# **Summary**

## Introduction

## **Materials and Methods**

Basic explanation of the models. We modeled a stage-structured population in two stages: immatures and matures. The demography is given by a transition matrix, with...

### Under constant environment, no plasticity

Using Lande (2009), under weak selection we have:

$$\Delta \overline{z} = \frac{d \ln \overline{\lambda}(\overline{z})}{d\overline{z}} = \frac{1}{\overline{\lambda}(\overline{z})} \frac{d\overline{\lambda}(\overline{z})}{d\overline{z}} \tag{1}$$

And we have:

$$\overline{\lambda}(\overline{z}) = \sum_{i,j} v_i u_j \overline{a_{ij}}$$

$$= v_I u_I \overline{a_{II}} + v_I u_M \overline{a_{IM}} + v_M u_I \overline{a_{MI}} + v_M u_M \overline{a_{MM}}$$

With  $\overline{a_{ij}}$  the expected values of the coefficient of the transition matrix. Thus,

$$\overline{\lambda}(\overline{z}) = v_I u_I \left[ \overline{f_1}(\overline{z}) m s_0 + (1 - m) \overline{s_I}(\overline{z}) \right] + v_I u_M s_0 \overline{f_2}(\overline{z}) 
+ v_M u_I m s_M + v_M u_M s_M$$
(2)

$$\frac{d\overline{\lambda}(\overline{z})}{d\overline{z}} = v_I u_I \left[ \frac{d\overline{f_1}(\overline{z})}{d\overline{z}} m s_0 + (1 - m) \frac{d\overline{s_I}(\overline{z})}{d\overline{z}} \right] + v_I u_M s_0 \frac{d\overline{f_2}(\overline{z})}{d\overline{z}}$$
(3)

Because  $f_i$  and  $s_I$  are gaussians we can write the population means  $\overline{f_i}$  and  $\overline{s_I}$  easily.

$$\overline{f_1}(\overline{z}) = f_1(\theta_f) \sqrt{\frac{\omega_f}{\omega_f + P_I}} \exp\left(-\frac{(\overline{z} - \theta_f)^2}{2(\omega_f + P_I)}\right)$$
(4a)

$$\overline{f_2}(\overline{z}) = f_2(\theta_f) \sqrt{\frac{\omega_f}{\omega_f + P_M}} \exp\left(-\frac{(\overline{z} - \theta_f)^2}{2(\omega_f + P_M)}\right)$$
(4b)

$$\overline{s_I}(\overline{z}) = s_I(\theta_s) \sqrt{\frac{\omega_s}{\omega_s + P_I}} \exp\left(-\frac{(\overline{z} - \theta_s)^2}{2(\omega_s + P_I)}\right)$$
(4c)

Thus we can derive these expression with respect to  $\overline{z}$ :

$$\frac{\partial \overline{f_1}(\overline{z})}{\partial \overline{z}} = f_1(\theta_f) \sqrt{\frac{\omega_f}{\omega_f + P_I}} \frac{\partial \exp\left(-\frac{(\overline{z} - \theta_f)^2}{2(\omega_f + P_I)}\right)}{\partial \overline{z}}$$

$$= f_1(\theta_f) \sqrt{\frac{\omega_f}{\omega_f + P_I}} \exp\left(-\frac{(\overline{z} - \theta_f)^2}{2(\omega_f + P_I)}\right) \frac{\theta_f - \overline{z}}{\omega_f + P_I}$$

$$= \overline{f_1}(\overline{z}) \frac{\theta_f - \overline{z}}{\omega_f + P_I}$$
(5)

We obtain similar formulas for  $\overline{f_2}$  and  $\overline{s_I}$ . Plugging (5) into (3) we have:

$$\frac{d\overline{\lambda}(\overline{z})}{d\overline{z}} = v_I u_I \left[ \frac{\theta_f - \overline{z}}{\omega_f + P_I} m s_0 + (1 - m) \frac{\theta_s - \overline{z}}{\omega_s + P_I} \right] + v_I u_M s_0 \frac{\theta_f - \overline{z}}{\omega_f + P_M}$$
 (6)

Using (6) into (1) gives us after rearranging: We have for variations of phenotype, under weak selection:

$$\Delta \overline{z} = (\theta_f - \overline{z}) \left[ \frac{v_I u_I G_I s_0 m \overline{f_1}}{\lambda (P_I + \omega_f)} + \frac{v_I u_M G_M s_0 \overline{f_2}}{\lambda (P_M + \omega_f)} \right] + (\theta_s - \overline{z}) \left[ \frac{v_I u_I G_I \overline{s_I} (1 - m)}{\lambda (P_I + \omega_s)} \right]$$
(7)

