# **Day 3 Frequentist statistics**

- 3.1 Interval estimation
- 3.2 Hypothesis testing
- 3.3 Some common tests
- 3.4 Summary and additional questions

#### Recap of last week...

- $\bullet$  Common distributions: Bernoulli, Binomial, Poisson, Normal, Gamma and  $\chi^2$
- Characteristic functions  $\phi_X(t) = \langle e^{itX} \rangle$
- Statistics of a sample  $X_1, X_2, \cdots, X_n$  from a population with  $\mu$  and  $\sigma^2$

sample mean 
$$\overline{X}=\frac{1}{n}\sum_{k=1}^n X_n$$
 where  $\langle \overline{X} \rangle = \mu$  sample variance  $s_u^2=\frac{1}{n-1}\sum_{k=1}^n (X_k-\overline{X})^2$  where  $\langle s_u^2 \rangle = \sigma^2$ 

- ullet Law of large numbers  $\overline{X}_n o \mu$  for many samples
- Central limit theorem: distribution of  $\overline{X}_n$  tends to a normal with mean  $\mu$  and variance  $\sigma^2/n$ . NB  $\sigma/n^{1/2}$  is the **standard error on mean**.

#### 3.1 Interval estimation

- 3.1.1 Analysis of house-price data
- 3.1.2 Confidence intervals
- 3.1.3 Question on confidence intervals

#### 3.1.1 Analysis of house-price data

- UK House price data. 105,238 properties less than or equal to £2M.
- Population mean=£290.6k and std=£229.2k so highly skewed distribution.
- Imagine we draw a sample of n=1000 which is  $\sim 1\%$  of the data.
- For a particular sample we get point esimations: sample mean  $\mu_s = £281.0$ k, sample std=£217.0k and sem  $\sigma_{sem} = £6.9$ k
- How can we estimate a range for the population mean?
- Let the 95% confidence interval be the range encompassing the central 95% of the normal distribution with  $\mu_{s}$ ,  $\sigma_{s}$
- **Question:** Use invlogcdf in the Distributions package to find this range.

#### 3.1.2 Confidence intervals

- NB Confidence intervals are random quantities depend on a sample.
- Frequently misunderstood! For example, a 95% confidence interval...
- Does **not** mean a 95% chance the population mean is in that range.
- It means that 95% of confidence intervals will include the population mean.

#### 3.1.3 Question: confidence intervals

- 100 samples from a chocolate-bar machine had weights with mean  $\mu_s=101{\rm g}$  and standard deviation  $\sigma_s$ =5g.
- $\bullet$  Give 99% confidence limits for the mean weight of all bars made.

## 3.2 Hypothesis testing

- 3.2.1 Type I and Type II errors
- 3.2.2 Significance and p-values
- 3.2.3 Type II errors and operator curves

#### 3.2.1 Type I and Type II errors

- ullet Want to distinguish between a null hypothesis  $H_0$  and other hypothesis  $H_1$  etc .
- Frequentist statistics: concerned with decisions made about the null hypothesis.
- Four possibilities: null can be true or false and we can accept or reject it.

	true	false	
accept	success	type-II error	
reject	type-I error	success	

- Type I error: probability reject a true null hypothesis  $P(\text{reject} \mid \text{null true})$ Use  $\alpha$  for type-I error "rate".
- Type II error: we accept a false null hypothesis  $P(\text{accept} \mid \text{null false})$  Use  $\beta$  for type-II error "rate".
- Type I errors tend to be main focus.
   Related to falsifiability of an existing theory.
   Fits with Popper's idea of good theories being falsifiable.

### 3.2.1 Type I and II error example

- Imagine there is a new blood test for stroke.
- Stroke or not is then confirmed clinically, later.
- 100 people suspected of having a stroke are tested.
- Hypothesis: patient has had a stroke

	true	false
accept	88	4
reject	2	6

• Question: What is the type-I error rate?

• Question: What is the type-II error rate?

#### 3.2.2 Significance and p-values

**Example** Test if a coin is fair - this is the **null** hypothesis.

- We imagine doing one experiment with *n* flips.
- If coin not fair anticipate result will look unlikely from null statistics.
- We fix ahead a probability level  $\alpha$  the **significance** type I error rate.

This the level at which we reject the null because result unlikely if null true.

Typically this is set at 5% or 1%.

• Do an experiment and measure the p-value. The probability

#### **Experiment**

Coin flip gives 7 heads out of 8 flips. What is the p-value? Is it significant at 5%?

heads	0	1	2	3	4	5	6	7	8
prob	0.004	0.0312	0.1094	0.2188	0.2734	0.2188	0.1094	0.0312	0.0

- As extreme is 0,1 7 or 8 heads so p-value is  $2 \times (0.0039 + 0.0312) = 7\%$ .
- Call this **Not sigificant** as greater that then  $\alpha = 5\%$  cut-off.
- This is a two sided test.
- Imagine if we were testing if the coin was biased towards heads.

