

## Lecture 15 (Ch. 5-7)

We arrived at The Central Limit Theorem (CLT):

Weak version: If  $x \sim N(\mu, \sigma)$ , Then  $\bar{x} \sim N(\mu, \frac{\sigma}{\sqrt{n}})$

Strong version: If  $x \sim$  any dist. with mean  $= \mu_x$ , Var.  $= \sigma_x^2$   
Then  $\bar{x} \sim N(\mu_x, \frac{\sigma_x}{\sqrt{n}})$  for large  $n$ .  
 $\mu_x \leftarrow \mu_{\bar{x}}$   
 $\sigma_x \leftarrow \sigma_{\bar{x}}$

In English: For any pop. with mean  $\mu_x$  and Variance  $\sigma_x^2$ , The sampling dist. of The sample means is Normal with  $\mu = \mu_x$ ,  $\sigma = \sigma_x / \sqrt{n}$ .  
 $\mu_x \leftarrow \mu_{\bar{x}}$   
 $\sigma_x \leftarrow \sigma_{\bar{x}}$

So, if we know The pop. (ie.  $f(x)$ ,  $p(x)$ ), Then we can compute The prob. That a random sample mean will be somewhere.

E.g.  $pr(a < \bar{x} < b)$ . This is how:

or  $\int f(x) \dots$

1) Compute  $\mu_x, \sigma_x$ :  $\mu_x \equiv E[x] = \sum_x x p(x)$ ,  $\sigma_x^2 \equiv V[x] = \sum_x (x - \mu_x)^2 p(x)$ .

2) From CLT we know  $\bar{x} \sim N(\mu_x, \frac{\sigma_x}{\sqrt{n}})$

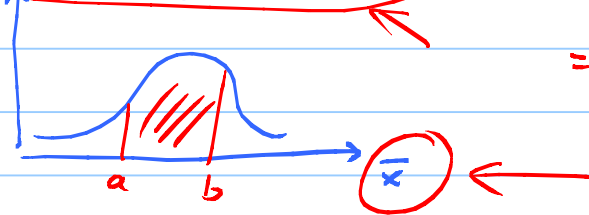
3) Then standardize:  $z \equiv \frac{\bar{x} - \mu_x}{\sigma_{\bar{x}}} = \frac{\bar{x} - \mu_x}{\sigma_x / \sqrt{n}} \sim N(0, 1)$

4) Finally  $pr(a < \bar{x} < b) = pr(a - \mu_x < \bar{x} - \mu_x < b - \mu_x)$

$$= pr\left(\frac{a - \mu_x}{\sigma_x / \sqrt{n}} < \frac{\bar{x} - \mu_x}{\sigma_x / \sqrt{n}} < \frac{b - \mu_x}{\sigma_x / \sqrt{n}}\right)$$

$$= pr(\text{''} < z < \text{''}) \Rightarrow \text{Table 1}$$

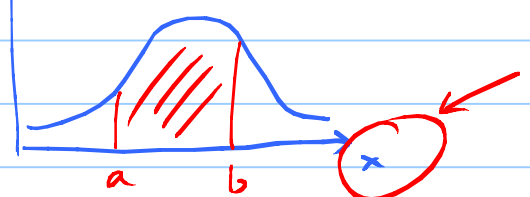
$f(\bar{x}) = \text{samp. dist. of } \bar{x}$



Compare with what we did in Ch. 2:

$$pr(a < x < b) = pr\left(\frac{a - \mu_x}{\sigma_x} < z < \frac{b - \mu_x}{\sigma_x}\right)$$

$f(x) = \text{pop.}$



E.g.

Suppose a sample of size 25 yields  $\bar{x}_{obs} = 3$ ,  $s = 1$ .

If the population is  $N(\mu = 2, \sigma = 1)$ , what's the prob. of getting an even larger sample mean?  $\mu_x = 2$   
 $\sigma_x = 1$

$$\text{prob}(\bar{x} > \bar{x}_{obs}) = \frac{\bar{x} - \mu_x}{\sigma_x / \sqrt{n}}$$

$$\text{prob}(\bar{x} > 3) = \text{prob}\left(\bar{z} > \frac{3-2}{1/\sqrt{25}}\right) = \text{prob}(z > 5) \approx 0!$$

This small prob suggests that  $\mu = 2$  is a bad assumption. In fact, we may even guess that  $\mu$  is greater than 2 (closer to 3)! We will formalize these qualitative conclusions, below.

