

APPENDIX

A MATHEMATICAL PROOFS

A.1 Preliminaries

We derive a lemma to bound the sum of n random variables. This lemma is similar to the Hoeffding bound but cannot be replaced by Hoeffding.

LEMMA 1. Let X_1, \dots, X_n be n random variables such that

$$X_i \in \{0, s_i\}, \quad \Pr(X_i = s_i \mid X_1, \dots, X_{i-1}) \leq p,$$

where $0 \leq s_i \leq 1$. Let $X = \sum_{i=1}^n X_i$, and $\mu = \sum_{i=1}^n p s_i = nmp$.

$$\Pr(X > (1 + \Delta)\mu) \leq e^{-(\Delta - (e-2))nmp}$$

PROOF. For any $t > 0$, by using the Markov inequality we have

$$\Pr(X > (1 + \Delta)\mu) = \Pr(e^X > e^{(1+\Delta)\mu}) \leq \frac{E(e^X)}{e^{(1+\Delta)\mu}}.$$

According to the conditions, we have

$$\begin{aligned} E(e^X) &= E\left(E(e^{\sum_{i=1}^n X_i} \mid X_1, \dots, X_{n-1})\right) \\ &= E\left(e^{\sum_{i=1}^{n-1} X_i} \cdot \Pr(X_n = 0 \mid X_1, \dots, X_{n-1})\right. \\ &\quad \left.+ e^{s_n + \sum_{i=1}^{n-1} X_i} \cdot \Pr(X_n = s_n \mid X_1, \dots, X_{n-1})\right) \\ &\leq E\left(e^{\sum_{i=1}^{n-1} X_i} \cdot (1 + p(e^{s_n} - 1))\right) \leq \dots \leq \prod_{i=1}^n (1 + p(e^{s_i} - 1)) \end{aligned}$$

Because of $1 + x < e^x$, we have

$$E(e^X) \leq \prod_{i=1}^n e^{p(e^{s_i} - 1)}.$$

Since for $s_i \leq 1$, there is $e^{s_i} - 1 \leq (e - 1)s_i$, so there is

$$E(e^X) \leq e^{\sum_{i=1}^n p(e-1)s_i} = e^{(e-1)nmp}.$$

That is

$$\Pr(X > (1 + \Delta)\mu) \leq \frac{e^{(e-1)nmp}}{e^{(1+\Delta)nmp}} = e^{-(\Delta - (e-2))nmp}$$

□

A.2 Definition of Symbols

- (1) \mathcal{S}_i : $\{e_1, \dots, e_{N_i}\}$, the set of keys entering the i -th layer, where $N_i = |\mathcal{S}_i|$.
- (2) $f_i(e)$: the number of times that key e enters the i -th layer.
- (3) \mathcal{S}_i^0 : $\{e \mid e \in \mathcal{S}_i \wedge \forall i' \leq i, f_{i'}(e) \leq \frac{\lambda_{i'}}{2}\}$, the set of mice keys.
- (4) \mathcal{S}_i^1 : $\{e \mid e \in \mathcal{S}_i \wedge \exists i' \leq i, f_{i'}(e) > \frac{\lambda_{i'}}{2}\}$, the set of elephant keys.
- (5) F_i : $\sum_{e \in \mathcal{S}_i^0} f_i(e)$, the total frequency of mice keys in \mathcal{S}_i^0 .
- (6) C_i : $|\mathcal{S}_i^1|$, the number of elephant keys in \mathcal{S}_i^1 .
- (7) $\mathcal{S}_{i,j}^0$: $\{e \mid e \in \mathcal{S}_i^0 \wedge h(e) = j\}$, the set of mice keys that are mapped to the j -th bucket.
- (8) $\mathcal{S}_{i,j}^1$: $\{e \mid e \in \mathcal{S}_i^1 \wedge h(e) = j\}$, the set of elephant keys that are mapped to the j -th bucket.
- (9) $F_{i,j}$: $\sum_{e \in \mathcal{S}_{i,j}^0} f_i(e)$, the total frequency of mice keys in $\mathcal{S}_{i,j}^0$.

(10) $C_{i,j}$: $|\mathcal{S}_{i,j}^1|$, the number of elephant keys in $\mathcal{S}_{i,j}^1$.

