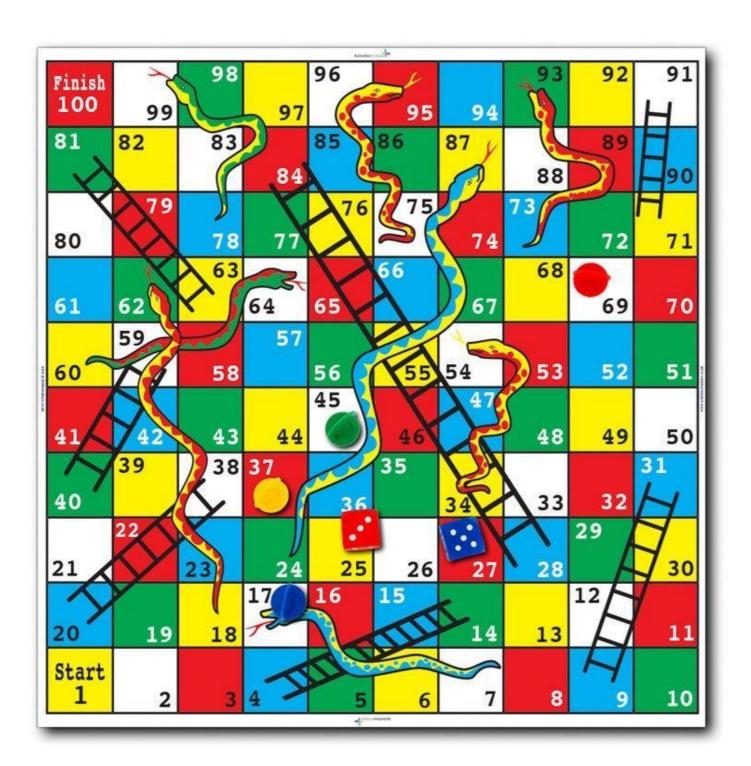
Workshop 16: Markov Chains

FIT 3139 Computational Modelling and Simulation





DISCRETE Markov chain

- Set of states $S = \{s_1, s_2, s_3, ..., s_r\}$
- Probability of going from s_i to s_j is given by p_{ij} (transition probability)

$$\sum_{j} p_{ij} = 1$$

• Transition matrix, $A = \{p_{ij}\}$

Conditional probability distribution of future states depends only on the present state

"Memoryless" process.

Conditional probability distribution of future states depends only on the present state

"Markovian"

Early 20th century...



https://en.wikipedia.org/wiki/Andrey_Markov

The Land of Oz is blessed by many things, but not by good weather.

They never have two nice days in a row. If they have a nice day, they are just as likely to have snow as rain the next day.

If they have snow or rain, they have an even chance of having the same the next day.

If there is change from snow or rain, only half of the time is this a change to a nice day.

Transition Matrix

R N S
$$R \begin{pmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix}$$

$$S \begin{pmatrix} \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix}$$

(stochastic matrix, because rows are probability distributions)

Simulation

- At every time step there is a discrete random variable that needs to be sampled.
- While stopping criteria is not met, decide what is the next state.

Simulation

If transitions are explicit....

Each state has an associated distribution, given by the corresponding row of the transition matrix.

$$X = \begin{cases} x_0 & \text{if } U < p_0 \\ x_1 & \text{if } p_0 \leq U < p_0 + p_1 \\ & \vdots \\ x_j & \text{if } \sum_{i=0}^{j-1} p_i \leq U < \sum_{i=0}^{j} p_i \\ & \vdots \\ & \text{with U, uniformly distributed on (0,1)} \end{cases}$$

R N S
$$R \begin{pmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix}$$

$$S \begin{pmatrix} \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix}$$

today tomorrow day after $R \longrightarrow R \longrightarrow S \quad \frac{1}{2} \cdot \frac{1}{4}$ R N S $R \rightarrow N \rightarrow S \frac{1}{4} \cdot \frac{1}{2}$

R N S

R
$$\left[\begin{array}{c|cccc} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ 1 & \frac{1}{2} & 0 & \frac{1}{2} \\ S & \left[\begin{array}{ccccc} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{array}\right] = P^2 = \begin{pmatrix} \frac{7}{16} & \frac{3}{16} & \frac{3}{8} \\ \frac{3}{8} & \frac{1}{4} & \frac{3}{8} \\ \frac{3}{8} & \frac{3}{16} & \frac{7}{16} \end{pmatrix}$$

If it is nice today, what is the probability that it is raining in two days?

