6. Probabilistic models

References

C. Bishop. Pattern Recognition and Machine Learning. Sect. 4.2, 4.3

Approaches that allows to estimate probability that given an instance we have a certain classes.

Two families of models:

Generative: estimate P(x|Ci) and then compute P(Ci|x) with Bayes

Discriminative: estimate P(Ci | x) directly

6.1 Probabilistic generative models

$$P(C_1|\mathbf{x}) = \frac{P(\mathbf{x}|C_1)P(C_1)}{P(\mathbf{x})} = \frac{P(\mathbf{x}|C_1)P(C_1)}{P(\mathbf{x}|C_1)P(C_1) + P(\mathbf{x}|C_2)P(C_2)}$$
$$= \frac{1}{1 + \exp(-a)} = \sigma(a)$$

with:

$$a = \ln \frac{p(x|C_1)P(C_1)}{p(x|C_2)P(C_2)}$$

and

$$\sigma(a) = \frac{1}{1 + \exp(-a)}$$
 the sigmoid function.

Multi-class

$$P(C_k|\mathbf{x}) = \frac{P(\mathbf{x}|C_k)P(C_k)}{\sum_j P(\mathbf{x}|C_j)P(C_j)} = \frac{\exp(a_k)}{\sum_j \exp(a_j)}$$

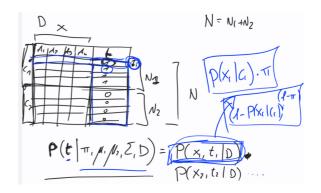
(normalized exponential or softmax function)

with
$$a_k = \ln P(\mathbf{x}|C_k)P(C_k)$$

6.1.1 Maximum likelihood

Likelihood function

$$P(\mathbf{t}|\pi,\mu_1,\mu_2,\mathbf{\Sigma},D) = \prod_{n=1}^{N} [\pi \mathcal{N}(\mathbf{x}_n;\mu_1,\mathbf{\Sigma})]^{\mathbf{Z}_n} [(1-\pi)\mathcal{N}(\mathbf{x}_n;\mu_2,\mathbf{\Sigma})]^{(1-t_n)}$$



For 2 classes, we obtain

$$\pi = \frac{N_1}{N_1} \sum_{n=1}^{N} t_n \mathbf{x}_n \qquad \mu_2 = \frac{1}{N_2} \sum_{n=1}^{N} (1 - t_n) \mathbf{x}_n$$

$$\mathbf{\Sigma} = \frac{N_1}{N} S_1 + \frac{N_2}{N} S_2$$
 with $S_i = \frac{1}{N_i} \sum_{n \in C_i} (\mathbf{x}_n - \mu_i) (\mathbf{x}_n - \mu_i)^T$, $i = 1, 2$

Posterior distributions with parametric models:

For two classes

$$P(C_1|\mathbf{x}) = \sigma(a)$$

For $k \ge 2$ classes

$$P(C_i|\mathbf{x}) = \frac{exp(a_k)}{\sum_j exp(a_j)}$$

$$a_k = \mathbf{w}^T \mathbf{x} + w_0$$

$$\mathbf{w}^{\mathsf{T}}\mathbf{x} + w_0 = (w_0 \ \mathbf{w}) \begin{pmatrix} 1 \\ \mathbf{x} \end{pmatrix}$$

$$ilde{\mathbf{w}} = \left(egin{array}{c} w_0 \\ \mathbf{w} \end{array}
ight), ilde{\mathbf{x}} = \left(egin{array}{c} 1 \\ \mathbf{x} \end{array}
ight)$$

$$\mathbf{a}_k = \mathbf{w}^T \mathbf{x} + w_0 = \tilde{\mathbf{w}}^T \tilde{\mathbf{x}}$$

Likelihood for a parametric model \mathcal{M}_{Θ} : $P(\mathbf{t}|\Theta, D)$, $D = \langle \mathbf{X}, \mathbf{t} \rangle$

