

Appendix 4: Test of the robustness of the T-statistics using simulations

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Abstract

Community ecologists are actively describing species by their functional traits, quantifying the functional structure of plant and animal assemblages, and inferring community assembly processes with null model analyses of trait distributions and functional diversity indices. Intraspecific variation in traits and effects of spatial scale are potentially important in these analyses.

// We introduce the R package *cati*, available on CRAN, for the analysis of community assembly with functional traits. The *cati* package (i) calculates a variety of single-trait and multi-trait indices from interspecific and intraspecific trait measures; (ii) partitions functional trait variation among spatial and taxonomic levels; (iii) implements a palette of flexible null models for detecting non-random patterns of functional traits. These patterns can be used to draw inferences about hypotheses of community assembly such as environmental filtering and negative species interactions.

// The basic input for *cati* is a data frame in which columns are traits, rows are species or individuals, and entries are the measured trait values. The *cati* package can also incorporate into analyses a square distance matrix, which could include phylogenetic or genetic distances among individuals or species. Users select from a variety of functional trait metrics and analyze them relative to a null model that specifies trait distributions in a regional source pool.

Key words: Functional space, functional structure, community assembly, ecological niche, environmental filter, individual differences, intraspecific variation, null model, trait, variance decomposition

To read a summary of this appendix, see directly section Summary 7. An up to date version of the *cati* tutorial is available [here](#).

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Contents

1	Introduction	4
1.1	R requirements	4
1.2	System and session informations	4
1.3	T-statistics and associated null models	5
1.4	Simulations parameters	7
2	No filter: calculation of error	10
2.1	Randomization outline	10
2.2	Results of the simulations without filtering	10
2.3	Assessing the statistical type I error of local T-statistics	11
2.4	Assessing the statistical type I error of global T-statistics	13
2.5	Conclusion on the type I error of T-statistics	17
3	Internal filter: assessing the statistical power of $T_{IP/IC}$	20
3.1	Randomization outline	20
3.2	Results of the simulations of internal filtering	22
3.2.1	Local p-values results	22
3.2.1.1	Calculation of beta-error mixing all strength of filtering	25
3.2.1.2	Local $T_{IP/IC}$ SES values against initial parameter values	25
3.2.1.3	Local $T_{IP/IC}$ SES values against modeled parameter values	28
3.2.2	Global p-values results for $T_{IP/IC}$	30
3.2.2.1	Global $T_{IP/IC}$ SES values against initial parameter values	30
3.2.2.2	Global $T_{IP/IC}$ SES values against modeled parameter values	30
3.3	Conclusion on the power of $T_{IP/IC}$ to detect internal filtering	33
4	External Filter	35
4.1	Randomization outline	35
4.2	Results of the simulations of external filtering	36
4.2.1	Local $T_{IC/IR}$ and $T_{PC/PR}$ results	37
4.2.1.1	Local $T_{IC/IR}$ and $T_{PC/PR}$ SES values against initial parameter values	37
4.2.1.2	Local $T_{IC/IR}$ and $T_{PC/PR}$ SES values against modeled parameter values	43
4.2.2	Global $T_{IC/IR}$ and $T_{PC/PR}$ results	44
4.2.2.1	Calculation of beta-error mixing all strength of filtering	47
4.2.2.2	Global $T_{IC/IR}$ and $T_{PC/PR}$ SES values against initial parameter values	47
4.2.2.3	Global $T_{IC/IR}$ and $T_{PC/PR}$ SES values against modeled parameter values	51
4.3	Conclusion on the power of $T_{IC/IR}$ to detect external filtering	51
4.4	Conclusion on the power of $T_{PC/PR}$ to detect external filtering	54
5	Internal and external filtering	58
5.1	Randomization outline	58
5.2	Results	60
6	Test of variance decomposition functions	66
6.1	Behavior of the function <code>partvar</code>	66
6.2	Behavior of the function <code>decompCTRE</code>	67

7	Summary	72
7.1	Type I error	72
7.2	Type II error	72
	Conclusion	72
	References	72
	List of Figures	73
	List of Tables	73

1 Introduction

This document presents the lack of bias in the functions `Tstats`, `partvar`, `decompCTRE` and a test of robustness for the T-statistics. A summary of the main results is available at the end of each section.

1.1 R requirements

First we need to install and load the package.

```
install.packages("cati", repos = "http://cran.rstudio.com/")
library("cati")
library("xtable")

# Save the graphical parameter
oldpar <- par(no.readonly = TRUE)
```

1.2 System and session informations

This document was created with R version 3.1.2 (2014-10-31) on Linux. See below for more information. The speed of computations of `cati` main functions are given in the [cati tutorial](#).

```
sessionInfo()

## R version 3.1.2 (2014-10-31)
## Platform: x86_64-pc-linux-gnu (64-bit)
##
## locale:
##  [1] LC_CTYPE=fr_FR.UTF-8      LC_NUMERIC=C
##  [3] LC_TIME=fr_FR.UTF-8      LC_COLLATE=fr_FR.UTF-8
##  [5] LC_MONETARY=fr_FR.UTF-8  LC_MESSAGES=fr_FR.UTF-8
##  [7] LC_PAPER=fr_FR.UTF-8     LC_NAME=C
##  [9] LC_ADDRESS=C             LC_TELEPHONE=C
## [11] LC_MEASUREMENT=fr_FR.UTF-8 LC_IDENTIFICATION=C
##
## attached base packages:
## [1] stats      graphics  grDevices  utils      datasets  methods   base
##
## other attached packages:
## [1] xtable_1.7-4 cati_0.96    ape_3.2     ade4_1.6-2  nlme_3.1-119
## [6] knitr_1.9
##
## loaded via a namespace (and not attached):
##  [1] class_7.3-11      cluster_2.0.1     e1071_1.6-4
##  [4] evaluate_0.5.5    FD_1.0-12         formatR_1.0
##  [7] geometry_0.3-5    grid_3.1.2        hexbin_1.27.0
## [10] highr_0.4         hypervolume_1.1.2  lattice_0.20-29
## [13] latticeExtra_0.6-26 magic_1.5-6        MASS_7.3-37
## [16] Matrix_1.1-5      mgcv_1.8-3        mice_2.22
## [19] nnet_7.3-8        parallel_3.1.2    pdist_1.2
## [22] permute_0.8-3     randomForest_4.6-10 raster_2.3-24
```

```
## [25] rasterVis_0.32      RColorBrewer_1.1-2  Rcpp_0.11.4
## [28] rgl_0.93.996        rpart_4.1-8         sp_1.0-17
## [31] stringr_0.6.2       tools_3.1.2         vegan_2.2-1
## [34] zoo_1.7-11
```

1.3 T-statistics and associated null models

Table 1: The four types of null models implemented in cati, their related null and alternative hypotheses, randomization design and associated T-statistics

	Null hypothesis	Randomization procedure	Unilateral alternative hypothesis	T-statistics
local	There is no internal filtering: the distribution of trait values of all individuals within a given community does not depend on species identity	Randomization of individual trait values within the community	Internal filtering significantly impacts the distribution of trait values within a given community: two individuals belonging to a population have more similar trait values than two individuals randomly-drawn in the community	$T_{IP/IC}$
regional.ind	There is no external filtering: the distribution of trait values of individuals within a given community is a random drawing from the regional pool	Draw without replacement of individual trait values belonging to the regional pool (keeping the actual number of individuals in each community)	Two individuals belonging to a community have more similar trait values than two individuals randomly-drawn in the regional pool	$T_{IC/IR}$
regional.pop & regional.pop.prab	There is no species-based external filtering: the distribution of mean trait values of species within a given community is a random draw from the regional pool	(i) Assignment of a population-level value to each individual and (ii) Draw without replacement of population-level trait values belonging to the regional pool (keeping the actual number of individuals in each community (regional.pop) or not (regional.pop .prab)).	Two individuals belonging to a community have more similar population-based trait values than two individuals randomly-drawn in the regional pool with (regional.pop) or without (regional.pop .prab) taking abundance into account.	$T_{PC/PR}$ (regional.pop)

1.4 Simulations parameters

Here we set the number of permutations for the entire analysis.

```
#Number of permutations for the analysis without filter
npermut <- 3

#Number of permutations for each parameters values in the analysis with filter(s)
N_repet_Param <- 2

#Number of values for each parameters
nb_param_val <- 3
```

Now we set the number of individuals, species and communities. The argument `sd log` allows to modify the species abundance distribution within communities.

```
#Total number of individuals
Nind <- 500

#Ten communities named A, B, C ... J
com <- c("A", "B", "C", "D", "E", "F", "G", "H", "I", "J")
Ncom <- length(com)

#Number of species (if Nsp > 676 you need to defined additional name)
Nsp <- 30
sp <- as.factor(paste("sp", paste(sort(rep(LETTERS, 26)), rep(LETTERS, 26), sep="")
                      [seq(1:Nsp)], sep="_"))
Nsp <- length(sp)

#Parameter of the log normal distribution for species abundances
#distribution within communities
sdlog = 1.5
```

We use toys models of 500 individuals belonging to 30 species occurring in 10 communities (sites). Each community contains 50 individuals and species abundances in each community follow a lognormal distribution of mean 0 and standard deviation `sdlog` of 1.5. Figure 1 depicts the general outline of the next fourth sections dealing with the statistical test of T-statistics from Violle *et al.* (2012).

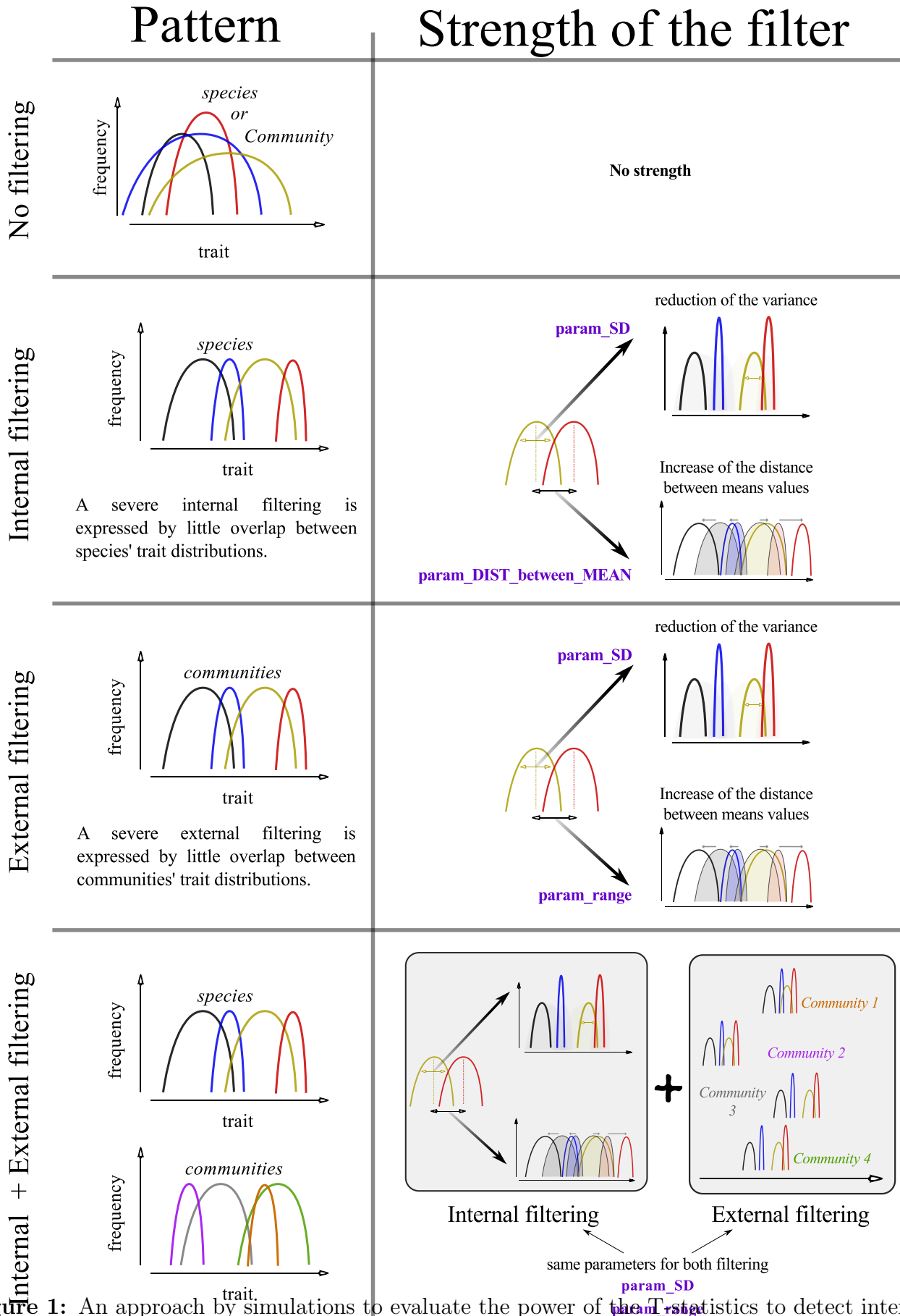


Figure 1: An approach by simulations to evaluate the power of the T-statistics to detect internal and external filtering. (Continue on next page)

Figure 1 (continue): The **lack of filtering** allows to test the type I error (alpha-error) of the three Tstatistics. In the absence of filtering the expected pattern is a random distribution of individual traits values among species and communities. **Internal filter** includes all of the assembly processes internal to the community. The alternative hypothesis is the non-overlap of trait distribution curves between species, as measured by the ratio $T_{IP/IC}$. We modulate the strength of the filter in the simulation using the variance of species' trait distributions (`param_SD`) and the distance between species trait means (`param_DIST_between_MEAN`). **External filter** includes all of the assembly processes external to the community. In this document we specifically treat the case of an environmental gradient influencing the individual traits values. Thus the expected pattern is a non-overlap between communities' trait distribution curves, as measured by the ratio $T_{IC/IR}$. We modulate the strength of external filter in the simulation using the variance of communities's trait distributions (`param_SD`) and the range of communities' trait means (the "extent" of the environmental gradient; (`param_range`)). Finally, we mix external and internal filter. First, we simulate the effect of internal filter alone. Second, we add to each individual trait value a number depending on its community. Again two parameters are used to modulate the strength of the filters (*cf* section 5 for more details).

2 No filter: calculation of error

2.1 Randomization outline

To calculate the **alpha error** (*i.e.* the probability to reject the null hypothesis H_0 while H_0 is true), we draw 3 random trait matrices irrespective of the species and community attributes of each individual.

```
#Start simulation
res.simul <- list() ; res.simul.pval <- list()

for(n in 1:npermut){#for each permutation

  ex.traits1 <- array(NA, dim = c(Nind, 2))
  colnames(ex.traits1) <- paste("trait", c("a", "b"), sep = " ")

  #trait a: normal distribution
  ex.traits1[, 1] <- rnorm(Nind, rlnorm(Nind, 0, 1), rlnorm(Nind, 0, 1))

  #trait b: uniform distribution
  ex.traits1[, 2] <- runif(Nind)

  # Draw communities using lognormal distribution of abundances
  ex.sp1 <- c()
  ex.com1 <- matrix(0, nrow = Ncom, ncol = Nsp)
  for(c in 1: Ncom){
    ex.com.interm <- table(sample(sp, size = Nind / Ncom,
                                   prob = rlnorm(Nsp, 0, sdlog), replace = T))

    ex.com1[c, sp %in% names(ex.com.interm)] <- ex.com.interm

    ex.sp1 <- c(ex.sp1, rep(sp, times = ex.com1[c,]))
  }
  ex.indplot1 <- sort(as.factor(rep(com, Nind / Ncom)))

  #Stock the results
  res.simul[[n]] <- Tstats(ex.traits1, ex.indplot1, ex.sp1)
  res.simul.pval[[n]] <- sum_Tstats(res.simul[[n]], type = "p.value")
  print(paste("----", round(n/npermut, 2) * 100, "%", sep = " "))
}#End of simulations
```

2.2 Results of the simulations without filtering

Let's see the result for one randomization. We can plot the distribution of traits values within species and/or communities thanks to the function `plotDistri` (Fig. 2). We can also plot the result of the T-statistics for one of the 3 permutations (Fig. 3).

```
par(mfrow=c(2, 2))
plotDistri(ex.traits1, rep("all_sp", times = dim(ex.traits1)[1]), ex.indplot1,
           plot.ask = F, multipanel = F, leg = c(T, F), main = c("a", "b"))
plotDistri(ex.traits1, rep("region", times = dim(ex.traits1)[1]), ex.sp1,
           plot.ask = F, multipanel = F, leg = c(T, F), main = c("c", "d"))
```

```
par(mfrow=c(1, 1))
```

```
plot(res.simu1[[1]])
```

2.3 Assessing the statistical type I error of local T-statistics

Here we call local p-values the p-values¹ corresponding to one index for one trait in one community². Now let's plot the ordered p.value for each T-statistics and each trait (Fig. 4)). In that case we have 20 p-values³ per T-statistics and per trait for each permutation. Using the defined parameters, we plot 60 points per T-statistics by trait (20 p.values * 3 permutations).

The alpha error on the figure are the proportion of p.values which are inferior to 0.025 (because performed a bilateral test).

```
par(mfrow = c(3, 2))
par(mar = c(3, 3.5, 2, 0.2))

xx <- log10(sort(unlist(lapply(res.simu1.pval, function(x) x[1:20, 1]))))
plot(xx, type = "l", main = "T_IP.IC norm", ylab = NA, xlab = NA)
abline(h = log10(0.025))
nbre_alpha_error_norm_loc_Tipic <- round((sum(xx < log10(0.025)) + 1) / (length(xx)+1), 5)
text(0, -0.3, labels = paste("alpha error", nbre_alpha_error_norm_loc_Tipic, sep = " = "),
     pos = 4)
mtext("log10 pvalue", 2, line = 2)

xx <- log10(sort(unlist(lapply(res.simu1.pval, function(x) x[1:20, 2]))))
plot(xx, type = "l", main = "T_IP.IC Uni", ylab = NA, xlab = NA)
abline(h = log10(0.025))
nbre_alpha_error_uni_loc_Tipic <- round((sum(xx < log10(0.025)) + 1) / (length(xx)+1), 5)
text(0, -0.3, labels = paste("alpha error", nbre_alpha_error_uni_loc_Tipic, sep = " = "),
     pos = 4)

xx <- log10(sort(unlist(lapply(res.simu1.pval, function(x) x[21:40, 1]))))
plot(xx, type = "l", main = "T_IC.IR norm", ylab = NA, xlab = NA)
abline(h = log10(0.025))
nbre_alpha_error_norm_loc_Ticir <- round((sum(xx < log10(0.025)) + 1) / (length(xx)+1), 5)
text(0, -0.3, labels = paste("alpha error", nbre_alpha_error_norm_loc_Ticir, sep = " = "),
     pos = 4)
mtext("log10 pvalue", 2, line = 2)

xx <- log10(sort(unlist(lapply(res.simu1.pval, function(x) x[21:40, 2]))))
plot(xx, type = "l", main = "T_IC.IR Uni", ylab = NA, xlab = NA)
abline(h = log10(0.025))
nbre_alpha_error_uni_loc_Ticir <- round((sum(xx < log10(0.025)) + 1) / (length(xx)+1), 5)
text(0, -0.3, labels = paste("alpha error", nbre_alpha_error_uni_loc_Ticir, sep = " = "),
     pos = 4)
```

¹Computed using the Tstats function

²In contrast with global statistics which aggregate the values of the metrics of all communities

³10 p-value, corresponding to 10 communities, multiplied by 2 because we performed a bilateral test

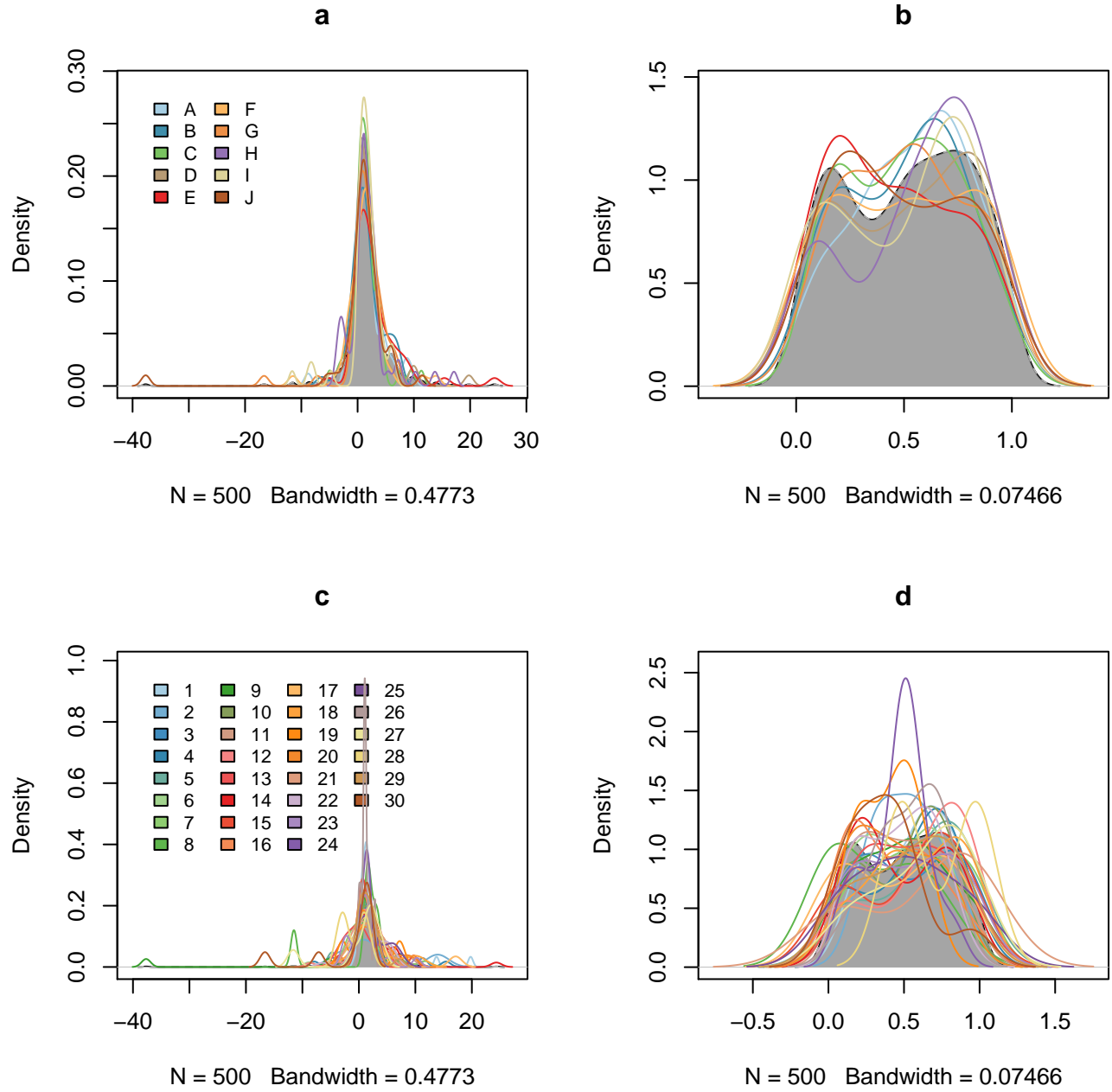


Figure 2: Distribution of trait values for one randomization without filtering: (a) Community-level trait distributions for the trait a (normal distribution); (b) Community-level trait distributions for the trait b (uniform distribution); (c) Species' trait distributions for the trait a; (d) Species' trait distributions for the trait b

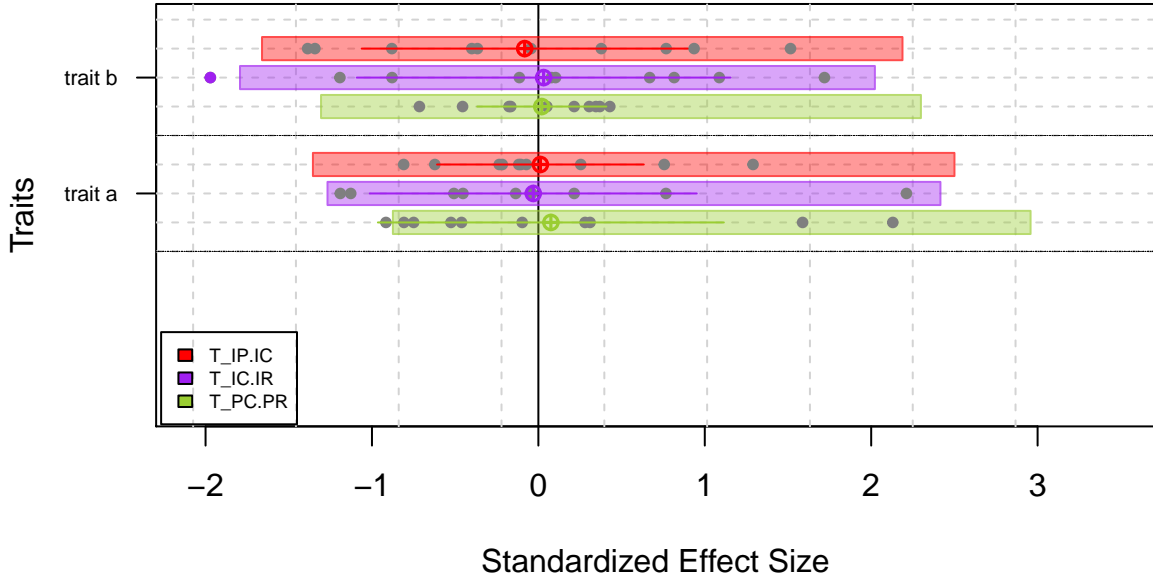


Figure 3: Results of T-statistics for one randomization without filter: $T_{IP/IC}$ in red, $T_{IC/IR}$ in purple and $T_{PC/PR}$ in green. All but three community-level metrics (colored dots) depart significantly from the null model. As expected under no filtering, the mean Standardized Effect Size (SES) values are closed to zero for all three T-statistics and are included in the area of the null model (colored boxes). Thus this plot is typically what we expected under the lack of both external and internal filtering.

```
xx <- log10(sort(unlist(lapply(res.simu1.pval, function(x) x[41:60, 1]))))
plot(xx, type = "l", main = "T_PC.PR norm", ylab = NA, xlab = NA)
abline(h = log10(0.025))
nbre_alpha_error_norm_loc_Tpcpr <- round((sum(xx < log10(0.025)) + 1) / (length(xx)+1), 5)
text(0, -0.3, labels = paste("alpha error", nbre_alpha_error_norm_loc_Tpcpr, sep = " = "),
     pos = 4)
mtext("log10 pvalue", 2, line = 2)
mtext("rank", 1, line = 2)

xx <- log10(sort(unlist(lapply(res.simu1.pval, function(x) x[41:60, 2]))))
plot(xx, type = "l", main = "T_PC.PR Uni", ylab = NA, xlab = NA)
abline(h = log10(0.025))
nbre_alpha_error_uni_loc_Tpcpr <- round((sum(xx < log10(0.025)) + 1) / (length(xx)+1), 5)
text(0, -0.3, labels = paste("alpha error", nbre_alpha_error_uni_loc_Tpcpr, sep = " = "),
     pos = 4)
mtext("rank", 1, line = 2)

par(oldpar)
```

2.4 Assessing the statistical type I error of global T-statistics

In contrast with local p-values, we call global p-values the p-values corresponding to one index for a given trait across all the communities.

