

COMP2610/COMP6261

Tutorial 4 Sample Solutions

Tutorial 4: Entropy and Information

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1. Suppose Y is a geometric random variable, $Y \sim \text{Geom}(p)$. i.e., Y has probability function

$$P(Y = y) = p(1 - p)^{y-1}, \quad y = 1, 2, \dots$$

Determine the mean and variance of the geometric random variable.

Solution.

The expectation of the geometric random variable can be calculated as:

$$\begin{aligned} E[Y] &= \sum_{y=1}^{\infty} y \cdot P(Y = y) \\ &= \sum_{y=1}^{\infty} y \cdot p(1 - p)^{y-1} \\ &= p \sum_{y=1}^{\infty} y(1 - p)^{y-1} \end{aligned}$$

$$E[Y] = p[1 + 2(1 - p) + 3(1 - p)^2 + \dots] \quad (1)$$

$$(1 - p)E[Y] = [(1 - p) + 2(1 - p)^2 + 3(1 - p)^3 + \dots] \quad (2)$$

$$E[Y] \cdot (1 - (1 - p)) = p[1 + (1 - p) + (1 - p)^2 + \dots] \quad (1) - (2)$$

$$E[Y] \cdot p = p \cdot \frac{1}{(1 - (1 - p))} \quad (*)$$

$$E[Y] = \frac{1}{p} \quad (3)$$

(*) Here we use the sum to infinity of geometric series, where $|p| < 1$,

$$\sum_{i=1}^{\infty} p^i = \frac{1}{1 - p} \quad (4)$$

To calculate the variance, we need to calculate $E[Y^2]$:

$$\begin{aligned}
E[Y^2] &= \sum_{y=1}^{\infty} y^2 \cdot P(Y = y) \\
&= \sum_{y=1}^{\infty} y^2 \cdot p(1-p)^{y-1} \\
&= \sum_{y=1}^{\infty} (y-1+1)^2 \cdot p(1-p)^{y-1} \\
&= \sum_{y=1}^{\infty} ((y-1)^2 + 2(y-1) + 1) \cdot p \cdot r^{y-1} && \text{let } r = 1 - p \\
&= \sum_{z=0}^{\infty} z^2 pr^z + 2 \sum_{z=0}^{\infty} zpr^z + \sum_{z=0}^{\infty} pr^z && \text{let } z = y - 1 \\
&= r \cdot \sum_{z=0}^{\infty} z^2 pr^{z-1} + 2r \cdot \sum_{z=0}^{\infty} zpr^{z-1} + p \sum_{z=0}^{\infty} r^z \\
&= r \cdot \sum_{z=1}^{\infty} z^2 pr^{z-1} + 2r \cdot \sum_{z=1}^{\infty} zpr^{z-1} + p \cdot \frac{1}{1 - (1-p)} && \text{using (4)} \\
E[Y^2] &= r \cdot E[Y^2] + 2r \cdot E[Y] + 1 \\
E[Y^2] &= \frac{1+r}{p^2} && (5)
\end{aligned}$$

\therefore the variance can be calculated as

$$\begin{aligned}
Var[Y] &= E[Y^2] - (E[Y])^2 \\
&= \frac{1+r}{p^2} - \left(\frac{1}{p}\right)^2 && \text{using (5)} \\
&= \frac{r}{p^2} \\
&= \frac{1-p}{p^2} && (6)
\end{aligned}$$

2. A standard deck of cards contains 4 *suits* — $\heartsuit, \diamondsuit, \clubsuit, \spadesuit$ (“hearts”, “diamonds”, “clubs”, “spades”) — each with 13 *values* — A, 2, 3, 4, 5, 6, 7, 8, 9, 10, J, Q, K (The A, J, Q, K are called “Ace”, “Jack”, “Queen”, “King”). Each card has a *colour*: hearts and diamonds are coloured red; clubs and spades are black. Cards with values J, Q, K are called *face cards*.