(11) $\mathcal{P}_{i,k}$: $\{e_1, \dots, e_k\}$, a subset of \mathcal{S}_i composed of the first k keys.

(12) $f_{i,k}^P$: $\sum_{\{e \in \mathcal{P}_{i,k-1} \cap \mathcal{S}_{i,h(e_k)}^0\}} f_i(e)$, the total frequency of mice keys with a smaller index that conflicts with key e_k .

(13) $c_{i,k}^P$: $|\{e \mid e \in \mathcal{P}_{i,k-1} \cap \mathcal{S}_{i,h(e_k)}^1\}|$, the number of elephant keys with a smaller index that conflicts with key e_k .

A.3 Properties in One Layer

This section aims to prove that only a small proportion of the keys inserted into the i -th layer will be inserted into the $(i+1)$ -th layer.

THEOREM A.1. (Theorem 1) Let

$$X_{i,k} = \begin{cases} 0 & C_{i,h(e_k)} = 0 \wedge f_{i,k}^P \leq \frac{\lambda_i}{2} \\ f_i(e_k) & C_{i,h(e_k)} = 0 \wedge f_{i,k}^P > \frac{\lambda_i}{2} \\ f_i(e_k) & C_{i,h(e_k)} > 0 \end{cases}, \quad X_i = \sum_{\{e_k \in \mathcal{S}_i^0\}} X_{i,k}.$$

The total frequency of the mice keys in the i -th layer leaving it does not exceed X_i , i.e.,

$$F_{i+1} \leq \sum_{\{e \in \mathcal{S}_i^0 \cap \mathcal{S}_{i+1}\}} f_{i+1}(e) \leq X_i.$$

PROOF. For the mice keys in the j -th bucket of the i -th layer, let the number of times they leave be $F'_{i,j} = \sum_{\{e \in \mathcal{S}_{i,j}^0 \cap \mathcal{S}_{i+1}^0\}} f_{i+1}(e)$.

Since a bucket can hold at least λ_i packets of the key, we have:

$$\begin{cases} F'_{i,j} = 0 & C_{i,j} = 0 \wedge F_{i,j} \leq \lambda_i \\ F'_{i,j} \leq F_{i,j} - \lambda_i & C_{i,j} = 0 \wedge F_{i,j} > \lambda_i \\ F'_{i,j} \leq F_{i,j} & C_{i,j} > 0 \end{cases}.$$

When $C_{i,j} = 0 \wedge F_{i,j} > \lambda_i$, exists k' satisfies

$$\sum_{\{e_k \in \mathcal{S}_{i,j}^0 \wedge k < k'\}} f_i(e_k) \leq \frac{\lambda_i}{2} \leq \sum_{\{e_k \in \mathcal{S}_{i,j}^0 \wedge k \leq k'\}} f_i(e_k) \leq \lambda_i.$$

Then for and only for any $e_k \in \mathcal{S}_{i,j}^0 \wedge k \leq k'$, there is $X_{i,k} = 0$, and

$$\begin{aligned} F'_{i,j} &\leq \left(\sum_{\{e_k \in \mathcal{S}_{i,j}^0 \wedge k \leq k'\}} f_i(e_k) + \sum_{\{e_k \in \mathcal{S}_{i,j}^0 \wedge k > k'\}} f_i(e_k) \right) - \lambda_i \\ &\leq 0 + \sum_{\{e_k \in \mathcal{S}_{i,j}^0 \wedge k > k'\}} f_i(e_k) \\ &\leq \sum_{\{e_k \in \mathcal{S}_{i,j}^0 \wedge k \leq k'\}} X_{i,k} + \sum_{\{e_k \in \mathcal{S}_{i,j}^0 \wedge k > k'\}} X_{i,k}. \end{aligned}$$

Then we have $F'_{i,j} \leq \sum_{\{e_k \in \mathcal{S}_{i,j}^0\}} X_{i,k}$, and

$$\sum_{\{e \in \mathcal{S}_i^0 \cap \mathcal{S}_{i+1}\}} f_{i+1}(e) = \sum_{j=1}^{w_i} F'_{i,j} \leq \sum_{j=1}^{w_i} \sum_{\{e_k \in \mathcal{S}_{i,j}^0\}} X_{i,k} = X_i.$$

□

Similarly, we have the following lemma.