As extreme are 7 or 8, so p-value= 3.5% which is significant at the 5% level.

#### 3.2.2 Question Significance and p-values

- A nationwide school test has mean mark  $\mu = 75$  with standard dev.  $\sigma = 7$ .
- A particular school with 30 children has a mean mark  $\mu_{s}=72$  .
- Question. Is the school significantly different from the national average?
   Give the p-value and compare to 5%.
- Hint 1: use cdf(Normal()) to calculate the area under the normal curve.
- Hint 2: you can use invlogcdf to go fro the cdf to x, when necessary.

#### 3.2.3 Type II errors and operator curves

- Decide coin fair: 100 flips there are between 40-60 heads (inclusive).
- The type I error rate is therefore  $\alpha = 0.0352$ .
- Imagine p = 0.6Question: what is the type II error  $\beta$ ?

#### 3.3 Some common tests

- 3.3.1 One-sample test
- 3.3.2 Two-sample test
- 3.3.3 Student's t test
- 3.3.4 Chi-square test

#### 3.3.1 One sample test

- Consider the sample mean  $\overline{X}$  of a set  $\{X_k\}$  of n random numbers.
- Assume *n* sufficiently large that the sample-mean distribution is normal.
- Get an unbiased estimate of population  $\sigma$  from data (n should be big for this).
- Null hypothesis is that the sample mean is  $\mu_0$  (often 0, as can subtract this out).
- This is called a **z-test**: examine z-statistic  $z=(\overline{X}-\mu_0)/(\sigma/\sqrt{n})$  for significance

```
In [7]:
        n=10
        X=randn(n) .+0.1;
        mx=mean(X); sx=std(X); semx=sx/sqrt(n);
        zscore=mx/semx; pvalue=2*cdf.(Normal(),-abs(zscore))
        println("X=$(round.(X;digits=2))")
        println("mean(X)=$(round(mx;digits=3)) and sem(X)=$(round(semx;di
        gits=3))")
        println("z-score=$(round(zscore;digits=3)) and pvalue=$(round(pva
        lue;digits=3))")
        println("-----")
        OneSampleZTest(X)
        X=[-1.29, -1.46, 0.26, 1.27, -1.12, -0.13, 0.58, 0.53, -0.71,
        1.7]
        mean(X) = -0.037 and sem(X) = 0.345
        z-score=-0.108 and pvalue=0.914
Out[7]: One sample z-test
        _____
        Population details:
            parameter of interest: Mean
            value under h 0:
            point estimate:
                                    -0.037256131373281855
            95% confidence interval: (-0.7143, 0.6398)
        Test summary:
            outcome with 95% confidence: fail to reject h_0
                                        0.9141
            two-sided p-value:
        Details:
            number of observations: 10
            z-statistic:
                                      -0.10784820875555436
            population standard error: 0.34544970012182147
```

#### 3.3.2 Two-sample test

- Now consider we have two sets of number  $\{X_k\}$  and  $\{Y_k\}$
- There are  $n_x$  and  $n_y$  of these, the sample means are  $\overline{X}$  and  $\overline{Y}$  and the population variance esimators  $\sigma_x$  and  $\sigma_y$ .
- $\bullet$  Are the sample means the same or different? Consider statistics of  $\overline{X} \overline{Y}$
- The variance of the difference is  $\sigma_x^2/n_x + \sigma_y^2/n_y$ .
- The test statistic is therefore  $z=\frac{(\overline{X}-\overline{Y})}{\sqrt{\sigma_x^2/n_x+\sigma_y^2/n_y}}$  which we test for significance.

```
In [8]:
        nx=10; X=randn(nx) .+0.1;
        ny=15; Y=1.2*randn(ny) .+0.2;
        UnequalVarianceZTest(X,Y)
Out[8]:
        Two sample z-test (unequal variance)
         Population details:
             parameter of interest: Mean difference
             value under h_0:
            point estimate: -0.052270158359419505
             95% confidence interval: (-0.7952, 0.6907)
         Test summary:
             outcome with 95% confidence: fail to reject h_0
             two-sided p-value:
                                         0.8903
         Details:
             number of observations: [10,15]
             z-statistic:
                                      -0.13789330589472157
             population standard error: 0.3790623338838986
```

