

Bayesian Statistics

Lecture 2

"PROBABILITY DOES NOT EXIST"

Subjectivists: De Finetti,
Savage, Ramsey

Example: coin tossing

- We want to learn θ from $\text{Bern}(\theta)$
- Specifying prior on $[0,1]$

1814 Laplace: uniform prior to express our ignorance. $\text{Beta}(1,1)$

Conjugate prior: Given a likelihood $p(x|\theta)$, we choose a family of prior distributions such that the posterior (and also integrals such as the Bayes marginal) can be analytically obtained.

Ex contⁿ. Coin is tossed t times,
 $z^t \sim \text{Bern}(\theta)$, n_1 ones, n_0 zeros.

$$p(z^t | \theta) = \theta^{n_1} (1-\theta)^{n_0}$$

\Rightarrow posterior $\text{Beta}(1+n_1, 1+n_0)$

\Rightarrow posterior mean: $E(\theta|z) = \frac{n_1+1}{n_1+n_0+2}$

(MCMC)

Exercise 1

$$p(\alpha, \beta | x) \propto \pi(\alpha, \beta) p(x | \alpha, \beta) \\ = \pi(\alpha) \pi(\beta) p(x | \alpha, \beta)$$

$$= (1-\alpha)^{b-1} \alpha^{a-1} (1-\beta)^{q-1} \beta^{p-1}$$

$$\alpha^r (1-\alpha)^{n-r} \left(\prod_{i=1}^{n-r} (1-\beta)^{x_i-1} \right) \beta^{n-r}$$

$$= \frac{\alpha^{r+a-1} (1-\alpha)^{b+n-r-1} \beta^{n+p-r-1}}{(1-\beta)^{\sum x_i + r + q + n - 1}}$$

$$\alpha: \text{Beta}(r+a, b+n-r) = \pi(\alpha | x)$$

$$B: \text{Beta}(n+p-r, \sum x_i + r+q-n)$$

$$= \pi(B|x)$$

Exercise 2

$$p(x|\theta) = \prod_{i=1}^3 p(x_i|\theta) \cdot P[x_4 > 100] \cdot P[x_5 > 100]$$

$$P[x_5 > 100] = \int_{100}^{\infty} \theta \exp\{-\theta x\} d\theta = e^{-100\theta}$$

$$\rightarrow p(x|\theta) = \theta^3 \exp\{-449\theta\}$$

prior

$$p(\theta) \sim \text{Ga}(2, 180) \propto \theta^{2-1} \exp\{-180\theta\}$$

posterior

$$p(\theta|x) \propto \theta^{2-1} \exp\{-180\theta\} \theta^3 \exp\{-449\theta\}$$

$$\propto \theta^{5-1} \exp\{-629\theta\} \rightarrow \text{Ga}(5, 629)$$

$$\text{mean: } 5/629 \quad \text{Var: } 5/629^2$$

Posterior mode : $\arg\max_{\theta} \pi(\theta|x)$

1. J. Good: "There are 46656 varieties of Bayesians." (1971)

Bayesianism

Central to Bayesian Statistics is the inter-

pretation of probability as degrees of belief.

1) Kolmogorov's (1933) axioms

2) Let S be some statement, then we start with a prior probability $P_{old}(S)$.

Upon acquiring new evidence E , we transform our prior to generate a posterior by conditionalization on E .

$$P_{new}(S) = P_{old}(S|E)$$

→ Bayes' Rule.

The origin of Priors

Subjectivists (De Finetti, Savage, Ramsey)

They take probability to be the expression of personal opinion.

Priors: no constraints (other than coherence, i.e. Kolmogorov's axioms).

Objectivists (Jeffreys, Jaynes)

Prior probabilities should be rationally constrained.

- no principles exist that uniquely determine rational priors in all cases.

Pragmatists most statisticians.

Choose their priors for pragmatic reasons

- mathematical convenience
- computational "
- special effects.