We test for the **alpha error** of global p-values in the same way as local p-values. We plot the ordered

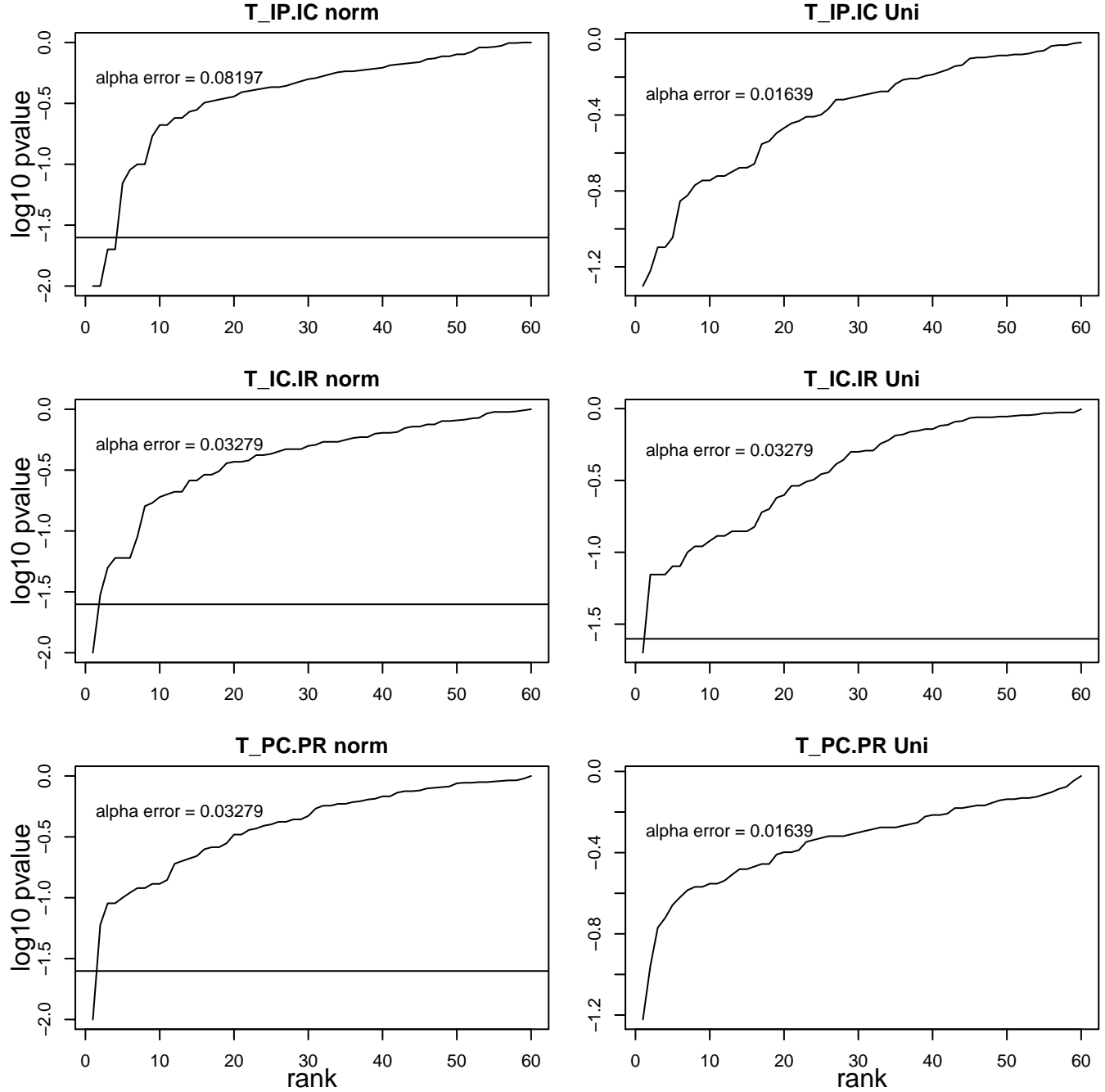


Figure 4: Local alpha errors: local p-values are log-transformed and ordered before plotting. The black line represent the value of 0.025. Consequently, p-values under this line are false-positive cases. Alpha-error is the proportion of this false-positive cases. Norm: normal distribution, Uni: uniform distribution.

mean⁴ standardized effect size (SES)⁵ and the 95% confidence interval in grey.

In that case we have one p-value by T-statistics by traits for each of the 3 permutations.

```
meanSES.1.T_IP.IC.distriNorm <- lapply(res.simu1, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_1$ses[,1], na.rm = T))

meanSES.1.T_IC.IR.distriNorm <- lapply(res.simu1, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_2$ses[,1], na.rm = T))

meanSES.1.T_PC.PR.distriNorm <- lapply(res.simu1, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_3$ses[,1], na.rm = T))

meanSES.1.T_IP.IC.distriUni <- lapply(res.simu1, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_1$ses[,2], na.rm = T))

meanSES.1.T_IC.IR.distriUni <- lapply(res.simu1, function(x)
  mean(ses.listofindex(as.listofindex(x))$
    index_1_2$ses[,2], na.rm = T))

meanSES.1.T_PC.PR.distriUni <- lapply(res.simu1, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_3$ses[,2], na.rm = T))
```

Now, plot the result as ordered SES values (Fig. 5). Horizontal lines represent the confidence area at 95%.

```
par(mfrow = c(3, 2))
par(mar = c(3, 3.5, 2, 0.2))

#### T_IP.IC
#T_IP.IC.distriNorm
xx <- sort(unlist(meanSES.1.T_IP.IC.distriNorm))
xx_lim.inf <- unlist(lapply(res.simu1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_1$ses.inf[,1], na.rm = T)))
xx_lim.sup <- unlist(lapply(res.simu1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_1$ses.sup[,1], na.rm = T)))
alpha1 <- round((sum(xx<xx_lim.inf | xx>xx_lim.sup)+1) / (length(xx)+1), 3)

plot(xx, type = "l", main = "T_IP.IC norm", ylim = c(-3, 3),
  ylab = NA, xlab = NA)
rect(-100, mean(xx_lim.inf), npermut*1.2, mean(xx_lim.sup), col = rgb(0, 0, 0, 0.2))
text(3, 0.8, labels = paste("alpha error", alpha1, sep = " = "), cex = 0.7, pos = 4)
mtext("Standardized Effect Size", 2, line = 2)

#T_IP.IC.distriUni
```

⁴mean of the 10 communities values

⁵compute as (the observed value - the mean value among simulation) / the standard error among simulations

```

xx <- sort(unlist(meanSES.1.T_IP.IC.distriUni))
xx_lim.inf <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_1$ses.inf[,2], na.rm = T)))
xx_lim.sup <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_1$ses.sup[,2], na.rm = T)))
alpha2 <- round((sum(xx<xx_lim.inf | xx>xx_lim.sup)+1) / (length(xx)+1), 3)

plot(xx, type = "l", main = "T_IP.IC uni", ylim = c(-3, 3),
  ylab = NA, xlab = NA)
rect(-100, mean(xx_lim.inf), npermut*1.2, mean(xx_lim.sup), col = rgb(0, 0, 0, 0.2))
text(3, 0.8, labels = paste("alpha error", alpha2, sep = " = "), cex = 0.7, pos = 4)

#-----
####   T_IC.IR
#T_IC.IR.distriNorm
xx <- sort(unlist(meanSES.1.T_IC.IR.distriNorm))

xx_lim.inf <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_2$ses.inf[,1], na.rm = T)))
xx_lim.sup <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_2$ses.sup[,1], na.rm = T)))
alpha3 <- round((sum(xx<xx_lim.inf | xx>xx_lim.sup)+1) / (length(xx)+1), 3)

plot(xx, type = "l", main = "T_IC.IR norm", ylim = c(-3, 3),
  ylab = NA, xlab = NA)
rect(-100, mean(xx_lim.inf), npermut*1.2, mean(xx_lim.sup), col = rgb(0, 0, 0, 0.2))
text(3, 0.8, labels = paste("alpha error", alpha3, sep = " = "), cex = 0.7, pos = 4)
mtext("Standardized Effect Size", 2, line = 2)

#T_IC.IR.distriUni
xx <- sort(unlist(meanSES.1.T_IC.IR.distriUni))

xx_lim.inf <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_2$ses.inf[,2], na.rm = T)))
xx_lim.sup <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_2$ses.sup[,2], na.rm = T)))
alpha4 <- round((sum(xx<xx_lim.inf | xx>xx_lim.sup)+1) / (length(xx)+1), 3)

plot(xx, type = "l", main = "T_IC.IR uni", ylim = c(-3, 3),
  ylab = NA, xlab = NA)
rect(-100, mean(xx_lim.inf), npermut*1.2, mean(xx_lim.sup), col = rgb(0, 0, 0, 0.2))
text(3, 0.8, labels = paste("alpha error", alpha4, sep = " = "), cex = 0.7, pos = 4)

#-----
####   T_PC.PR
#T_PC.PR.distriNorm
xx <- sort(unlist(meanSES.1.T_PC.PR.distriNorm))

xx_lim.inf <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))
  $index_1_3$ses.inf[,1], na.rm = T)))
xx_lim.sup <- unlist(lapply(res.simul1, function(x) mean(ses.listofindex(as.listofindex(x))

```



```

    $index_1_3$ses.sup[,1], na.rm = T)))
alpha5 <- round((sum(xx<xx_lim.inf | xx>xx_lim.sup)+1) / (length(xx)+1), 3)

plot(xx, type = "l", main = "T_PC.PR norm", ylim = c(-3, 3),
     ylab = NA, xlab = NA)
rect(-100, mean(xx_lim.inf), npermut*1.2, mean(xx_lim.sup), col = rgb(0, 0, 0, 0.2))
text(3, 0.8, labels = paste("alpha error", alpha5, sep = " = "), cex = 0.7, pos = 4)
mtext("Standardized Effect Size", 2, line = 2)
mtext("rank", 1, line = 2)

#T_PC.PR.distribUni
xx <- sort(unlist(meanSES.1.T_PC.PR.distribUni))
xx_lim.inf <- unlist(lapply(res.simu1, function(x) mean(ses.listofindex(as.listofindex(x))
    $index_1_3$ses.inf[,2], na.rm = T)))
xx_lim.sup <- unlist(lapply(res.simu1, function(x) mean(ses.listofindex(as.listofindex(x))
    $index_1_3$ses.sup[,2], na.rm = T)))
alpha6 <- round((sum(xx<xx_lim.inf | xx>xx_lim.sup)+1) / (length(xx)+1), 3)

plot(xx, type = "l", main = "T_PC.PR uni", ylim = c(-3, 3),
     ylab = NA, xlab = NA)
rect(-100, mean(xx_lim.inf), npermut*1.2, mean(xx_lim.sup), col = rgb(0, 0, 0, 0.2))
text(3, 0.8, labels = paste("alpha error", alpha6, sep = " = "), cex = 0.7, pos = 4)
mtext("rank", 1, line = 2)

par(oldpar)

```

2.5 Conclusion on the type I error of T-statistics

The 3 permutations allow the calculation of alpha error for two traits (the first normally distributed and the second uniformly distributed), three indices (the three T-statistics) and using local indices (one by community) or global indices (averaged between communities).

Table 2: Alpha-errors for T-statistics

Traits	Indices	Average between communities?	alpha-error
Trait normally distributed	$T_{IP/IC}$	no (local)	0.082
		yes (global)	0.25
	$T_{IC/IR}$	no (local)	0.033
		yes (global)	0.25
	$T_{PC/PR}$	no (local)	0.033
		yes (global)	0.25
Trait uniformly distributed	$T_{IP/IC}$	no (local)	0.016
		yes (global)	0.25
	$T_{IC/IR}$	no (local)	0.033
		yes (global)	0.25
	$T_{PC/PR}$	no (local)	0.016
		yes (global)	0.25

The results of table 2 attest of a good behavior of the T-statistics regardless of the distributions of the traits. Alpha-error of the local indices are always near the wanted alpha-error of 2.5%. Logically

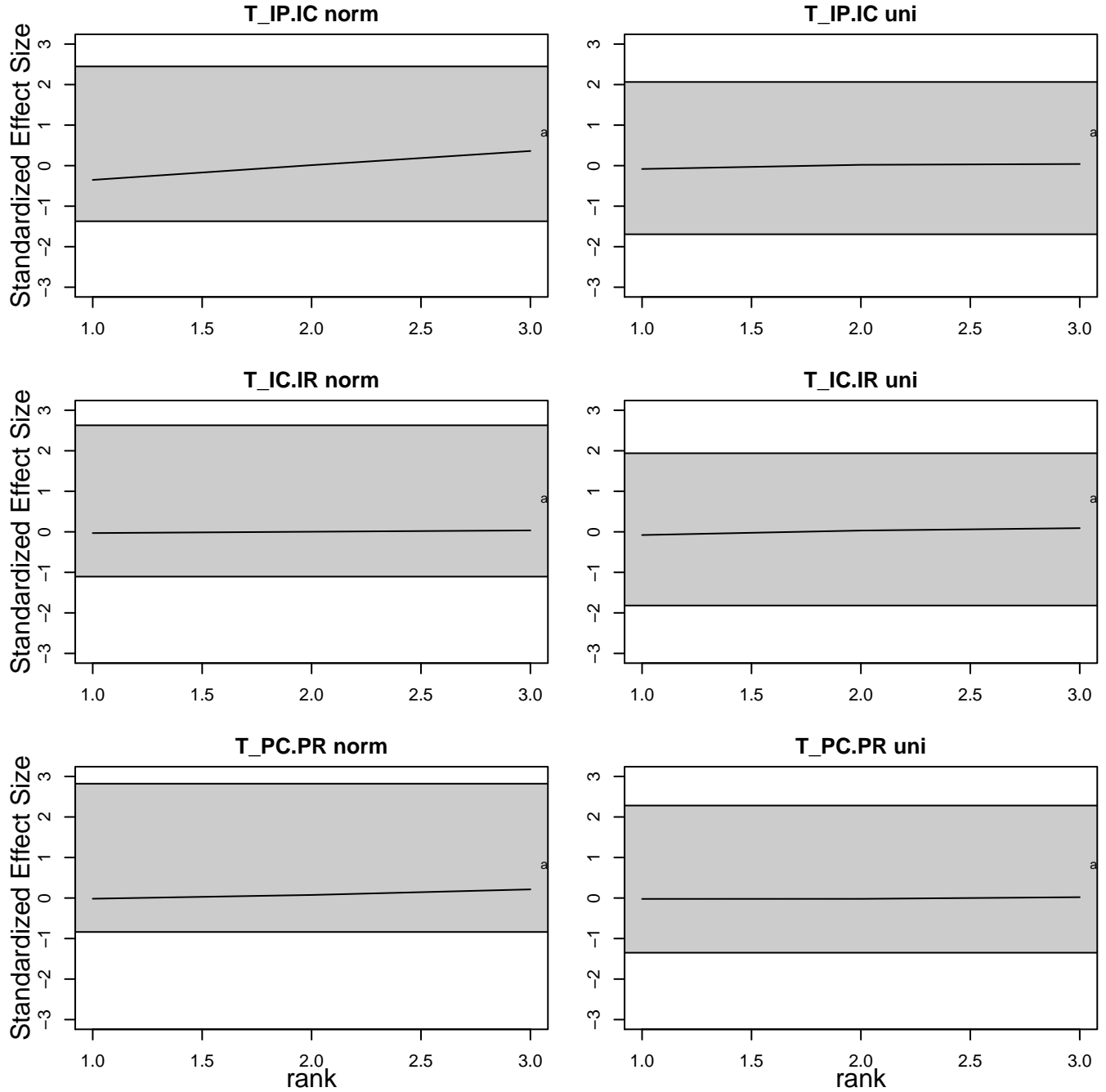


Figure 5: Global alpha errors: global Standardized Effect Size are ordered before plotting. The grey area represent the confidence interval at 2.5 percent on each side. Consequently, p-values outside this area are false-positive cases. Alpha-error is the proportion of this false-positive cases. Norm: normal distribution, Uni: uniforme distribution.

the global indices which averaged local indices are far more robust. Note that the higher alpha-error is obtained using $T_{PC/PR}$ on the trait normally distributed (0.033).

3 Internal filter: assessing the statistical power of $T_{IP/IC}$

3.1 Randomization outline

We decide to modulate the strength of the internal filter thanks to two parameters: we define (i) a gradient of distance between species means traits distributions and (ii) a gradient of the mean variance in species traits distributions. It is strongly linked to the definition of limiting similarity by May and Mac Arthur (1972) who defined a coefficient of competition as

$$a = \exp(-d^2/4w^2)$$

with d correspond to the distance between the species trait means and w correspond to the mean variance of species trait distribution.

```
# Number of values for each parameters: param_DIST_between_MEAN.init and param_SD
nb_param_val

## [1] 3

# Number of permutations is the product of the number of values for each parameter
#(nb_param_val) and the number of permutations by parameter (N_repet_Param)
nperm <- nb_param_val * N_repet_Param

# Maximum mean value for traits
max.value_traits <- 250

# Parameter for the distance between species mean trait values
param_DIST_between_MEAN.init <- round(sort(seq(10, 200, length.out = nb_param_val)), 2)
param_DIST_between_MEAN <- rep(param_DIST_between_MEAN.init, N_repet_Param)
mean_sd.param <- rep(10, nperm)

# Parameter for the variance in species mean trait values
param_SD <- seq(10, 100, length.out = nb_param_val)
sd_mean.param <- sample(rep(param_SD, N_repet_Param),
  size = length(rep(param_SD, N_repet_Param)), replace = F)
sd_sd.param <- rep(10, nperm)
```

Practically, we defined 3 values for two parameters:

(i) `param_DIST_between_MEAN.init` is a vector of length 3 which defined the minimum value for the mean traits values (the maximum value is 250). In decreasing the range of traits values while keeping the same number of species, there is a decrease of mean distance between species means of traits values. Thus the mean for each species is drawn from a normal distribution with standard deviation of 10 (parameter `mean_sd.param`) and means evenly distributed between `max.value_traits` - `param_DIST_between_MEAN` and `max.value_traits`.

(ii) `param_SD` is a vector of length 3 which defined the standard deviation of traits distributions for each species. In order to decorrelate the mean and standard deviation of species traits distributions, `param_SD` is permuted before the analysis.

For the trait "a" the trait value for each individual is drawn from a normal distribution with the mean depending on it species attribute. For the trait "b" the trait value for each individual is drawn

from a uniform distribution with the range depending on its species attribute. The range is defined as: *species mean value* - *species sd value* for the minimum and *species mean value* + *species sd value* for the maximum.

Now we can start the simulation with 3 parameters values and 6 permutations.

```
mean.sp_stock2 <- list() ; sd.sp_stock2 <- list() ; res.simu2 <- list()
res.simu2.pval <- list() ; res.simu.traits2 <- list() ; mean_dist_sp <- list()

for(n in 1:nperm){#for each permutation

  # Draw communities using lognormal distribution of abundances
  # 100 individuals by populations
  ex.sp2 <- c()
  ex.com2 <- matrix(0, nrow = Ncom, ncol = Nsp)
  for(i in 1:Ncom){
    ex.com.interm <- table(sample(sp, size = Nind/Ncom, prob = rlnorm(Nsp, 0, sdlog),
                                replace = T))
    ex.com2[i, sp%in% names(ex.com.interm)] <- ex.com.interm
    ex.sp2 <- c(ex.sp2, rep(sp, times = ex.com2[i,]))
  }

  ex.indplot2 <- sort(as.factor(rep(com, Nind/Ncom)))

  # Defining traits mean and sd by species
  mean_mean.param.interm <- seq(max.value_traits - param_DIST_between_MEAN[n],
                                max.value_traits, length.out =
                                length(unique(param_DIST_between_MEAN)))

  mean_mean.param <- rep(round(sort(mean_mean.param.interm), 2), N_repet_Param)

  mean.sp <- rnorm(length(unique(sp)), mean = mean_mean.param, sd = mean_sd.param[n])
  sd.sp <- rnorm(length(unique(sp)), mean = sd_mean.param[n], sd = sd_sd.param[n])

  ex.traits2 <- array(NA, dim = c(Nind, 2))
  colnames(ex.traits2) <- paste("trait", c("a", "b"), sep = " ")

  # Draw the individual traits depending on species attributes
  for(s in unique(ex.sp2)){
    #trait a : normal distribution
    ex.traits2[ex.sp2 == s, 1] <-
      rnorm(5*Nind/Ncom, rep(mean.sp[unique(ex.sp2) == s], 5*Nind/Ncom),
            rep(sd.sp[unique(ex.sp2) == s], 5*Nind/Ncom))[1:sum(ex.sp2 == s)]

    #trait b : uniform distribution
    ex.traits2[ex.sp2 == s, 2] <-
      runif(5*Nind/Ncom, min = rep(mean.sp[unique(ex.sp2) == s], 5*Nind/Ncom) -
            rep(sd.sp[unique(ex.sp2) == s], 5*Nind/Ncom),
            max = rep(mean.sp[unique(ex.sp2) == s], 5*Nind/Ncom) +
            rep(sd.sp[unique(ex.sp2) == s], 5*Nind/Ncom))[1:sum(ex.sp2 == s)]
  }

  #stock results
```

```

mean.sp_stock2[[n]] <- mean.sp
sd.sp_stock2[[n]] <- sd.sp

mean_dist_sp[[n]] <- c(mean(tapply(ex.traits2[, 1], ex.indplot2, function(x)
                                mean(dist(x), na.rm=T)), na.rm=T),
                      mean(tapply(ex.traits2[, 2], ex.indplot2, function(x)
                                mean(dist(x), na.rm=T)), na.rm=T))

res.simu.traits2[[n]] <- ex.traits2
res.simu2[[n]] <- Tstats(ex.traits2, ex.indplot2, ex.sp2, printprogress = FALSE)
res.simu2.pval[[n]] <- sum_Tstats(res.simu2[[n]], type = "p.value")
print(paste(round(n/nperm, 2) * Nind/Ncom, "%", sep = " "))
}#End of simulations

```

3.2 Results of the simulations of internal filtering

Let's see the result for one randomization. We can plot the distribution of traits values within species and/or communities thanks to the function `plotDistri` (Fig. 6).

```

par(mfrow=c(2, 2))
plotDistri(ex.traits2, rep("all_sp", times = dim(ex.traits2)[1]), ex.indplot2,
           plot.ask = F, multipanel = F, cex.leg = 0.6, main= c("a", "b"))
plotDistri(ex.traits2, rep("region", times = dim(ex.traits2)[1]), ex.sp2, plot.ask = F,
           multipanel = F, ylim = c(0,0.04), cex.leg = 0.6, main= c("c", "d"))
par(mfrow=c(1, 1))

```

We can also plot the result of the T-statistics for two contrasting cases (Fig. 7).

```

par(mfrow=c(2, 1))
plot(res.simu2[[1]], main="a")
plot(res.simu2[[n]], main="b")
par(mfrow=c(1, 1))

```

For the power analysis of internal filtering we present the result in two parts: (i) the test of robustness of each T-statistics values using local p-values and (ii) the test of community-aggregated T-statistics using global p-values.

3.2.1 Local p-values results

Here we call local p-values the p-values corresponding to one index for one trait in one community. First, we can represent the ordered p.value in decimal logarithms and calculate the power of the $T_{IP/IC}$ index using the beta-error⁶. The less the beta-error, the more the index is powerful. Contrary to the test without filter, we choose a unilateral test because we only want to test if $T_{IP/IC}$ is smaller than it should be under the associated null hypothesis.

We can plot the local p-values in different ways. First, we can represent the ordered p.value in decimal logarithms and calculate the power of the $T_{IP/IC}$ index using the beta-error. The less the beta-error, the more the index is powerful. Contrary to the test without filter, we choose a unilateral test because we only want to test if $T_{IP/IC}$ is smaller than it should be under the associated null hypothesis.

⁶The proportion of case where we do not reject the null hypothesis while this hypothesis is wrong.