R N S
$$R \begin{pmatrix} \frac{7}{16} & \frac{3}{16} & \frac{3}{8} \\ N & \frac{3}{8} & \frac{1}{4} & \frac{3}{8} \\ S & \frac{3}{8} & \frac{3}{16} & \frac{7}{16} \end{pmatrix} = P^2$$

Weather in two days...

given weather today (pick a row)

$$P^{2} \cdot P = \begin{pmatrix} \frac{7}{16} & \frac{3}{16} & \frac{3}{8} \\ \frac{3}{8} & \frac{1}{4} & \frac{3}{8} \\ \frac{3}{8} & \frac{3}{16} & \frac{7}{16} \end{pmatrix} \begin{pmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix} = \begin{pmatrix} \frac{11}{64} & \frac{5}{64} & \frac{1}{8} \\ \frac{3}{16} & \frac{1}{8} & \frac{3}{16} \\ \frac{1}{8} & \frac{5}{64} & \frac{11}{64} \end{pmatrix} \begin{array}{c} R \\ N = P^{3} \\ S \end{pmatrix}$$

Weather in three days...



For a Markov chain given by transition matrix \mathbf{P} the probability to go from state s_i to state s_j in n steps, is given by the ij-th entry of \mathbf{P}^n Entries of this matrix are denoted as: $p_{ij}^{(n)}$

Long-term behaviour:

$$\mathbf{P}^* = \lim_{k \to \infty} \mathbf{P}^k$$

$$(1,0,0) \cdot P^5 = \begin{pmatrix} \frac{205}{512} & \frac{205}{1024} & \frac{409}{1024} \end{pmatrix}$$

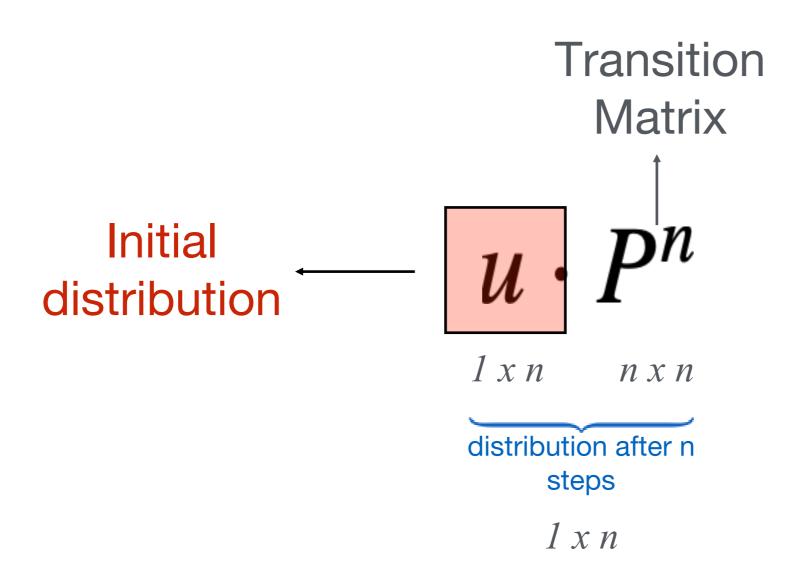
$$(0,1,0) \cdot P^5 = \begin{pmatrix} \frac{205}{512} & \frac{51}{256} & \frac{205}{512} \end{pmatrix}$$

$$(0,0,1) \cdot P^5 = \begin{pmatrix} \frac{409}{1024} & \frac{205}{1024} & \frac{205}{512} \end{pmatrix}$$

Initial
$$\left(\frac{1}{3}, \frac{1}{3}, \frac{1}{3}\right) \cdot P^5 = \left(\frac{205}{512} \quad \frac{205}{1024} \quad \frac{409}{1024}\right)$$