THEOREM A.2. Let

$$Y_{i,k} = \begin{cases} 0 & c_{i,k}^P = 0 \wedge F_{i,h(e_k)} \leq \lambda_i, \\ 2 & c_{i,k}^P = 0 \wedge F_{i,h(e_k)} > \lambda_i, \\ 2 & c_{i,k}^P > 0. \end{cases} \quad Y_i = \sum_{e_k \in \mathcal{S}_i^1} Y_{i,k}.$$

The number of distinct elephant keys in the i -th layer leaving it does not exceed Y_i , i.e.,

$$|\mathcal{S}_i^1 \cap \mathcal{S}_{i+1}^1| \leq Y_i.$$

PROOF. For the elephant keys in the j -th bucket of the i -th layer, $\sum_{\{e_k \in \mathcal{S}_{i,j}^1\}} Y_{i,k} < C_{i,j}$ if and only if $C_{i,j} = 1 \wedge F_{i,j} \leq \lambda_i$. In this case, the number of collisions in the bucket does not exceed λ_i , and no key enters the $(i+1)$ -th layer. Thus we have $|\mathcal{S}_{i,j}^1 \cap \mathcal{S}_{i+1}^1| \leq \sum_{\{e_k \in \mathcal{S}_{i,j}^1\}} Y_{i,j}$, and

$$|\mathcal{S}_i^1 \cap \mathcal{S}_{i+1}^1| = \sum_{j=1}^{w_i} |\mathcal{S}_{i,j}^1 \cap \mathcal{S}_{i+1}^1| \leq \sum_{j=1}^{w_i} \sum_{\{e_k \in \mathcal{S}_{i,j}^1\}} Y_{i,j} = Y_i.$$

□

THEOREM A.3. Let $W = \frac{4N(R_w R_\lambda)^6}{\Lambda(R_w - 1)(R_\lambda - 1)}$, $\alpha_i = \frac{\|F\|_1}{(R_w R_\lambda)^{i-1}}$, $\beta_i = \frac{\alpha_i}{\frac{\lambda_i}{2}}$, $\gamma_i = (R_w R_\lambda)^{(2^{i-1}-1)}$, and $p_i = (R_w R_\lambda)^{-(2^{i-1}+4)}$. Under the conditions of $F_i \leq \frac{\alpha_i}{\gamma_i}$ and $C_i \leq \frac{\beta_i}{\gamma_i}$, we have:

$$\begin{aligned} \Pr(X_{i,k} > 0 \mid X_{i,1}, \dots, X_{i,k-1}) &\leq p_i, \quad \forall e_k \in \mathcal{S}_i^0. \\ \Pr(Y_{i,k} > 0 \mid Y_{i,1}, \dots, Y_{i,k-1}) &\leq \frac{3}{4} p_i, \quad \forall e_k \in \mathcal{S}_i^1. \end{aligned}$$

PROOF. By using Markov's inequality, we have

$$\begin{aligned} &\Pr(X_{i,k} > 0 \mid X_{i,1}, \dots, X_{i,k-1}) \\ &= \Pr\left(\left(C_{i,h(e_k)} = 0 \wedge f_{i,k}^P > \frac{\lambda_i}{2}\right) \mid X_{i,1}, \dots, X_{i,k-1}\right) \\ &\quad \vee \quad C_{i,h(e_k)} > 0 \\ &\leq \Pr\left(C_{i,h(e_k)} > 0 \mid X_{i,1}, \dots, X_{i,k-1}\right) \\ &\quad + \Pr\left(F_{i,h(e_k)} - f_i(e_k) > \frac{\lambda_i}{2} \mid X_{i,1}, \dots, X_{i,k-1}\right) \\ &\leq \frac{E(C_{i,h(e_k)} \mid X_{i,1}, \dots, X_{i,k-1})}{1} \\ &\quad + \frac{E(F_{i,h(e_k)} - f_i(e_k) \mid X_{i,1}, \dots, X_{i,k-1})}{\frac{\lambda_i}{2}} \\ &\leq \frac{C_i}{w_i} + \frac{2F_i}{\lambda_i w_i} \end{aligned}$$

