

## §. Stochastic Integration

### 4.2 Quadratic Variation

Def. 4.2.1. Let  $M$  be a stochastic process. Then

$$[M, M]_T(\omega) := \lim_{\|\Pi\| \rightarrow 0} \sum_{i=0}^{n-1} (\Delta M_i(\omega))^2$$

where  $\Delta M_i := M_{t_{i+1}} - M_{t_i}$  and  $\Pi = \{0 = t_0 < t_1 < \dots < t_n = T\}$

for Brownian motion  $W$ ,  $[W, W]_t = t$

Lemma. 4.2.1  $M_t = W_t^2 - t$  is a martingale

proof: • Adapted:  $M_t$  is  $\mathcal{F}_t$ -measurable because  $W_t$  is  $\mathcal{F}_t$ -measurable and  $M_t$  is a function of  $W_t$  and  $t$ .

• Martingale property: check  $E[M_t | \mathcal{F}_s] = M_s \quad \forall t \geq s$

$$E[M_t | \mathcal{F}_s] = E[W_t^2 - t | \mathcal{F}_s]$$

$$= E[(W_t - W_s + W_s)^2 - t | \mathcal{F}_s]$$

$$= E[(W_t - W_s)^2 + 2(W_t - W_s)W_s + W_s^2 - t | \mathcal{F}_s]$$

$$= E[\underbrace{(W_t - W_s)^2}_{\text{ind. of } \mathcal{F}_s} | \mathcal{F}_s] + E[\underbrace{2(W_t - W_s) \cdot W_s}_{\mathcal{F}_s\text{-measurable}} | \mathcal{F}_s]$$

$$+ \underbrace{E[W_s^2 | \mathcal{F}_s]}_{\mathcal{F}_s\text{-measurable}} - \underbrace{E[t | \mathcal{F}_s]}_{\mathcal{F}_s\text{-measurable}}$$

$$= E[(W_t - W_s)^2] + 2W_s E[W_t - W_s]$$

$$+ W_s^2 - t$$

$$= t - s + W_s^2 - t$$

$$= W_s^2 - s$$

$$= M_s$$

Theorem 4.2.1 Let  $M$  be a martingale w.r.t. a filtration  $(\mathcal{F}_t)_{t \geq 0}$

Then for all  $t \geq 0$ ,  $E[M_t^2] < \infty \Leftrightarrow E[ [M, M]_t ] < \infty$

In this case,  $(M_t^2 - [M, M]_t)_{t \geq 0}$  is a martingale w.r.t. the same filtration, and  $E[M_t^2] - E[M_0^2] = E[ [M, M]_t ]$

Example:  $M = W$  is a martingale  $\Rightarrow (W_t^2 - t)_{t \geq 0}$  is a martingale

Theorem 4.2.2. If  $M$  is a martingale and  $A_t$  is a continuous adapted increasing stochastic process s.t.  $A_0 = 0$  and  $(M_t^2 - A_t)_{t \geq 0}$  is a martingale, then

$$A_t = [M, M]_t \quad (\text{proof is omitted})$$

Example:  $M = W$ ,  $A_t = t$

$A_t$  is continuous, adapted,  $\mathcal{F}_0$ -measurable, increasing, and  $A_0 = 0$

$M_t^2 - A_t = W_t^2 - t$  is a martingale  $\Rightarrow A_t = [W, W]_t$

Remark:

▷ intuition for first variation and quadratic variation:

divide the interval  $[0, T]$  into  $T/(\delta t)$  intervals of size  $\delta t$ .

if  $X$  has finite first variation, then on each subinterval

$(k\delta t, (k+1)\delta t)$  the increment of  $X$  should be of order  $\delta t$ .

Similarly,  $X$  has finite quadratic variation  $\Rightarrow$  increment

▷ If a continuous process has finite first variation,  $\propto \sqrt{\delta t}$   
its quadratic variation will necessarily be zero.

If a continuous process has finite and non-zero quadratic variation, its first variation will necessarily be infinite

### 4.3. Construction of Itô integral.

Let  $W$  be a standard BM,  $(\mathcal{F}_t)_{t \geq 0}$  be the Brownian filtration and  $D$  be an adapted process.