Maximum likelihood solution:

$$\boldsymbol{\Theta}^* = \operatorname*{argmax}_{\boldsymbol{\Theta}} \ln P(\mathbf{t}|\boldsymbol{\Theta}, \mathbf{X})$$

When \mathcal{M}_{Θ} belongs to the exponential family, likelihood $P(\mathbf{t}|\Theta,\mathbf{X})$ can be expressed in the form $P(\mathbf{t}|\tilde{\mathbf{w}},\mathbf{X})$, with maximum likelihood

$$\tilde{\boldsymbol{w}}^* = \operatorname*{argmax}_{\tilde{\boldsymbol{w}}} \ln P(\boldsymbol{t}|\tilde{\boldsymbol{w}}, \boldsymbol{X})$$

6.2 Probabilistic discriminative models

Estimate dirrectly

$$P(C_i|\tilde{\mathbf{x}},D) = \frac{exp(a_k)}{\sum_j exp(a_j)}$$
 $a_k = \tilde{\mathbf{w}}^T \tilde{\mathbf{x}}$

with maximum likelihood

$$\tilde{\mathbf{w}}^* = \underset{\tilde{\mathbf{w}}}{\operatorname{argmax}} \ln P(\mathbf{t}|\tilde{\mathbf{w}}, \mathbf{X})$$

without estimating the model parameters.

Simplified notation (dataset omitted): $P(\mathbf{t}|\tilde{\mathbf{w}})$

6.2.1 Logistic regression

For 2 classes, likelihood function:

Note:tn: value in the data set corresponding toxn,yn: posterior prediction of the current model wforxn.

$$p(\mathbf{t}|\mathbf{w}) = \prod_{n=1}^{N} y_n^{t_n} (1 - y_n)^{1 - t_n}$$
 with $y_n = p(C_1|\mathbf{x}_n) = \sigma(\mathbf{w}^T \mathbf{x}_n)$
$$E(\mathbf{w}) \equiv -\ln p(\mathbf{t}|\mathbf{w}) = -\sum_{n=1}^{N} [t_n \ln y_n + (1 - t_n) \ln(1 - y_n)]$$

Solution concept: solve the optimization problem

$$\mathbf{w}^* = \operatorname*{argmin}_{\mathbf{w}} E(\mathbf{w})$$

Cross-entropy error function:

To minimize error we apply Newton-Raphson:

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$$\nabla E(\mathbf{w}) = \sum_{n=1}^{N} (y_n - t_n) \mathbf{x}_n$$

Gradient descent step

$$\mathbf{w} \leftarrow \mathbf{w} - \mathbf{H}^{-1} \nabla E(\mathbf{w})$$

 $\mathbf{H} = \nabla \nabla E(\mathbf{w})$ is the Hessian matrix of $E(\mathbf{w})$ (second derivatives with respect to \mathbf{w}).

Given $\tilde{\mathbf{X}} = \begin{pmatrix} \tilde{\mathbf{x}}_1^T \\ \dots \\ \tilde{\mathbf{x}}_N^T \end{pmatrix} \quad \mathbf{t} = \begin{pmatrix} t_1 \\ \dots \\ t_N \end{pmatrix}$,

 $\mathbf{y}(\tilde{\mathbf{w}}) = (y_1, \dots, y_n)^T$ posterior predictions of model $\tilde{\mathbf{w}}$

 $oldsymbol{R}(ilde{oldsymbol{w}})$: diagonal matrix with $R_{nn}=y_n(1-y_n)$

we have

$$abla E(\tilde{\mathbf{w}}) = \tilde{\mathbf{X}}^T (\mathbf{y}(\tilde{\mathbf{w}}) - \mathbf{t})$$

$$\mathbf{H}(\tilde{\mathbf{w}}) = \nabla \nabla E(\tilde{\mathbf{w}}) = \sum_{n=1}^{N} y_n (1 - y_n) \tilde{\mathbf{x}}_n \tilde{\mathbf{x}}_n^T = \tilde{\mathbf{X}}^T \mathbf{R}(\tilde{\mathbf{w}}) \tilde{\mathbf{X}}$$

