# Lecture 13: Central Limit Theorem + Confidence Intervals

Chapter 4.4 + 4.2

# Goals for Today

- Discuss the Central Limit Theorem
- Introduce confidence intervals
- Interpretation

## Recap

▶ Point estimates are based on a sample  $x_1, ..., x_n$  and are used to estimate population parameters.

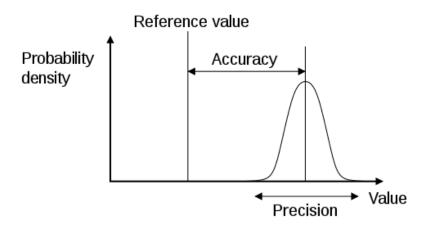
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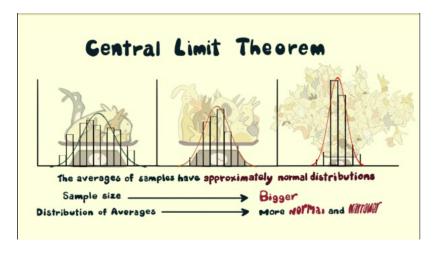
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- ▶ The sampling distribution characterizes the (random) behavior of point estimates (like  $\overline{x}$ ).
- The standard deviation of a sampling distribution is the standard error: it quantifies the uncertainty/variability of point estimates.

# Illustrative Image





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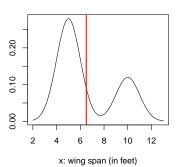
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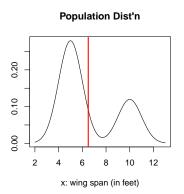
#### Population Dist'n

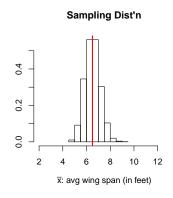


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Answer: So we can use the Normal table on p.409 of the book to calculate probabilities. We call this using the normal model.

	Second decimal place of $Z$									
Z	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7704	0.7734	0.7764	0.7794	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
<u>:</u>	:	:	:	:	÷	:	:	:	:	·

#### **Definition**

For a sample  $x_1, \ldots, x_n$  of independent observations, if n is "large" enough to counteract the skew of the population distribution, then the sampling distribution of  $\overline{x}$  is approximately Normal with

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Recall: If we don't know  $\sigma$ , we can plug in its point estimate s if the two conditions are satisfied.

This translates to the following conditions to be able to use the Normal model with s in place of  $\sigma$ , as stated in the book:

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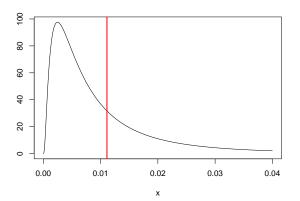
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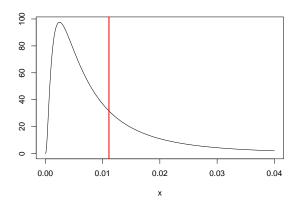
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- n ≥ 30.
   Comment: This is a rule of thumb that works for most cases.
   In some situations you might need less, in others more.
- 3. The population distribution is not strongly skewed. Comment: This is related to the *n* necessary. The larger the *n*, the more lenient we can be with the skew assumption.

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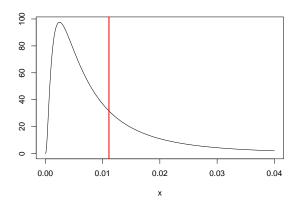


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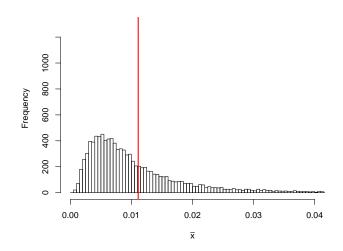
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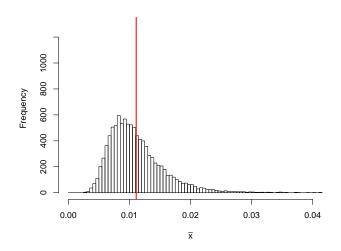
This is where your individual observations  $x_i$  come from. Now compare 10000 values of  $\overline{x}$ 's based on different n: 2, 10, 30, 75.