Let  $(\Delta_t)_{t \geq 0}$  be our position in  $S$  at time  $t$ , that is

invest  $\Delta_t S_t$  at time  $t$  and the value of portfolio

time  $t+1$  is  $\Delta_t \cdot S_{t+1}$  however, almost any continuous martingale  $S$  will not have finite first variation, thus we need the Itô integral

$$\Rightarrow P_n I = \sum_{i=0}^{n-1} \Delta_i (S_{i+1} - S_i) \xrightarrow{n \rightarrow \infty} \int \Delta_t \cdot dS_t$$

Lemma 4.3.1 let  $\Pi = \{0 = t_0 < t_1 < \dots < t_n\}$  be an increasing sequence of times and assume that  $D$  is constant on  $[t_i, t_{i+1}) \forall i$  i.e. the asset is only traded at time  $t_0, \dots, t_n$

$$\text{let } I_T^\Pi = \sum_{i=0}^{n-1} D_{t_i} \Delta W_i + D_{t_n} (W_T - W_{t_n}) \text{ if } T \in [t_n, t_{n+1})$$

$$\text{where } \Delta W_i = W_{t_{i+1}} - W_{t_i}$$

denote the cumulative earnings up to time  $T$ , then

$$E[(I_T^\Pi)^2] = E\left[\sum_{i=0}^{n-1} D_{t_i}^2 (t_{i+1} - t_i) + D_{t_n}^2 (T - t_n)\right] \text{ if } T \in [t_n, t_{n+1})$$

Moreover,  $I^\Pi$  is a martingale and.

$$[I^\Pi, I^\Pi]_T = \sum_{i=0}^{n-1} D_{t_i}^2 (t_{i+1} - t_i) + D_{t_n}^2 (T - t_n) \text{ if } T \in [t_n, t_{n+1})$$

Theorem 4.3.1 If  $\int_0^T D_t^2 dt < \infty$ , then as  $\|\Pi\| \rightarrow 0$ , the process

$I^\Pi$  converge to a cost process  $I$  given by

$$I_T := \lim_{\|\Pi\| \rightarrow 0} I_T^\Pi = \int_0^T D_t dW_t$$

Note that  $D$  should be adapted, and is sampled at the left endpoint of the time interval, i.e. terms in the sum are  $D_{t_i}(W_{t_{i+1}} - W_{t_i})$

This is called the Itô integral of  $D$  w.r.t.  $W$ .

If further,  $E[\int_0^T D_t^2 dt] < \infty$ , then the process  $I$  is a martingale and the quadratic variation  $[I, I]$  satisfies  $[I, I]_T = \int_0^T D_t^2 dt$  almost surely. Besides,  $E[\int_0^T D_t dW_t] = 0$

#### Property 4.3.1 (linearity)

If  $D^1, D^2$  are two adapted processes,  $\lambda \in \mathbb{R}$ , then

$$\int_0^T (D_t^1 + \lambda D_t^2) dW_t = \int_0^T D_t^1 dW_t + \lambda \int_0^T D_t^2 dW_t$$

#### Itô Isometry

If  $E[\int_0^T D_t^2 dt] < \infty$ , then

$$E[(\int_0^T D_t dW_t)^2] = E[\int_0^T D_t^2 dt]$$

Example:  $D_t = 1$ , then

$$E[(\underbrace{\int_0^T D_t dW_t}_{W_T - W_0})^2] = E[\underbrace{\int_0^T 1^2 dt}_T]$$

$$\Rightarrow E[W_T^2] = T$$

#### Remark:

positivity is not preserved by Itô integrals. Namely, if  $D^1 \leq D^2$ , there is no reason to expect  $\int_0^T D_t^1 dW_t \leq \int_0^T D_t^2 dW_t$ .