- satisfying specific criteria

"Frequentist Bayes"

Priors: often exhibit both subjective and objective elements

Computation before 1990: conjugate priors

after 1990: MCMC - Markov Chain Monte Carlo

Exercise 3

$$X^n := X_1, \dots, X_n \quad p(\delta) \propto \delta^{-1}$$

$$L(x_1, \dots, x_n | \theta=2, \delta) = \prod_{i=1}^n 2\delta^2 x_i = 2^n \delta^{2n} \prod_{i=1}^n x_i$$

$$p(\delta | x_1, \dots, x_n) \propto \delta^{-1} \cdot \underbrace{2^n \prod_{i=1}^n x_i}_C \delta^{2n} \propto \underbrace{C \delta^{2n-1}}_{\uparrow}$$

$$\int C \delta^{2n-1} d\delta = 1$$

$$\downarrow$$
$$C \int_0^t \delta^{2n-1} d\delta = 1$$

$$\downarrow$$
$$\frac{\delta^{2n}}{2n} \Big|_0^t = \frac{t^{2n}}{2n}$$

$$\rightarrow C = \frac{2n}{t^{2n}} \rightarrow p(\delta | x_1, \dots, x_n) =$$

$$= \frac{2n \delta^{2n-1}}{t^{2n}} \quad 0 < \delta < t$$

Jeffreys' prior

uninformative priors

Example

$$X \sim \text{Bern}(\theta)$$

$$\rightarrow \pi(\theta) = 1 \quad 0 \leq \theta \leq 1$$

log odds ratio

$$\eta = \log \frac{\theta}{1-\theta}$$

$$\rightarrow \pi(\eta) = \frac{e^\eta}{(1+e^\eta)^2}$$

'Flat priors' are not parameterization-invariant.

Jethneys: par.-invariant prior

$$\pi(\theta) \propto \sqrt{|I(\theta)|}$$

$$= \sqrt{E\left[\left(\frac{\partial}{\partial \theta} \log f(x; \theta)\right)^2 \mid \theta\right]}$$

if $f(x; \theta)$ is twice differentiable w.r.t θ and under certain regularity conditions.

$$I(\theta) = -E\left[\frac{\partial^2}{\partial \theta^2} \log f(x; \theta) \mid \theta\right]$$

Why?

$$\frac{\partial^2}{\partial \theta^2} \log f(x; \theta) = \frac{\frac{\partial^2}{\partial \theta^2} f(x; \theta)}{f(x; \theta)} -$$

$$\left(\frac{\frac{\partial}{\partial \theta} f(x; \theta)}{f(x; \theta)}\right)^2$$

$$= \frac{\frac{\partial^2}{\partial \theta^2} f(x; \theta)}{f(x; \theta)} - \left(\frac{\partial}{\partial \theta} \log f(x; \theta) \right)^2$$

and

$$E \left[\frac{\frac{\partial^2}{\partial \theta^2} f(x; \theta)}{f(x; \theta)} \mid \theta \right] = \frac{\partial^2}{\partial \theta^2} \int_{\mathbb{R}} f(x; \theta) dx = 0$$

Note: ~~the~~ only holds for 1-dimension.



Jeffreys' prior for
Beta-Bernoulli model

Beta($\frac{1}{2}, \frac{1}{2}$) prior

par. invariant \rightarrow this does not mean that
it is 'uninformative'

\Rightarrow A general principle for making
non-informative priors turns out not to exist.

Exercise 4

$$p(x|\delta) = \theta \delta^\theta x$$

$$\log p(x|\delta) = \log \theta + \theta \log \delta + \log x$$

$$\frac{d \log p}{d \delta} = \frac{\theta}{\delta}$$

$$\frac{d^2 \log p}{d\delta^2} = -\frac{\partial}{\delta^2}$$

$$I(\delta) = - \int p(x|\delta) \frac{d^2 \log p}{d\delta^2} dx$$

$$= \frac{\partial}{\delta^2} \underbrace{\int p(x|\delta) dx}_{=1}$$

$$= \frac{\partial}{\delta^2}$$

$$\Rightarrow \text{fwhm prior} \propto \sqrt{|I(\delta)|}$$

$$\propto \frac{1}{\delta}$$

(since ∂ is known)

Exercise 5

$$p(\theta) \propto \sqrt{|I(\theta)|}$$

$$f(x|n, \theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}$$

$$\log f(x|n, \theta) = \log \binom{n}{x} + x \log \theta + (n-x) \log_{(1-\theta)}$$

$$\frac{d}{d\theta} f(x|\theta) = \frac{x}{\theta} - (n-x) \frac{1}{1-\theta}$$

$$\frac{\partial}{\partial \theta^2} f(x|\theta) = -\frac{x}{\theta^2} - \frac{n-1}{(1-\theta)^2}$$