$$\begin{aligned} &\Pr(Y_{i,k} > 0 \mid Y_{i,1}, \dots, Y_{i,k-1}) \\ &= \Pr\left(\left(c_{i,k}^P = 0 \wedge F_{i,h(e_k)} > \lambda_i\right) \vee c_{i,k}^P > 0 \mid Y_{i,1}, \dots, Y_{i,k-1}\right) \\ &\leq \Pr\left(C_{i,h(e_k)} - 1 > 0 \mid Y_{i,1}, \dots, Y_{i,k-1}\right) \\ &\quad + \Pr\left(F_{i,h(e_k)} > \lambda_i \mid Y_{i,1}, \dots, Y_{i,k-1}\right) \\ &\leq \frac{E(C_{i,h(e_k)} - 1 \mid Y_{i,1}, \dots, Y_{i,k-1})}{1} \\ &\quad + \frac{E(F_{i,h(e_k)} \mid Y_{i,1}, \dots, Y_{i,k-1})}{\lambda_i} \\ &\leq \frac{C_i}{w_i} + \frac{F_i}{\lambda_i w_i}. \end{aligned}$$

Recall that $w_i = \lceil \frac{W(R_w - 1)}{R_\lambda^i} \rceil$ and $\lambda_i = \frac{\Lambda(R_\lambda - 1)}{R_\lambda^i}$, under the conditions of $F_i \leq \frac{\alpha_i}{\gamma_i}$ and $C_i \leq \frac{\beta_i}{\gamma_i}$, we have

$$\begin{aligned} &\Pr(X_{i,k} > 0 \mid X_{i,1}, \dots, X_{i,k-1}) \\ &\leq \frac{\beta_i}{\gamma_i w_i} + \frac{2\alpha_i}{\gamma_i \lambda_i w_i} = \frac{4\alpha_i}{\gamma_i \lambda_i w_i} \leq \frac{1}{(R_w R_\lambda)^{2^{i-1}+4}} = p_i. \\ &\Pr(Y_{i,k} > 0 \mid Y_{i,1}, \dots, Y_{i,k-1}) \\ &\leq \frac{\beta_i}{\gamma_i w_i} + \frac{\alpha_i}{\gamma_i \lambda_i w_i} = \frac{3\alpha_i}{\gamma_i \lambda_i w_i} \leq \frac{3}{4(R_w R_\lambda)^{2^{i-1}+4}} \leq \frac{3}{4} p_i. \end{aligned}$$

□

THEOREM A.4. (Theorem 2) Let $W = \frac{4N(R_w R_\lambda)^6}{\Lambda(R_w - 1)(R_\lambda - 1)}$, $\alpha_i = \frac{\|F\|_1}{(R_w R_\lambda)^{i-1}}$, $\beta_i = \frac{\alpha_i}{\frac{\lambda_i}{2}}$, $\gamma_i = (R_w R_\lambda)^{(2^{i-1}-1)}$, and $p_i = (R_w R_\lambda)^{-(2^{i-1}+4)}$. Under the conditions of $F_i \leq \frac{\alpha_i}{\gamma_i}$ and $C_i \leq \frac{\beta_i}{\gamma_i}$, we have

$$\Pr\left(X_i > (1 + \Delta) \frac{p_i \alpha_i}{\gamma_i}\right) \leq \exp\left(-(\Delta - (e - 2)) \frac{2p_i \alpha_i}{\lambda_i \gamma_i}\right).$$

and

$$\Pr\left(Y_i > (1 + \Delta) \frac{3}{2} \frac{p_i \beta_i}{\gamma_i}\right) \leq \exp\left(-(\Delta - (e - 2)) \frac{3p_i \beta_i}{4\gamma_i}\right).$$

PROOF. According to Theorem A.3,

$$\begin{aligned} &\Pr\left(\frac{X_{i,k}}{\frac{\lambda_i}{2}} = \frac{f_i(e_k)}{\frac{\lambda_i}{2}} \mid \frac{X_{i,1}}{\frac{\lambda_i}{2}}, \dots, \frac{X_{i,k-1}}{\frac{\lambda_i}{2}}\right) \leq p_i. \\ &\Pr\left(\frac{Y_{i,k}}{2} = 1 \mid \frac{Y_{i,1}}{2}, \dots, \frac{Y_{i,k-1}}{2}\right) \leq \frac{3}{4} p_i. \end{aligned}$$