Def (GBM). We define Geometric Brownian Motion  $S$  as

$$dS_t = \sigma S_t dW_t + \lambda S_t dt, \quad \sigma, \lambda \in \mathbb{R}$$

$\Rightarrow$  stochastic differential equation (SDE)

$$\int_0^T 1 dS_t = \int_0^T \sigma S_t dW_t + \int_0^T \lambda S_t dt$$

$$\Rightarrow S_T - S_0 = \int_0^T \sigma S_t dW_t + \int_0^T \lambda S_t dt$$

#### 4.4. Itô formula

Goal: Compute  $\int_0^T W_t dW_t = ?$

Def 4.4.1 Let  $b, \sigma$  be adapted process. Then a process  $X$  defined as

$$X_T = X_0 + \int_0^T b_t dt + \int_0^T \sigma_t dW_t \quad X_0 \in \mathbb{R}$$

Riemann integralItô integral

is called an Itô process if  $X_0$  is deterministic (not random)

and for all  $T \geq 0$   $E[\int_0^T \sigma_t^2 dt] < \infty$  and  $\int_0^T |b_t| dt < \infty$

Remark:

the equation above is equivalent to

$$dX_t = b_t dt + \sigma_t dW_t$$

property 4.4.1 The quadratic variation of  $X$  is

$$[X, X]_T = \int_0^T \sigma_t^2 dt$$

Def 4.4.2. Process  $X$  which can be decomposed as  $X = A + M$   
(semi-MG)

where  $M$  is a martingale and  $A$  has finite variation  
are called semi-martingale. (first)

▷  $M$  is called the martingale part of  $X$

▷  $A$  is called the finite variation part of  $X$

prop 4.4.2 The semi-MG decomposition is unique, that is

if  $X = A_1 + M_1 = A_2 + M_2^{(*)}$ , then

$$A_1 = A_2, M_1 = M_2$$

(where  $A_1, A_2$  are finite variation processes,  $M_1, M_2$  are martingales)

proof: by (\*) we have

$$\underbrace{A_1 - A_2}_{\text{finite variation}} = \underbrace{M_1 - M_2}_{\text{MG}} := M$$

thus  $M$  is a MG with finite Variation  $\Rightarrow [M, M]_T = 0$

$$\stackrel{4.2.1}{\Rightarrow} E[M_t^2] = E[[M, M]_T] = 0$$

$(M^2 - [M, M])_t$   
is a M.G.)

$\Rightarrow M_t^2 \geq 0$  thus  $M_t = 0$ , that is  $A_1 = A_2$   
 $M_1 = M_2$

Prop 4.4.3 Let  $X$  be an Itô process, then

$X$  is a MG  $\Leftrightarrow b_t = 0 \quad \forall t \geq 0$  (i.e.  $X_T = X_0 + \int_0^T \sigma_t dW_t$ )

proof: If  $b_t = 0 \quad \forall t \geq 0 \Rightarrow X$  is a M.G.

Suppose  $X$  is a M.G. Define

$$A_T := \int_0^T b_t dt = \underbrace{X_T - X_0}_{\text{M.G.}} - \underbrace{\int_0^T \sigma_t dW_t}_{\text{M.G.}}$$

$\Rightarrow A$  is a martingale. (also a semi-M.G.)

$$A = \underbrace{A + 0}_{\uparrow} = 0 + A \quad \text{thus } A = 0 \Leftrightarrow b_t = 0$$

Def 4.4.3. We define the integral of  $D$  w.r.t.  $X$  by

$$\int_0^T D_t dX_t := \int_0^T D_t \cdot b_t dt + \int_0^T D_t \sigma_t dW_t$$

where  $dX_t = b_t dt + \sigma_t dW_t$

[ Given an adapted process  $D$ , we interpret  $X$  as the price of an asset. and  $D$  as our position in it, which can be positive or negative).

Theorem 4.4 (Itô formula)

Recall that

if  $f: \mathbb{R} \rightarrow \mathbb{R}$  is continuously differentiable

$$f(y) - f(x) = \int_x^y \frac{\partial f}{\partial x}(z) dz$$

If  $f: [0, \infty) \times \mathbb{R} \rightarrow \mathbb{R}$  is such that

$t \mapsto f(t, x)$  is cont. differentiable  $\forall x \in \mathbb{R}$   $\partial_t f(t, x)$  exists

$x \mapsto f(t, x)$  is twice cont. differentiable.

