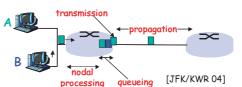
# A few Slides on Foundations of Performance Analysis of Data Networks



The intention here is to consolidate some useful results (not to teach the underlying material).

#### Slide 1

#### **Outline**

- 1. Overview of random variables
- 2. Some useful discrete r.v. distributions
- 3. Some useful continuous r.v. distributions
- 4. Little's Theorem
- 5. The M/M/1 Queue

## 1. Overview of random variables

- 1. RVs arise in 3 possible ways:
  - direct observation of a random experiment
  - $\blacksquare$  mapping from a sample space to real numbers (e.g., HHT  $\rightarrow$  2)
  - functions of other random variables

#### 2. Examples:

- $\mathbf{X} = \mathbf{X} = \mathbf{X}$ 
  - ullet  $S_X$  denotes the range of the r.v. X
  - $S_X = \{0, 1, \ldots\}$
- lacksquare Y= time delay seen by a packet while waiting in a queue
  - $S_Y \in [0, \infty]$
- $Y = (X \mu_X)^2$ , here Y is derived from X (in general, we write Y = g(X), for some function g())
- 3. We describe a discrete r.v. X by a probability mass function (PMF)  $P_X(x)$ ,

defined as:

$$P_X(x) = Prob[X = x]$$

e.g.,

$$P_X(x) = \begin{cases} 1/4 & x = 0 \\ 3/4 & x = 1 \\ 0 & \text{otherwise} \end{cases}$$

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4.  $P_X(.)$  satisfies

For every event  $B \subseteq S_X$ :  $Prob[B] = \sum_{x \in B} P_X(x)$ 

5. X can also be described by the *Cumulative Distribution Function* (CDF)  $F_X(x)$ , defined as:

$$F_X(x) = Prob[X \le x]$$

6. A continuous r.v. X can be described by its CDF  $F_X(x)$ , or its *probability* density function (PDF), denoted  $f_X(x)$ :

Prob
$$[x_1 < X \le x_2] = \int_{x_1}^{x_2} f_X(u) du$$

$$F_X(x) = \int_{-\infty}^x f_X(u) du$$

7. Mean (denoted E[X] or  $\mu_X$ ):

discrete:  $E[X] = \sum_{x \in S_X} x P_X(x)$ 

continuous:  $E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$ 

8. Variance:  $Var[X] = E[(X - \mu_X)^2] = E[X^2] - \mu_X^2$ 

standard deviation:  $\sigma_X = \sqrt{Var[X]}$ 

## 2. Some Useful Discrete Distributions

**2.1** Discrete Uniform  $[k, k+1, \ldots, l]$ 

lacksquare X is equally likely to take one the possible values

PMF:

$$P_X(x) = \begin{cases} \frac{1}{l-k+1} & x = k, k+1, k+2, \dots, l \\ 0 & \text{otherwise} \end{cases}$$

- $E[X] = \frac{k+l}{2}$
- $Var[X] = \frac{(l-k)(l-k+2)}{12}$

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#### 2.2 Bernoulli

lacksquare X is binary (success/fail); so  $S_X = \{0,1\}$ 

$$P_X(x) = \begin{cases} p & x = 1\\ 1 - p & x = 0\\ 0 & \text{otherwise} \end{cases}$$

- E[X] = 1.p + 0.(1 p) = p
- Var[X] = p(1-p)

## 2.3 Geometric

- Perform a sequence of independent Bernoulli (success/fail) trials until we get the first success.
- lack X= the number of independent trials performed until we get the first success. (So,  $S_X=\{1,2,3,\ldots\}$ )

$$P_X(x) = \begin{cases} p(1-p)^{x-1} & x = 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

- $E[X] = \frac{1}{p}$  (intuition?)

## 2.4 Binomial

Perform a sequence of  $n, n \ge 1$ , independent Bernoulli (success/fail) trials,

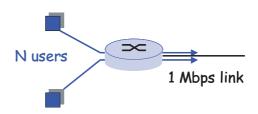
X = the number of successes in n trials

PMF:

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$$P_X(x) = \begin{cases} \binom{n}{x} p^x (1-p)^{n-x} & x = 0, 1, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

- E[X] = np
- Var[X] = np(1-p)
- Application: justifying statistical multiplexing

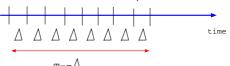


- Assume: 1 Mbps link, N=35 users, each user sends at 100 kbps when active, each user is active 10% of time
- Circuit-switching: can accommodate at most 10 users
- Packet-switching: What is the prob. that the aggregate traffic demand exceeds the link capacity?
  - \* Model user activity during a short period as a Bernoulli trial with Prob[active]= p = 0.1 (q = 1 - p = 0.9)
  - $\ast\,$  So, Prob [# of active users during a short period  $\geq 11$ ]  $=\sum_{i=11}^{\infty} \binom{N}{i} p^i q^{N-i}$

$$=\sum_{i=11}^{i=11}$$
 (= .0004

# 2.5 Poisson (with parameter $\alpha$ , where $\alpha = \lambda T$ )

- First, let's introduce the Poisson process :
  - events: arrivals, departures, births, deaths, hits to a WEB server, etc.
  - event rate:  $\lambda$  events/sec
  - view time as slotted: each slot is  $\Delta$  sec. ( $\Delta$  is infinitesimal)