According to Lemma 1,

$$\begin{aligned}
& \Pr \left(X_i > (1 + \Delta) \frac{p_i \alpha_i}{\gamma_i} \right) \leq \Pr \left(X_i > (1 + \Delta) p_i F_i \mid F_i = \frac{\alpha_i}{\gamma_i} \right) \\
&= \Pr \left(\sum_{\{e_k \in \mathcal{S}_i^0\}} \frac{X_{i,k}}{\frac{\lambda_i}{2}} > (1 + \Delta) p_i \sum_{\{e_k \in \mathcal{S}_i^0\}} \frac{f_i(e_k)}{\frac{\lambda_i}{2}} \mid F_i = \frac{\alpha_i}{\gamma_i} \right) \\
&\leq \exp \left(-(\Delta - (e - 2)) \frac{\alpha_i}{\gamma_i \frac{\lambda_i}{2}} p_i \right) = \exp \left(-(\Delta - (e - 2)) \frac{2p_i \alpha_i}{\lambda_i \gamma_i} \right). \\
&\Pr \left(Y_i > (1 + \Delta) \frac{3}{2} \frac{p_i \beta_i}{\gamma_i} \right) \leq \Pr \left(Y_i > (1 + \Delta) \frac{3}{2} p_i C_i \mid C_i = \frac{\beta_i}{\gamma_i} \right) \\
&= \Pr \left(\sum_{\{e_k \in \mathcal{S}_i^1\}} \frac{Y_{i,k}}{2} > (1 + \Delta) \frac{3}{4} p_i \sum_{\{e_k \in \mathcal{S}_i^1\}} \frac{2}{2} \mid C_i = \frac{\beta_i}{\gamma_i} \right) \\
&\leq \exp \left(-(\Delta - (e - 2)) \frac{\beta_i}{\gamma_i} \frac{3}{4} p_i \right) = \exp \left(-(\Delta - (e - 2)) \frac{3p_i \beta_i}{4\gamma_i} \right).
\end{aligned}$$

□

THEOREM A.5. (Theorem 3) Let $R_w R_\lambda \geq 2$, $W = \frac{4N(R_w R_\lambda)^6}{\Lambda(R_w - 1)(R_\lambda - 1)}$, $\alpha_i = \frac{\|F\|_1}{(R_w R_\lambda)^{i-1}}$, $\beta_i = \frac{\alpha_i}{\frac{\lambda_i}{2}}$, $\gamma_i = (R_w R_\lambda)^{(2^{i-1}-1)}$, and $p_i = (R_w R_\lambda)^{-(2^{i-1}+4)}$. We have

$$\begin{aligned}
& \Pr \left(F_{i+1} > \frac{\alpha_{i+1}}{\gamma_{i+1}} \mid F_i \leq \frac{\alpha_i}{\gamma_i} \wedge C_i \leq \frac{\beta_i}{\gamma_i} \right) \\
&\leq \exp \left(-(9 - e) \frac{2p_i \alpha_i}{\lambda_i \gamma_i} \right). \\
&\Pr \left(C_{i+1} > \frac{\beta_{i+1}}{\gamma_{i+1}} \mid F_i \leq \frac{\alpha_i}{\gamma_i} \wedge C_i \leq \frac{\beta_i}{\gamma_i} \right) \\
&\leq \exp \left(-(5 - e) \frac{2p_i \alpha_i}{\lambda_i \gamma_i} \right) + \exp \left(-\left(\frac{11}{3} - e\right) \frac{3p_i \beta_i}{4\gamma_i} \right).
\end{aligned}$$

PROOF. According to settings, we have

$$\begin{aligned}
p_i \frac{\alpha_i}{\gamma_i} &= \frac{\|F\|_1}{(R_w R_\lambda)^{(2^i+i+2)}} \leq \frac{1}{8} \frac{\alpha_{i+1}}{\gamma_{i+1}} \\
p_i \frac{\beta_i}{\gamma_i} &= p_i \frac{\alpha_i}{\gamma_i \frac{\lambda_i}{2}} \leq \frac{1}{8} \frac{\alpha_{i+1}}{\gamma_{i+1} \frac{\lambda_{i+1}}{2}} = \frac{1}{8} \frac{\beta_{i+1}}{\gamma_{i+1}}.
\end{aligned}$$