$\forall t \in \mathbb{R}$ .  $\partial_x f(t, x)$ ,  $\partial_x^2 f(t, x)$  exist

and if  $X$  is an Itô process, then

$$f(T, X_T) - f(0, X_0) = \int_0^T \partial_t f(t, X_t) dt + \int_0^T \partial_x f(t, X_t) dX_t \\ + \underbrace{\frac{1}{2} \int_0^T \partial_x^2 f(t, X_t) d[X, X]_t}_{\text{Itô correction term}}$$

Remark:

- ▷  $\partial_t f(t, X_t)$  stands for taking derivative of  $f(t, X_t)$  w.r.t.  $t$  and then substitute  $X_t$ , Similar for  $\partial_x f(t, X_t)$ ,  $\partial_x^2 f(t, X_t)$
- ▷ The Itô formula is simply a version of the chain rule for Stochastic processes.

Stochastic form:

$$df(t, X_t) = \partial_t f(t, X_t) dt + \partial_x f(t, X_t) dX_t \\ + \frac{1}{2} \partial_x^2 f(t, X_t) d[X, X]_t$$

Substitute (the Itô process)  $\begin{cases} dX_t = b_t dt + \sigma_t dW_t \\ d[X, X]_t = \sigma_t^2 dt \end{cases}$  we have

$$\Rightarrow df(t, X_t) = \partial_t f(t, X_t) dt + \partial_x f(t, X_t) [b_t dt + \sigma_t dW_t] \\ + \frac{1}{2} \partial_x^2 f(t, X_t) \cdot \sigma_t^2 dt \\ = [\partial_t f(t, X_t) + \partial_x f(t, X_t) \cdot b_t + \frac{1}{2} \partial_x^2 f(t, X_t) \sigma_t^2] dt \\ + \partial_x f(t, X_t) \cdot \sigma_t dW_t$$

if the term before  $dt$  is zero,  $f(t, X_t)$  is a martingale.  
otherwise not

## 4.5 Examples

ex 1. write  $W^2$  as a sum of  $dt$  and  $dW_t$ -integral

use the Itô formula and substitute  $f(t, X) = X^2$ ,  $X = W_t$

$$\Rightarrow \partial_t f(t, X) = 0, \partial_X f(t, X) = 2X, \partial_X^2 f(t, X) = 2$$

$$\Rightarrow [W, W]_t = t$$

By Itô formula

$$\begin{aligned} d(W_t^2) &= d(f(t, W_t)) = \overset{\rightarrow 0}{\partial_t f(t, W_t)} dt + \overset{\rightarrow 2W_t}{\partial_X f(t, W_t)} dW_t \\ &\quad + \frac{1}{2} \underbrace{\partial_X^2 f(t, W_t)}_{\rightarrow 2} \underbrace{d[W, W]_t}_t \\ &= 2W_t dW_t + dt \end{aligned}$$

calculating the integral from 0 to  $T$

$$\begin{aligned} W_T^2 - 0 &= \int_0^T 2W_t dW_t + \int_0^T dt = \int_0^T 2W_t dW_t + T \\ \Rightarrow \int_0^T W_t dW_t &= \frac{1}{2} (W_T^2 - T) \end{aligned}$$

Q: calculate  $[W^2, W^2]_T$

$$\begin{aligned} [W^2, W^2]_T &= \int_0^T (2W_t)^2 dt = 4 \int_0^T W_t^2 dt \\ &\quad \begin{array}{l} \rightarrow \text{not constant,} \\ \text{a r.v.} \end{array} \end{aligned}$$

ex 2. Let  $M_t = W_t$ ,  $N_t = W_t^2 - t$ , is  $M, N$  a M.G.?

first method is to verify  $E[M_t \cdot N_t | \mathcal{F}_s] = M_s \cdot N_s \quad \forall t \geq s$

another method is to use the Itô formula.