Probability
of being in a state
after 5 steps
...given initial state

Initial distribution



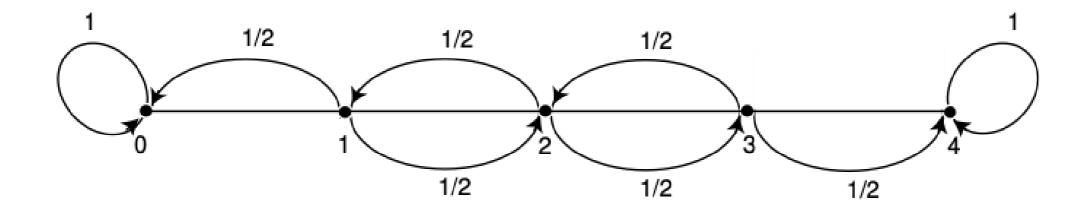
Drunkard's walk

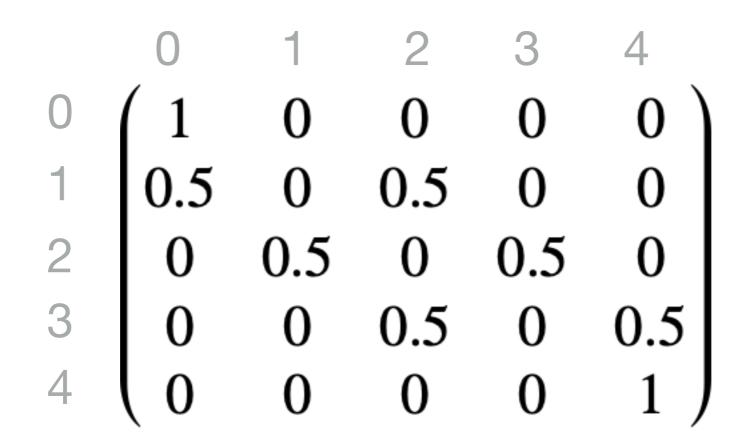
A man walks across a four block stretch of Sydney road.

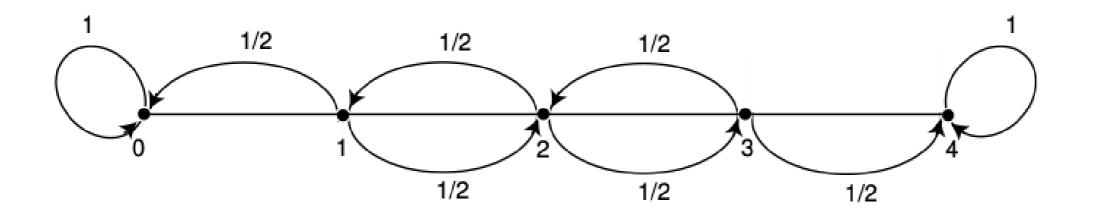
If he's at corner 1, 2, or 3 he walks to the left of right with equal probability.

He continues until he reaches corner 4, a hip bar... or corner 0, which is home.

If he reaches either home, or the bar he stays there.







What happens in the long term...



Absorbing vs Ergodic

Absorbing chains.

- A state is absorbing if it is impossible to leave (i.e., $p_{ii} = 1$).
- A <u>chain is absorbing</u>, if it has at least one absorbing state, and if from every state it is possible to go to an absorbing state (possibly in several steps)
- A state that is non-absorbing is also called <u>transient</u>.

Which ones are absorbing?

$$A = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0.5 & 0 & 0.5 & 0 & 0 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ 0 & 0 & 0.5 & 0 & 0.5 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix} \qquad B = \begin{pmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix}$$