Recall that $C_{i+1} = |\mathcal{S}_{i+1}^1 \cap \mathcal{S}_i^0| + |\mathcal{S}_{i+1}^1 \cap \mathcal{S}_i^1|$, and

$$|\mathcal{S}_{i+1}^1 \cap \mathcal{S}_i^0| \leq \frac{\sum_{\{e \in \mathcal{S}_i^0 \cap \mathcal{S}_{i+1}^1\}} f_i(e)}{\frac{\lambda_{i+1}}{2}} \leq \frac{X_i}{\frac{\lambda_{i+1}}{2}}$$

Let $\Gamma_i = \left(F_i \leq \frac{\alpha_i}{\gamma_i} \wedge C_i \leq \frac{\beta_i}{\gamma_i} \right)$, according to Theorem A.1 and Theorem 2,

$$\begin{aligned}
\Pr(F_{i+1} > \frac{\alpha_{i+1}}{\gamma_{i+1}} \mid \Gamma_i) &\leq \Pr \left(X_i > 8p_i \frac{\alpha_i}{\gamma_i} \mid \Gamma_i \right) \\
&\leq \exp \left(-(9 - e) \frac{2p_i \alpha_i}{\lambda_i \gamma_i} \right).
\end{aligned}$$

According to Theorem A.2 and Theorem 2,

$$\begin{aligned}
& \Pr(C_{i+1} > \frac{\beta_{i+1}}{\gamma_{i+1}} \mid \Gamma_i) \\
&= \Pr(|\mathcal{S}_{i+1}^1 \cap \mathcal{S}_i^0| + |\mathcal{S}_{i+1}^1 \cap \mathcal{S}_i^1| > \frac{\beta_{i+1}}{\gamma_{i+1}} \mid \Gamma_i) \\
&\leq \Pr(|\mathcal{S}_{i+1}^1 \cap \mathcal{S}_i^0| > \frac{\beta_{i+1}}{2\gamma_{i+1}} \vee |\mathcal{S}_{i+1}^1 \cap \mathcal{S}_i^1| > \frac{\beta_{i+1}}{2\gamma_{i+1}} \mid \Gamma_i) \\
&\leq \Pr \left(\frac{X_i}{\frac{\lambda_{i+1}}{2}} > \frac{\beta_{i+1}}{2\gamma_{i+1}} \mid \Gamma_i \right) + \Pr(Y_i > \frac{\beta_{i+1}}{2\gamma_{i+1}} \mid \Gamma_i) \\
&\leq \Pr(X_i > 4p_i \frac{\alpha_i}{\gamma_i} \mid \Gamma_i) + \Pr(Y_i > 4p_i \frac{\beta_i}{\gamma_i} \mid \Gamma_i) \\
&\leq \exp \left(-(5 - e) \frac{2p_i \alpha_i}{\lambda_i \gamma_i} \right) + \exp \left(-\left(\frac{11}{3} - e\right) \frac{3p_i \beta_i}{4\gamma_i} \right).
\end{aligned}$$

□

A.4 Space and Time Complexity

THEOREM A.6. (Theorem 4) Let $R_w R_\lambda \geq 2$, $W = \frac{4N(R_w R_\lambda)^6}{\Lambda(R_w - 1)(R_\lambda - 1)}$, $\alpha_i = \frac{\|F\|_1}{(R_w R_\lambda)^{i-1}}$, $\beta_i = \frac{\alpha_i}{\frac{\lambda_i}{2}}$, $\gamma_i = (R_w R_\lambda)^{(2^{i-1}-1)}$, and $p_i = (R_w R_\lambda)^{-(2^{i-1}+4)}$.

For given Λ and $\Delta < \frac{1}{4}$, let d be the root of the following equation

$$\frac{R_\lambda^d}{(R_w R_\lambda)^{(2^d+d)}} = \Delta_1 \frac{\Lambda}{N} \ln \left(\frac{1}{\Delta} \right).$$

And use an SpaceSaving of size $\Delta_2 \ln(\frac{1}{\Delta})$ (as the $(d+1)$ -layer), then

$$\Pr \left(\forall \text{ item } e, \left| \hat{f}(e) - f(e) \right| \leq \Lambda \right) \geq 1 - \Delta,$$

where

$$\Delta_1 = 2R_w^2 R_\lambda^2 (R_\lambda - 1), \quad \Delta_2 = 3 \left(\frac{R_w R_\lambda^2}{R_\lambda - 1} \right) \Delta_1 = 6R_w^3 R_\lambda^4.$$

