

Exercise 5.1

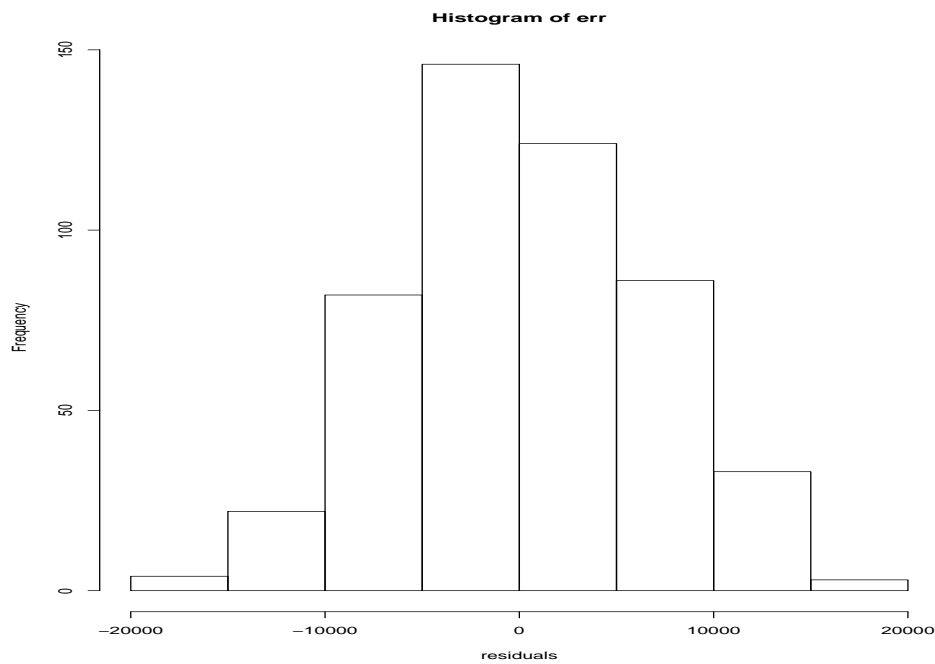


Figure 1: plot of residuals

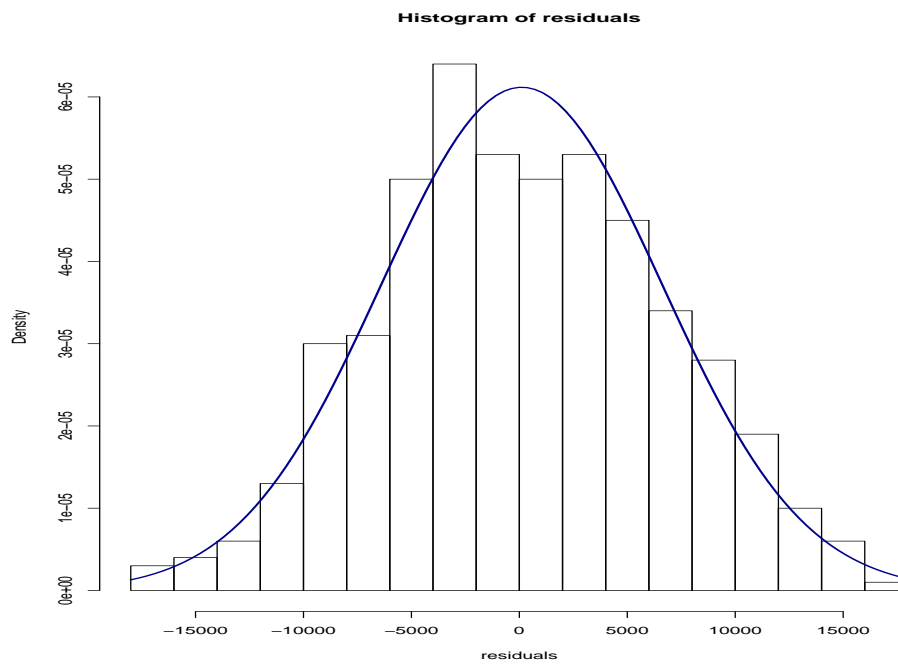


Figure 2: plot of residuals with Normal density curve

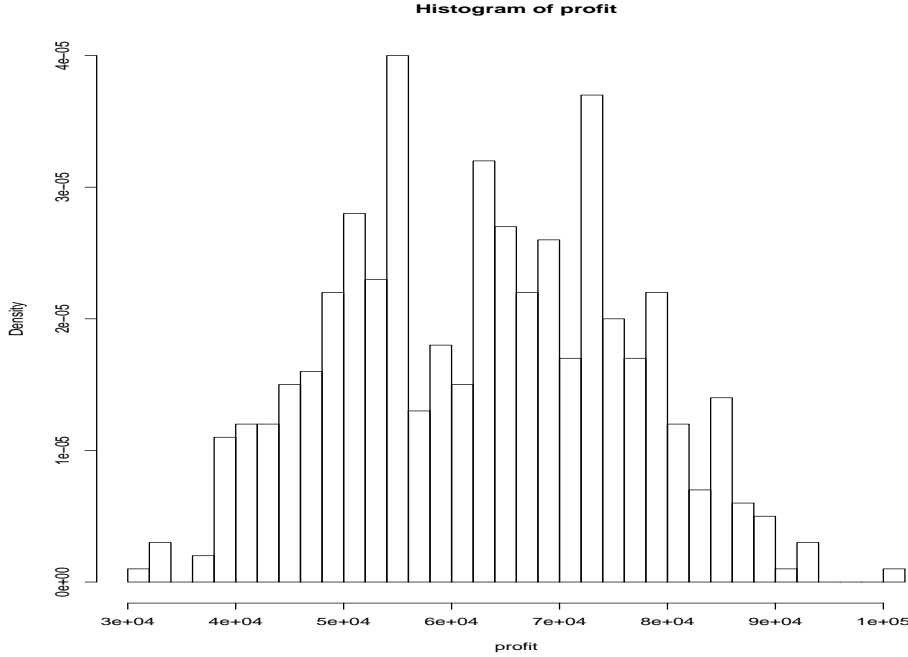


Figure 3: plot of raw data

It seems from Figure 1 and Figure 2 plots that the residuals resemble the nature of a Normal distribution. Figure 3 plot represents the distribution of raw profit data. It seems that the plot does not exactly resemble the nature of a Normal distribution but it is close enough.

Exercise 5.2

Model: $g(x|\theta) = \pi N(x|\mu_1, \sigma^2) + (1-\pi)N(x|\mu_2, \sigma^2)$ with priors on $\lambda = \frac{1}{\sigma^2} \sim \text{Gamma}(1, 1)$, $f(\mu_1) = N(0, 100)$ and $f(\mu_2) = N(0, 100)$ and $\pi = 0.5$. The latent variable here is the d which determines which group each observation came from. So $d = 0$ indicates sample from $N(x|\mu_1, \sigma^2)$ and $d = 1$ indicates sample from $N(x|\mu_2, \sigma^2)$.

$$g(x, 0|\theta) = \pi N(x|\mu_1, \sigma^2), g(x, 1|\theta) = (1 - \pi)N(x|\mu_2, \sigma^2)$$

We derive all the posterior distributions for d, μ_1, μ_2, λ .