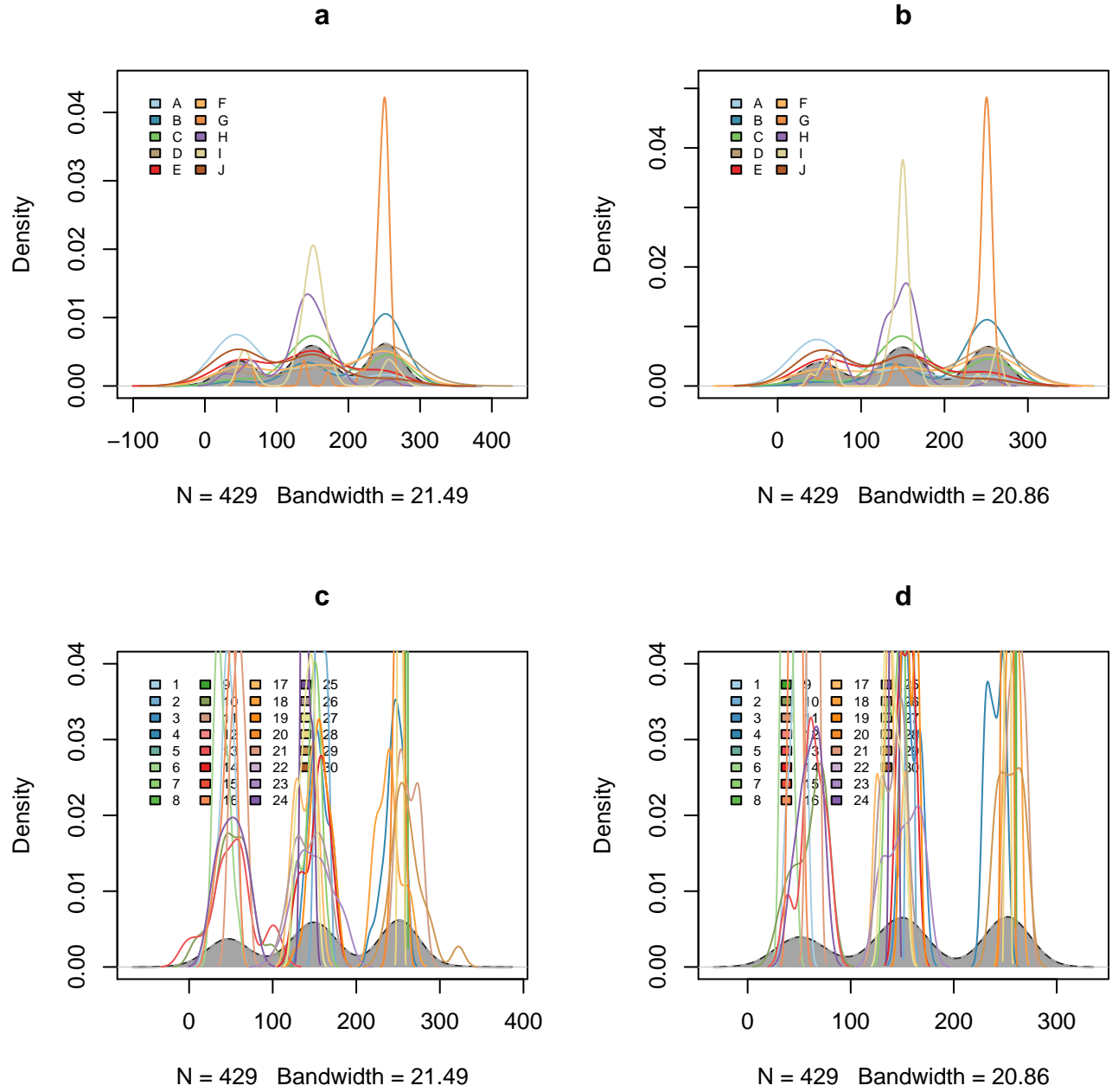


Figure 6: Distribution of trait values for one randomization with internal filtering: (a) Community-level trait distributions for the trait a (normal distribution); (b) Community-level trait distributions for the trait b (uniform distribution); (c) Species' trait distributions for the trait a; (d) Species' trait distributions for the trait b

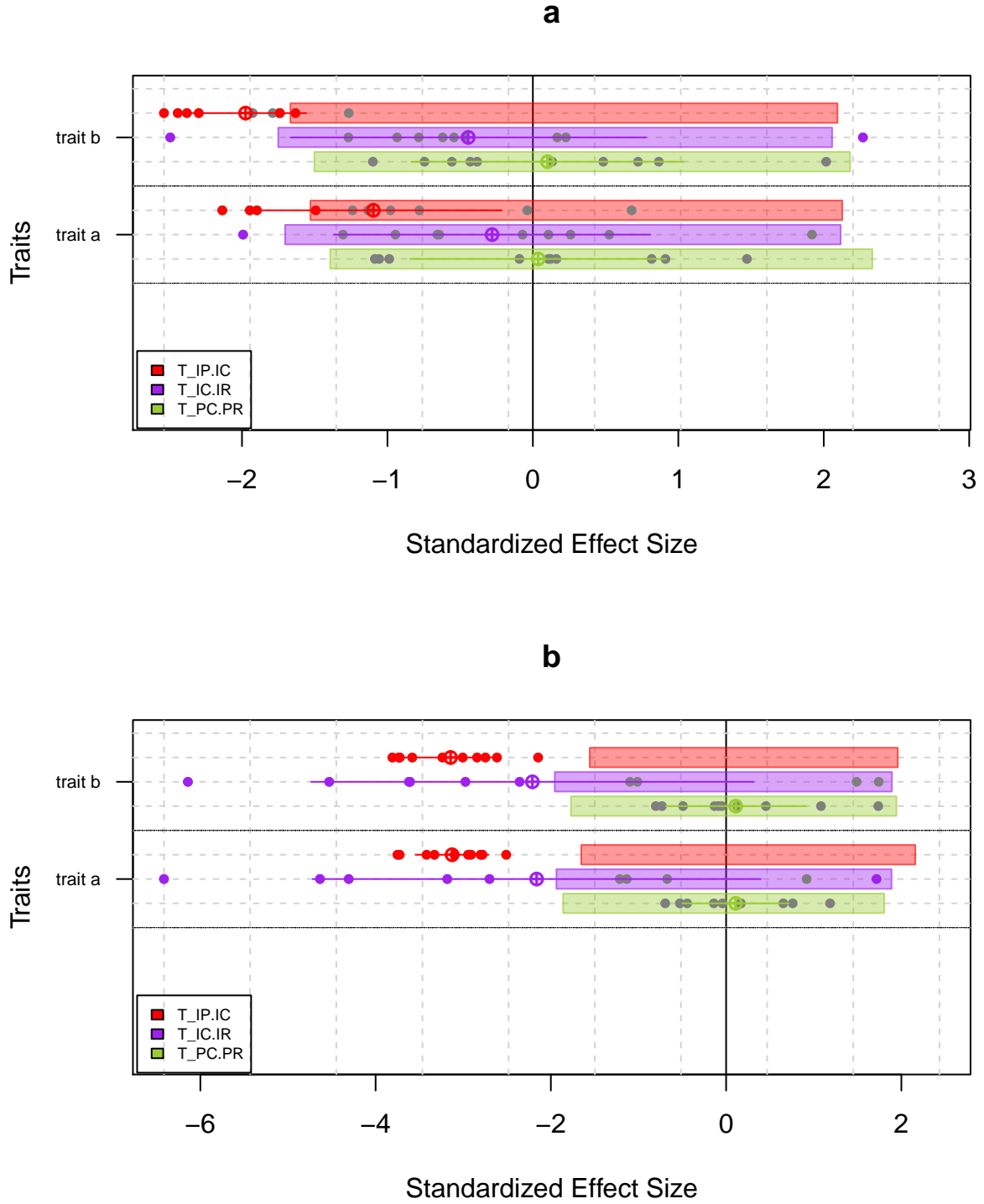


Figure 7: Results of T-statistics for two randomization with contrasted strength of internal filtering: $T_{IP/IC}$ in red, $T_{IC/IR}$ in purple and $T_{PC/PR}$ in green. (a) Lower strength of internal filtering and (b) higher strength of internal filtering.

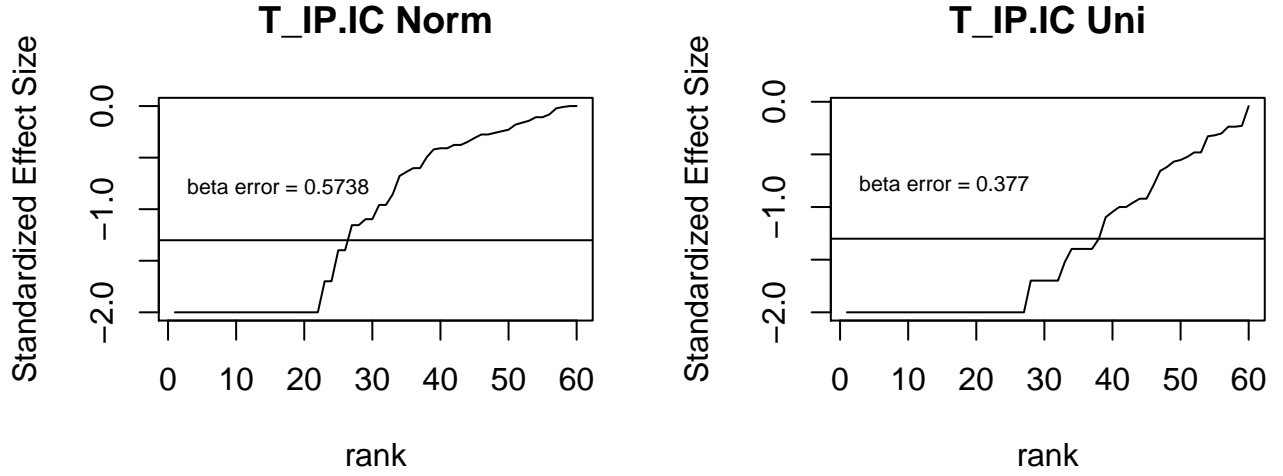


Figure 8: Beta-error mixing all strength of internal filtering: Standardized effect size of $T_{IP/IC}$ are ordered and plot. The black line represent the value of 0.05. Consequently, p-values above this line are false-negative cases. Norm: normal distribution, Uni: uniforme distribution.

3.2.1.1 Calculation of beta-error mixing all strength of filtering

```
par(mfrow = c(1, 2))
xx <- log10(sort(unlist(lapply(res.simu2.pval, function(x) x[1:10, 1]))))
plot(xx, type = "l", main = "T_IP.IC Norm",
      xlab = "rank", ylab = "Standardized Effect Size")
abline(h = log10(0.05))
nbre_beta_error <- round((sum(xx > log10(0.05)) + 1) / (length(xx) + 1), 4)
text(0, -0.8, labels = paste("beta error", nbre_beta_error, sep = " = "),
     cex = 0.7, pos = 4)

xx <- log10(sort(unlist(lapply(res.simu2.pval, function(x) x[1:10, 2]))))
plot(xx, type = "l", main = "T_IP.IC Uni",
      xlab = "rank", ylab = "Standardized Effect Size")
abline(h = log10(0.05))
nbre_beta_error <- round((sum(xx > log10(0.05)) + 1) / (length(xx) + 1), 4)
text(0, -0.8, labels = paste("beta error", nbre_beta_error, sep = " = "),
     cex = 0.7, pos = 4)
par(oldpar)
```

The beta-error is very high and thus the power of $T_{IP/IC}$ can *prima facie* seen as problematic (Fig. 8). But here we mingle the p-value for several parameters values. Thus it is far more informative to see the power of the test in relation to the initial parameters values. We represent standardized effect sizes values (hereafter called SES) in relation to the strength of internal filter to determine the parameters values which allow a satisfactory power of $T_{IP/IC}$.

As we add stochasticity around the initial parameters values, we can measure the strength of internal filtering by (i) the initial parameter values or (ii) by calculating the modeled parameter values⁷.

3.2.1.2 Local $T_{IP/IC}$ SES values against initial parameter values

⁷*i.e.* the parameter values after adding stochasticity

```

meanSES.2loc.norm_Tipic <- unlist(lapply(res.simu2, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses[, 1]))

meanSES.2loc.uni_Tipic <- unlist(lapply(res.simu2, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses[, 2]))

SES.inf.MEAN.norm_Tipic <- unlist(lapply(res.simu2, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses.inf[, 1]))

SES.inf.MEAN.uni_Tipic <- unlist(lapply(res.simu2, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses.inf[, 2]))

```

Now, we can plot SES of $T_{IP/IC}$ (Fig. 9). The trait "a" normally distributed is in black and the uniform traits "b" is in purple. The colored rectangles represent the null model area with $\alpha = 5\%$. Thus, when a point is outside these area, the modeled parameters are strong enough to detect the internal filter with a high power (beta-error < 0.05).

```

init_param <- param_DIST_between_MEAN / sd_mean.param
init_param.loc <- rep(init_param, each = Ncom)

plot(meanSES.2loc.norm_Tipic, init_param.loc, pch = 16, col = rgb(0, 0, 0, 0.6),
  xlim = c(min(c(unlist(meanSES.2loc.norm_Tipic), unlist(meanSES.2loc.uni_Tipic))),
  na.rm = T), 0),
  main = "Local T_IP.IC",
  xlab = "standardized effect size of T_IP.IC",
  ylab = "initial parameters: range/sd")
abline(v = mean(SES.inf.MEAN.norm_Tipic, na.rm = T))

points(meanSES.2loc.uni_Tipic, init_param.loc, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(SES.inf.MEAN.uni_Tipic, na.rm = T), col = "purple")
rect(mean(SES.inf.MEAN.norm_Tipic, na.rm = T), -1, 0, max(init_param) + 1,
  col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(SES.inf.MEAN.uni_Tipic, na.rm = T), 0, 0, max(init_param),
  col = rgb(0.5, 0, 1, 0.3), border = NA)

```

For each initial parameters values we can calculate the proportion of case where we do not reject the null hypothesis while this hypothesis is wrong (*i.e.* the Beta-error). The next chunk only shows the calculation for the $T_{IC/IR}$ metrics on the trait "a".

```

beta_sd_norm_Tipic <-
  (unlist(lapply(by(cbind(meanSES.2loc.norm_Tipic, SES.inf.MEAN.norm_Tipic),
    rep(sd_mean.param, each = Ncom), function(x) x[,1] > x[,2]),
    function(x) sum(x, na.rm = T)))+1)/(10 * N_repet_Param + 1)
beta_range_norm_Tipic <-
  (unlist(lapply(by(cbind(meanSES.2loc.norm_Tipic, SES.inf.MEAN.norm_Tipic),
    rep(param_DIST_between_MEAN, each = Ncom), function(x) x[,1] > x[,2]),
    function(x) sum(x, na.rm = T)))+1)/(10 * N_repet_Param + 1)

```

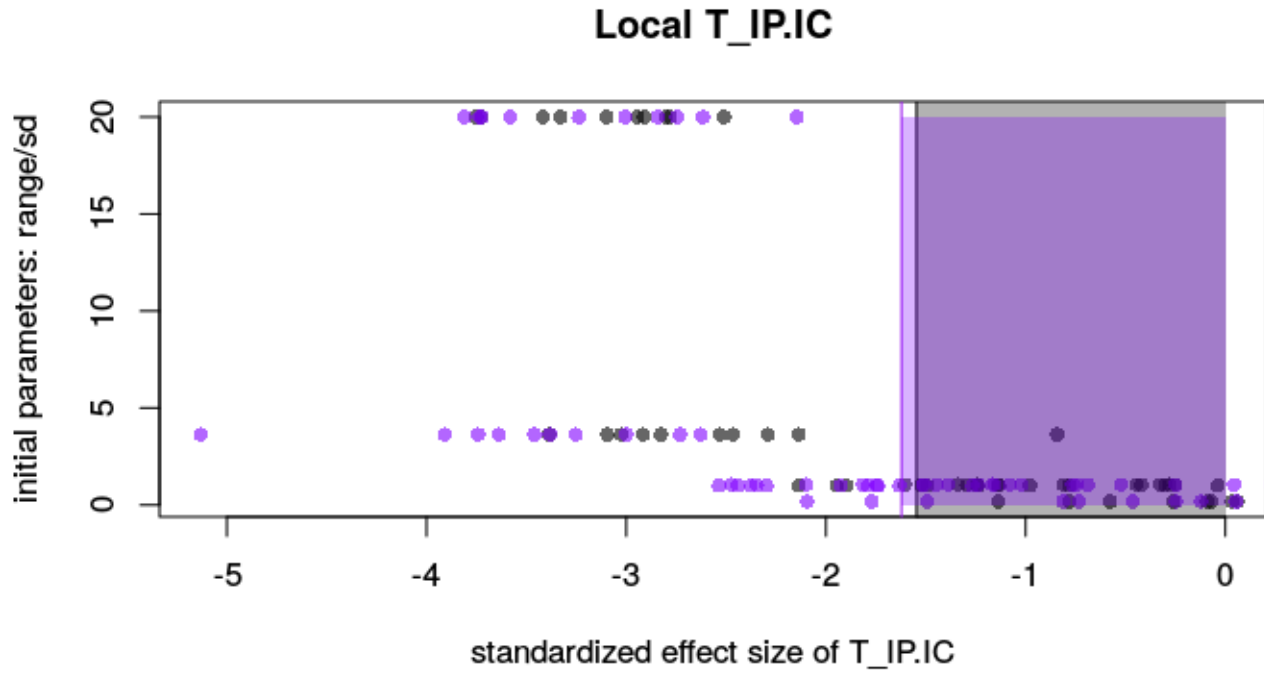


Figure 9: Local $T_{IP.IC}$ SES and initial parameters values: Standardized effect size of $T_{IP.IC}$ in function of the strength of internal filter defined by the ratio of the initial range parameter by the initial standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

```

beta_sd_uni_Tipic <-
  (unlist(lapply(by(cbind(meanSES.2loc.uni_Tipic, SES.inf.MEAN.uni_Tipic),
    rep(sd_mean.param, each = Ncom), function(x) x[,1] > x[,2]),
    function(x) sum(x, na.rm = T)))+1)/(10 * N_repet_Param + 1)
beta_range_uni_Tipic <-
  (unlist(lapply(by(cbind(meanSES.2loc.uni_Tipic, SES.inf.MEAN.uni_Tipic),
    rep(param_DIST_between_MEAN, each = Ncom), function(x) x[,1] > x[,2]),
    function(x) sum(x, na.rm = T)))+1)/(10 * N_repet_Param + 1)

res_beta_Tipic <- rbind(rev(beta_sd_norm_Tipic), rev(beta_sd_uni_Tipic),
  beta_range_norm_Tipic, beta_range_uni_Tipic)
colnames(res_beta_Tipic) <- paste("str", 1:ncol(res_beta_Tipic))
rownames(res_beta_Tipic) <- c("effect of sd (a)", "effect of sd (b)",
  "effect of range (a)", "effect of range (b)")

```

```

print(xtable(res_beta_Tipic, caption = 'Local beta-error of  $T_{IP/IC}$  in function
  of the strength of internal filtering. str: strength. Trait a is
  normally distributed and trait b is uniformly distributed.',
  label = "tab:local_beta_error_xtable"),
  caption.placement = "top", size = "small")

```

Table 3: Local beta-error of $T_{IP/IC}$ in function of the strength of internal filtering. str: strength. Trait a is normally distributed and trait b is uniformly distributed.

	str 1	str 2	str 3
effect of sd (a)	0.95	0.57	0.33
effect of sd (b)	0.71	0.43	0.24
effect of range (a)	0.81	0.95	0.10
effect of range (b)	0.62	0.71	0.05

Now we can see the beta-error for each strength in table 3. A high value of range and a low value of standard error defined a high strength of filtering. For example, in the case of a trait normally distributed, the proportion of false negative is 95.24% for the higher value of standard errors and thus for the lower strength of internal filtering. The good behavior of the function `Tstats` is ascertained by the decrease of beta-error along increasing strength of the filter.

3.2.1.3 Local $T_{IP/IC}$ SES values against modeled parameter values

First we need to calculate the modeled parameter values. Here again we calculate the strength of internal filtering by dividing the range of values in communities by their standard error.

```

mean_dist_sp.interm <- t(matrix(unlist(lapply(mean_dist_sp, function(x) x)), nrow = 2))
mean_sd_of_com <- unlist(lapply(sd.sp_stock2, function(x) mean(x)))

modeled_param_norm <- mean_dist_sp.interm[, 1] / mean_sd_of_com
modeled_param_uni <- mean_dist_sp.interm[, 2] / mean_sd_of_com
modeled_param_norm.loc <- rep(modeled_param_norm, each = Ncom)
modeled_param_uni.loc <- rep(modeled_param_uni, each = Ncom)

```

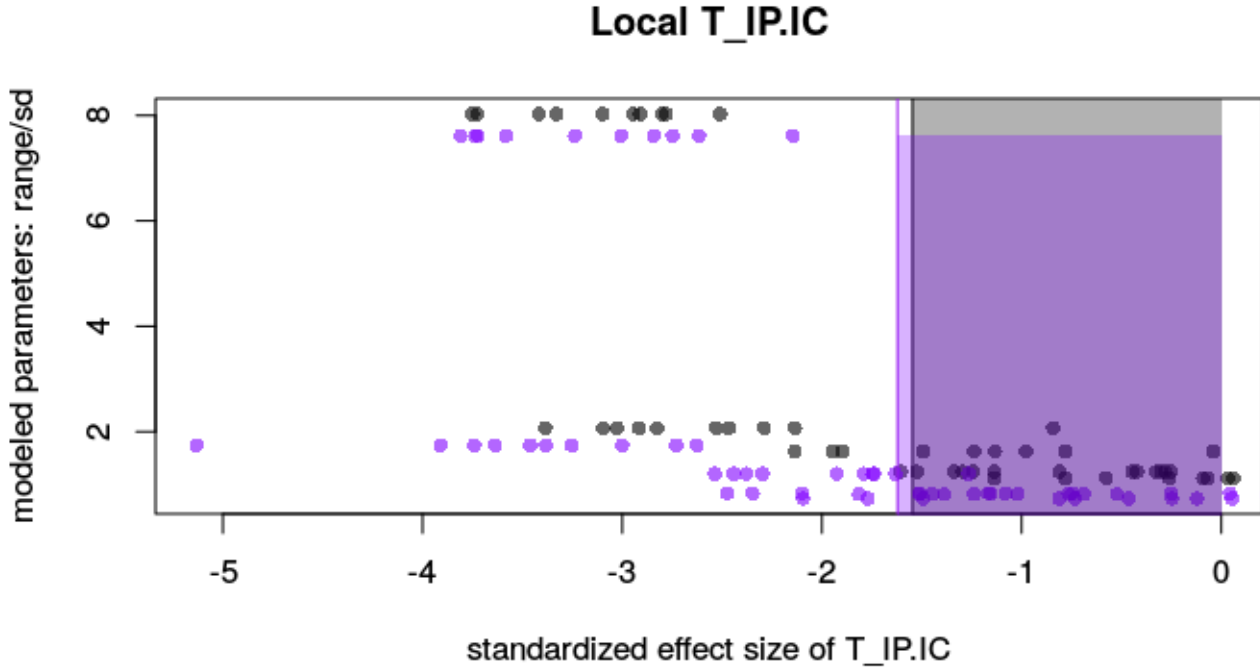


Figure 10: Local $T_{IP/IC}$ SES and modeled parameters values: Standardized effect size of $T_{IP/IC}$ in function of the strength of internal filter defined by the ratio of the modeled range parameter by the modeled standard error parameter. The colored rectangles represent the mean confidence interval for SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

```
plot(meanSES.2loc.norm_Tipic, modeled_param_norm.loc, pch = 16, col = rgb(0, 0, 0, 0.6),
     xlim = c(min(c(unlist(meanSES.2loc.norm_Tipic), unlist(meanSES.2loc.uni_Tipic)),
                  na.rm = T), 0),
     ylim = c(min(c(modeled_param_norm.loc, modeled_param_uni.loc), na.rm = T),
              max(c(modeled_param_norm.loc, modeled_param_uni.loc), na.rm = T)),
     main = "Local T_IP.IC",
     xlab = "standardized effect size of T_IP.IC",
     ylab = "modeled parameters: range/sd")
abline(v = mean(SSES.inf.MEAN.norm_Tipic, na.rm = T))

points(meanSES.2loc.uni_Tipic, modeled_param_uni.loc, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(SSES.inf.MEAN.uni_Tipic, na.rm = T), col = "purple")
rect(mean(SSES.inf.MEAN.norm_Tipic, na.rm = T), -1, 0, max(modeled_param_norm.loc) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(SSES.inf.MEAN.uni_Tipic, na.rm = T), 0, 0, max(modeled_param_uni.loc),
     col = rgb(0.5, 0, 1, 0.3), border = NA)
```

There is a very high correlation coefficient between initial and modeled parameter either for the range (trait a: Pearson correlation = 0.608; trait b: Pearson correlation = 0.615) and for the standard error (trait a and b: Pearson correlation = 1). Consequently the consistency between the figures 9 and 10 is not surprising.

3.2.2 Global p-values results for $T_{IP/IC}$

In contrast with local p-values, we call global p-values the p-values corresponding to one index for one trait across all the communities.

3.2.2.1 Global $T_{IP/IC}$ SES values against initial parameter values

Again, this first results mix different initial parameter values. Now, we can plot the SES values in relation to the strength of the internal filter assessed by the two parameters (either the initial values: `mean_range_between_com` and `mean_sd_of_com` or the modeled values `mean_range_com` and `sd.com_stock2`⁸).

First, we need to compute the SES values from simulations.

```
meanSES.2glob.norm_Tipic <- unlist(lapply(res.simu2, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_1$ses[,1], na.rm = T)))
meanSES.2glob.uni_Tipic <- unlist(lapply(res.simu2, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_1$ses[,2], na.rm = T)))

meanSES.INF_glob.norm_Tipic <- unlist(lapply(res.simu2, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_1$ses.inf[, 1], na.rm = T)))
meanSES.INF_glob.uni_Tipic <- unlist(lapply(res.simu2, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_1$ses.inf[, 2], na.rm = T)))
```

Plot the result against initial parameters (Fig. 11).

```
plot(meanSES.2glob.norm_Tipic, init_param, pch = 16, col = rgb(0, 0, 0, 0.6),
  main = "Global T_IP.IC",
  xlim = c(min(c(meanSES.2glob.norm_Tipic, meanSES.2glob.uni_Tipic),
    na.rm = T), 0),
  xlab = "standardized effect size of T_IP.IC",
  ylab = "initial parameters: range/sd")
points(meanSES.2glob.uni_Tipic, init_param, pch = 16, col = rgb(0.5, 0, 1, 0.6),)
abline(v = mean(meanSES.INF_glob.norm_Tipic, na.rm = T))
abline(v = mean(meanSES.INF_glob.uni_Tipic, na.rm = T), col = "purple")
rect(mean(meanSES.INF_glob.norm_Tipic, na.rm = T), -1, 0, max(modeled_param_norm.loc) + 1,
  col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(meanSES.INF_glob.uni_Tipic, na.rm = T), 0, 0, max(modeled_param_uni.loc),
  col = rgb(0.5, 0, 1, 0.3), border = NA)
```

3.2.2.2 Global $T_{IP/IC}$ SES values against modeled parameter values

We can verify the consistency of our results by plotting SES values against the modeled parameters values (Fig. 12).

⁸This modeled values are stochastic versions of the initial values.