For 10000 values of  $\overline{x}$  based on samples of size n=2, the sampling distribution is:

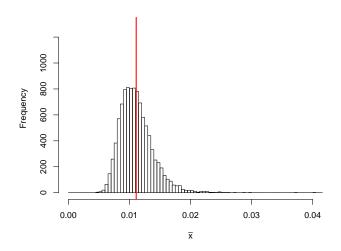
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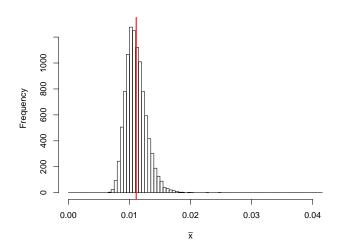
For 10000 values of  $\overline{x}$  based on samples of size n = 10, the sampling distribution is:



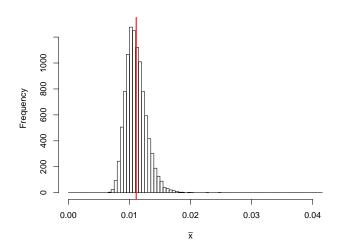
For 10000 values of  $\overline{x}$  based on samples of size n=30, the sampling distribution is:



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Using just the point estimate:

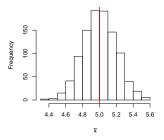
Using a confidence interval:



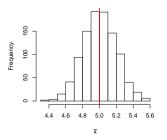


Recall the example of 1000 instances of  $\overline{x}$  based on n=100. Each observation came from a population distribution that was Normal with  $\mu=5$  &  $\sigma=2$ .

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We observed the sampling distribution

- $\blacktriangleright$  is centered at  $\mu$
- ▶ has spread  $SE = \frac{\sigma}{\sqrt{n}} = \frac{2}{\sqrt{100}} = 0.2$

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- ▶ the SE is the standard deviation of the sampling distribution
- roughly 95% of the time  $\overline{x}$  will be within 2 SE of  $\mu$  if the sampling distribution is normal

If the interval spreads out 2 SE from  $\overline{x}$ , we can be roughly "95% confident" that we have captured the true parameter  $\mu$ .

A 95% confidence interval is (no more using rule of thumb  $2 \times SD$ ):

$$\overline{x} \pm 1.96SE = [\overline{x} - 1.96SE, \overline{x} + 1.96SE]$$

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If we don't know  $\sigma$ , assuming the conditions hold, plug in s

$$\overline{x} \pm 1.96 \frac{s}{\sqrt{n}} = \left[ \overline{x} - 1.96 \frac{s}{\sqrt{n}}, \ \overline{x} + 1.96 \frac{s}{\sqrt{n}} \right]$$

### Confidence Intervals

In general a confidence interval will be

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Ex: For 95% confidence  $z^* = 1.96$ . For 99% confidence  $z^* = 2.58$ 

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- ▶ Wrong, yet common, interpretation: There is a 95% chance that the C.I. captures the true population mean  $\mu$ . The probability is 0 or 1: either it does or it doesn't.
- ▶ Correct, interpretation: If we were to repeat this sampling procedure 100 times, we expect 95 (i.e. 95%) of calculated C.I.'s to capture the true  $\mu$

### Illustration: How to Interpret a Confidence Interval

In Chapter 4 there is an example of finish times (in minutes) from the 2012 Cherry Blossom 10 mile run with n=16,924 participants. In this case, we can compute the true population mean  $\mu=94.52$ .

### Illustration: How to Interpret a Confidence Interval

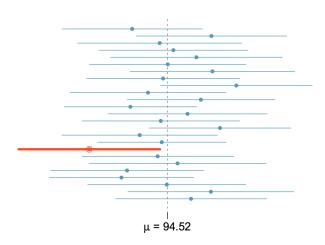
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Say we take 25 (random) samples of size n = 100 and for each sample we compute:

- **▶** S
- ▶ and hence the 95% CI:  $\left[\overline{x} 1.96 \times \frac{s}{\sqrt{n}}, \ \overline{x} + 1.96 \times \frac{s}{\sqrt{n}}\right]$

# How to Interpret a Confidence Interval

Of the 25 Cl's based on 25 different samples of size n=100, one of them (in red) did not capture the true population mean  $\mu$ :



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- ▶ The margin of error of  $\pm 3.4\%$  indicates that 95% C.I. is:

$$45 \pm 3.4\% = [41.6, 48.4]$$

Intrepretation: the interpretation is not that there is a 95% chance that [41.6, 48.4] captures the true %'age. Rather, that if we were to take 20 such polls, 19 of them would capture the true %'age.

### Next Time

Hypothesis Testing: we can perform statistical tests on population parameters such as  $\mu$ :

#### Define:

- Null and alternative hypotheses.
- Testing hypotheses using confidence intervals.
- Types of errors