Computation of posteriors:

$$d_i | \mu_1, \mu_2, \lambda, x_1, \dots, x_n \sim \text{Bernoulli}(p) \text{ where } p = \frac{\pi N(x_i | \mu_1, \sigma^2)}{\pi N(x_i | \mu_1, \sigma^2) + (1-\pi)N(x_i | \mu_2, \sigma^2)}.$$

$$f(\mu_1 | \dots) \propto \exp\left\{-\frac{1}{2\sigma^2} \sum_{d_i=0} (x_i - \mu_1)^2\right\} f(\mu_1)$$

$$\mu_1 | \dots \sim N\left(\frac{\lambda \sum_{d_i=0} x_i}{n_1 \lambda + 0.01}, (n_1 \lambda + 0.01)^{-1}\right), \sum x_i \text{ over } d_i = 0.$$

$$\text{Similarly, in this case, } \mu_2 | \dots \sim N\left(\frac{\lambda \sum_{d_i=1} x_i}{n_2 \lambda + 0.01}, (n_2 \lambda + 0.01)^{-1}\right), \sum x_i \text{ over } d_i = 1.$$

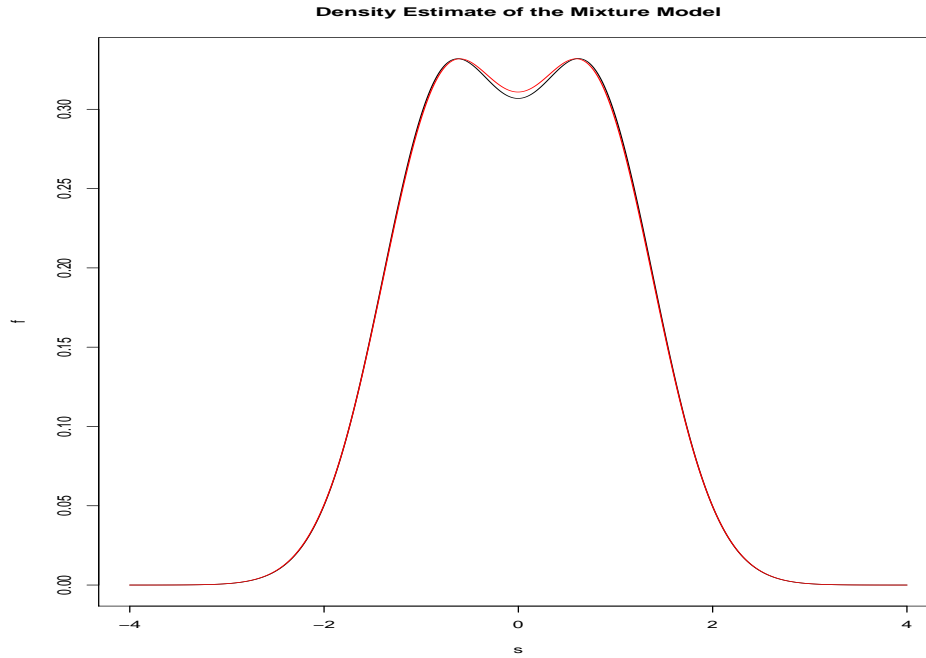


Figure 4: plot of estimated density assuming λ is same for μ_1 and μ_2

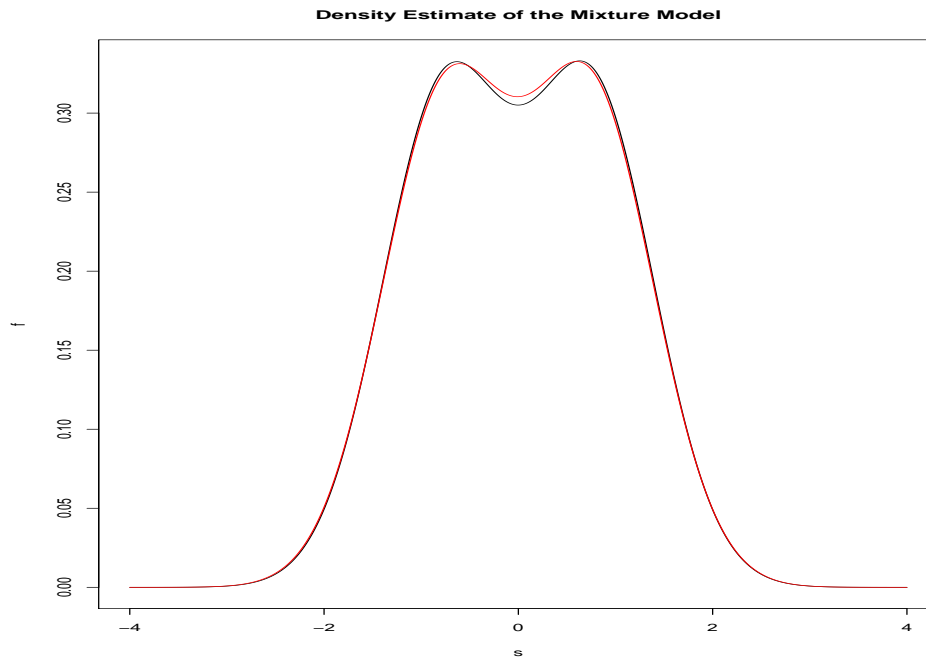


Figure 5: plot of estimated density assuming λ is not same for μ_1 and μ_2

$$f(\lambda|\dots) \propto f(\lambda) \prod N(x_i|\mu_{d_i}, 1/\lambda)$$

$$\lambda|\dots \sim \text{Gamma}(\frac{n}{2} + 1, 1 + \frac{1}{2} \sum (x_i - \mu_{d_i})^2)$$

where $\sum (x_i - \mu_{d_i})^2 = \sum x_i^2 - 2(\mu_1 \sum x_i + \mu_2 \sum x_i) + n_1 \mu_1^2 + n_2 \mu_2^2$

After getting all the full conditionals, I run Gibbs sampling with 5000 simulations. Next, I run the sampler using the given dataset. The two plots show the estimated densities.