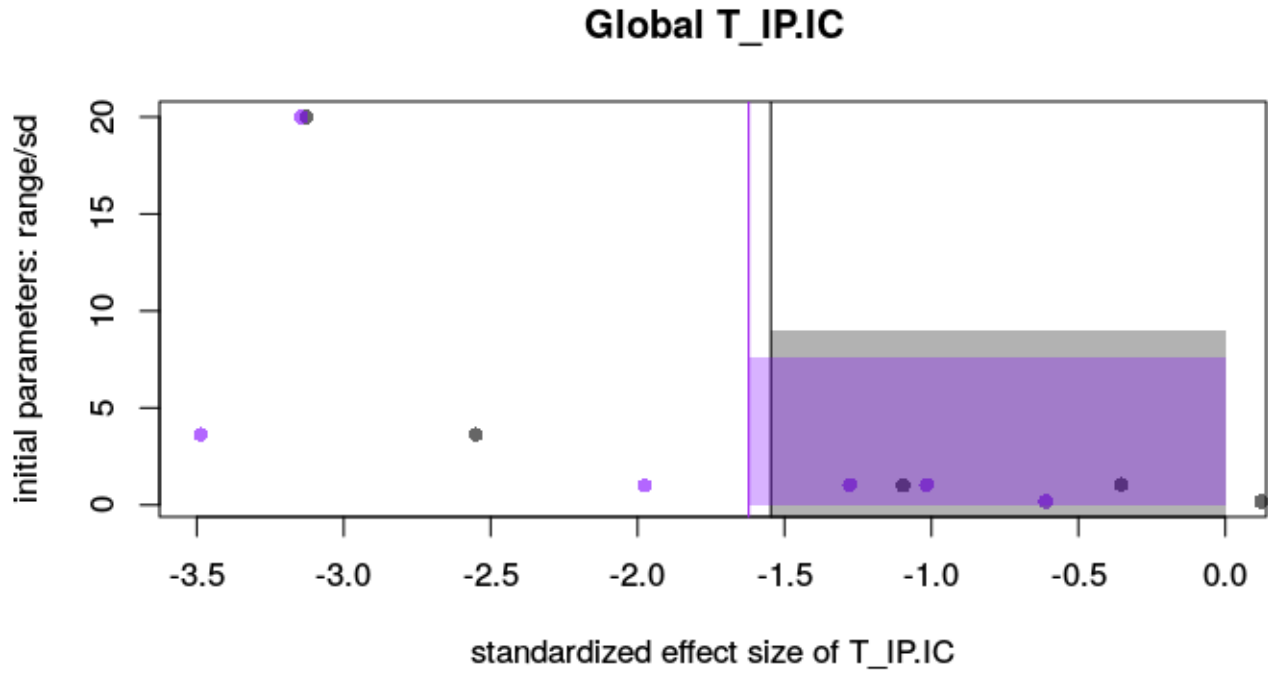


Figure 11: Global $T_{IP/IC}$ SES and initial parameters values: Standardized effect size of $T_{IP/IC}$ in function of the strength of internal filter defined by the ratio of the initial range parameter by the initial standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

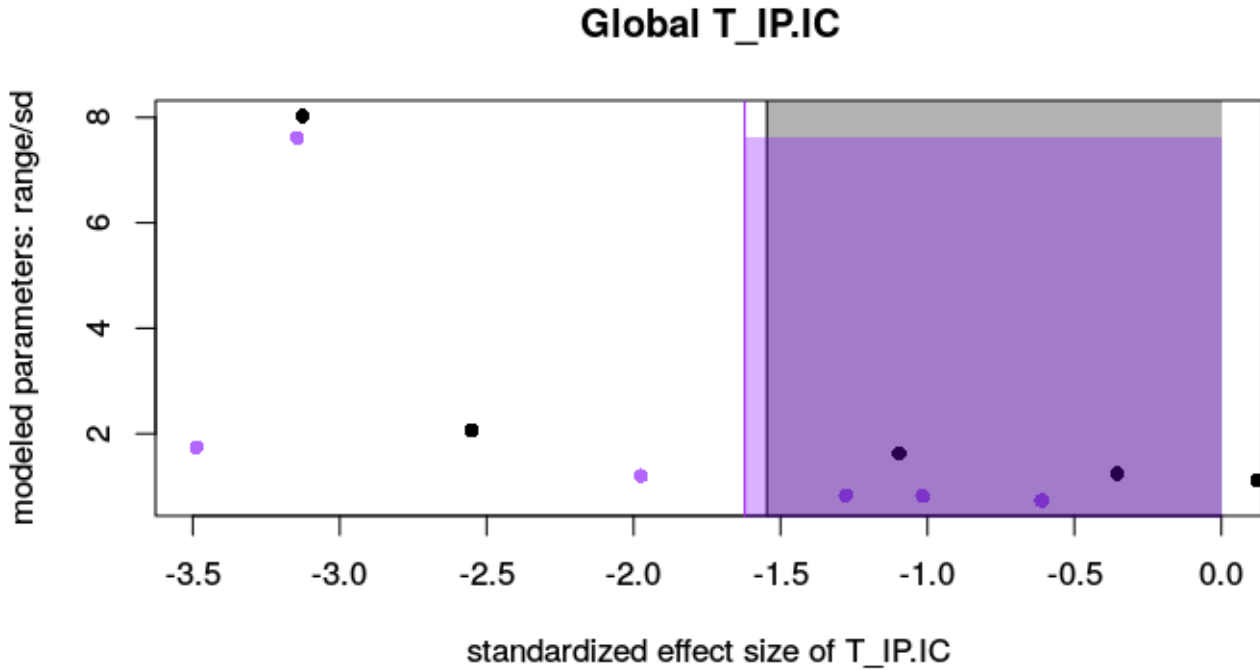


Figure 12: Global $T_{IP/IC}$ SES and modeled parameters values: Standardized effect size of $T_{IP/IC}$ in function of the strength of internal filter defined by the ratio of the modeled range parameter by the modeled standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

```
mean_dist_sp.interm <- t(matrix(unlist(lapply(mean_dist_sp, function(x) x)), nrow = 2))
mean_sd_of_com <- unlist(lapply(sd.sp_stock2, function(x) mean(x)))

modeled_param_norm <- mean_dist_sp.interm[, 1] / mean_sd_of_com
modeled_param_uni <- mean_dist_sp.interm[, 2] / mean_sd_of_com

plot(meanSES.2glob.norm_Tipic, modeled_param_norm, pch = 16,
     main = "Global T_IP.IC",
     xlim = c(min(c(meanSES.2glob.norm_Tipic, meanSES.2glob.uni_Tipic),
                    na.rm = T), 0),
     ylim = c(min(c(modeled_param_norm, modeled_param_uni), na.rm = T),
              max(c(modeled_param_norm, modeled_param_uni), na.rm = T)),
     xlab = "standardized effect size of T_IP.IC",
     ylab = "modeled parameters: range/sd")
points(meanSES.2glob.uni_Tipic, modeled_param_uni, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(meanSES.INF_glob.norm_Tipic, na.rm = T))
abline(v = mean(meanSES.INF_glob.uni_Tipic, na.rm = T), col = "purple")
rect(mean(meanSES.INF_glob.norm_Tipic, na.rm = T), -1, 0, max(modeled_param_norm) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(meanSES.INF_glob.uni_Tipic, na.rm = T), 0, 0, max(modeled_param_uni),
     col = rgb(0.5, 0, 1, 0.3), border = NA)
```


3.3 Conclusion on the power of $T_{IP/IC}$ to detect internal filtering

To conclude on the power of $T_{IP/IC}$ to detect internal filtering, we compute exponential linear model and identify the strength from which the beta-error is inferior to 0.05. This strength is defined by the two ratio of the initial parameters `param_DIST_between_MEAN` and `sd_mean.param`.

```
plot(meanSES.2loc.norm_Tipic ~ init_param.loc, col = rgb(0, 0, 0, 0.2),
     ylim = c(min(c(meanSES.2loc.norm_Tipic, meanSES.2loc.uni_Tipic,
                    meanSES.2glob.norm_Tipic, meanSES.2glob.uni_Tipic), na.rm = T), 0),
     main = "T_IP.IC",
     ylab = "standardized effect size of T_IP.IC",
     xlab = "initial parameters: range/sd",
     type = "n")
#points(meanSES.2loc.uni_Tipic ~ init_param.loc, col = rgb(0.5, 0, 1, 0.2))

points(meanSES.2glob.norm_Tipic ~ init_param, pch = 16, col = rgb(0, 0, 0, 0.5),
       cex = 1.2)
lm.norm <- lm(meanSES.2glob.norm_Tipic ~ log(init_param))
lm.norm_conf <- confint(lm.norm, level = 0.90)
curve(lm.norm$coef[1] + log(x) * (lm.norm$coef[2]), add = T, lwd = 3)
curve(lm.norm_conf [1, 1] + log(x) * lm.norm_conf [2, 1], add = T, lty = 2)
curve(lm.norm_conf [1, 2] + log(x) * lm.norm_conf [2, 2], add = T, lty = 2)

points(meanSES.2glob.uni_Tipic ~ init_param, pch = 16, col = rgb(0.5, 0, 1, 0.5),
       cex = 1.2)
lm.uni <- lm(meanSES.2glob.uni_Tipic ~ log(init_param))
lm.uni_conf <- confint(lm.uni, level = 0.90)
curve(lm.uni$coef[1] + log(x) * (lm.uni$coef[2]), add = T, lwd = 3,
      col = rgb(0.5, 0, 1, 1))
curve(lm.uni_conf [1, 1] + log(x) * lm.uni_conf [2, 1], add = T, lty = 2,
      col = rgb(0.5, 0, 1, 1))
curve(lm.uni_conf [1, 2] + log(x) * lm.uni_conf [2, 2], add = T, lty = 2,
      col = rgb(0.5, 0, 1, 1))

rect(-1, mean(c(SSES.inf.MEAN.uni_Tipic, meanSES.INF_glob.uni_Tipic), na.rm = T),
     max(init_param), 10, col = rgb(0.5, 0, 1, 0.3), border = NA)

rect(-1, mean(c(SSES.inf.MEAN.norm_Tipic, meanSES.INF_glob.norm_Tipic), na.rm = T),
     max(init_param) + 10, 10, col = rgb(0, 0, 0, 0.3), border = NA)

yy.norm <- mean(c(SSES.inf.MEAN.norm_Tipic, meanSES.INF_glob.norm_Tipic), na.rm = T)
yy.uni <- mean(c(SSES.inf.MEAN.uni_Tipic, meanSES.INF_glob.uni_Tipic), na.rm = T)

param_beta0.05_Tipic.norm <- exp((yy.norm - lm.norm_conf [1, 2]) / lm.norm_conf [2, 2])
param_beta0.05_Tipic.uni <- exp((yy.uni - lm.uni_conf [1, 2]) / lm.uni_conf [2, 2])
```

For the trait "a" normally distributed, the power of $T_{IP/IC}$ is satisfactory if the ratio of the initial parameter is superior to 86.07. For the trait "b" uniformly distributed, this value is 17.35. Further investigation on real data with different strength of internal filtering are needed to complete these simulated values.

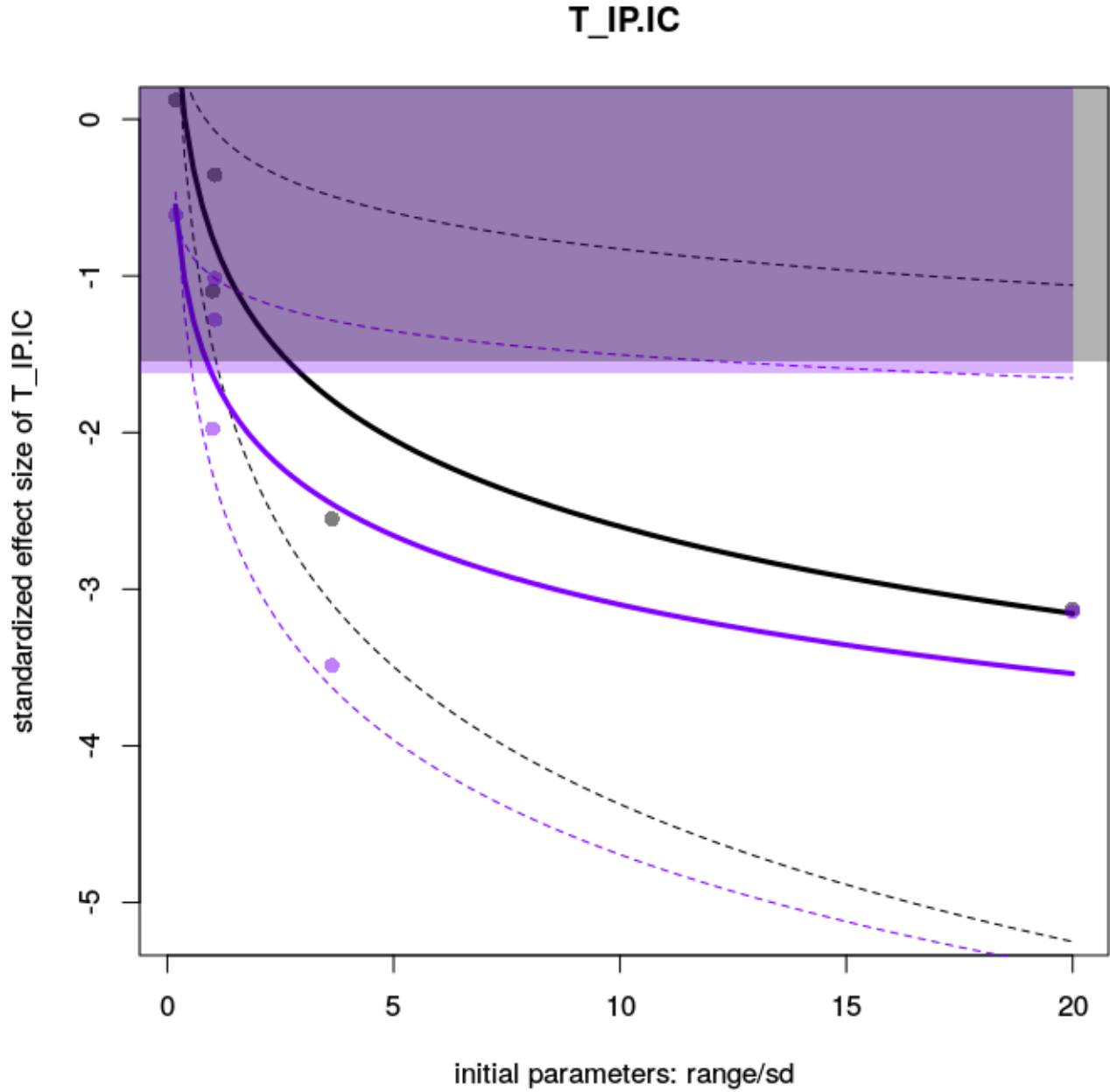


Figure 13: Power of $T_{IP/IC}$ to detect internal filtering: Standardised Effect Size (SES) of $T_{IP/IC}$ in relation to the strength of internal filtering. Trait 'a' normally distributed is shown in black, traits 'b' uniformly distributed in purple. Open circle represent local SES values whereas close one represent global SES values. The exponential linear model is presented with 5 percents confidence interval on both side. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$)

4 External Filter

4.1 Randomization outline

As for internal filter we modulate the strength of the external filter thanks to two parameters: we define (i) a gradient of range between communities means traits distributions and (ii) a gradient of the mean variance in communities traits distributions. We expect the external filtering to increase (i) when the range in communities means increase (*i.e.* the environmental gradient is larger) and (ii) when the variance in the communities traits distributions decrease (*i.e.* the environmental gradient is stronger).

In analogy with the internal filtering model, we defined 3 values for two parameters:

(i) `param_RANGE.init` is a vector of length 3 which defined the above range for the mean traits values (the maximum value for the mean is assessed to 250). In decreasing the range of traits values while keeping the same number of communities, there is a decrease in the overlap of the communities traits distributions and consequently an decrease of the impact of external filtering in shaping traits distributions. Thus the mean for each communities is drawn from a normal distribution with standard deviation of 10 (parameter `mean_sd.param`) and means evenly ⁹ of 10 distributed between `max.value_traits - param_RANGE` and `max.value_traits`.

(ii) `param_SD` is a vector of length 3 which defined the standard deviation of traits distributions for each species. In order to decorrelate the mean and standard deviation of species traits distributions, `param_SD` is permuted before the analysis.

Thus for the trait "a" the trait value for each individual is drawn from a normal distribution with the mean depending on it community attribute. For the trait "b" the trait value for each individual is drawn from a uniform distribution with the range depending on it community attribute.

```
# Parameter for the range of communities traits distributions
param_range.init <- round(sort(seq(10, 200, length.out = nb_param_val)), 2)
param_range <- rep(param_range.init, N_repet_Param)
mean_sd.param <- rep(10, nperm)
```

```
# Parameter for the variance in species mean trait values
param_SD <- seq(10, 100, length.out = nb_param_val)
sd_mean.param <- sample(rep(param_SD, N_repet_Param),
  size = length(rep(param_SD, N_repet_Param)), replace = F)
sd_sd.param <- rep(10, nperm)
```

```
nperm <- N_repet_Param*length(param_SD)

mean_range_com <- list() ; sd.com_stock3 <- list() ; res.simu3 <- list()
res.simu3.pval <- list() ; res.simu.traits3 <- list()

for(n in 1:nperm){#for each permutation

  ex.sp3 <- c()
```

⁹In fact, we draw this values in a normal distribution with means evenly distributed and with a standard error equal to `mean_sd.param`

```

ex.com3 <- matrix(0, nrow = Ncom, ncol = Nsp)
for(i in 1: 10){
  ex.com.interm <- table(sample(sp, size = Nind/Ncom, prob = rlnorm(Nsp, 0, sdlog), replace = T))
  ex.com3[i, sp%in% names(ex.com.interm)] <- ex.com.interm
  ex.sp3 <- c(ex.sp3, rep(sp, times = ex.com3[i,]))
}

ex.indplot3 <- sort(as.factor(rep(com, Nind/Ncom)))

#Defining mean and sd by community
mean_mean.param.interm <- seq(max.value_traits - param_range[n], max.value_traits,
                              length.out = length(unique(param_range)))

mean_mean.param <- rep(round(sort(mean_mean.param.interm), 2), N_repet_Param)

mean.com <- rnorm(length(unique(com)), mean = mean_mean.param, sd = mean_sd.param[n])
sd.com <- rnorm(length(unique(com)), mean = sd_mean.param[n], sd = sd_sd.param[n])

ex.traits3 <- array(NA, dim = c(Nind, 2))
colnames(ex.traits3) <- paste("trait", c("a", "b"), sep = " ")

for(c in unique(ex.indplot3)){
  #trait a : normal distribution
  ex.traits3[ex.indplot3 == c, 1] <-
  rnorm(500, rep(mean.com[unique(ex.indplot3) == c], 500),
        rep(sd.com[unique(ex.indplot3) == c], 500))[1:sum(ex.indplot3 == c)]

  #trait b : uniform distribution
  ex.traits3[ex.indplot3 == c, 2] <-
  runif(500, min = rep(mean.com[unique(ex.indplot3) == c], 500) -
        rep(sd.com[unique(ex.indplot3) == c], 500),
        max = rep(mean.com[unique(ex.indplot3) == c], 500) +
        rep(sd.com[unique(ex.indplot3) == c], 500))[1:sum(ex.indplot3 == c)]
}

#stock results
mean_range_com[[n]] <- c(max(ex.traits3[, 1], na.rm = T) - min(ex.traits3[, 1], na.rm = T),
                        max(ex.traits3[, 2], na.rm = T) - min(ex.traits3[, 2], na.rm = T))
sd.com_stock3[[n]] <- sd.com

res.simu.traits3[[n]] <- ex.traits3
res.simu3[[n]] <- Tstats(ex.traits3, ex.indplot3, ex.sp3)
res.simu3.pval[[n]] <- sum_Tstats(res.simu3[[n]], type = "p.value")
print(paste("----", round(n/nperm, 2) * 100, "%", sep = " "))
}#End of simulations

```

4.2 Results of the simulations of external filtering

Let's see the result for one randomization. We can plot the distribution of traits values within species and/or communities thanks to the function `plotDistri` (Fig. 14).

```

par(mfrow=c(2, 2))
plotDistri(ex.traits3, rep("all_sp", times = dim(ex.traits3)[1]), ex.indplot3,
           plot.ask = F, multipanel = F, main = c("a", "b"), leg = c(T, F))
plotDistri(ex.traits3, rep("region", times = dim(ex.traits3)[1]), ex.sp3, plot.ask = F,
           multipanel = F, main = c("c", "d"), leg = c(T, F))
par(mfrow=c(1, 1))

```

We can also plot the result of the T-statistics for two contrasting case (Fig. 15).

```

par(mfrow=c(2, 1))
plot(res.simu3[[1]], main = "a")
plot(res.simu3[[n]], main = "b")
par(mfrow=c(1, 1))

```

4.2.1 Local $T_{IC/IR}$ and $T_{PC/PR}$ results

As for internal filtering, we mingle the p-value for different parameters values. Consequently we plot the standard effect size (SES) of $T_{IC/IR}$ and $T_{PC/PR}$ in relation to the initial parameters values.

4.2.1.1 Local $T_{IC/IR}$ and $T_{PC/PR}$ SES values against initial parameter values

```

#T_IC.IR
meanSES.3loc.norm_Ticir <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses[, 1])))

meanSES.3loc.uni_Ticir <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses[, 2])))

SES.inf.MEAN.norm_Ticir <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses.inf[, 1])))

SES.inf.MEAN.uni_Ticir <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses.inf[, 2])))

#T_PC.PR
meanSES.3loc.norm_Tpcpr <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_3$ses[, 1])))

meanSES.3loc.uni_Tpcpr <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_3$ses[, 2])))

SES.inf.MEAN.norm_Tpcpr <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))

```

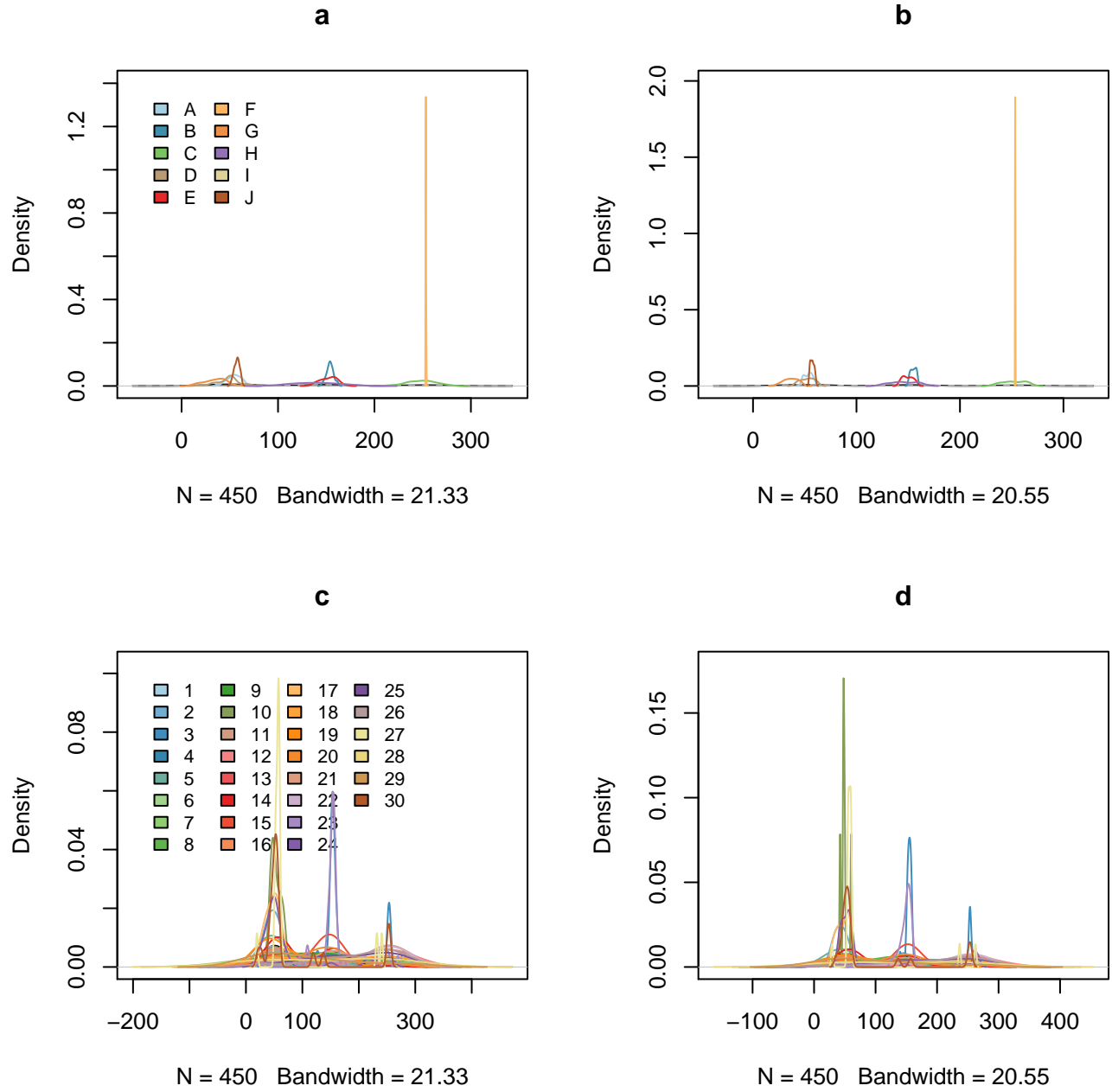


Figure 14: Distribution of trait values for one randomization with external filtering: (a) Community-level trait distributions for the trait a (normal distribution); (b) Community-level trait distributions for the trait b (uniform distribution); (c) Species' trait distributions for the trait a; (d) Species' trait distributions for the trait b