Each of the 52 cards in a deck is identified by its value v and suit s and denoted vs . For example, $2\heartsuit$, $J\clubsuit$, and $7\spadesuit$ are the “two of hearts”, “Jack of clubs”, and “7 of spades”, respectively. The variable c will be used to denote a card’s colour. Let $f = 1$ if a card is a face card and $f = 0$ otherwise.

A card is drawn at random from a thoroughly shuffled deck. Calculate:

- The information $h(c = \text{red}, v = K)$ in observing a red King
- The conditional information $h(v = K | f = 1)$ in observing a King given a face card was drawn.
- The entropies $H(S)$ and $H(V, S)$.
- The mutual information $I(V; S)$ between V and S .
- The mutual information $I(V; C)$ between the value and colour of a card using the last result and the *data processing inequality*.

Solution.

$$(a) \ h(c = \text{red}, v = K) = \log_2 \frac{1}{P(c=\text{red}, v=K)} = \log_2 \frac{1}{1/26} = 4.7004 \text{ bits.}$$

$$(b) \ h(v = K | f = 1) = \log_2 \frac{1}{P(v=K | f=1)} = \log_2 \frac{1}{1/3} = 1.585 \text{ bits.}$$

(c) We have

$$i. \ H(S) = \sum_s p(s) \log_2 \frac{1}{p(s)} = 4 \times \frac{1}{4} \times \log_2 \frac{1}{1/4} = 2 \text{ bits.}$$

$$ii. \ H(V, S) = \sum_{v,s} p(v, s) \log_2 \frac{1}{p(v, s)} = 52 \times \frac{1}{52} \log_2 \frac{1}{1/52} = 5.7 \text{ bits.}$$

(d) Since V and S are independent we have $I(V; S) = 0$ bits.

(e) Since C is a function of S and by the data processing inequality $I(V; C) \leq I(V; S) = 0$. However, mutual information must be nonnegative so we must have $I(V; C) = 0$ bits.

3. Recall that for a random variable X , its variance is

$$\text{Var}[X] = E[X^2] - (E[X])^2.$$

Using Jensen’s inequality, show that the variance must always be non-negative.

Solution.

This is a direct application of Jensen’s inequality to the convex function $g(x) = x^2$.

4. Let X and Y be independent random variables with possible outcomes $\{0, 1\}$, each having a Bernoulli distribution with parameter $\frac{1}{2}$, i.e.

$$p(X = 0) = p(X = 1) = \frac{1}{2}$$

$$p(Y = 0) = p(Y = 1) = \frac{1}{2}.$$

- (a) Compute $I(X; Y)$.
- (b) Let $Z = X + Y$. Compute $I(X; Y|Z)$.
- (c) Do the above quantities contradict the data-processing inequality? Explain your answer.

Solution.

- (a) We see that $I(X; Y) = 0$ as $X \perp\!\!\!\perp Y$.
- (b) To compute $I(X; Y|Z)$ we apply the definition of conditional mutual information:

$$I(X; Y|Z) = H(X|Z) - H(X|Y, Z)$$

Now, X is fully determined by Y and Z . In other words, given Y and Z there is only one state of X that is possible, i.e it has probability 1. Therefore the entropy $H(X|Y, Z) = 0$. We have that:

$$I(X; Y|Z) = H(X|Z)$$

To determine this value we look at the distribution $p(X|Z)$, which is computed by considering the following possibilities:

X	Y	Z
0	0	0
0	1	1
1	0	1
1	1	2

Therefore:

$$\mathbf{p}(X|Z = 0) = (1, 0)$$

$$\mathbf{p}(X|Z = 1) = (1/2, 1/2)$$

$$\mathbf{p}(X|Z = 2) = (0, 1)$$

From this, we obtain: $H(X|Z = 0) = 0$, $H(X|Z = 2) = 0$, $H(X|Z = 1) = 1$ bit. Therefore:

$$I(X; Y|Z) = p(Z = 1)H(X|Z = 1) = (1/2)(1) = 0.5 \text{ bits.}$$

- (c) This does not contradict the data-processing inequality (or more specifically the “conditioning on a downstream variable” corollary): the random variables in this example do not form a Markov chain. In fact, Z depends on both X and Y .