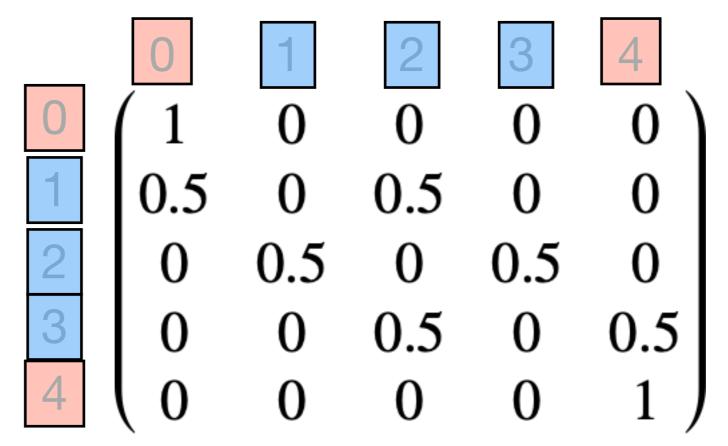
$$\mathsf{B} = \begin{pmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix}$$

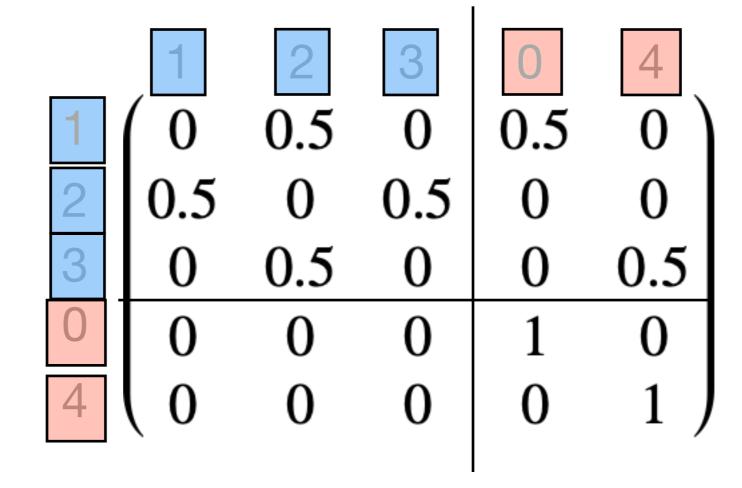
$$C = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 0.4 & 0.6 \\ 0 & 0.2 & 0.8 \end{pmatrix}$$

$$D = \begin{pmatrix} 1 & 0 & 0 \\ 0.3 & 0.4 & 0.3 \\ 0.2 & 0.1 & 0.7 \end{pmatrix}$$

Absorbing

Transient

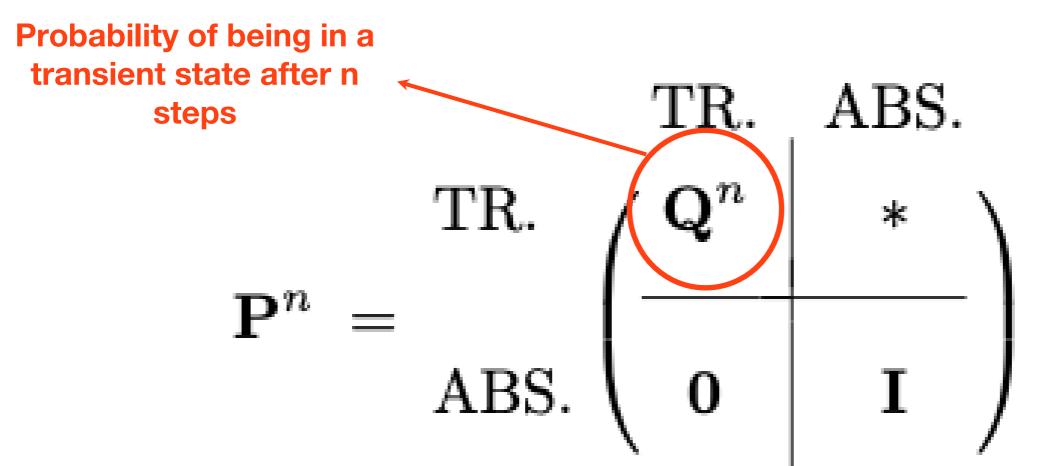




Canonical form

$$\mathbf{P} = \begin{bmatrix} \mathbf{TR.} & \mathbf{ABS.} \\ \mathbf{TR.} & \mathbf{Q} & \mathbf{R} \\ \mathbf{ABS.} & \mathbf{0} & \mathbf{I} \end{bmatrix}$$