#### 3.3.3 Student's t test

- Developed for small sample where typically  $n \leq 30$  is quoted.
- The z-score comes from a standardised normal  $z = (\bar{X} \mu)/(\sigma/\sqrt{n})$
- Previously assumed that esimate  $s_u^2$  for population  $\sigma^2$  is sharp
- But  $s_u^2$  is a random number too increasing uncertainty distribution  $t = (\bar{X} \mu)/(s_u/\sqrt{n})$  is **not** normal
- Follows t-distribution (if samples normal).  $Y = Y_0(1 + t^2/(n-1))^{-n/2}$ . similar to normal, but tails fatter
- t-tests are preformed exactly like z-tests.

```
In [11]: # example
    N,n=100000,6
    M=randn(N,n);
    Sm=mean(M,dims=2)
    Sv=var(M,dims=2)
    t=Sm./sqrt.(Sv/n);

figure(figsize=(5,3))
    x=-5:0.1:5
    y1=pdf.(Normal(),x)
    y2=pdf.(TDist(n-1),x)
    plt[:hist](t,400,normed=5,color="lightgreen");
    plot(x,y1,"k:",x,y2,"k-"); xlabel("t")
    axis([-6,6,0,0.45]); title("t-distribution with $(n-1) degrees of freedom")
```

### 

Out[11]: PyObject Text(0.5, 1.0, 't-distribution with 5 degrees of free dom')

#### 3.3.4 Chi-square test

- This test is used for seeing if a distribution of numbers is as expected.
- Let  $a_k$  where  $k=1\cdots n$  be the measured frequency and  $b_k$  that expected.
- The test statistics is  $\chi^2 = \sum_{k=1}^n \frac{(a_k b_k)^2}{b_k}$ .
- It has a sampling distribution is  $Y(\chi^2) = \frac{1}{2^{\nu/2}\Gamma(\nu/2)}(\chi^2)^{(\nu-2)/2}e^{-\chi^2/2}$  where  $\nu=n-1$ . This is a  $\chi^2$  distribution with n-1 degrees of freedom.
- Question: A six-sided die is rolled 120 times with frequencies:

-	1	2	3	4	5	6
-	17	14	19	15	21	34

 $\bullet$  Is the die a fair one? Test at 95% and 99% and give the p-value.

# 3.4 Summary and additional questions

#### Day 3 Frequentist approach

- 3.1 Interval estimation
- 3.2 Hypothesis testing
- 3.3 Additional topics
- 3.4 Summary and homework questions

#### **Questions**

Make sure you have understood and done all the questions in the lectures.

The questions below are to be handed in for marking by 10am on Monday 26th November 2018

**NB** there will also be questions from the Day 4 lectures to be handed in on the 26th.

- Q3.1 Spotting fake financial data
- Q3.2 One-sided and two-sided test differences
- Q3.3 Limit of Student's t-distribution.

#### Q3.1 Spotting fake data

- The first digits of many real-world data sets do not follow a uniform distribution.
- This observation is called Benford's law, where  $P(d) = \log_{10}(1 + 1/d)$  for  $d = 1 \cdots 9$
- When data is faked a uniform random number generator is often used.
- Download the following csv file from the Nasdaq archive: www.nasdaq.com/screening/companies-byindustry.aspx?exchange=NASDAQ&render=download
- The third column is the price of the stock. The data has some entries that are "n/a" which you need to clean up.
- Part (a) For a random sample of 90 stocks, plot the distribution of first digits of stock prices.
- Part (b) Use the Chi-square test to check if the distribution is significantly different from a uniform distribution and from Benford's law. Give the p-value for these two tests.
- Part (c) Increase the sample to 900. Comment on what happens to the significance test in respect to Benford's law.

# Q3.2 One-sided and two-sided test differences.

- Refer back to the question in section 3.2.2 on the nationwide school test.
- A nationwide school test has mean mark  $\mu = 75$  with standard dev.  $\sigma = 7$ .
- $\bullet$  A particular school with 30 children has a mean mark  $\overline{X}=72$  .
- Test the data for the question:
   Is the school underpreforming?
   What is the p-value and compare it to 5% or 1% significances.

#### Q3.3 Limit of Student's t-distribution

• **Part (a)** Demontrate that in the limit of large *n* Student's t-distribution tends to a standard normal. The distribution is

$$f(t) = \frac{1}{\sqrt{\nu\pi}} \frac{\Gamma(\frac{\nu+1}{2})}{\Gamma(\frac{\nu}{2})} \frac{1}{(1+t^2/\nu)^{(\nu+1)/2}}$$

- Where  $\nu = n 1$  is the number of degrees of freedom.
- Part(b) Show that in this limit the leading order correction to the variance is  $\sigma^2 \simeq 1 + 2/\nu$  and therefore larger than that of the standard normal.