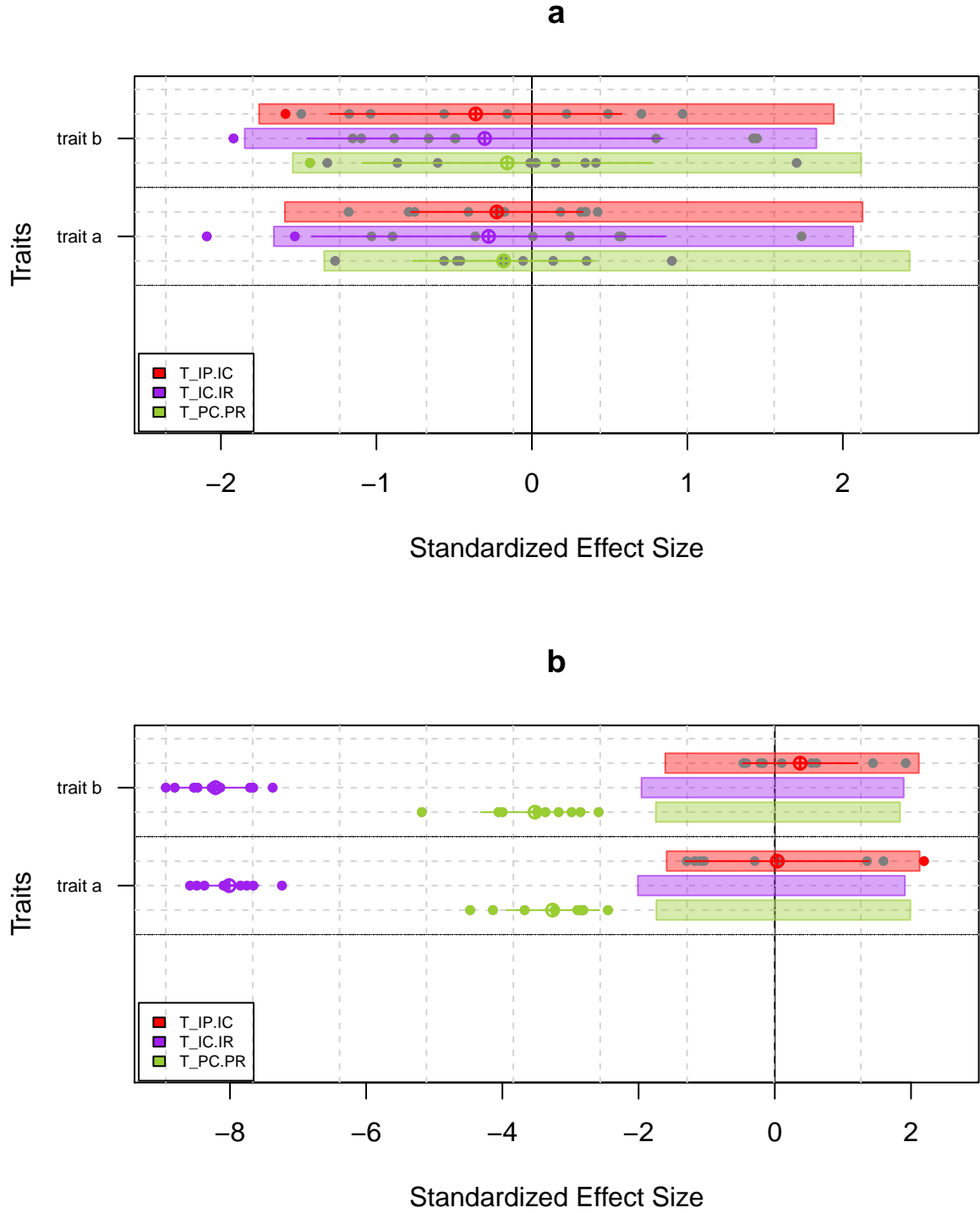


Figure 15: Results of T-statistics for two randomization with contrasted strength of external filtering: $T_{IP/IC}$ in red, $T_{IC/IR}$ in purple and $T_{PC/PR}$ in green. (a) Lower strength of external filtering and (b) higher strength of external filtering.

```

    $index_1_3$ses.inf[, 1]))
SES.inf.MEAN.uni_Tpcpr <- unlist(lapply(res.simu3, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_3$ses.inf[, 2]))

```

Now, we can plot SES of $T_{IC/IR}$ (Fig. 16) and $T_{PC/PR}$ (Fig. 17). The trait "a" normally distributed is in black and the uniform traits "b" is in purple. The colored rectangles represent the null model area with $\alpha = 5\%$. Thus, when a point is outside these area, the modeled parameters are strong enough to detect the external filter with a high power (beta-error < 0.05).

```

init_param <- param_range / sd_mean.param
#each value is replicate Ncom times (there is Ncom communities for each parameter values)
init_param.loc <- rep(init_param, each = Ncom)

plot(meanSES.3loc.norm_Ticir, init_param.loc, pch = 16, col = rgb(0, 0, 0, 0.6),
     xlim = c(min(c(unlist(meanSES.3loc.norm_Ticir), unlist(meanSES.3loc.uni_Ticir))),
     na.rm = T), 0),
     main = "Local T_IC.IR",
     xlab = "standardized effect size of T_IC.IR",
     ylab = "initial parameters: range/sd")
abline(v = mean(SSES.inf.MEAN.norm_Ticir, na.rm = T))

points(meanSES.3loc.uni_Ticir, init_param.loc, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(SSES.inf.MEAN.uni_Ticir, na.rm = T), col = "purple")
rect(mean(SSES.inf.MEAN.norm_Ticir, na.rm = T), -1, 0, max(init_param) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(SSES.inf.MEAN.uni_Ticir, na.rm = T), 0, 0, max(init_param),
     col = rgb(0.5, 0, 1, 0.3), border = NA)

plot(meanSES.3loc.norm_Tpcpr, init_param.loc, pch = 16, col = rgb(0, 0, 0, 0.6),
     xlim = c(min(c(unlist(meanSES.3loc.norm_Tpcpr), unlist(meanSES.3loc.uni_Tpcpr))),
     na.rm = T), 0),
     main = "Local T_PC.PR",
     xlab = "standardized effect size of T_PC.PR",
     ylab = "initial parameters: range/sd")
abline(v = mean(SSES.inf.MEAN.norm_Tpcpr, na.rm = T))

points(meanSES.3loc.uni_Tpcpr, init_param.loc, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(SSES.inf.MEAN.uni_Tpcpr, na.rm = T), col = "purple")
rect(mean(SSES.inf.MEAN.norm_Tpcpr, na.rm = T), -1, 0, max(init_param) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(SSES.inf.MEAN.uni_Tpcpr, na.rm = T), 0, 0, max(init_param),
     col = rgb(0.5, 0, 1, 0.3), border = NA)

```

For each initial parameters values we can calculate the proportion of case where we do not reject the null hypothesis while this hypothesis is wrong (*i.e.* the Beta-error). The next chunk only shows the calculation for the $T_{IC/IR}$ metrics on the trait "a".

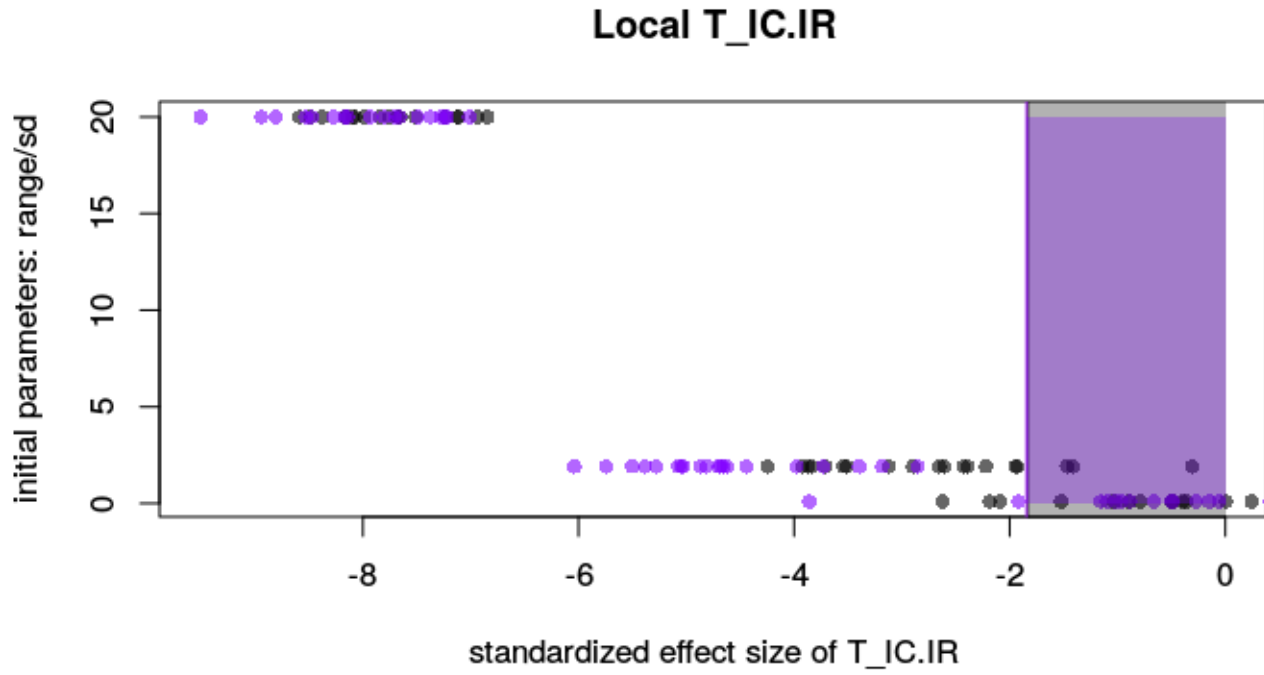


Figure 16: Local $T_{IC/IR}$ SES and initial parameters values: Standardized effect size of $T_{IC/IR}$ in function of the strength of external filter defined by the ratio of the initial range parameter by the initial standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

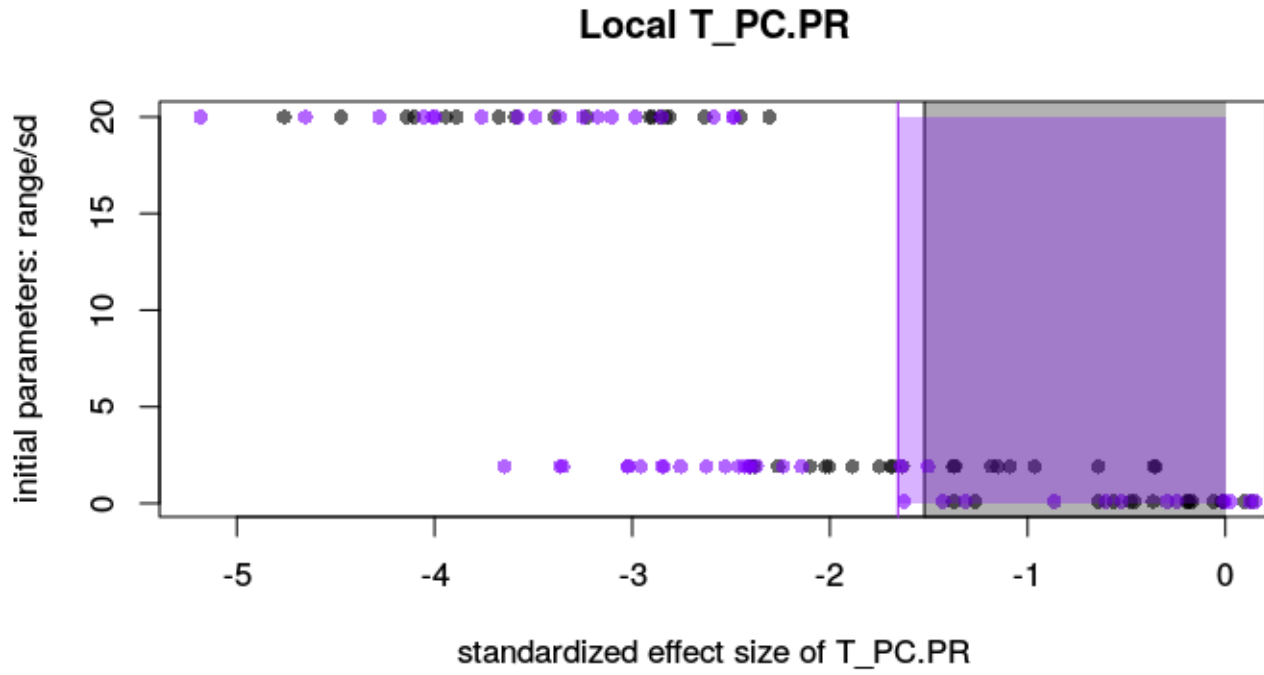


Figure 17: Local $T_{PC/PR}$ SES and initial parameters values: Standardized effect size of $T_{PC/PR}$ in function of the strength of external filter defined by the ratio of the initial range parameter by the initial standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

```

beta_range_norm_Ticir <-
  (unlist(lapply(by(cbind(meanSES.3loc.norm_Ticir, SES.inf.MEAN.norm_Ticir),
    rep(param_range, each = Ncom), function(x) x[,1] > x[,2]),
    function(x) sum(x, na.rm = T)))+1)/(10 * N_repet_Param + 1)

print(xtable(res_beta_Ticir, caption = 'Local beta-error of  $T_{IC/IR}$  in function
  of the strength of external filtering. str: strength. Trait a is
  normally distributed and trait b is uniformly distributed.',
  label = "tab:local_beta_error_xtable_ticir"),
  caption.placement = "top", size = "small")

```

Table 4: Local beta-error of $T_{IC/IR}$ in function of the strength of external filtering. str: strength. Trait a is normally distributed and trait b is uniformly distributed.

	str 1	str 2	str 3
effect of sd (a)	0.81	0.24	0.05
effect of sd (b)	0.90	0.05	0.05
effect of range (a)	0.81	0.24	0.05
effect of range (b)	0.90	0.05	0.05

```

print(xtable(res_beta_Ticir, caption = 'Local beta-error of  $T_{PC/PR}$  in function
  of the strength of external filtering. str: strength. Trait a is
  normally distributed and trait b is uniformly distributed.',
  label = "tab:local_beta_error_xtable_tpcpr"),
  caption.placement = "top", size = "small")

```

Table 5: Local beta-error of $T_{PC/PR}$ in function of the strength of external filtering. str: strength. Trait a is normally distributed and trait b is uniformly distributed.

	str 1	str 2	str 3
effect of sd (a)	0.81	0.24	0.05
effect of sd (b)	0.90	0.05	0.05
effect of range (a)	0.81	0.24	0.05
effect of range (b)	0.90	0.05	0.05

Now we can see the beta-error for each strength (table 4 and 5). A high value of range and a low value of standard error defined a high strength of filtering. For example, in the case of a trait normally distributed, the proportion of false negative is 80.95% for the higher value of standard errors and thus for the lower strength of external filtering.

4.2.1.2 Local $T_{IC/IR}$ and $T_{PC/PR}$ SES values against modeled parameter values

First we need to calculate the modeled parameter values. Here again we calculate the strength of external filtering by dividing the range of values in communities by their standard error.

```

mean_range_com.interm <- t(matrix(unlist(lapply(mean_range_com, function(x) x)), nrow = 2))
mean_sd_of_com <- unlist(lapply(sd.com_stock3, function(x) mean(x)))

modeled_param_norm <- mean_range_com.interm[, 1] / mean_sd_of_com
modeled_param_uni <- mean_range_com.interm[, 2] / mean_sd_of_com
modeled_param_norm.loc <- rep(modeled_param_norm, each = Ncom)
modeled_param_uni.loc <- rep(modeled_param_uni, each = Ncom)

```

```

plot(meanSES.3loc.norm_Ticir, modeled_param_norm.loc, pch = 16, col = rgb(0, 0, 0, 0.6),
     xlim = c(min(c(unlist(meanSES.3loc.norm_Ticir), unlist(meanSES.3loc.uni_Ticir)),
                 na.rm = T), 0),
     ylim = c(0, min(max(c(modeled_param_norm.loc, modeled_param_uni.loc), na.rm = T),
                      100)), main = "Local T_IC.IR",
     xlab = "standardized effect size of T_IC.IR",
     ylab = "modeled parameters: range/sd")
abline(v = mean(SSES.inf.MEAN.norm_Ticir, na.rm = T))

points(meanSES.3loc.uni_Ticir, modeled_param_uni.loc, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(SSES.inf.MEAN.uni_Ticir, na.rm = T), col = "purple")

rect(mean(SSES.inf.MEAN.norm_Ticir, na.rm = T), -1, 0, max(modeled_param_norm.loc) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(SSES.inf.MEAN.uni_Ticir, na.rm = T), 0, 0, max(modeled_param_uni.loc),
     col = rgb(0.5, 0, 1, 0.3), border = NA)

```

```

plot(meanSES.3loc.norm_Tpcpr, modeled_param_norm.loc, pch = 16, col = rgb(0, 0, 0, 0.6),
     xlim = c(min(c(unlist(meanSES.3loc.norm_Tpcpr), unlist(meanSES.3loc.uni_Tpcpr)),
                 na.rm = T), 0),
     ylim = c(0, min(max(c(modeled_param_norm.loc, modeled_param_uni.loc), na.rm = T),
                      100)),
     main = "Local T_PC.PR",
     xlab = "standardized effect size of T_PC.PR",
     ylab = "modeled parameters: range/sd")
abline(v = mean(SSES.inf.MEAN.norm_Tpcpr, na.rm = T))

points(meanSES.3loc.uni_Tpcpr, modeled_param_uni.loc, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(SSES.inf.MEAN.uni_Tpcpr, na.rm = T), col = "purple")
rect(mean(SSES.inf.MEAN.norm_Tpcpr, na.rm = T), -1, 0, max(modeled_param_norm.loc) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(SSES.inf.MEAN.uni_Tpcpr, na.rm = T), 0, 0, max(modeled_param_uni.loc),
     col = rgb(0.5, 0, 1, 0.3), border = NA)

```

There is a very high correlation coefficient between initial and modeled parameter either for the range (trait a: Pearson correlation = 0.883; trait b: Pearson correlation = 0.9) and for the standard error (trait a and b: Pearson correlation = 0.998). Consequently the consistency between the figures 16 - 17 and 18 - 19 is not surprising.

4.2.2 Global $T_{IC/IR}$ and $T_{PC/PR}$ results

In contrast with local p-values, we call global p-values the p-values corresponding to one index for one trait across all the communities.

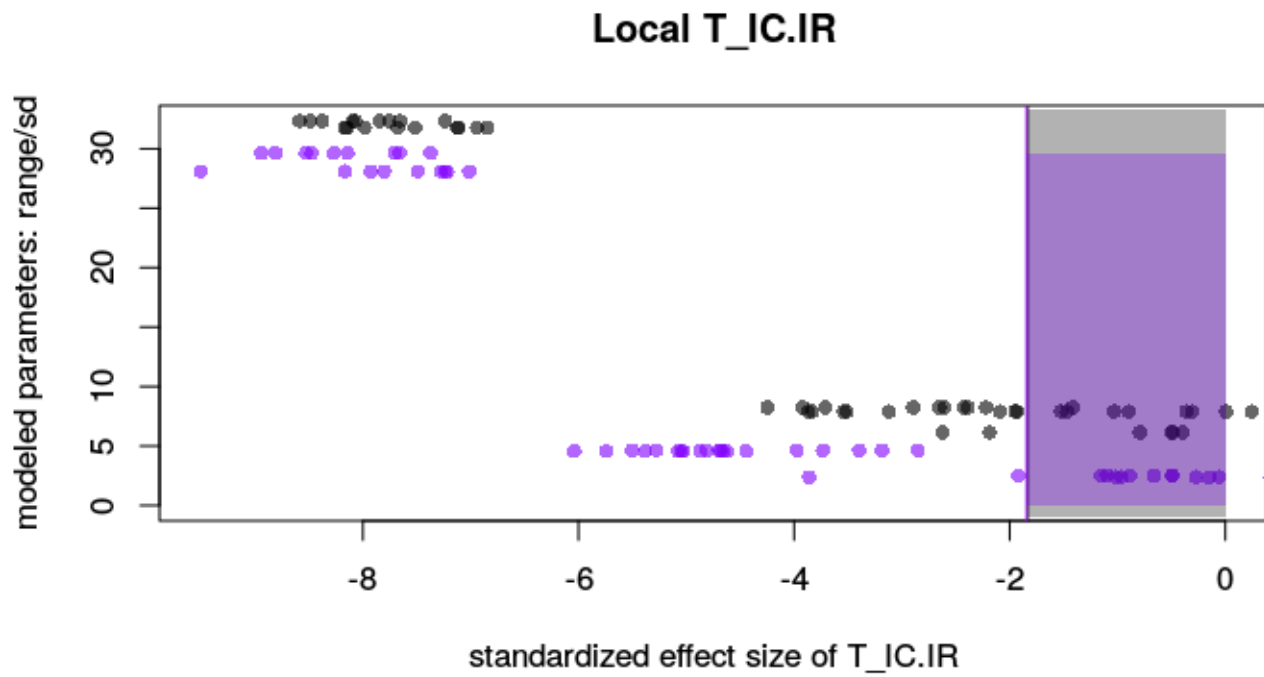


Figure 18: Local $T_{IC/IR}$ SES and modeled parameters values: Standardized effect size of $T_{IC/IR}$ in function of the strength of external filter defined by the ratio of the modeled range parameter by the modeled standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

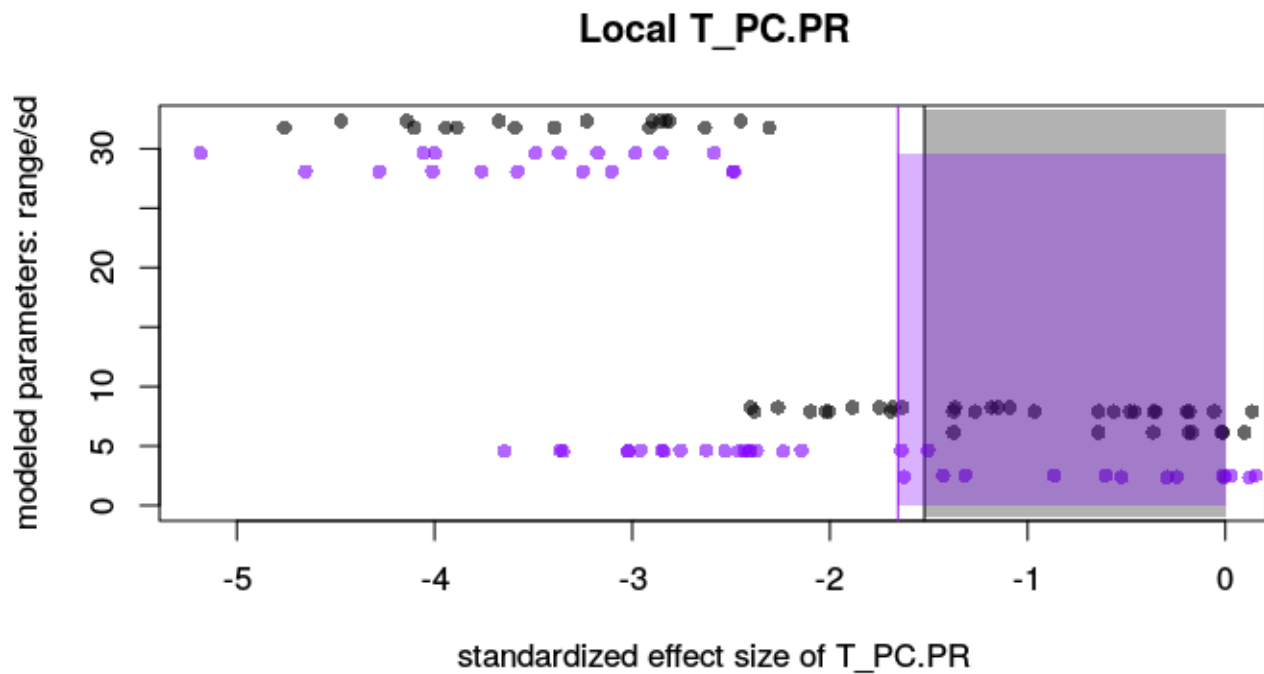


Figure 19: Local $T_{PC/PR}$ SES and modeled parameters values: Standardized effect size of $T_{PC/PR}$ in function of the strength of external filter defined by the ratio of the modeled range parameter by the modeled standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

4.2.2.1 Calculation of beta-error mixing all strength of filtering

To illustrate the global power of $T_{IC/IR}$ and $T_{PC/PR}$ you can plot the ordered p-values and calculate a beta-error thanks to the code below. It's just an illustration because we mingle very different strength of external filter with different initial parameter values. Plot results are not shown in this document.

```
par(mfrow = c(2, 2))
xx <- log10(sort(unlist(lapply(res.simu3.pval, function(x) x[21:30, 1]))))
plot(xx, type = "l", main = "T_IC.IR_distribNorm")
abline(h = log10(0.05))
nbre_beta_error <- round((sum(xx>log10(0.05))+1)/(length(xx)+1), 4)
text(0, -0.8, labels = paste("beta error", nbre_beta_error, sep = " = "), cex = 0.7, pos = 4)

xx <- log10(sort(unlist(lapply(res.simu3.pval, function(x) x[41:50, 1]))))
plot(xx, type = "l", main = "T_PC.PR_distribNorm")
abline(h = log10(0.05))
nbre_beta_error <- round((sum(xx>log10(0.05))+1)/(length(xx)+1), 4)
text(0, -0.8, labels = paste("beta error", nbre_beta_error, sep = " = "), cex = 0.7, pos = 4)

xx <- log10(sort(unlist(lapply(res.simu3.pval, function(x) x[21:30, 2]))))
plot(xx, type = "l", main = "T_IC.IR_distribUni")
abline(h = log10(0.05))
nbre_beta_error <- round((sum(xx>log10(0.05))+1)/(length(xx)+1), 4)
text(0, -0.8, labels = paste("beta error", nbre_beta_error, sep = " = "), cex = 0.7, pos = 4)

xx <- log10(sort(unlist(lapply(res.simu3.pval, function(x) x[41:50, 2]))))
plot(xx, type = "l", main = "T_PC.PR_distribUni")
abline(h = log10(0.05))
nbre_beta_error <- round((sum(xx>log10(0.05))+1)/(length(xx)+1), 4)
text(0, -0.8, labels = paste("beta error", nbre_beta_error, sep = " = "), cex = 0.7, pos = 4)

par(mfrow = c(1, 1))
```

4.2.2.2 Global $T_{IC/IR}$ and $T_{PC/PR}$ SES values against initial parameter values

Again, this first results mix different initial parameter values. Now, we can plot the SES values in relation to the strength of the external filter assessed by the two parameters (either the initial values: `mean_range_between_com` and `mean_sd_of_com` or the modeled values `mean_range_com` and `sd.com_stock3`¹⁰).