Iterative method:

- Initialize w
- 2. Repeat until termination condition

$$\mathbf{w} \leftarrow \mathbf{w} - (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{y} - \mathbf{t})$$

Multiclass:

K classes

$$P(C_k|\tilde{\mathbf{x}}) = \frac{exp(a_k)}{\sum_j exp(a_j)}$$
 $a_k = \tilde{\mathbf{w}}_k^T \tilde{\mathbf{x}}$ $k = 1, ..., K$

$$\tilde{\mathbf{X}} = \begin{pmatrix} \tilde{\mathbf{x}}_1^T \\ \dots \\ \tilde{\mathbf{x}}_N^T \end{pmatrix}$$
 $\mathbf{T} = \begin{pmatrix} t_1^T \\ \dots \\ t_N^T \end{pmatrix}$ 1-of- K encoding of labels

 $\mathbf{y}_n^T = (y_{n1} \dots y_{nK})^T$ posterior prediction of $\tilde{\mathbf{x}}_n$ for model $\tilde{\mathbf{w}}_1, \dots, \tilde{\mathbf{w}}_K$

$$\mathbf{Y}(\tilde{\mathbf{w}}_1,\ldots,\tilde{\mathbf{w}}_K) = \begin{pmatrix} \mathbf{y}_1^T \\ \ldots \\ \mathbf{y}_N^T \end{pmatrix}$$
 posterior predictions of model $\tilde{\mathbf{w}}_1,\ldots,\tilde{\mathbf{w}}_K$

Discriminative model

$$P(\mathbf{T}|\tilde{\mathbf{w}}_1,\ldots,\tilde{\mathbf{w}}_K) = \prod_{n=1}^N \prod_{k=1}^K P(C_k|\tilde{\mathbf{x}}_n)^{t_{nk}} = \prod_{n=1}^N \prod_{k=1}^K y_{nk}^{t_{nk}}$$

with
$$y_{nk} = \mathbf{Y}[n, k]$$
 and $t_{nk} = \mathbf{T}[n, k]$.

Cross-entropy error function

$$E(\tilde{\mathbf{w}}_1, \dots \tilde{\mathbf{w}}_K) = -\ln P(\mathbf{T}|\tilde{\mathbf{w}}_1, \dots \tilde{\mathbf{w}}_K) = -\sum_{n=1}^N \sum_{k=1}^K t_{nk} \ln y_{nk}$$

Iterative algorithm

gradient
$$abla_{ ilde{oldsymbol{w}}_j} E(ilde{oldsymbol{w}}_1, \dots ilde{oldsymbol{w}}_{\mathcal{K}}) = \dots$$

Hessian matrix
$$\nabla_{\tilde{\mathbf{w}}_k} \nabla_{\tilde{\mathbf{w}}_j} E(\tilde{\mathbf{w}}_1, \dots \tilde{\mathbf{w}}_K) = \dots$$

SUMMARY

Given a target function $f: X \to C$, and data set D

assume a parametric model for the posterior probability $P(C_k | \tilde{\mathbf{x}}, \tilde{\mathbf{w}})$ $\sigma(\tilde{\mathbf{w}}^T \tilde{\mathbf{x}})$ (2 classes) or $\frac{exp(\tilde{\mathbf{w}}_k^T \tilde{\mathbf{x}})}{\sum_{j=1}^K exp(\tilde{\mathbf{w}}_j^T \tilde{\mathbf{x}})}$ (k classes)

Define an error function $E(\tilde{\mathbf{w}})$ (negative log likelihood)

Solve the optimization problem

$$\tilde{\boldsymbol{w}}^* = \operatorname*{argmin}_{\tilde{\boldsymbol{w}}} E(\tilde{\boldsymbol{w}})$$

Classify new sample $\tilde{\mathbf{x}}'$ as C_{k^*} where $k^* = \operatorname{argmax}_{k=1,\dots,K} P(C_k | \tilde{\mathbf{x}}', \tilde{\mathbf{w}}^*)$

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