Re-number states so that transient states come first



In an absorbing Markov chain the probability that the process will be absorbed is 1

$$\lim_{n\to\infty}\mathbf{Q}^n=\mathbf{0}$$



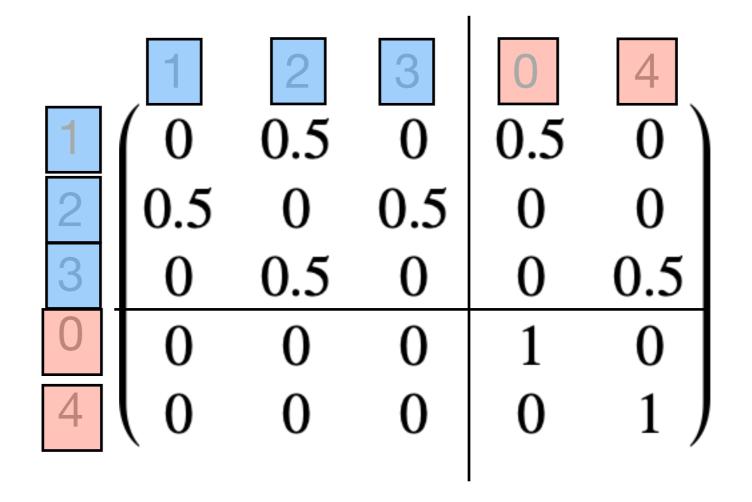
```
# Define M
   M = np.array([[0, 0.5, 0, 0.5, 0],
                [0.5, 0, 0.5, 0,
                [0, 0.5, 0, 0, 0.5],
                [0, 0, 0, 1, 0],
 5
 6
                [0,
                                   1 ]])
                    0, 0, 0,
   print(M)
[[0. 0.5 0. 0.5 0.]
[0.5 0. 0.5 0. 0.]
[0. 0.5 0. 0. 0.5]
[0. 0. 0. 1. 0.]
[0. 0. 0. 0. 1.]]
 1 # M to some power
 2 print(np.linalg.matrix_power(M, 8))
                0.03125 0.71875 0.21875]
 [0.03125 0.
        0.0625 0.
                      0.46875 0.46875]
 [0.
                0.03125 0.21875 0.71875]
 [0.03125 0.
0.
        0.
                0.
                              0.
                                     ]]
[0.
        0.
                0.
                       0.
                              1.
 1 # Q to some power
   print(np.linalg.matrix_power(M[0:3, 0:3], 8))
[[0.03125 0. 0.03125]
[0.
        0.0625 0.
[0.03125 0.
                0.03125]]
```

In an absorbing Markov chain the probability that the process will be absorbed is 1

$$\lim_{n\to\infty}\mathbf{Q}^n=\mathbf{0}$$

- Let m_j be the minimum number of steps required to reach an absorbing state, starting in transient state j
- Let p_j be the probability that starting in j, the process will not reach an absorbing state in m_i steps. Thus, $p_i < 1$.
- Let m be the largest m_j , and p be the largest p_j
- The <u>probability of not being absorbed</u> in m steps is at most p
- The probability of not being absorbed in mk steps is at most p^k
- The probability of not being absorbed goes to θ as k increases.

```
# Define M
   M = np.array([[0, 0.5, 0, 0.5, 0],
 3
                [0.5, 0, 0.5, 0, 0],
                [0, 0.5, 0, 0, 0.5],
 4
                [0, 0, 0, 1, 0],
 5
 6
                [0,
                     0, 0, 0, 1 ]])
   Q = M[0:3, 0:3]
   print(Q)
[[0. 0.5 0.]
[0.5 0. 0.5]
[0. 0.5 0. ]]
   print(np.linalg.matrix_power(Q, 15))
[[0. 0.00390625 0.
[0.00390625 0.
                     0.00390625]
[0.
    0.00390625 0.
                              ]]
   print(np.linalg.matrix_power(Q, 150))
[[1.32348898e-23 0.0000000e+00 1.32348898e-23]
[0.00000000e+00 2.64697796e-23 0.00000000e+00]
[1.32348898e-23 0.00000000e+00 1.32348898e-23]]
```



Remaining interesting questions?