PROOF. Recall that $\Gamma_i = \left(F_i \leq \frac{\alpha_i}{\gamma_i} \wedge C_i \leq \frac{\beta_i}{\gamma_i} \right)$, When all conditions Γ_i (including Γ_{d+1}) are true, we have

$$C_{d+1} \leq \frac{\beta_{d+1}}{\gamma_{d+1}} = \frac{2NR_\lambda^{d+1}}{(R_w R_\lambda)^{(2^d+d-1)}(R_\lambda - 1)\Lambda} = \left(\frac{2R_w R_\lambda^2}{R_\lambda - 1} \right) \Delta_1 \ln \left(\frac{1}{\Delta} \right).$$

$$F_{d+1} \leq \frac{\alpha_{d+1}}{\gamma_{d+1}} = \frac{\lambda_{d+1}}{2} \frac{\beta_{d+1}}{\gamma_{d+1}} = \lambda_{d+1} \left(\frac{R_w R_\lambda^2}{R_\lambda - 1} \right) \Delta_1 \ln \left(\frac{1}{\Delta} \right).$$

Since we use an SpaceSaving of size $\Delta_2 \ln(\frac{1}{\Delta}) > C_{d+1}$, it can record all elephant keys without error, and the estimation error for mice keys does not exceed

$$\frac{F_{d+1}}{\Delta_2 \ln(\frac{1}{\Delta}) - C_{d+1}} \leq \frac{\lambda_{d+1} \left(\frac{R_w R_\lambda^2}{R_\lambda - 1} \right) \Delta_1 \ln \left(\frac{1}{\Delta} \right)}{\Delta_2 \ln(\frac{1}{\Delta}) - \left(\frac{2R_w R_\lambda^2}{R_\lambda - 1} \right) \Delta_1 \ln \left(\frac{1}{\Delta} \right)} = \lambda_{d+1}$$

Therefore, for any item e ,

$$\left| \hat{f}(e) - f(e) \right| = \sum_{i=1}^d \lambda_i \leq \sum_{i=1}^{\infty} \frac{\Lambda(R_\lambda - 1)}{R_\lambda^i} = \Lambda$$

Next, we deduce the probability that at least one condition Γ_i is false. Note that

$$\left. \begin{aligned} & \left(\frac{11}{3} - e \right) \frac{3p_i\beta_i}{4\gamma_i} \\ & (9 - e) \frac{2p_i\alpha_i}{\lambda_i\gamma_i} \\ & (5 - e) \frac{2p_i\alpha_i}{\lambda_i\gamma_i} \end{aligned} \right\} \geq \frac{p_i\alpha_i}{\lambda_i\gamma_i}.$$

Then According to Theorem A.5, we have

$$\begin{aligned} \Pr \left(\neg \left(\bigwedge_{i=1}^d \Gamma_{i+1} \right) \right) &= \Pr \left(\bigvee_{i=1}^d \neg \Gamma_{i+1} \right) = \Pr \left(\bigvee_{i=1}^d \left(\bigwedge_{j=1}^i \Gamma_j \wedge \neg \Gamma_{i+1} \right) \right) \\ &\leq \sum_{i=1}^d \Pr (\Gamma_i \wedge \neg \Gamma_{i+1}) \leq \sum_{i=1}^d \Pr (\neg \Gamma_{i+1} \mid \Gamma_i) \\ &\leq \sum_{i=1}^d \left(\exp \left(- \left(\frac{11}{3} - e \right) \frac{3p_i\beta_i}{4\gamma_i} \right) \right. \\ &\quad \left. + \exp \left(- (9 - e) \frac{2p_i\alpha_i}{\lambda_i\gamma_i} \right) + \exp \left(- (5 - e) \frac{2p_i\alpha_i}{\lambda_i\gamma_i} \right) \right) \\ &\leq \sum_{i=1}^d 3 \exp \left(- \frac{p_i\alpha_i}{\lambda_i\gamma_i} \right). \end{aligned}$$