$$\text{Note } M_t \cdot N_t = W_t(W_t^2 - t) = W_t^3 - W_t \cdot t$$

$$f(t, X) = X^3 - tX \quad X_t = W_t \quad [X, X]_t = t$$



$$\Rightarrow \partial_x f(t, x) = 3x^2 - t \quad \partial_t f(t, x) = -x \quad \partial_x^2 f(t, x) = 6x$$

$$\begin{aligned} \Rightarrow df(t, x) &= \partial_t f(t, x_t) dt + \partial_x f(t, x_t) dx_t \\ &\quad + \frac{1}{2} \partial_x^2 f(t, x_t) d[x, x]_t \\ &= -x_t dt + (3x_t^2 - t) dx_t + 3x_t dt \\ &= 2x_t dt + (3x_t^2 - t) dx_t \\ &= 2W_t dt + (3W_t^2 - t) dW_t \end{aligned}$$

$$M_T N_T = \int_0^T 2W_t dt + \int_0^T (3W_t^2 - t) dW_t$$

Since the finite (first) Variation part is not zero,  
 $M_T N_T$  is not a martingale.

ex3. Let  $X_t = t \sin(W_t)$ . Is  $X^2 - [X, X]$  a martingale?

$$\text{let } f(t, x) = t \sin(x) \quad \partial_t f = \sin x \quad \partial_x f = t \cos x$$

$$\partial_x^2 f = -t \sin x$$

$$\begin{aligned} \Rightarrow dX_t &= \sin W_t dt + t \cos W_t dW_t - \frac{1}{2} t \sin W_t dt \\ &= \left( \sin W_t - \frac{1}{2} t \sin W_t \right) dt + t \cos W_t dW_t \end{aligned}$$

$$(\text{then, } d[X, X] = t^2 \cos^2 W_t dt)$$

$$dX_t^2 = 2X_t dX_t + d[X, X]_t$$

$$\begin{aligned} \Rightarrow d[X^2 - [X, X]] &= 2X_t dX_t \\ &= 2t \sin(W_t) \cdot \left[ \sin(W_t) - \frac{1}{2} t \sin(W_t) \right] dt \\ &\quad + 2t^2 \sin(W_t) \cos(W_t) dW_t \end{aligned}$$

Since the  $dt$  term above is not 0, thus  $X^2 - [X, X]$  is not a martingale

## 4.6.1 Multidimensional Itô

### Def 4.6.1 (quadratic covariation)

Let  $X, Y$  be two Itô processes. Define

$$[X, Y]_T = \lim_{\|\pi\| \rightarrow 0} \sum_{i=0}^{n-1} (X_{t_{i+1}} - X_{t_i})(Y_{t_{i+1}} - Y_{t_i})$$

$$\pi = \{0 = t_0 < \dots < t_n = T\}$$

the quadratic covariation

lemma:  $\forall a, b \in \mathbb{R} \quad ab = \frac{1}{4}[(a+b)^2 - (a-b)^2]$

$$\Rightarrow \text{let } a = (X_{t_{i+1}} - X_{t_i}) \quad b = (Y_{t_{i+1}} - Y_{t_i})$$

$$\text{thus } [X, Y]_T = \frac{1}{4} [ [X+Y, X+Y]_T - [X-Y, X-Y]_T ]$$

Prop 4.6.1 (Product Rule) For Itô processes  $X, Y$ , we have

$$d(XY) = XdY + YdX + d[X, Y]$$

Prop 4.6.2 If  $X$  is an Itô process and  $A$  is adapted

process of finite variation, then  $[X, A] = 0$

(Note that  $[X \pm A, X \pm A] = [X, X]$ )

Prop 4.6.4. If  $X, Y, Z$  are Itô processes and  $\alpha \in \mathbb{R}$ , then  
(Bi-linearity)  $[X, Y + \alpha Z] = [X, Y] + \alpha [X, Z]$