Exercise 5.3

$$\begin{aligned}
 p(Z_i | Z_{-i}, X_i, \theta, \alpha, \beta) &= \int p(Z_i | X_i, \pi, \theta) p(\pi | Z_{-i}, \alpha, \beta) d\pi \\
 &= \int \pi^{Z_i} (1 - \pi)^{1 - Z_i} N(X_i | \mu(Z_i), \sigma^2(Z_i)) \text{Beta}(\alpha + \sum_{j \neq i} Z_j, \beta + n - \sum_{j \neq i} Z_j) d\pi \\
 \text{Let } \alpha + \sum_{j \neq i} Z_j &= \alpha^* \text{ and } \beta + n - \sum_{j \neq i} Z_j = \beta^*. \\
 &= N(X_i | \mu(Z_i), \sigma^2(Z_i)) \int \frac{\pi^{\alpha^* + Z_i - 1} (1 - \pi)^{\beta^* + 1 - Z_i - 1}}{\text{Beta}(\alpha^*, \beta^*)} d\pi \\
 &= N(X_i | \mu(Z_i), \sigma^2(Z_i)) \frac{\Gamma(\alpha^* + Z_i) \Gamma(\beta^* + 1 - Z_i)}{(\alpha^* + \beta^*) \Gamma(\alpha^*) \Gamma(\beta^*)} \\
 \text{Let } p &= \frac{\alpha^*}{\alpha^* + \beta^*} \\
 &= N(X_i | \mu(Z_i), \sigma^2(Z_i)) p^{Z_i} (1 - p)^{1 - Z_i}
 \end{aligned}$$

Exercise 5.6

Let $p_1, \dots, p_k \sim \text{Dir}(\alpha_1, \dots, \alpha_k)$
 $y_1, \dots, y_k \sim \text{Mult}(p_1, \dots, p_k)$

$$\begin{aligned}
 f(p_1, \dots, p_k | \dots) &\propto f(y_1, \dots, y_k | p_1, \dots, p_k) f(p_1, \dots, p_k | \alpha_1, \dots, \alpha_k) \\
 &\propto \prod p_j^{\alpha_j - 1 + \sum_i I(y_i = j)}
 \end{aligned}$$

Thus, we see that the above density is exactly a Dirichlet distribution with $\alpha_j^* = \alpha_j + \sum_i I(y_i = j)$