Note that

$$\begin{aligned} \exp \left(- \frac{p_d\alpha_d}{\lambda_d\gamma_d} \right) &= \exp \left(- \frac{NR_\lambda^d}{(R_w R_\lambda)^{(2^d+d+2)} \Lambda(R_\lambda - 1)} \right) \\ &= \exp \left(- \frac{1}{R_w^2 R_\lambda^2 (R_\lambda - 1)} \Delta_1 \ln \left(\frac{1}{\Delta} \right) \right) \\ &= \Delta \left(\frac{1}{R_w^2 R_\lambda^2 (R_\lambda - 1)} \Delta_1 \right) = \Delta^2. \end{aligned}$$

Since $\Delta \leq 1$, and the monotonicity of $\exp \left(- \frac{p_d\alpha_d}{\lambda_d\gamma_d} \right)$, we have

$$\begin{aligned} \exp \left(- \frac{p_i\alpha_i}{\lambda_i\gamma_i} \right) &= \exp \left(- \frac{p_{i+1}\alpha_{i+1}}{\lambda_{i+1}\gamma_{i+1}} \cdot R_w^{(2^i+1)} R_\lambda^{(2^i)} \right) \\ &\leq \exp \left(- \frac{p_{i+1}\alpha_{i+1}}{\lambda_{i+1}\gamma_{i+1}} \right)^{R_w R_\lambda} \leq \exp \left(- \frac{p_{i+1}\alpha_{i+1}}{\lambda_{i+1}\gamma_{i+1}} \right)^2 \\ &\leq \Delta^2 \exp \left(- \frac{p_{i+1}\alpha_{i+1}}{\lambda_{i+1}\gamma_{i+1}} \right) \end{aligned}$$

Therefore, we have

$$\sum_{i=1}^d 3 \exp \left(- \frac{p_i\alpha_i}{\lambda_i\gamma_i} \right) \leq 3 \sum_{i=1}^d \Delta^{2i} \leq \left(\frac{3\Delta}{1 - \Delta^2} \right) \Delta \leq \Delta.$$

In other words,

$$\Pr \left(\forall \text{ item } e, \left| \hat{f}(e) - f(e) \right| \leq \Delta \right) \geq 1 - \Delta,$$

which leads to a weaker conclusion,

$$\forall \text{ item } e, \Pr \left(\left| \hat{f}(e) - f(e) \right| \leq \Delta \right) \geq 1 - \Delta.$$

□

THEOREM A.7. *Using the same settings as Theorem A.6, the space complexity of the algorithm is $O(\frac{N}{\Lambda} + \ln(\frac{1}{\Delta}))$, and the time complexity of the algorithm is amortized $O(1 + \Delta \ln \ln(\frac{N}{\Lambda}))$.*

PROOF. Recall that d is the root of the equation

$$\frac{R_\lambda^d}{(R_w R_\lambda)^{(2^d+d)}} = \Delta_1 \frac{\Lambda}{N} \ln \left(\frac{1}{\Delta} \right),$$

which means $d = O \left(\ln \ln \left(\frac{N}{\Lambda} \right) \right)$. Therefore, total space used by the data structure is

$$\begin{aligned} \sum_{i=1}^d w_i + \Delta_1 \ln \left(\frac{1}{\Delta} \right) &= \sum_{i=1}^d \left\lceil \frac{W(R_w - 1)}{R_w^i} \right\rceil + O(\ln \left(\frac{1}{\Delta} \right)) \\ &\leq \frac{4N(R_w R_\lambda)^6}{\Lambda(R_w - 1)(R_\lambda - 1)} + d + O(\ln \left(\frac{1}{\Delta} \right)) \\ &= O \left(\frac{N}{\Lambda} + \ln \left(\frac{1}{\Delta} \right) \right) \end{aligned}$$

Next, we analyze the time complexity. When all condition Γ_i are true, for a new item $e \notin \mathcal{S}_1$, the probability that it enters the $(i+1)$ -th

layer from the i -th layer is $\frac{\frac{F_i}{\lambda_i^2} + C_i}{w_i} \leq p_i$. Thus the time complexity of insert item e does not exceed

$$(1 - \Delta) \cdot \left(1 + \sum_{i=1}^d p_i \right) + \Delta \cdot d = O \left(1 + \Delta \ln \ln \left(\frac{N}{\Lambda} \right) \right).$$

□