First, we need to compute the SES values from simulations.

```
meanSES.3glob.norm_Ticir <- unlist(lapply(res.simu3, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_2$ses[,1], na.rm = T)))
meanSES.3glob.uni_Ticir <- unlist(lapply(res.simu3, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_2$ses[,2], na.rm = T)))
meanSES.3glob.norm_Tpcpr <- unlist(lapply(res.simu3, function(x)
```

¹⁰This modeled values are stochastic versions of the initial values.

```

      mean(ses.listofindex(as.listofindex(x))
        $index_1_3$ses[,1], na.rm = T)))
meanSES.3glob.uni_Tpcpr <- unlist(lapply(res.simu3, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_3$ses[,2], na.rm = T)))

meanSES.INF_glob.norm_Ticir <- unlist(lapply(res.simu3, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_2$ses.inf[, 1], na.rm = T)))
meanSES.INF_glob.uni_Ticir <- unlist(lapply(res.simu3, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_2$ses.inf[, 2], na.rm = T)))
meanSES.INF_glob.norm_Tpcpr <- unlist(lapply(res.simu3, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_3$ses.inf[, 1], na.rm = T)))
meanSES.INF_glob.uni_Tpcpr <- unlist(lapply(res.simu3, function(x)
  mean(ses.listofindex(as.listofindex(x))
    $index_1_3$ses.inf[, 2], na.rm = T)))

```

Plot the result against initial parameters (Fig. 20 and 21).

```

init_param <- param_range / sd_mean.param

plot(meanSES.3glob.norm_Ticir, init_param, pch = 16, col = rgb(0, 0, 0, 0.6),
  main = "Global T_IC.IR",
  xlim = c(min(c(meanSES.3glob.norm_Ticir, meanSES.3glob.uni_Ticir),
    na.rm = T), 0),
  xlab = "standardized effect size of T_IC.IR",
  ylab = "initial parameters: range/sd")
points(meanSES.3glob.uni_Ticir, init_param, pch = 16, col = rgb(0.5, 0, 1, 0.6),)
abline(v = mean(meanSES.INF_glob.norm_Ticir, na.rm = T))
abline(v = mean(meanSES.INF_glob.uni_Ticir, na.rm = T), col = "purple")
rect(mean(meanSES.INF_glob.norm_Ticir, na.rm = T), -1, 0, max(modeled_param_norm.loc) + 1,
  col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(meanSES.INF_glob.uni_Ticir, na.rm = T), 0, 0, max(modeled_param_uni.loc),
  col = rgb(0.5, 0, 1, 0.3), border = NA)

```

```

plot(meanSES.3glob.norm_Tpcpr, init_param, pch = 16, col = rgb(0, 0, 0, 0.6),
  main = "Global T_PC.PR",
  xlim = c(min(c(meanSES.3glob.norm_Tpcpr, meanSES.3glob.uni_Tpcpr),
    na.rm = T), 0),
  xlab = "standardized effect size of T_PC.PR",
  ylab = "initial parameters: range/sd")
points(meanSES.3glob.uni_Tpcpr, init_param, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(meanSES.INF_glob.norm_Tpcpr, na.rm = T))
abline(v = mean(meanSES.INF_glob.uni_Tpcpr, na.rm = T), col = "purple")
rect(mean(meanSES.INF_glob.norm_Tpcpr, na.rm = T), -1, 0, max(modeled_param_norm.loc) + 1,
  col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(meanSES.INF_glob.uni_Tpcpr, na.rm = T), 0, 0, max(modeled_param_uni.loc),
  col = rgb(0.5, 0, 1, 0.3), border = NA)

```

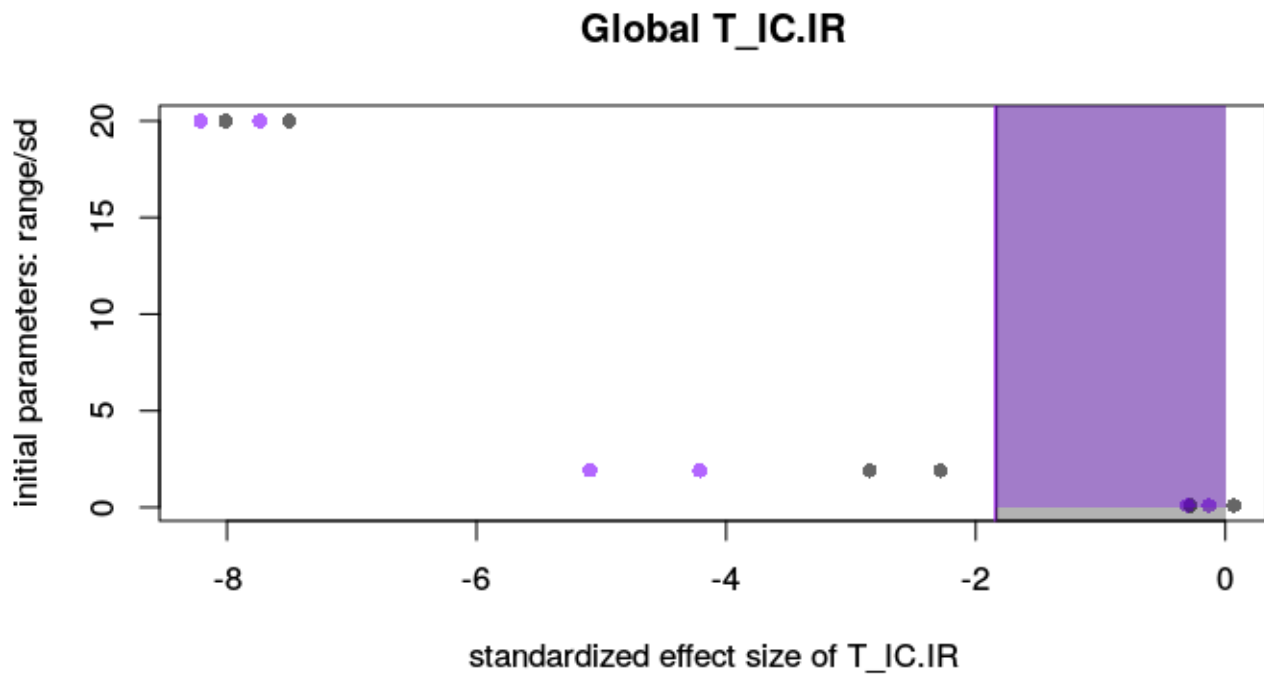



Figure 20: Global $T_{IC/IR}$ SES and initial parameters values: Standardized effect size of $T_{IC/IR}$ in function of the strength of external filter defined by the ratio of the initial range parameter by the initial standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

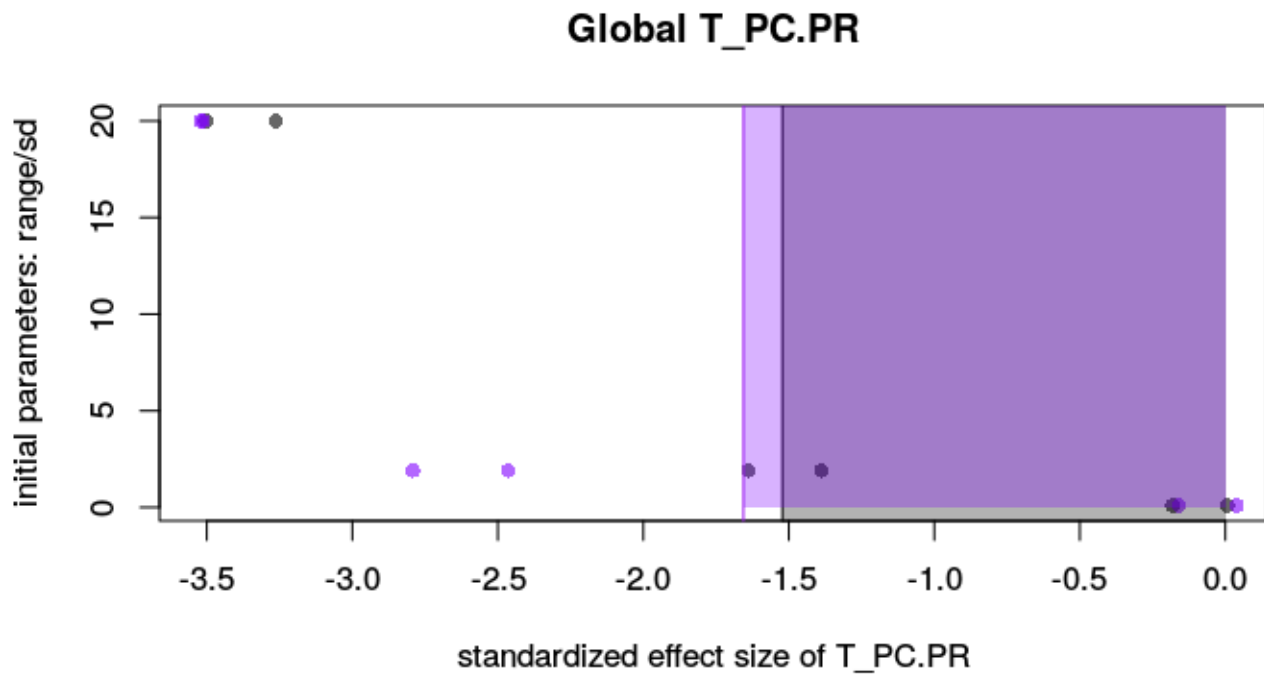


Figure 21: Global $T_{PC/PR}$ SES and initial parameters values: Standardized effect size of $T_{PC/PR}$ in function of the strength of external filter defined by the ratio of the initial range parameter by the initial standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

4.2.2.3 Global $T_{IC/IR}$ and $T_{PC/PR}$ SES values against modeled parameter values

We can verify the consistency of our results by plotting SES values against the modeled parameters values (Fig. 22 and 23).

```
mean_range_com.interm <- t(matrix(unlist(lapply(mean_range_com, function(x) x)), nrow = 2))
mean_sd_of_com <- unlist(lapply(sd.com_stock3, function(x) mean(x)))

modeled_param_norm <- mean_range_com.interm[, 1] / mean_sd_of_com
modeled_param_uni <- mean_range_com.interm[, 2] / mean_sd_of_com

plot(meanSES.3glob.norm_Ticir, modeled_param_norm, pch = 16,
     main = "Global T_IC.IR",
     xlim = c(min(c(meanSES.3glob.norm_Ticir, meanSES.3glob.uni_Ticir), na.rm = T), 0),
     ylim = c(0, min(max(c(modeled_param_norm, modeled_param_uni), na.rm = T), 100)),
     xlab = "standardized effect size of T_IC.IR",
     ylab = "modeled parameters: range/sd")
points(meanSES.3glob.uni_Ticir, modeled_param_uni, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(meanSES.INF_glob.norm_Ticir, na.rm = T))
abline(v = mean(meanSES.INF_glob.uni_Ticir, na.rm = T), col = "purple")
rect(mean(meanSES.INF_glob.norm_Ticir, na.rm = T), -1, 0, max(modeled_param_norm) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(meanSES.INF_glob.uni_Ticir, na.rm = T), 0, 0, max(modeled_param_uni),
     col = rgb(0.5, 0, 1, 0.3), border = NA)
```

```
plot(meanSES.3glob.norm_Tpcpr, modeled_param_norm, pch = 16,
     main = "Global T_PC.PR",
     xlim = c(min(c(meanSES.3glob.norm_Tpcpr, meanSES.3glob.uni_Tpcpr),
                    na.rm = T), 0),
     ylim = c(0, min(max(c(modeled_param_norm, modeled_param_uni), na.rm = T), 100)),
     xlab = "standardized effect size of T_PC.PR",
     ylab = "modeled parameters: range/sd")
points(meanSES.3glob.uni_Tpcpr, modeled_param_uni, pch = 16, col = rgb(0.5, 0, 1, 0.6))
abline(v = mean(meanSES.INF_glob.norm_Tpcpr, na.rm = T))
abline(v = mean(meanSES.INF_glob.uni_Tpcpr, na.rm = T), col = "purple")
rect(mean(meanSES.INF_glob.norm_Tpcpr, na.rm = T), -1, 0, max(modeled_param_norm) + 1,
     col = rgb(0, 0, 0, 0.3), border = NA)
rect(mean(meanSES.INF_glob.uni_Tpcpr, na.rm = T), 0, 0, max(modeled_param_uni),
     col = rgb(0.5, 0, 1, 0.3), border = NA)
```

4.3 Conclusion on the power of $T_{IC/IR}$ to detect external filtering

To conclude on the power of $T_{IC/IR}$ to detect external filtering, we compute exponential linear model and identify the strength from which the beta-error is inferior to 0.05. This strength is define by the two ratio of the initial parameters `param_range` and `sd.mean.param`.

```
plot(meanSES.3loc.norm_Ticir ~ init_param.loc, col = rgb(0, 0, 0, 0.2),
     ylim = c(min(c(meanSES.3loc.norm_Ticir, meanSES.3loc.uni_Ticir,
                    meanSES.3glob.norm_Ticir, meanSES.3glob.uni_Ticir), na.rm = T), 0),
     main = "Global T_IC.IR",
```

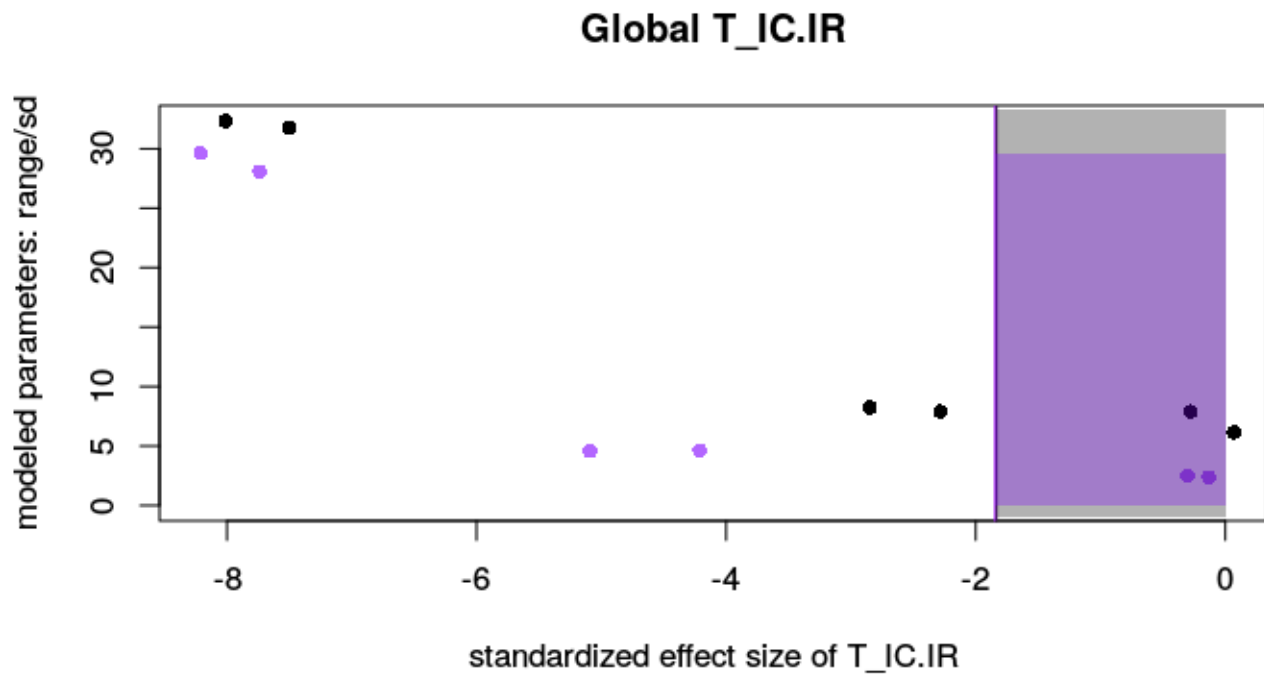


Figure 22: Global $T_{IC/IR}$ SES and modeled parameters values: Standardized effect size of $T_{IC/IR}$ in function of the strength of external filter defined by the ratio of the modeled range parameter by the modeled standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

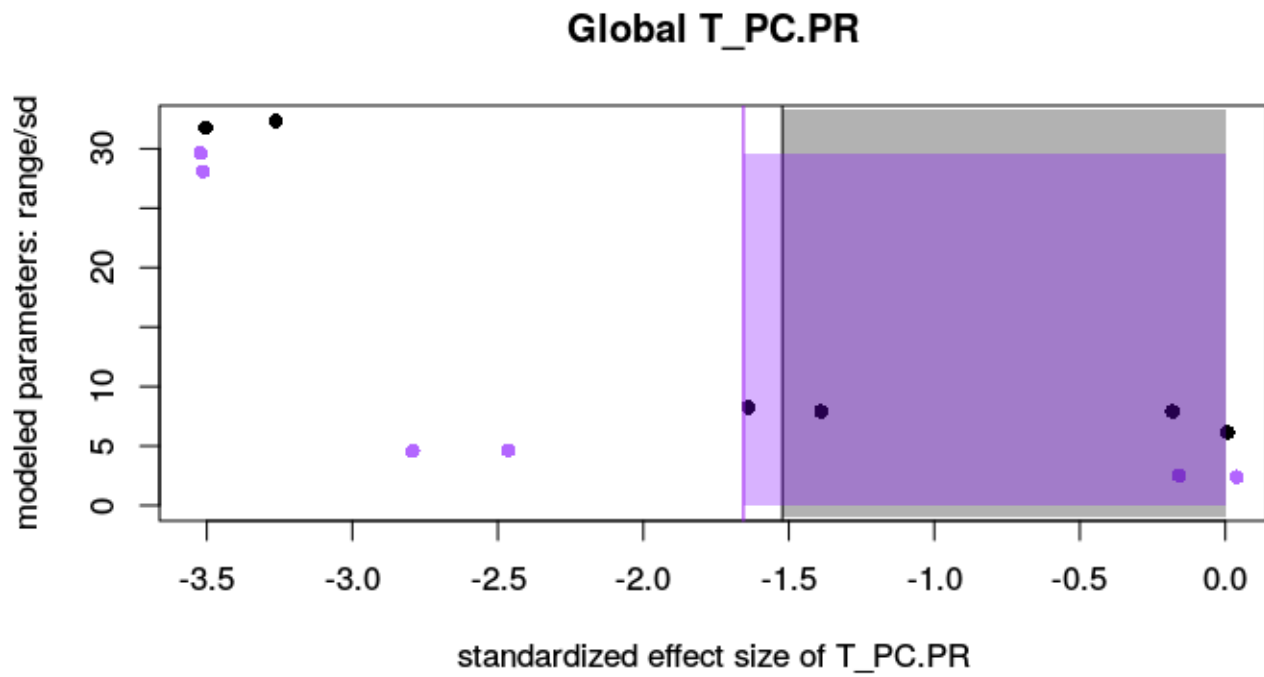


Figure 23: Global $T_{PC/PR}$ SES and modeled parameters values: Standardized effect size of $T_{PC/PR}$ in function of the strength of external filter defined by the ratio of the modeled range parameter by the modeled standard error parameter. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$). Trait a (normal) in black and b (uniform) in purple.

```

ylab = "standardized effect size of T_IC.IR",
xlab = "initial parameters: range/sd",
type = "n")
#points(meanSES.3loc.uni_Ticir ~ init_param.loc, col = rgb(0.5, 0, 1, 0.2))

points(meanSES.3glob.norm_Ticir ~ init_param, pch = 16, col = rgb(0, 0, 0, 0.5), cex = 1.2)
lm.norm <- lm(meanSES.3glob.norm_Ticir ~ log(init_param))
lm.norm_conf <- confint(lm.norm, level = 0.90)
curve(lm.norm$coef[1] + log(x) * (lm.norm$coef[2]), add = T, lwd = 3, col = rgb(0, 0, 0, 1))
curve(lm.norm_conf [1, 1] + log(x) * lm.norm_conf [2, 1], add = T, lty = 2, col = rgb(0, 0, 0, 1))
curve(lm.norm_conf [1, 2] + log(x) * lm.norm_conf [2, 2], add = T, lty = 2, col = rgb(0, 0, 0, 1))

points(meanSES.3glob.uni_Ticir ~ init_param, pch = 16, col = rgb(0.5, 0, 1, 0.5), cex = 1.2)
lm.uni <- lm(meanSES.3glob.uni_Ticir ~ log(init_param))
lm.uni_conf <- confint(lm.uni, level = 0.90)
curve(lm.uni$coef[1] + log(x) * (lm.uni$coef[2]), add = T, lwd = 3, col = rgb(0.5, 0, 1, 1))
curve(lm.uni_conf [1, 1] + log(x) * lm.uni_conf [2, 1], add = T, lty = 2, col = rgb(0.5, 0, 1, 1))
curve(lm.uni_conf [1, 2] + log(x) * lm.uni_conf [2, 2], add = T, lty = 2, col = rgb(0.5, 0, 1, 1))

rect(-1, mean(c(SSES.inf.MEAN.uni_Ticir, meanSES.INF_glob.uni_Ticir), na.rm = T), max(init_param), 10,
     col = rgb(0.5, 0, 1, 0.3), border = NA)

rect(-1, mean(c(SSES.inf.MEAN.norm_Ticir, meanSES.INF_glob.norm_Ticir), na.rm = T), max(init_param) + 10,
     col = rgb(0, 0, 0, 0.3), border = NA)

```

```

yy.norm <- mean(c(SSES.inf.MEAN.norm_Ticir, meanSES.INF_glob.norm_Ticir), na.rm = T)
yy.uni <- mean(c(SSES.inf.MEAN.uni_Ticir, meanSES.INF_glob.uni_Ticir), na.rm = T)

param_beta0.05_Ticir.norm <- exp( (yy.norm - lm.norm_conf [1, 2]) / lm.norm_conf [2, 2] )
param_beta0.05_Ticir.uni <- exp( (yy.uni - lm.uni_conf [1, 2]) / lm.uni_conf [2, 2] )

```

For the trait "a" normally distributed, the power of $T_{IC/IR}$ is satisfactory if the ratio of the initial parameter is superior to 0.95. For the trait "b" uniformly distributed, this value is 0.33. Further investigation on real data with different strength of external filtering are needed to complete these simulated values.