- How many times do I visit a particular transient state before absorption?
- How long until absorption?
- Which of the absorbing states am I more likely to end up in?

Fundamental matrix

$$N = (I-Q)^{-1} = I + Q^1 + Q^2 + Q^3 + \dots$$
 fundamental matrix

The entry ij of N, is the expected number of times the chain visits state j, starting in state i... before absorption.

The entry ij of N, is the expected number of times the chain visits state j, starting in state i... before absorption.

Absorption time

$$t = N \cdot \begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \\ 1 \end{pmatrix}$$

vector with as many components as transient states

 t_i is the **expected time before absorption** from transient state i

(intuition: sum the times I visit all transient states before absorption)



```
# Define M
   M = np.array([[0, 0.5, 0, 0.5, 0],
 3
                 [0.5, 0, 0.5, 0, 0],
                  [0, 0.5, 0, 0, 0.5],
 4
                  [0, 0, 0, 1, 0],
[0, 0, 0, 0, 1]])
 5
 6
    Q = M[0:3, 0:3]
   print(Q)
[[0. 0.5 0.]
[0.5 \ 0. \ 0.5]
[0. 0.5 0. ]]
 1 # Find fundamental matrix
 2 N = np.linalg.inv((np.eye(3) - Q))
 3 # Calculate Absorption times
 4 t = N.dot(np.array([1, 1, 1]))
```

[3. 4. 3.]

5 print(t)

Absorption probability

$$\mathbf{P} = egin{array}{c|c} \mathbf{TR.} & \mathbf{ABS.} \\ \mathbf{P} & \mathbf{R} \\ \mathbf{ABS.} & \mathbf{O} & \mathbf{I} \\ \end{array}$$

Remember R

Absorption probability

Let b_{ij} be the probability that an absorbing chain will be absorbed in the absorbing state j if it starts in the transient state i. Let B be the matrix with entries b_{ij} . Then B is a t-by-r matrix and:

$$B = NR$$

$$B_{ij} = \sum_{n} \sum_{k} q_{ik}^{(n)} r_{kj}$$
 $q_{ik}^{(n)} = \sum_{i}^{i} \sum_{k} e^{(n)} r_{kj}$ $e^{(n)} = \sum_{i}^{i} \sum_{k} e^{(n)} r_{kj}$ $e^{$



```
1 # Define M
   M = np.array([[0, 0.5, 0, 0.5, 0],
                 [0.5, 0, 0.5, 0, 0],
                 [0, 0.5, 0, 0, 0.5],
 4
                 [0, 0, 0, 1, 0],
 5
                 [0,
                     0, 0, 0, 1 ]])
 7 | Q = M[0:3, 0:3]
 8 print(Q)
 9 R = M[0:3,3:]
    print(R)
[[0. 0.5 0.]
[0.5 0. 0.5]
[0. 0.5 0. ]]
[[0.5 0.]
[0. 0.]
[0. 0.5]]
 1 # Find fundamental matrix
 2 N = np.linalg.inv((np.eye(3) - Q))
 3 # Calculate Absorption probabilities
 4 \mid B = N.dot(R)
 5 print(B)
[[0.75 0.25]
[0.5 0.5]
[0.25 0.75]]
```

Absorbing vs Ergodic

Ergodic chains.

 A Markov chain is called <u>ergodic</u> if it is possible to go from every state to every state (not necessarily in one step).

 Ergodic chains have a <u>unique long-run equilibrium</u> that does not depend on initial conditions.

 Intuitively, given enough time, the process forgets about the initial configuration.