\* Recall "prob" = proportion of samples (of size  $n$ ) taken from the population, in the long-run (e.g. out of  $10^8$  samples) prob works on random variables:

I.e.  $\text{prob}(a < \bar{x} < b)$  is computable. ✓

$\text{prob}(a < \bar{x}_{obs} < b)$  is NOT ✗

Note: in these calculations we are assuming we know the pop. But we don't. Intuitively, these probs give us a sense of how likely it would be to get a random sample mean somewhere, IF the pop. is given. In ch. 7, 8 we will come up with 2 ways of turning things around to say something about pop. from data.

Recall our symbols:

statistics

pop. parameters

$\bar{x}$  (sample mean) is a point estimate of  $\mu_x$  (pop. mean)

$s$  ("std. dev.") " " " "  $\sigma_x$  ("std. dev.")

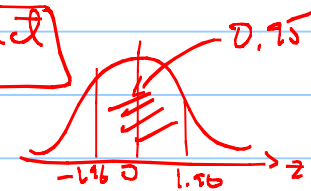
$p$  ("prop.") " " " "  $\pi_x$  ("prop.")

$n$  ("size") is NOT related to pop. size. ← for us =  $\infty$

The 1<sup>st</sup> way is to build a Confidence Interval (CI) for  $\mu_x$ :

The procedure is to start with  $pr(a < z < b) = \text{blah}$ , with specific values of  $a, b$ , and  $\text{blah}$ . E.g.

self-evident fact

$$pr(-1.96 < z < 1.96) = 0.95$$


$$pr\left(-1.96 < \frac{\bar{x} - \mu_x}{\frac{\sigma_x}{\sqrt{n}}} < 1.96\right) = 0.95$$

$\mu_x$   
 $\sigma_x$   
 $\sigma_x/\sqrt{n}$

$$pr\left(-1.96 \frac{\sigma_x}{\sqrt{n}} < \bar{x} - \mu_x < +1.96 \frac{\sigma_x}{\sqrt{n}}\right) = 0.95$$

← "Confidence level"

$\bar{x} - 1.96 \frac{\sigma_x}{\sqrt{n}} < \mu_x < \bar{x} + 1.96 \frac{\sigma_x}{\sqrt{n}}$   
 NB Prob.      fixed random

$\therefore$  95% C.I. for  $\mu_x$ :  $\bar{x} \pm 1.96 \frac{\sigma_x}{\sqrt{n}}$

pop. mean

This is a random C.I., because  $\bar{x}$  is random (how else would it have a sampling dist?!) For now, approximate this with sample std. dev.

The (observed) 95% C.I. for  $\mu_x$  is

$$\bar{x}_{obs} \pm 1.96 \frac{\sigma_x}{\sqrt{n}}$$

1 Interpretation: We are 95% Confident That  $\mu_x$  is in here.  
 2<sup>nd</sup> " : Below.

Often we forget saying "observed".

It's up to you to find out if we're talking about a random CI or The observed CI.

E.g. Suppose a sample of size 25 yields  $\bar{x}_{obs} = 3$ ,  $s = 1$ .  
What can we say about the pop. mean?

prov. eg. Suppose pop is normal ( $\mu_x = 2$ ,  $\sigma_x = 1$ ). What's the prob of getting an even larger sample mean?  
 $prob(\bar{x} > \bar{x}_{obs}) =$   
 $prob(\bar{x} > 3) = prob(z > \frac{3-2}{1/\sqrt{25}}) = prob(z > 5) \approx 0!$  (ch. 8)  
 $\bar{x} > 3$  is unlikely, if  $\mu_x = 2$ .

(observed) 95% C.I. for  $\mu_x$ :  $\bar{x}_{obs} \pm 1.96 \frac{\sigma_x}{\sqrt{n}}$  (estimate with  $s_x$ )  
 $3 \pm 1.96 \frac{1}{\sqrt{25}} = 3 \pm .392 = (2.6, 3.4)$

1 Interp.: We can be 95% Confident that the True mean is in here.

Note that we have actually made it to our goal of being able to say something about a pop. mean, from a sample.  
Review how we needed everything we've done since Ch. 1.  
Go and celebrate!