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- Prob[exactly n events in one  $\Delta$ ] =  $\begin{cases} \lambda\Delta & n=1\\ 1-\lambda\Delta & n=0\\ 0 & \text{otherwise} \end{cases}$
- ullet So, the number of possible events in one  $\Delta$  is a Bernoulli r.v.
- Number of events in non-overlapping intervals are independent r.v.s
- Now, suppose we take an interval of T sec, and let

X = number of events in T secs

PMF: let  $\alpha = \lambda T$  (average number of events in T)

$$P_X(x) = \begin{cases} \frac{\alpha^x e^{-\alpha}}{x!} & x = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

- $E[X] = \alpha$
- $Var[X] = \alpha$
- **Example.** Suppose the number of hits to a WEB server in a 10-sec interval is a Poisson r.v. K, with  $\alpha=5$  hits.
  - 1. Find Prob[no hit in 10 sec].
  - 2. Find Prob [at least 2 hits occur in 2 seconds].

Solution.

1. 
$$P_K(k) = \begin{cases} \frac{5^k e^{-5}}{k!} & k = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

Prob[no hit in 10 sec]= Prob[K=0]=  $\frac{e^{-5}}{0!}=.0067$ 

2. By definition,  $\alpha=\lambda T$ , so  $\lambda=\frac{5}{10}=0.5$  hits/sec.

For T=2 sec,  $\alpha=\lambda T=0.5\times 2=1.0$  hit.

Let N be the number of hits in a 2-sec interval then

$$P_N(n) = \left\{ egin{array}{ll} rac{e^{-1}}{n!} & n=0,1,2,\dots \\ 0 & ext{otherwise} \end{array} 
ight.$$

Prob[at least 2 hits in 2 secs] =  $1-P_N(0)-P_N(1)=0.264$ 

#### 3. Some Useful Continuous Distributions

# 3.1 Uniform in the interval $\left[a,b\right]$

PDF

CDF

$$f_X(x) = \begin{cases} \frac{1}{b-a} & a \le x < b \\ 0 & \text{otherwise} \end{cases}$$

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$$F_X(x) = Prob[X \le x] = \begin{cases} 0 & x \le a \\ \frac{x-a}{b-a} & a < x \le b \\ 1 & x > b \end{cases}$$

- E[X] = (b+a)/2
- $Var[X] = (b-a)^2/12$

# 3.2 Exponential

Consider a Poisson process with rate  $\lambda$  events/sec, let X= the time between two consecutive events (inter-arrival time):

Events: Poisson with rate  $\lambda$  events/sec  $\leftarrow$  inter-arrival time  $\times$ 

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 $\mathbf{Q}$ : What is X's distribution?

A: X is exponentially distributed. CDF:

$$F_X(x) = Prob[X \le x] = \begin{cases} 1 - e^{-\lambda x} & x > 0 \\ 0 & \text{otherwise.} \end{cases}$$

■ Why?

$$\begin{array}{rcl} Prob[X \leq x] & = & Prob \ [\text{at least one event in } x] \\ & = & 1 - Prob \ [\text{zero event in } x] \\ & = & 1 - \left[\frac{(\lambda x)^n e^{-\lambda x}}{n!}\right]_{n=0} \\ & = & 1 - e^{-\lambda x}. \end{array}$$

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- $E[X] = 1/\lambda$   $Var[X] = 1/\lambda^2$
- **Example.** The call duration T in some cellular network is exponentially distributed with an average of 3 minutes per call. Find the probability that a call lasts between 2 and 4 minutes.

Solution.

1.  $\lambda = 1/3$  minutes/call

2. 
$$P[2 \le T \le 4] = F_T(4) - F_T(2) = e^{-2/3} - e^{-4/3} = 0.25.$$

# 4. J. Little's Theorem [1961]



N= avg. # of customers in the system
T= avg. delay per customer in the system

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- Theorem: If the limits N,  $\lambda$ , and T exist, and the system is in equilibrium (avg. arrival rate = avg. departure rate) then  $N = \lambda T$ .
- Example (The Burger King Spy).
  Spying on McDonald's, the spy recorded:
  - Arrivals:  $\lambda = 32$  customer's/hour
  - Each customer exits after 12 minutes on the average
  - ${\bf Q}$ : What is N (the avg. # of customers at McDonald's)?

#### Solution

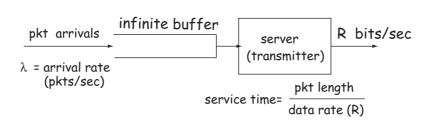
- 1.  $\lambda=32$  customers/hour  $=\frac{32}{60}\approx0.53$  customers/minute
- 2. Assuming equilibrium, by Little's Theorem:

$$N = \lambda T = 0.53 \times 12 \approx 6.4$$
 customers.

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## 5. The M/M/1 (Infinite Buffer) Queue

- Assumptions:
  - ullet arrival process: Poisson with rate  $\lambda$  pkt/sec
  - service time: exponentially distributed with average  $\frac{1}{\mu}$  sec. per pkt (that is,  $\mu$  pkts/sec can be viewed as the departure rate )



lacksquare A steady state solution exists only if  $\lambda < \mu$ .

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avg. # of pkts (queued 
$$+$$
 in-service)  $= \frac{\lambda}{\mu - \lambda}$  pkts.

By Little's Theorem:

Then

avg. pkt delay in system 
$$= \frac{1}{\mu - \lambda}.$$

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