Long-term behaviour:

$$\mathbf{P}^* = \lim_{k \to \infty} \mathbf{P}^k$$

$$\begin{pmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix}$$

$$\begin{pmatrix} 0.4 & 0.2 & 0.4 \\ 0.4 & 0.2 & 0.4 \\ 0.4 & 0.2 & 0.4 \end{pmatrix}$$

$$\mathbf{P}^* = \{p_{ij}^*\}$$
 if history does not matter $p_{ij}^* = p_{i'j}^*$ i.e. all rows of \mathbf{P}^* must be the same $\mathbf{PP}^* = \mathbf{P}\lim_{k \to \infty} \mathbf{P}^k = \lim_{k \to \infty} \mathbf{P}^{k+1} = \mathbf{P}^*$

$$\mathbf{u} = (u_1, u_2, ..., u_n)$$
 is the common row of \mathbf{P}^*

$$u_j = \lim_{m \to \infty} p_{ij}^{(m)}$$

stationary distribution



```
# Define M
   M = np.array([[0.5, 0.25, 0.25],
                  [0.5, 0, 0.5],
                  [0.25, 0.25, 0.5]
    print(np.linalg.matrix_power(M, 5))
[[0.40039062 0.20019531 0.39941406]
 [0.40039062 0.19921875 0.40039062]
 [0.39941406 0.20019531 0.40039062]]
   print(np.linalg.matrix_power(M, 20))
[[0.4 \ 0.2 \ 0.4]
[0.4 \ 0.2 \ 0.4]
[0.4 0.2 0.4]]
```

Consider:
$$M = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

Is M ergodic?

Yes, it is possible to go from every state to every other state

What happens when we raise M to some power?

$$M^{2k} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \qquad \qquad M^{2k+1} = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

The rows don't converge – what's the stationary distribution?

This method only works for Regular Markov Chains

A Markov Chain is called a *regular* chain if some power of the transition matrix has only positive elements.

$$M = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$
 is not regular

So how do we find the stationary distribution for *M*?

$\mathbf{u} = \mathbf{u}\mathbf{P}$

u is the <u>left</u> eigenvector for eigenvalue 1 of **P**

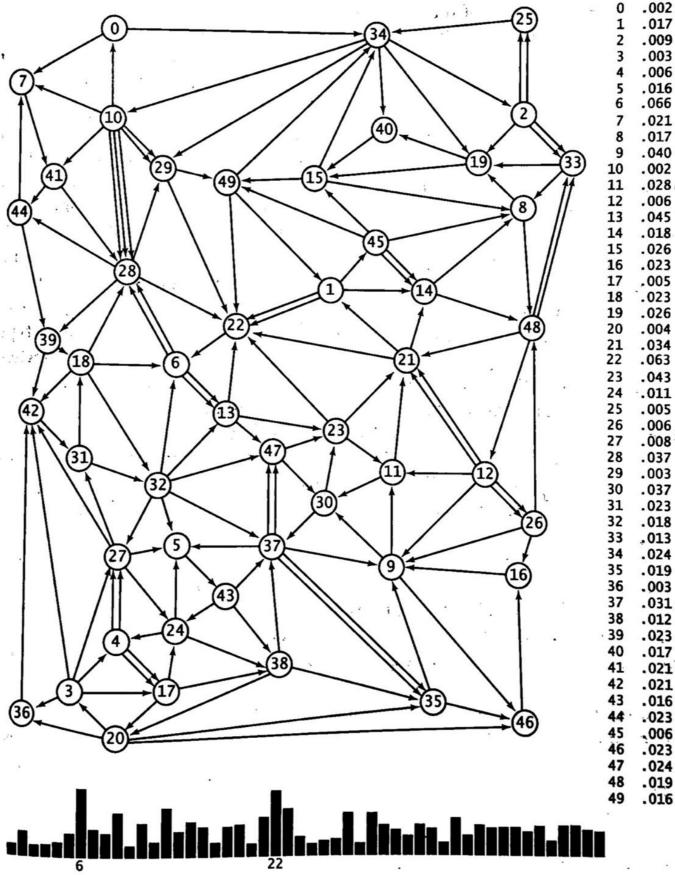
stationary distribution

What problem do we have with **u** when calculated this way?



```
1 | # Define M
 2 M = np.array([[0.5, 0.25, 0.25],
                  [0.5, 0, 0.5],
                  [0.25, 0.25, 0.5]
   lambda_, v = np.linalg.eig(M.T)
   print(lambda_)
[ 1. 0.25 -0.25]
 1 print(v[:,0]/sum(v[:,0]))
[0.4 \ 0.2 \ 0.4]
```

The ergodic chain of a random web surfer



Page ranks with histogram for a larger example