[Q1:] For the above e.g. which of the following is correct.

- A) The prob that  $2.6 < \mu_x < 3.4$  is 95%  $\mu_x = \text{fixed}$
- B) "  $2.6 < \bar{x}_{obs} < 3.4$  "  $\bar{x}_{obs} = \text{fixed}$
- C) "  $2.6 < \bar{x} < 3.4$  "  $pr(\dots < \bar{x} < \mu_x + 1.96 \frac{\sigma_x}{\sqrt{n}}) = 0.95$   
 $\uparrow$   
 not  $\bar{x}_{obs}$ .
- D) none of the above.

See Below for an interpretation that does involve prob.

Note That in The last step of The derivation of The C.I. for  $\mu_x$ , I dropped The pr. That is because  $\text{pr}(\dots > \mu_x > \dots)$  does not exist, because  $\mu_x$  is fixed, not random.

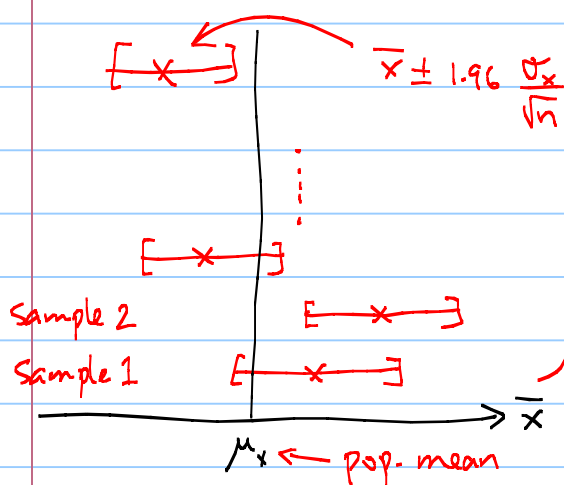
There is a way of squeezing "probability" into The conclusions, but it has to pertain to The random C.I.

We are 95% confident that the pop. mean is in the interval  $\bar{x} \pm 1.96 \frac{\sigma_x}{\sqrt{n}}$ .

↕ Equivalent interpretations of C.I.

There is a 95% prob that a random sample will yield a C.I.  $(\bar{x} \pm 1.96 \frac{\sigma_x}{\sqrt{n}})$  that covers  $\mu_x$ .

Look at the derivation of CI; This is obvious.



→ 95% of these intervals cover  $\mu_x$ .

→ I.e. The prob. that a random C.I.  $(\bar{x} \pm 1.96 \frac{\sigma_x}{\sqrt{n}})$  will include  $\mu_x$  is 0.95.

→ If you want to say something directly about  $\mu_x$ , use "Confidence". Not prob.

C.I.'s are all about coverage ;  
a 95% C.I. for  $\mu_x$  is designed to cover  $\mu_x$  in 95% of samples.

For The above example : (Observed) 95% CI (2.6, 3.4)

2<sup>nd</sup> interp.: There is 95% prob that a random CI will cover  $\mu_x$ .

### hw-lect 15-1

A sample of size 36 from a Normal pop. yields  $\bar{x}=3, s=1$ .

- Under the assumption that  $\mu_x=2.5, \sigma_x=2$ , what's the prob of a sample mean larger than the one observed.
  - Under the assumption that  $\mu_x=2.5, \sigma_x=2$ , what's the prob of a sample mean smaller than the one observed.
  - Under the assumption that  $\mu_x=3.5, \sigma_x=2$ , what's the prob of a sample mean larger than the one observed.
  - Under the assumption that  $\mu_x=3.5, \sigma_x=2$ , what's the prob of a sample mean smaller than the one observed.
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### hw-lect 15-2

- It turns out that the sample std. dev.,  $s$ , has a Normal distr. with parameters  $\sigma_x$  and  $\sigma_x/\sqrt{2n}$ , where  $\sigma_x$  is the pop. std. dev. Now, follow the procedure we have developed, starting from a "self-evident fact" to develop a C.I. formula for  $\sigma_x$ .
- Suppose for a specific data set based on a sample of size 169, we have found the sample std. dev. of 3.73. Compute the 95% CI for the pop. std. dev.
- provide 2 interpretations.

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