4.4 Conclusion on the power of $T_{PC/PR}$ to detect external filtering

To conclude on the power of $T_{PC/PR}$ to detect external filtering, we compute exponential linear model and identify the strength from which the beta-error is inferior to 0.05. This strength is define by the two ratio of the initial parameters `param_range` and `sd_mean.param`.

```

plot(meanSES.3loc.norm_Tpcpr ~ init_param.loc, col = rgb(0, 0, 0, 0.2),
     ylim = c(min(c(meanSES.3loc.norm_Tpcpr , meanSES.3loc.uni_Tpcpr,
                     meanSES.3glob.norm_Tpcpr, meanSES.3glob.uni_Tpcpr), na.rm = T), 0),
     main = "Global T_PC.IR",
     ylab = "standardized effect size of T_PC.IR",
     xlab = "initial parameters: range/sd",
     type = "n")

```

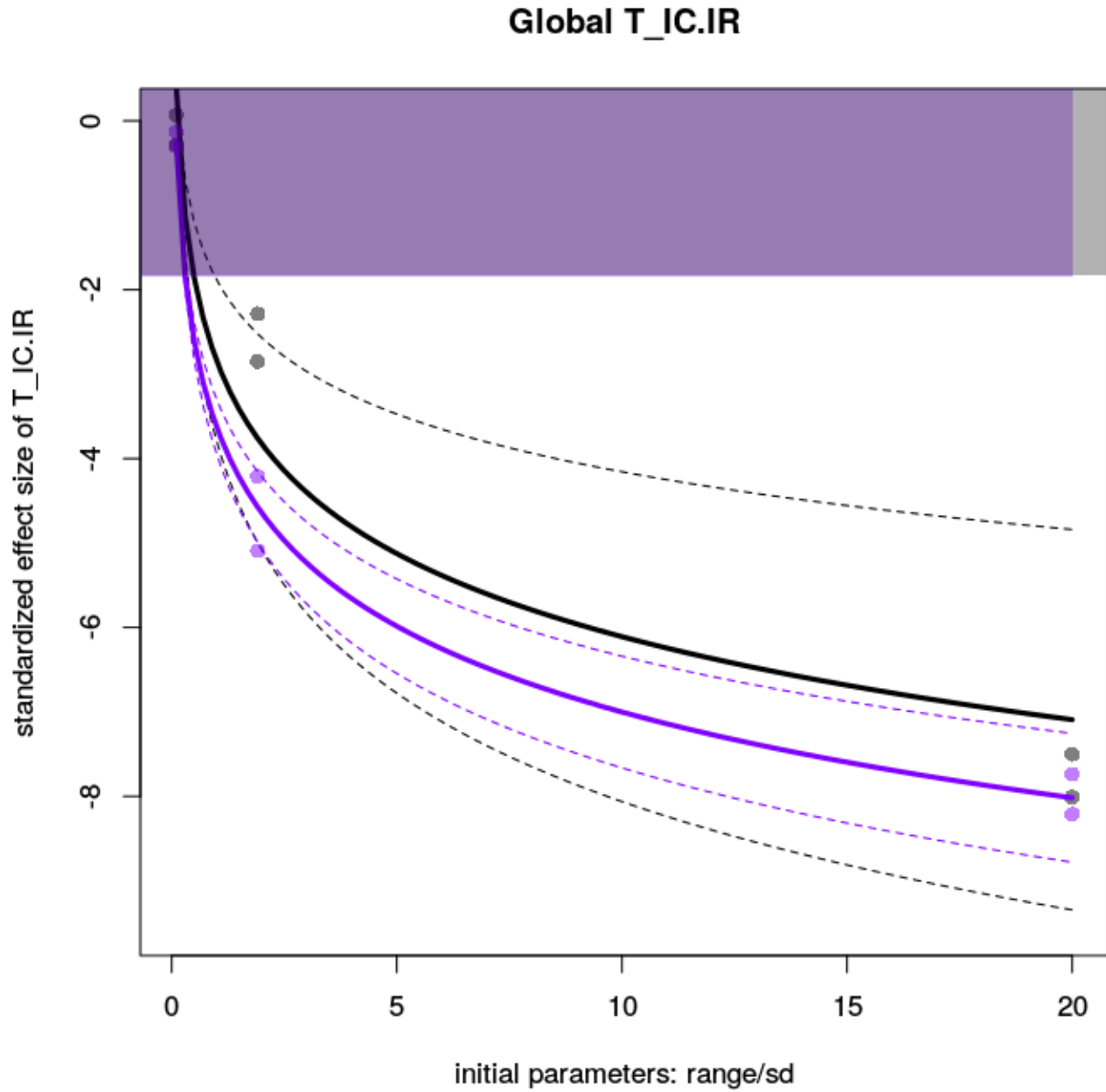


Figure 24: Power of $T_{IC/IR}$ to detect external filtering: Standardised Effect Size (SES) of $T_{IC/IR}$ in relation to the strength of external filtering. Trait 'a' normally distributed is shown in black, traits 'b' uniformly distributed in purple. Close circles represent global SES values. The exponential linear model is presented with 5 percents confidence interval on both side. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$)

```

#points(meanSES.3glob.uni_Tpcpr ~ init_param.loc, col = rgb(0.5, 0, 1, 0.2))

points(meanSES.3glob.norm_Tpcpr ~ init_param, pch = 16, col = rgb(0, 0, 0, 0.5), cex = 1.2)
lm.norm <- lm(meanSES.3glob.norm_Tpcpr ~ log(init_param))
lm.norm_conf <- confint(lm.norm, level = 0.90)
curve(lm.norm$coef[1] + log(x) * (lm.norm$coef[2]), add = T, lwd = 3, col = rgb(0, 0, 0, 1))
curve(lm.norm_conf [1, 1] + log(x) * lm.norm_conf [2, 1], add = T, lty = 2, col = rgb(0, 0, 0, 1))
curve(lm.norm_conf [1, 2] + log(x) * lm.norm_conf [2, 2], add = T, lty = 2, col = rgb(0, 0, 0, 1))

points(meanSES.3glob.uni_Tpcpr ~ init_param, pch = 16, col = rgb(0.5, 0, 1, 0.5), cex = 1.2)
lm.uni <- lm(meanSES.3glob.uni_Tpcpr ~ log(init_param))
lm.uni_conf <- confint(lm.uni, level = 0.90)
curve(lm.uni$coef[1] + log(x) * (lm.uni$coef[2]), add = T, lwd = 3, col = rgb(0.5, 0, 1, 1))
curve(lm.uni_conf [1, 1] + log(x) * lm.uni_conf [2, 1], add = T, lty = 2, col = rgb(0.5, 0, 1, 1))
curve(lm.uni_conf [1, 2] + log(x) * lm.uni_conf [2, 2], add = T, lty = 2, col = rgb(0.5, 0, 1, 1))

rect(-1, mean(c(SSES.inf.MEAN.uni_Tpcpr, meanSES.INF_glob.uni_Tpcpr), na.rm = T), max(init_param), 10,
     col = rgb(0.5, 0, 1, 0.3), border = NA)

rect(-1, mean(c(SSES.inf.MEAN.norm_Tpcpr, meanSES.INF_glob.norm_Tpcpr), na.rm = T), max(init_param), 10,
     col = rgb(0, 0, 0, 0.3), border = NA)

yy.norm <- mean(c(SSES.inf.MEAN.norm_Tpcpr, meanSES.INF_glob.norm_Tpcpr), na.rm = T)
yy.uni <- mean(c(SSES.inf.MEAN.uni_Tpcpr, meanSES.INF_glob.uni_Tpcpr), na.rm = T)

param_beta0.05_Tpcpr.norm <- exp( (yy.norm - lm.norm_conf [1, 2]) / lm.norm_conf [2, 2] )
param_beta0.05_Tpcpr.uni <- exp( (yy.uni - lm.uni_conf [1, 2]) / lm.uni_conf [2, 2] )

```

For the trait "a" normally distributed, the power of $T_{PC/PR}$ is satisfactory if the ratio of the initial parameter is superior to 2.11. For the trait "b" uniformly distributed, this value is 1.59. Further investigation on real data with different strength of external filtering are needed to complete these simulated values. As expected, $T_{PC/PR}$ is far less powerfull than $T_{IC/IR}$ to detect external filtering either with traits "a" (2.11 *vs* 0.95) and "b" (1.59 *vs* 0.33). This is due to the loss of information in $T_{PC/PR}$ when we compute the mean by population and exclude intra-populationnal variation from the test of external filter.

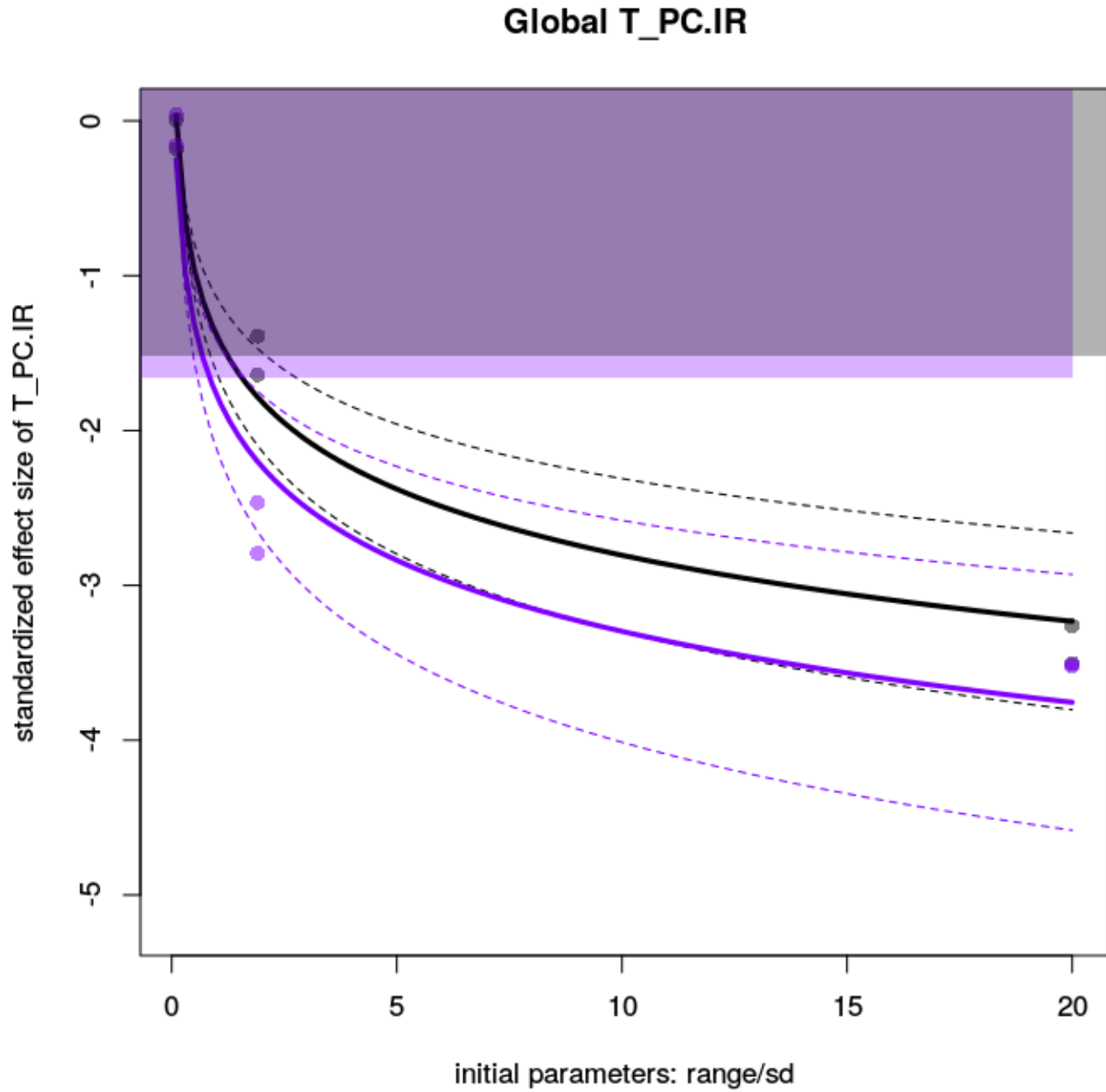


Figure 25: Power of $T_{PC/PR}$ to detect external filtering: Standardised Effect Size (SES) of $T_{PC/PR}$ in relation to the strength of external filtering. Trait 'a' normally distributed is shown in black, traits 'b' uniformly distributed in purple. Close circles represent global SES values. The exponential linear model is represented with 5 percents confidence interval on both side. Colored area represent the mean confidence interval of SES values ($\alpha = 0.95$)

5 Internal and external filtering

5.1 Randomization outline

We use the same two parameters as for internal and external filtering analysis. The internal filtering is computed as in the section dedicated to internal filter. To add the effect of external filtering, we add a value to each individual belonging to a community, this values differing among communities.

```
# Parameter for the distance between species mean trait values
param_DIST_between_MEAN.init <- round(sort(seq(10, 200, length.out = nb_param_val)), 2)
param_DIST_between_MEAN <- rep(param_DIST_between_MEAN.init, N_repet_Param)
mean_sd.param <- rep(10, nperm)

# Parameter for the variance in species mean trait values
param_SD <- seq(10, 100, length.out = nb_param_val)
sd_mean.param <- sample(rep(param_SD, N_repet_Param),
  size = length(rep(param_SD, N_repet_Param)), replace = F)
sd_sd.param <- rep(10, nperm)

nperm <- N_repet_Param*length(param_SD)

#Start simulation
mean.sp_stock4 <- list()
sd.sp_stock4 <- list()
mean.com_stock4 <- list()
sd.com_stock4 <- list()
res.simu4 <- list()
res.simu4.pval <- list()
res.simu.traits4 <- list()

for(n in 1:nperm){#for each permutation

ex.sp4 <- c()
ex.com4 <- matrix(0, nrow = Ncom, ncol = Nsp)
for(i in 1: 10){
  ex.com.interm <- table(sample(sp, size = Nind/Ncom, prob = rlnorm(Nsp, 0, sdlog), replace = T))
  ex.com4[i, sp%in% names(ex.com.interm)] <- ex.com.interm
  ex.sp4 <- c(ex.sp4, rep(sp, times = ex.com4[i,]))
}

ex.indplot4 <- sort(as.factor(rep(com, Nind/Ncom)))

#Defining traits mean and sd by species
mean_mean.param.interm <- seq(max.value_traits - param_DIST_between_MEAN[n],
  max.value_traits, length.out =
  length(unique(param_DIST_between_MEAN)))

mean_mean.param <- rep(round(sort(mean_mean.param.interm), 2),
  N_repet_Param)
```

```

mean.sp <- rnorm(length(unique(sp)), mean = mean_mean.param, sd = mean_sd.param[n])
sd.sp <- rnorm(length(unique(sp)), mean = sd_mean.param[n], sd = sd_sd.param[n])

ex.traits4 <- array(NA, dim = c(Nind, 2))
colnames(ex.traits4) <- paste("trait", c("a", "b"), sep = " ")

for(s in unique(ex.sp4)){
  #trait a : normal distribution
  ex.traits4[ex.sp4 == s, 1] <- rnorm(500, rep(mean.sp[unique(ex.sp4) == s], 500),
    rep(sd.sp[unique(ex.sp4) == s], 500))[1:sum(ex.sp4 == s)]

  #trait b : uniform distribution
  ex.traits4[ex.sp4 == s, 2] <- runif(500, min = rep(mean.sp[unique(ex.sp4) == s], 500) -
    rep(sd.sp[unique(ex.sp4) == s], 500),
    max = rep(mean.sp[unique(ex.sp4) == s], 500) +
    rep(sd.sp[unique(ex.sp4) == s], 500))[1:sum(ex.sp4 == s)]
}

mean.com.add <- rnorm(length(unique(com)), mean = mean_mean.param, sd = mean_sd.param[n])
sd.com.add <- rnorm(length(unique(com)), mean = sd_mean.param[n], sd = sd_sd.param[n])

for(c in unique(ex.indplot4)){
  #trait a : normal distribution
  ex.traits4[ex.indplot4 == c, 1] <-
  ex.traits4[ex.indplot4 == c, 1] +
  rnorm(500, rep(mean.com.add[unique(ex.indplot4) == c], 500),
  rep(sd.com.add [unique(ex.indplot4) == c], 500))[1:sum(ex.indplot4 == c)]

  #trait b : uniform distribution
  ex.traits4[ex.indplot4 == c, 2] <-
  ex.traits4[ex.indplot4 == c, 2] +
  runif(500, min = rep(mean.com.add[unique(ex.indplot4) == c], 500) -
  rep(sd.com.add[unique(ex.indplot4) == c], 500),
  max = rep(mean.com.add [unique(ex.indplot4) == c], 500) +
  rep(sd.com.add [unique(ex.indplot4) == c], 500))[1:sum(ex.indplot4 == c)]
}

#stock results
mean.sp_stock4[[n]] <- mean.sp
sd.sp_stock4[[n]] <- sd.sp

mean.com_stock4[[n]] <- mean.com.add
sd.com_stock4[[n]] <- sd.com.add

res.simu.traits4[[n]] <- ex.traits4
res.simu4[[n]] <- Tstats(ex.traits4, ex.indplot4, ex.sp4)
res.simu4.pval[[n]] <- sum_Tstats(res.simu4[[n]], type = "p.value")
print(paste("----", round(n/nperm, 2) * 100, "%", sep = " "))
}#End of simulations

```

5.2 Results

```
par(mfrow=c(2, 2))
plotDistri(ex.traits4, rep("all_sp", times = dim(ex.traits1)[1]), ex.indplot4,
  plot.ask = F, multipanel = F, leg =c(T, F), main = c("a", "b"))
plotDistri(ex.traits4, rep("region", times = dim(ex.traits1)[1]), ex.sp4,
  plot.ask = F, multipanel = F, leg =c(T, F), main = c("c", "d"),
  ylim = c(0, 0.01))
par(mfrow=c(1, 1))
```

```
par(mfrow=c(3, 1))
plot(res.simu4[[1]], main = "a")
plot(res.simu4[[nperm/2]], main = "b")
plot(res.simu4[[nperm]], main = "c")
par(mfrow=c(1, 1))
```

```
meanSES.4loc.norm_Tipic <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses[, 1]))
meanSES.4loc.uni_Tipic <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses[, 2]))
SES.inf.MEAN.norm_Tipic <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses.inf[, 1]))
SES.inf.MEAN.uni_Tipic <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_1$ses.inf[, 2]))

meanSES.4loc.norm_Ticir <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses[, 1]))
meanSES.4loc.uni_Ticir <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses[, 2]))
SES.inf.MEAN.norm_Ticir <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses.inf[, 1]))
SES.inf.MEAN.uni_Ticir <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_2$ses.inf[, 2]))

meanSES.4loc.norm_Tpcpr <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_3$ses[, 1]))
meanSES.4loc.uni_Tpcpr <- unlist(lapply(res.simu4, function(x)
  ses.listofindex(as.listofindex(x))
  $index_1_3$ses[, 2]))
SES.inf.MEAN.norm_Tpcpr <- unlist(lapply(res.simu4, function(x)
```

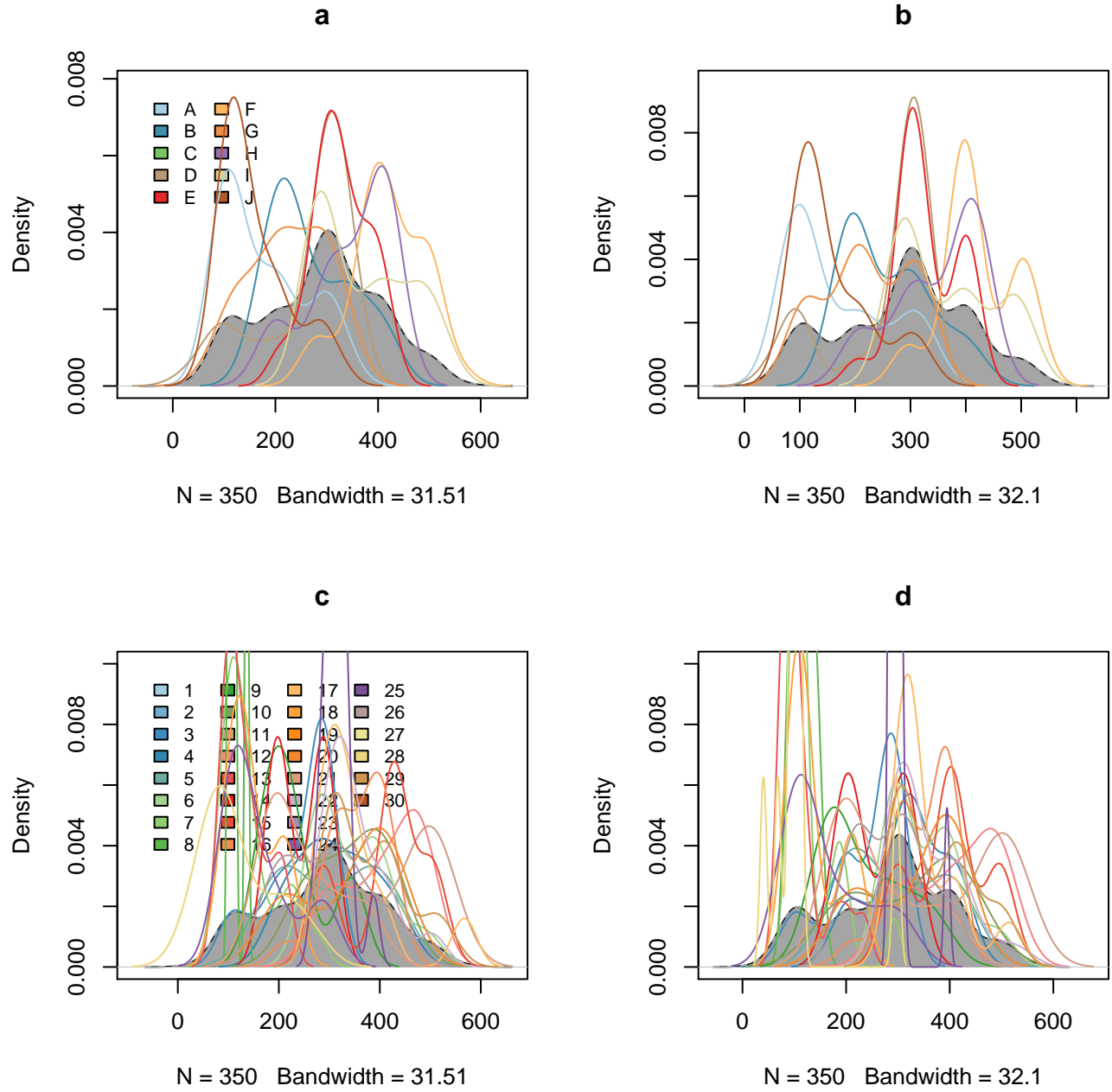


Figure 26: Distribution of trait values for one randomization with both internal and external filter: (a) Community-level trait distributions for the trait a (normal distribution); (b) Community-level trait distributions for the trait b (uniform distribution); (c) Species' trait distributions for the trait a; (d) Species' trait distributions for the trait b

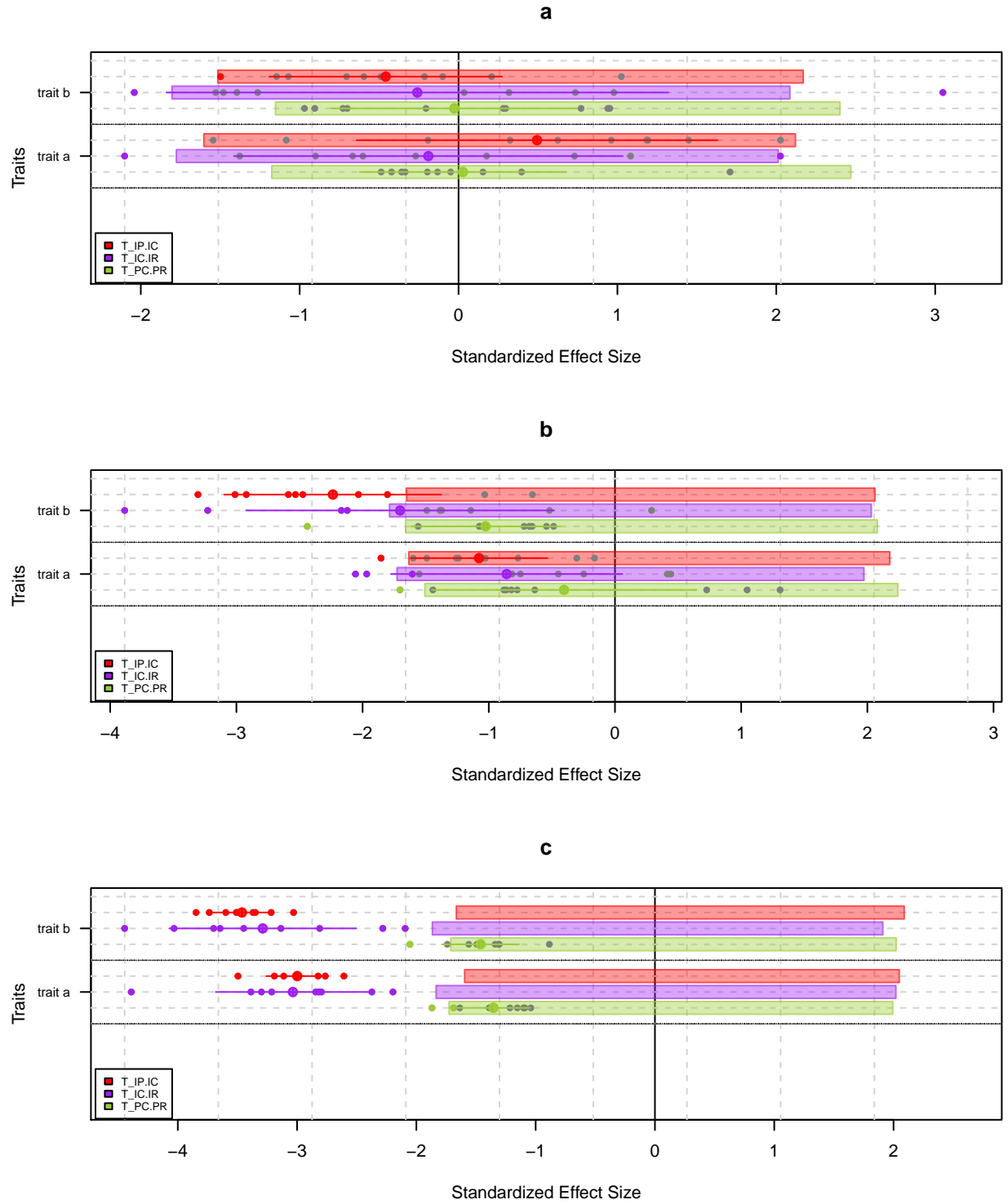


Figure 27: Results of T-statistics with both internal and external filter: $T_{IP/IC}$ in red, $T_{IC/IR}$ in purple and $T_{PC/PR}$ in green. (a) Lower strength of internal filtering 5 (b) higher strength of internal filtering.

```

      ses.listofindex(as.listofindex(x))
      $index_1_3$ses.inf[, 1]))
SES.inf.MEAN.uni_Tpcpr <- unlist(lapply(res.simu4, function(x)
      ses.listofindex(as.listofindex(x))
      $index_1_3$ses.inf[, 2]))

meanSES.4glob.norm_Tipic <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_1$ses[,1], na.rm = T)))
meanSES.4glob.uni_Tipic <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_1$ses[,2], na.rm = T)))
meanSES.INF_glob.norm_Tipic <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_1$ses.inf[, 1], na.rm = T)))
meanSES.INF_glob.uni_Tipic <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_1$ses.inf[, 2], na.rm = T)))

meanSES.4glob.norm_Ticir <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_2$ses[,1], na.rm = T)))
meanSES.4glob.uni_Ticir <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_2$ses[,2], na.rm = T)))
meanSES.INF_glob.norm_Ticir <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_2$ses.inf[, 1], na.rm = T)))
meanSES.INF_glob.uni_Ticir <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_2$ses.inf[, 2], na.rm = T)))

meanSES.4glob.norm_Tpcpr <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_3$ses[,1], na.rm = T)))
meanSES.4glob.uni_Tpcpr <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_3$ses[,2], na.rm = T)))
meanSES.INF_glob.norm_Tpcpr <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_3$ses.inf[, 1], na.rm = T)))
meanSES.INF_glob.uni_Tpcpr <- unlist(lapply(res.simu4, function(x)
      mean(ses.listofindex(as.listofindex(x))
      $index_1_3$ses.inf[, 2], na.rm = T)))

```

```

# trait a
lm_glob_norm_Tipic <- lm(meanSES.4glob.norm_Tipic ~ log(init_param))
lm_glob_norm_Tipic.conf <- confint(lm_glob_norm_Tipic , level = 0.90)
lm_glob_norm_Ticir <- lm(meanSES.4glob.norm_Ticir ~ log(init_param))
lm_glob_norm_Ticir.conf <- confint(lm_glob_norm_Ticir , level = 0.90)

