = \sum_1^K M_i^n$ .

Calculation by ch.f:  $(M_i^n \stackrel{\text{i.i.d.}}{\sim} M_i)$

$\Rightarrow \tilde{H}_i \stackrel{\Delta}{\sim} \text{Exp}(p_i)$

3) Reset  $\tilde{\pi}_i = p_i$ ,  $\tilde{P}_{ii} = 0$ ,  $\tilde{P}_{ij} = P_{ij} / p$ .

Then we obtain the reduced form.

### (3) PASTA:

"PASTA" refers to "Poisson Arrival See Time Average".

Consider:  $M/M/1$  queue model:

Denote:  $Z_j^a$  is long-run proportion of a arrival customer finding there're  $j$  customers in the system, i.e.

$$Z_j^a = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_1^n \mathbb{I}_{\{X(t_n) = j\}}. (t_n) \text{ is the arrival times.}$$

We can see:  $\lambda Z_j^* = \lambda P_j$ .

$B_j$ : LHS is long-run rate of  $j \rightarrow j+1$ . And  $RHS$  is long-run rate of arrival happens when  $j$  customers are in the system.

$\Rightarrow$  Proportion of Poisson Arrivals who see  $j$  customers in the system is equi. with the proportion of time when there're  $j$  customers in

Def: A Poisson process  $\varphi = \{t_n\}$  satisfies LAC (Law of Anticipation Condition) if for  $N(t)$ :  
 $\{N(t+s) - N(t)\}_{s \geq 0}$  indep't with  $\{N(u), X(u)\}_{0 \leq u \leq t}$ .

Thm. If poisson process  $\varphi = \{t_n\}$  satisfies LAC for  $\{X(t)\}_{t \geq 0}$ . Then we have a.s.:

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n f(X(t_{n-i})) = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t f(X(s)) ds. \text{ if}$$

one of limits exists, finite.

Rmk: Set  $f(x) = I_{\{x=j\}}$ . Then we obtain

the conclusion in M/M/1 above.