Prop 4.6.6 Let  $X, Y$  be two continuous martingales (e.g. Itô processes) w.r.t. a common filtration  $(\mathcal{F}_t)_{t \geq 0}$  such that  $E[X_t^2] < \infty$  and  $E[Y_t^2] < \infty$ , if  $X, Y$  are independent, then  $[X, Y] = 0$

the converse is not true, for example  $X_t = \int_0^t 1_{\{ws > 0\}} dW_s$

$Y_t = \int_0^t 1_{\{ws < 0\}} dW_s$ ,  $[X, Y] = 0$  but  $X \not\perp Y$

### Theorem 4.6.1 Multidimensional Itô process

Let  $X^1, X^2, \dots, X^n$  be Itô processes and  $X = (X^1, X^2, \dots, X^n)$

Let  $f: [0, +\infty) \times \mathbb{R}^n \rightarrow \mathbb{R} \quad (t, x) \mapsto f(t, x)$

be  $C^1$  in  $t$  ( $\partial_t f$  exists)

be  $C^2$  in  $X^i$  ( $\partial_{X^i} f = \partial_i f$ ,  $\partial_{X^i X^j} f = \partial_i \partial_j f$  exist)

Then

$$f(T, X_T) = f(0, X_0) + \int_0^T \partial_t f(t, X_t) dt + \sum_{i=1}^n \int_0^T \partial_i f(t, X_t) dX_t^i \\ + \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n \int_0^T \partial_i \partial_j f(t, X_t) d[X^i, X^j]_t$$

or

$$df(t, X_t) = \partial_t f(t, X_t) dt + \sum_{i=1}^n \partial_i f(t, X_t) dX_t^i \\ + \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n \partial_i \partial_j f(t, X_t) d[X^i, X^j]_t$$

Remark: we most often use the two dimensions case

$$df(t, X_t, Y_t) = \partial_t f dt + \partial_x f dX_t + \partial_y f dY_t \\ + \frac{1}{2} [\partial_x^2 f d[X, X]_t + 2 \partial_x \partial_y f d[X, Y]_t + \partial_y^2 f d[Y, Y]_t]$$

Prop 4.6.3. Let  $M, N$  be Martingales. Then

(1)  $M \cdot N - [M, N]$  is a MG

(when  $M=N$ ,  $M^2 - [M, M]$  is a MG, which is mentioned before)

(2) If  $A$  is adapted process of finite variation such that  $A_0=0$  and  $MN - A$  is a MG, then  $A = [M, N]$

Example:  $M=W$ ,  $N=-W$   $[W, -W]_t = -[W, W]_t = -t$   
(Typically,  $M, N$  are MG  $\nRightarrow M \cdot N$  is MG, eg  $M=N=W$ )

Prop 4.6.5. let  $X^1, X^2$  be Itô processes,  $\sigma^1, \sigma^2$  be adapted processes and  $I_t^j = \int_0^t \sigma_s^j dx_s^j \quad j=1,2$ .

$$\text{Then } [I^1, I^2]_t = \int_0^t \sigma_s^1 \sigma_s^2 d[X^1, X^2]_s$$

Def 4.6.2. We say that  $W = (W^1, W^2, \dots, W^n)$  is a  $n$ -dimensional standard BM if

- ① each  $W^j$  is a standard BM  $\forall j=1, \dots, n$
- ②  $\forall i \neq j \quad W^i$  and  $W^j$  are independent.

Example.

o for a 2-dimensional SBM  $W = (W^1, W^2)$

$$[W^1, W^2]_t = 0 \quad [W^1, W^1]_t = [W^2, W^2]_t = t$$

$$df(W^1, W^2) = \partial_1 f(W^1, W^2) dW^1 + \partial_2 f(W^1, W^2) dW^2 + \frac{1}{2} [\partial_1^2 f(W^1, W^2) dt + \partial_2^2 f(W^1, W^2) dt]$$

Theorem (Lévy)

Assume that  $M$  is a MG,  $M_0 = 0$  and

$$d[M^i, M^j]_t = \begin{cases} dt & i=j \\ 0 & i \neq j \end{cases}$$

Then  $M$  is  $n$ -dimensional Brownian motion

Remark:

$$\underbrace{\text{Cov}(X_t, Y_t)}_{\text{a scalar}} \neq \underbrace{[X, Y]_t}_{\text{a random variable}}$$