```

```

lm_glob_norm_Tpcpr <- lm(meanSES.4glob.norm_Tpcpr ~ log(init_param))
lm_glob_norm_Tpcpr.conf <- confint(lm_glob_norm_Tpcpr , level = 0.90)
yy_glob_norm_Tipic <- mean(meanSES.INF_glob.norm_Tipic, na.rm = T)
yy_glob_norm_Ticir <- mean(meanSES.INF_glob.norm_Ticir, na.rm = T)
yy_glob_norm_Tpcpr <- mean(meanSES.INF_glob.norm_Tpcpr, na.rm = T)
param_B_glob_norm_Tipic <- exp( (yy_glob_norm_Tipic - lm_glob_norm_Tipic.conf [1, 2]) /
                                lm_glob_norm_Tipic.conf [2, 2] )
param_B_glob_norm_Ticir <- exp( (yy_glob_norm_Ticir - lm_glob_norm_Ticir.conf [1, 2]) /
                                lm_glob_norm_Ticir.conf [2, 2] )
param_B_glob_norm_Tpcpr <- exp( (yy_glob_norm_Tpcpr - lm_glob_norm_Tpcpr.conf [1, 2]) /
                                lm_glob_norm_Tpcpr.conf [2, 2] )

lm_loc_norm_Tipic <- lm(meanSES.4loc.norm_Tipic ~ log(init_param.loc))
lm_loc_norm_Tipic.conf <- confint(lm_loc_norm_Tipic, level = 0.90)
lm_loc_norm_Ticir <- lm(meanSES.4loc.norm_Ticir ~ log(init_param.loc))
lm_loc_norm_Ticir.conf <- confint(lm_loc_norm_Ticir, level = 0.90)
lm_loc_norm_Tpcpr <- lm(meanSES.4loc.norm_Tpcpr ~ log(init_param.loc))
lm_loc_norm_Tpcpr.conf <- confint(lm_loc_norm_Tpcpr, level = 0.90)
yy_loc_norm_Tipic <- mean(SSES.inf.MEAN.norm_Tipic, na.rm = T)
yy_loc_norm_Ticir <- mean(SSES.inf.MEAN.norm_Ticir, na.rm = T)
yy_loc_norm_Tpcpr <- mean(SSES.inf.MEAN.norm_Tpcpr, na.rm = T)
param_B_loc_norm_Tipic <- exp( (yy_loc_norm_Tipic - lm_loc_norm_Tipic.conf [1, 2]) /
                                lm_loc_norm_Tipic.conf [2, 2] )
param_B_loc_norm_Ticir <- exp( (yy_loc_norm_Ticir - lm_loc_norm_Ticir.conf [1, 2]) /
                                lm_loc_norm_Ticir.conf [2, 2] )
param_B_loc_norm_Tpcpr <- exp( (yy_loc_norm_Tpcpr - lm_loc_norm_Tpcpr.conf [1, 2]) /
                                lm_loc_norm_Tpcpr.conf [2, 2] )

# trait b
lm_glob_uni_Tipic <- lm(meanSES.4glob.uni_Tipic ~ log(init_param))
lm_glob_uni_Tipic.conf <- confint(lm_glob_uni_Tipic , level = 0.90)
lm_glob_uni_Ticir <- lm(meanSES.4glob.uni_Ticir ~ log(init_param))
lm_glob_uni_Ticir.conf <- confint(lm_glob_uni_Ticir , level = 0.90)
lm_glob_uni_Tpcpr <- lm(meanSES.4glob.uni_Tpcpr ~ log(init_param))
lm_glob_uni_Tpcpr.conf <- confint(lm_glob_uni_Tpcpr , level = 0.90)
yy_glob_uni_Tipic <- mean(meanSES.INF_glob.uni_Tipic, na.rm = T)
yy_glob_uni_Ticir <- mean(meanSES.INF_glob.uni_Ticir, na.rm = T)
yy_glob_uni_Tpcpr <- mean(meanSES.INF_glob.uni_Tpcpr, na.rm = T)
param_B_glob_uni_Tipic <- exp( (yy_glob_uni_Tipic - lm_glob_uni_Tipic.conf [1, 2]) /
                                lm_glob_uni_Tipic.conf [2, 2] )
param_B_glob_uni_Ticir <- exp( (yy_glob_uni_Ticir - lm_glob_uni_Ticir.conf [1, 2]) /
                                lm_glob_uni_Ticir.conf [2, 2] )
param_B_glob_uni_Tpcpr <- exp( (yy_glob_uni_Tpcpr - lm_glob_uni_Tpcpr.conf [1, 2]) /
                                lm_glob_uni_Tpcpr.conf [2, 2] )

lm_loc_uni_Tipic <- lm(meanSES.4loc.uni_Tipic ~ log(init_param.loc))
lm_loc_uni_Tipic.conf <- confint(lm_loc_uni_Tipic, level = 0.90)
lm_loc_uni_Ticir <- lm(meanSES.4loc.uni_Ticir ~ log(init_param.loc))
lm_loc_uni_Ticir.conf <- confint(lm_loc_uni_Ticir, level = 0.90)
lm_loc_uni_Tpcpr <- lm(meanSES.4loc.uni_Tpcpr ~ log(init_param.loc))
lm_loc_uni_Tpcpr.conf <- confint(lm_loc_uni_Tpcpr, level = 0.90)

```



```

yy.loc_uni_Tipic <- mean(SES.inf.MEAN.uni_Tipic, na.rm = T)
yy.loc_uni_Ticir <- mean(SES.inf.MEAN.uni_Ticir, na.rm = T)
yy.loc_uni_Tpcpr <- mean(SES.inf.MEAN.uni_Tpcpr, na.rm = T)
param_B_loc_uni_Tipic <- exp( (yy.loc_uni_Tipic - lm_loc_uni_Tipic.conf [1, 2]) /
                              lm_loc_uni_Tipic.conf [2, 2] )
param_B_loc_uni_Ticir <- exp( (yy.loc_uni_Ticir - lm_loc_uni_Ticir.conf [1, 2]) /
                              lm_loc_uni_Ticir.conf [2, 2] )
param_B_loc_uni_Tpcpr <- exp( (yy.loc_uni_Tpcpr - lm_loc_uni_Tpcpr.conf [1, 2]) /
                              lm_loc_uni_Tpcpr.conf [2, 2] )

```

For the analysis with both internal and external filter we only present the type II error using the initial parameter ratio allowing a beta-error < 0.05 .

Table 6: Initial parameter ratio allowing beta-error < 0.05 for T-statistics under both external and internal filtering

Traits	Indices	Average between communities?	Initial parameter ratio
Trait a (normal)	$T_{IP.IC}$	no (local)	24.494
		yes (global)	0
	$T_{IC.IR}$	no (local)	190.863
		yes (global)	0
	$T_{PC.PR}$	no (local)	5.7651721×10^7
		yes (global)	0
Trait b (uniform)	$T_{IP.IC}$	no (local)	3.029
		yes (global)	1860.474
	$T_{IC.IR}$	no (local)	10.943
		yes (global)	4.8966336×10^5
	$T_{PC.PR}$	no (local)	1675.525
		yes (global)	2.9678684×10^5

6 Test of variance decomposition functions

In addition to the `cati` vignette which test this function on the Darwin finch data, we test the behavior of the function `partvar` and `decompCTRE` by building two toy models associated to each functions.

6.1 Behavior of the function `partvar`

First to test the behavior of the `partvar` function we assemble the same data as in the test of the T-statistics. 1000 individuals belonging to 20 species occurring in 10 communities. But here we define the trait value for one individual as the addition of three values, each one draw in a normal distribution of mean 0 and standard error 1: (i) a value depending on the species the individual belong to, (ii) a value depending on the community the individual belong to and (iii) a independent value measuring the individual variance.

```
nperm <- npermut

sdlog
table(com)
table(sp)

res.partvar_toymodel <- list()
for(n in 1:nperm){#for each permutation

  # Draw communities using lognormal distribution of abundances
  # 100 individuals by populations
  ex.sp5 <- c()
  ex.com5 <- matrix(0, nrow = Ncom, ncol = Nsp)
  for(i in 1:Ncom){
    ex.com.interm <- table(sample(sp, size = Nind/Ncom, prob = rlnorm(Nsp, 0, sdlog), replace = T))
    ex.com5[i, sp%in% names(ex.com.interm)] <- ex.com.interm
    ex.sp5 <- c(ex.sp5, rep(sp, times = ex.com5[i,]))
  }

  x1 <- c()
  ex.sp5 <- as.factor(ex.sp5)
  for(s in 1:Nsp){
    x1[ex.sp5 == levels(ex.sp5)[s]] <- rnorm(1)
  }

  ex.indplot5 <- sort(as.factor(rep(com, Nind/Ncom)))
  x2 <- c()
  x3 <- c()
  for(c in 1:Ncom){
    x2[ex.indplot5 == levels(ex.indplot5)[c]] <- rnorm(1)
  }
  X <- rnorm(Nind) + x1 + x2

  res.partvar_toymodel[[n]] <- partvar(X, factors = cbind(sites = ex.indplot5,
    species = ex.sp5))
}
```

```
## The partvar function decompose the variance accross nested scales. Thus choose the order
of the factors very carefully!
## The partvar function decompose the variance accross nested scales. Thus choose the order
of the factors very carefully!
## The partvar function decompose the variance accross nested scales. Thus choose the order
of the factors very carefully!
```

Now we can plot the result as the density of the three component of variation in individual traits distributions. We are waiting for a mean of approximately 0.33 for each component in view of our model.

```
res.simu.partvar_toymodel <- lapply(res.partvar_toymodel, function (x) x[, 1])
res.simu.partvar_toymodel <- t(matrix(unlist(res.simu.partvar_toymodel), nrow = 3))
colnames(res.simu.partvar_toymodel) <- c("sites", "species", "within")
col.funk <- funky.col(3)

plot(density(res.simu.partvar_toymodel[, 3]), col = col.funk[3],
     pch = 16, xlim = c(0, 0.9), lwd = 2)
lines(density(res.simu.partvar_toymodel[, 1]), col = col.funk[1], lwd = 2)
lines(density(res.simu.partvar_toymodel[, 2]), col = col.funk[2], lwd = 2)
abline(v = 0.33)
points(apply(res.simu.partvar_toymodel, 2, mean),
       c(max(density(res.simu.partvar_toymodel[, 1])$y)/2,
         max(density(res.simu.partvar_toymodel[, 2])$y)/2,
         max(density(res.simu.partvar_toymodel[, 3])$y)/2),
       col = col.funk)
segments(apply(res.simu.partvar_toymodel, 2, mean) - apply(res.simu.partvar_toymodel, 2, sd),
         c(max(density(res.simu.partvar_toymodel[, 1])$y)/2,
           max(density(res.simu.partvar_toymodel[, 2])$y)/2,
           max(density(res.simu.partvar_toymodel[, 3])$y)/2),
         apply(res.simu.partvar_toymodel, 2, mean) + apply(res.simu.partvar_toymodel, 2, sd),
         c(max(density(res.simu.partvar_toymodel[, 1])$y)/2,
           max(density(res.simu.partvar_toymodel[, 2])$y)/2,
           max(density(res.simu.partvar_toymodel[, 3])$y)/2),
         col = col.funk)
legend(0.6, 4, legend = c("sites", "species", "within"), fill = col.funk)
text(0.33, 0.5, pos = 4, "0.33")
```

The mean contribution of each component is close to 0.33. Consequently the function `partvar` doesn't show apparent dysfunction.

6.2 Behavior of the function `decompCTRE`

To test the function `decompCTRE`, we adopt the same model but we add a new step to control the strength of the turnover in defining the community weighted mean. We define several value of turnover strength by deleting individuals from the data set if there their species-components trait and their community-component trait are to different. To be more precise, when `turnover_strength = 1`, we delete all individual whose the square of the species value minus the community value is superior o one.

```
sdlog
turnover_strength <- rep(c(0.01, 0.1, 0.2, 0.5, 1, 2, 3.5, 5, 7.5, 10, 12.5, 15),
                        N_repet_Param)
```

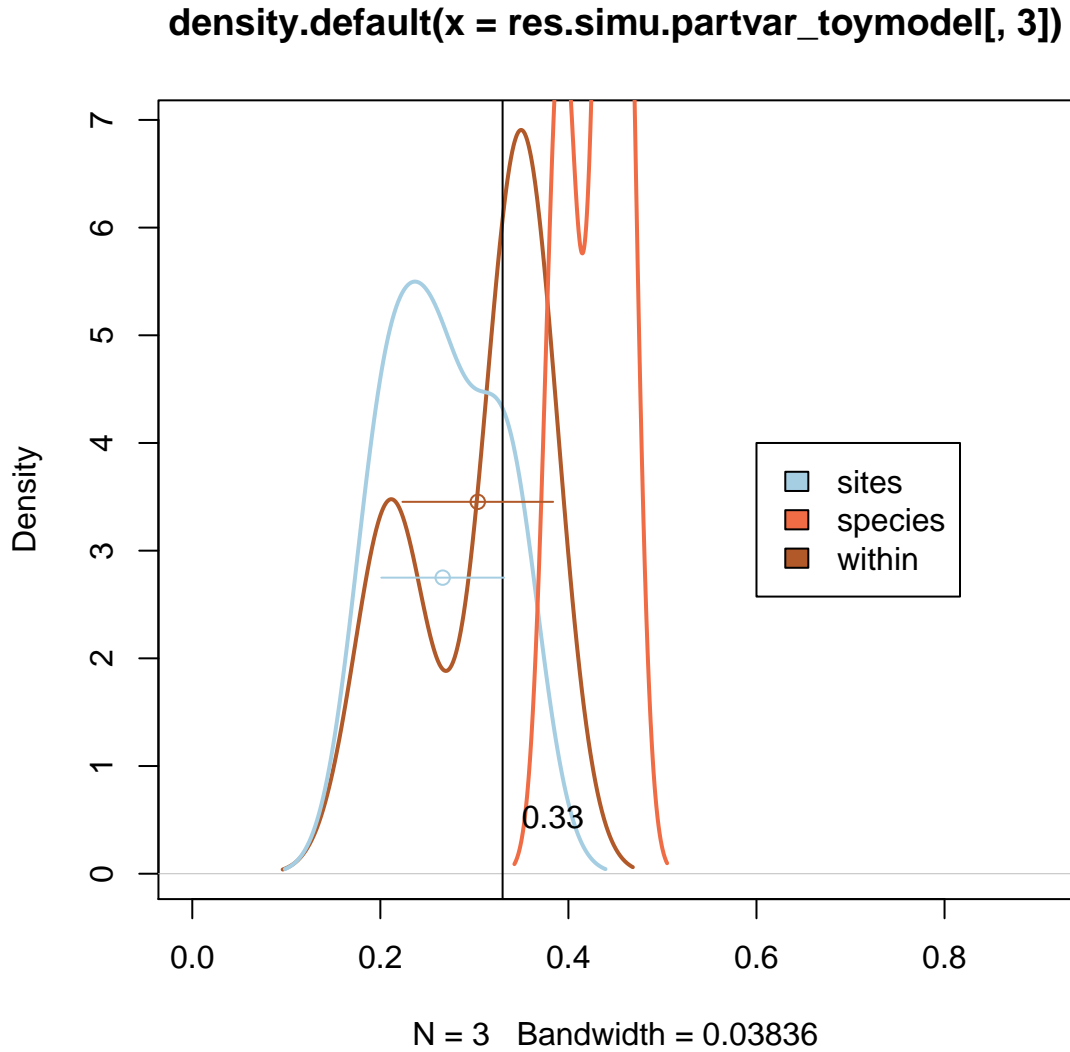


Figure 28: Variance partitioning accross nested scales: each colored lines represent the density of permutations results (sites in blue, species in orange and the remaining variance in brown). Points represent the mean and segments the standard deviation. Vertical line figure the value of 0.33.

```

nperm <- length(turnover_strength)
res.decompCTRE_toymodel <- list()
for(n in 1:nperm){#for each permutation

  # Draw communities using lognormal distribution of abundances
  # 100 individuals by populations
  ex.sp5 <- c()
  ex.com5 <- matrix(0, nrow = Ncom, ncol = Nsp)
  for(i in 1:Ncom){
    ex.com.interm <- table(sample(sp, size = Nind/Ncom, prob = rlnorm(Nsp, 0, sdlog), replace = T))
    ex.com5[i, sp%in% names(ex.com.interm)] <- ex.com.interm
    ex.sp5 <- c(ex.sp5, rep(sp, times = ex.com5[i,]))
  }

  x1 <- c()
  ex.sp5 <- as.factor(ex.sp5)
  for(s in 1:Nsp){
    x1[ex.sp5 == levels(ex.sp5)[s]] <- rnorm(1)
  }

  ex.indplot5 <- sort(as.factor(rep(com, Nind/Ncom)))
  x2 <- c()
  for(c in 1:Ncom){
    x2[ex.indplot5 == levels(ex.indplot5)[c]] <- rnorm(1)
  }

  X <- rnorm(Nind) + x1 + x2

  #We delete the individual whose species values and communities values are to different
  #For this we use the argument turnover_strength
  X2 <- X[(x1-x2)^2 < turnover_strength[n]]
  ex.sp5.bis <- ex.sp5[(x1-x2)^2 < turnover_strength[n]]
  ex.indplot5.bis <- ex.indplot5[(x1-x2)^2 < turnover_strength[n]]

  res.decompCTRE_toymodel[[n]] <- decompCTRE(traits = cbind(X2, X2), sp = ex.sp5.bis,
                                             ind.plot = ex.indplot5.bis)
}#End of simulations

```

Now we can plot the relation between the modeled strength of the turnover and the resulting contribution of turnover from the function `decompCTRE` (Fig. 29).

```

par(mfrow=c(1, 2))
plot(res.decompCTRE_toymodel[[1]]$X2)
plot(res.decompCTRE_toymodel[[n]]$X2)
par(mfrow=c(1, 1))

```

```

res.simu.decompCTRE_toymodel <- unlist(lapply(res.decompCTRE_toymodel, function(x)
                                             x$X2$RelSumSq[1]))
plot(res.simu.decompCTRE_toymodel, -turnover_strength, pch = 16,

```

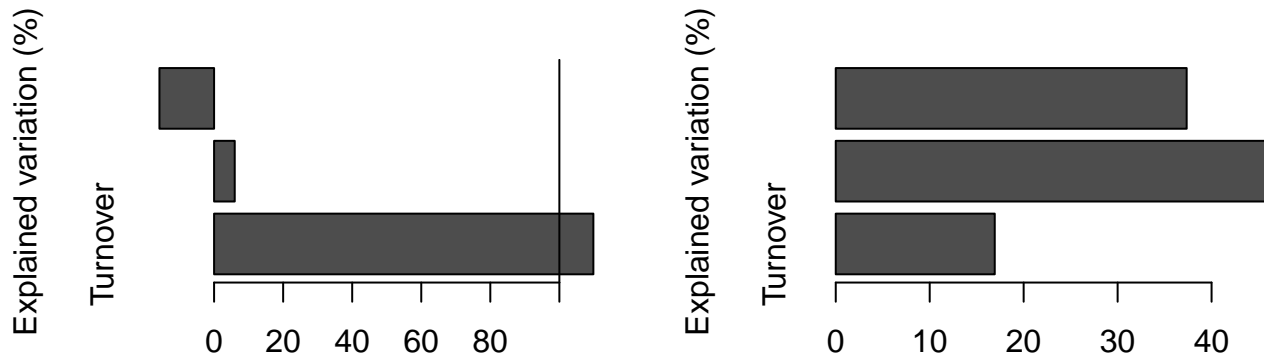


Figure 29: Decomposition of the variation in community trait composition for two contrasted cases: on the left almost all the variation is due to turnover whereas on the right case intraspecific variation explain most of the variation.

```
col = rgb(0, 0, 0, 0.5), xlim=c(0, 10),
ylab = "Strength of the turnover (maximum strength = 0)",
xlab = "Turnover relative contribution to community mean traits")
```

The correlation between the modeled strength of the turnover and the turnover relative contribution result is very high. Consequently the function `decompCTRE` doesn't show apparent dysfunction (Fig. 30).

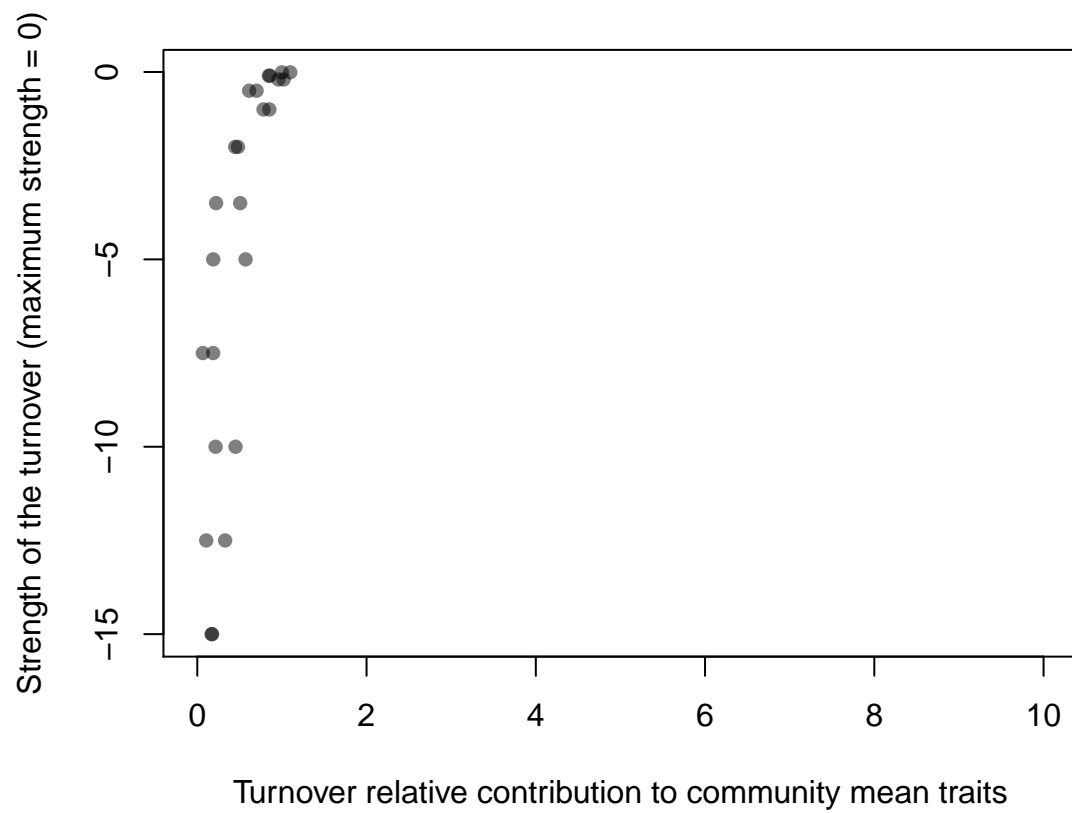


Figure 30: Modeled against obtained turnover contribution in community trait variance

7 Summary

7.1 Type I error

Table 2 summarize the alpha error for T-statistics. The three global ¹¹ T-statistics are very robust regardless of the distribution of the trait (alpha error ranging from 0.25 to 0.25). Local ¹² T-statistics present lower but nevertheless satisfactory alpha-error with a maximum of 0.03279 for $T_{PC/PR}$ on a trait normally distributed. We therefore recommend to use $T_{IC/IR}$ to detect external filtering and more generally to prefer global T-statistics which are more robust and more powerful (see the next section).

7.2 Type II error

Table 7: Summary of initial parameter ratio allowing beta-error < 0.05 for global T-statistics

Traits	Indices	Initial parameter ratio with internal filtering	Initial parameter ratio with external filtering	Initial parameter ratio with internal and external filtering
Trait a (normal)	$T_{IP.IC}$	86.07		0
	$T_{IC.IR}$		0.95	0
	$T_{PC.PR}$		2.11	0
Traits	Indices	Initial parameter ratio with internal filtering	Initial parameter ratio with external filtering	Initial parameter ratio with internal and external filtering
Trait b (uniform)	$T_{IP.IC}$	17.35		1860.47
	$T_{IC.IR}$		0.33	4.8966336×10^5
	$T_{PC.PR}$		1.59	2.9678684×10^5

Conclusion

References

1. May RM, Mac Arthur RH (1972) Niche overlap as a function of environmental variability. Proceedings of the National Academy of Sciences 69: 1109-1113.
2. Violle C, Enquist BJ, McGill BJ, Jiang L, Albert CH, et al. (2012) The return of the variance: intraspecific variability in community ecology. Trends in Ecology and Evolution 27: 244-252.

List of Figures

1	General outline of the simulation concerning the Tstatistics	8
2	Distribution of trait values for one randomization without filtering	12
3	Results of T-statistics for one randomization without filter	13

¹¹one index for all the dataset

¹²one index for each community of the dataset

4	Local alpha errors	14
5	Global alpha errors	18
6	Distribution of trait values for one randomization with internal filtering	23
7	Results of T-statistics for two randomization with contrasted strength of internal filtering	24
8	Beta-error mixing all strength of internal filtering	25
9	Local $T_{IP/IC}$ SES and initial parameters values	27
10	Local $T_{IP/IC}$ SES and modeled parameters values	29
11	Global $T_{IP/IC}$ SES and initial parameters values	31
12	Global $T_{IP/IC}$ SES and modeled parameters values	32
13	Power of $T_{IP/IC}$ to detect internal filtering	34
14	Distribution of trait values for one randomization with external filtering	38
15	Results of T-statistics for two randomization with contrasted strength of external filtering	39
16	Local $T_{IC/IR}$ SES and initial parameters values	41
17	Local $T_{PC/PR}$ SES and initial parameters values	42
18	Local $T_{IC/IR}$ SES and modeled parameters values	45
19	Local $T_{PC/PR}$ SES and modeled parameters values	46
20	Global $T_{IC/IR}$ SES and initial parameters values	49
21	Global $T_{PC/PR}$ SES and initial parameters values	50
22	Global $T_{IC/IR}$ SES and modeled parameters values	52
23	Global $T_{PC/PR}$ SES and modeled parameters values	53
24	Power of $T_{IC/IR}$ to detect external filtering	55
25	Power of $T_{PC/PR}$ to detect external filtering	57
26	Distribution of trait values for one randomization with both internal and external filter	61
27	Results of T-statistics with both internal and external filter	62
28	Variance partitioning accross nested scales	68
29	Decomposition of the variation in community trait composition for two contrasted cases	70
30	Modeled against obtained turnover contribution in community trait variance	71

List of Tables

1	T-statistics and associated null models	6
2	Alpha-errors for T-statistics	17
3	Local beta-error of $T_{IP/IC}$ in function of the strength of internal filtering. str: strength. Trait a is normally distributed and trait b is uniformly distributed.	28
4	Local beta-error of $T_{IC/IR}$ in function of the strength of external filtering. str: strength. Trait a is normally distributed and trait b is uniformly distributed.	43
5	Local beta-error of $T_{PC/PR}$ in function of the strength of external filtering. str: strength. Trait a is normally distributed and trait b is uniformly distributed.	43
6	Initial parameter ratio allowing beta-error < 0.05 for T-statistics under both external and internal filtering	65
7	Summary of initial parameter ratio allowing beta-error < 0.05 for global T-statistics	72