Within the square brackets, we see weighting average of fecundity and survival. Thus, we define them as  $\gamma_f$  and  $\gamma_s$  such as:

$$\gamma_f = \frac{v_I u_I s_0 m \overline{f_1}}{\lambda (P_I + \omega_f)} + \frac{v_I u_M \frac{G_M}{G_I} s_0 \overline{f_2}}{\lambda (P_M + \omega_f)}$$
(8a)

and

$$\gamma_s = \frac{v_I u_I \overline{s_I} (1 - m)}{\lambda (P_I + \omega_s)} \tag{8b}$$

We end up having a simpler expression for  $\Delta \overline{z}$  under constant environment:

$$\Delta \overline{z} = -G_I \left[ \gamma_f (\overline{z} - \theta_f) + \gamma_s (\overline{z} - \theta_s) \right] 
\Delta \overline{z} = -G_I \gamma (\overline{z} - \theta_v)$$
(9)

with

$$\gamma = \gamma_f + \gamma_s \tag{10}$$

$$\theta_v = \frac{\frac{\gamma_f}{\gamma_s}\theta_f + \theta_s}{\frac{\gamma_f}{\gamma_s} + 1} \tag{11}$$

# Under varying environment, without plasticity

From Engen et al. (2011), we derived equations for mean variation of phenotype on our model. We supposed an auto-correlated fluctuating environment  $\epsilon_t$  influencing optimums  $\theta_i$  such as:

$$\begin{cases} \theta_i(t) = \overline{\theta}_i + \alpha_i \epsilon_t \\ \epsilon_{t+1} = (1 - \rho)\overline{\epsilon} + \rho \epsilon_t + \xi \end{cases}$$
 (12)

with  $\alpha_i$  the dependence factor of the optimum on the environment,  $\rho$  the auto-correlation coefficient of the environment,  $\bar{\epsilon}$  the expected environment and  $\xi$  a gaussian noise vector with variance  $\sigma_{\xi}^2$  and mean 0. We chose  $\bar{\epsilon} = 0$  to simplify the calculations so that  $\epsilon_{t+1} = \rho \epsilon_t + \xi$ , we can see that:

$$\theta_{i}(t+1) = \overline{\theta}_{i} + \alpha_{i}\epsilon_{t+1}$$

$$= \overline{\theta}_{i} + \alpha_{i}(\rho\epsilon_{t} + \xi)$$

$$= \overline{\theta}_{i} + \alpha_{i}\rho(\frac{\theta_{i}(t) - \overline{\theta}_{i}}{\alpha_{i}}) + \alpha_{i}\xi$$

$$\theta_{i}(t+1) = \overline{\theta}_{i}(1-\rho) + \rho\theta_{i}(t) + \alpha_{i}\xi$$
(13)

The auto-correlation in the environment  $\epsilon_t$  causes  $\theta_i$  to be auto-correlated with the same correlation coefficient  $\rho$ .

Using the same approach as in a constant environment, under weak selection, we end up having a similar equation than (7) but with optimum depending on environment:

$$\Delta \overline{z}_t = (\theta_f(t) - \overline{z}_t) \left[ \frac{v_I u_I G_I s_0 m \overline{f_1}}{\lambda_t (P_I + \omega_f)} + \frac{v_I u_M G_M s_0 \overline{f_2}}{\lambda_t (P_M + \omega_f)} \right] + (\theta_s(t) - \overline{z}_t) \left[ \frac{v_I u_I G_I \overline{s_I} (1 - m)}{\lambda_t (P_I + \omega_s)} \right]$$
(14)

Plugging (12) in (14) we obtain

$$\Delta \overline{z}_{t} = (\overline{\theta}_{f} + \alpha_{f} \epsilon_{t} - \overline{z_{t}}) \left[ \frac{v_{I} u_{I} G_{I} s_{0} m \overline{f_{1}}}{\lambda_{t} (P_{I} + \omega_{f})} + \frac{v_{I} u_{M} G_{M} s_{0} \overline{f_{2}}}{\lambda_{t} (P_{M} + \omega_{f})} \right] + (\overline{\theta}_{s} + \alpha_{s} \epsilon_{t} - \overline{z_{t}}) \left[ \frac{v_{I} u_{I} G_{I} \overline{s_{I}} (1 - m)}{\lambda_{t} (P_{I} + \omega_{s})} \right]$$

$$(15a)$$

Defining the same  $\gamma_f$ ,  $\gamma_s$  and  $\theta_v$  as in (8a), (8b) and (11), respectively:

$$\Delta \overline{z}_{t} = G_{I} \left[ (\overline{\theta}_{f} + \alpha_{f} \epsilon_{t} - \overline{z_{t}}) \gamma_{f} + (\overline{\theta}_{s} + \alpha_{s} \epsilon_{t} - \overline{z_{t}}) \gamma_{s} \right] 
= -G_{I} \left[ \gamma_{f} (\overline{z_{t}} - \theta_{f}) + \gamma_{s} (\overline{z_{t}} - \theta_{s}) \right] - G_{I} \epsilon_{t} (\gamma_{f} \alpha_{f} + \gamma_{s} \alpha_{s}) 
\Delta \overline{z}_{t} = \Delta \overline{z} + G_{I} \epsilon_{t} (\gamma_{f} \alpha_{f} + \gamma_{s} \alpha_{s}) 
= -G_{I} \gamma (\overline{z_{t}} - \theta_{v}) + G_{I} \epsilon_{t} (\gamma_{f} \alpha_{f} + \gamma_{s} \alpha_{s}) 
= -G_{I} \gamma \left( \overline{z_{t}} - \theta_{v} - \epsilon_{t} \frac{\gamma_{f} \alpha_{f} + \gamma_{s} \alpha_{s}}{\gamma} \right) 
\Delta \overline{z}_{t} = -G_{I} \gamma (\overline{z_{t}} - \theta_{v} - \alpha_{v} \epsilon_{t})$$
(15c)

With  $\alpha_v$  a weighted component between  $\alpha_f$  and  $\alpha_s$ , defined in similar fashion as  $\theta_v$  in (11):

$$\alpha_v = \frac{\gamma_f \alpha_f + \gamma_s \alpha_s}{\gamma}$$

$$= \frac{\gamma_f \alpha_f + \gamma_s \alpha_s}{\gamma_f + \gamma_s}$$

$$\alpha_v = \frac{\frac{\gamma_f}{\gamma_s} \alpha_f + \alpha_s}{\frac{\gamma_f}{\gamma_s} + 1}$$
(16)

#### Estimating variance of $\overline{z}_t$

Taking (15c) we can estimate variance of  $\overline{z}_t$ . Using the same process as Engen et al. (2011). Indeed, (15c) has the form:

$$\Delta A_t = -DA_t + e_t \tag{17}$$

with  $A_t = \overline{z_t} - \theta_v$ ,  $D = G_I \gamma$  and  $e_t = D\alpha_v \epsilon_t$ . (17) has a stationary solution:

$$A_{t+1} = (1-D)^{t+1}A_0 + \sum_{r=0}^{t} (1-D)^r e_{t-r}$$
(18a)

If we consider the evolution of  $A_t$  over a long time, (18a) becomes, because (1 - D) < 1:

$$A_t = \sum_{r=0}^{\infty} e_{t-r} (1 - D)^r$$
 (18b)

We want to estimate how will  $\bar{z}_t$  move away from the mean because of environmental fluctuations, that is why we compute its variance. From (18a):

$$\operatorname{Var}(A_{t}) = \operatorname{Var}\left[\sum_{r=0}^{\infty} e_{t-r}(1-D)^{r}\right]$$

$$= \sum_{r=0}^{\infty} \operatorname{Var}\left[e_{t-r}(1-D)^{r}\right]$$

$$= \sum_{r=0}^{\infty} \operatorname{Var}\left[\epsilon_{t-r}D\alpha_{v}(1-D)^{r}\right]$$

$$\stackrel{\text{def}}{=} \sigma_{\epsilon}^{2}D^{2}\alpha_{v}^{2} \sum_{r=0}^{\infty} (1-D)^{2r}$$

$$\operatorname{Var}(A_{t}) = \sigma_{\epsilon}^{2}D^{2}\alpha_{v}^{2} \frac{1}{1-(1-D)^{2}}$$

$$\operatorname{Var}(\overline{z_{t}}) \stackrel{\text{def}}{=} \sigma_{\epsilon}^{2}G_{I}^{2}\gamma^{2}\alpha_{v}^{2} \frac{1}{G_{I}\gamma(2-G_{I}\gamma)}$$

$$\operatorname{Var}(\overline{z_{t}}) \stackrel{\text{def}}{=} \frac{1}{2}G_{I}\gamma\sigma_{\epsilon}^{2}\alpha_{v}^{2}$$

$$\operatorname{Var}(\overline{z_{t}}) \stackrel{\text{def}}{=} \frac{1}{2}G_{I}\gamma\sigma_{\epsilon}^{2}\alpha_{v}^{2}$$

### **Results**

Subheading1

Subheading2

#### **Discussion**

# **Authors Contributions and Acknowledgments**

#### References

(Barfield et al., 2011)

# References

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