EE7330: Network Information Theory

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Lecture Notes 11: Tools for proving converse results

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11.1 Fano's inequality

Consider M and \hat{M} are jointly distributed and probability of error $P_e \approx Pr[M \neq \hat{M}]$. And M & $\hat{M} \in \mathbb{M}$.

$$H(M/\hat{M}) \leqslant H_2(P_e) + P_e \log|\mathbb{M}| \tag{11.1}$$

Proof:

Consider a indicator function E as:

$$E = \begin{cases} 1 & \text{if } \hat{M} \neq M \\ 0 & \text{if } \hat{M} = M \end{cases}$$
 (11.2)

So, we have $P_E(1) = P_e, P_E(0) = 1 - P_e$, and we can write $H(E) = H_2(P_e)$. Now

$$H\left(M, \frac{E}{\hat{M}}\right) = H\left(\frac{M}{\hat{M}}\right) + H\left(\frac{E}{M}, \hat{M}\right) \text{ (chain rule of entropy)}$$
 (11.3)

$$H\left(\frac{M}{\hat{M}}\right) = H\left(\frac{E}{\hat{M}}\right) + H\left(\frac{M}{E}, \hat{M}\right) - H\left(\frac{E}{M\hat{M}}\right) \tag{11.4}$$

Since E is independent from M and M, we can write,

$$H\left(\frac{M}{\hat{M}}\right) = H\left(\frac{E}{\hat{M}}\right) + H\left(\frac{M}{E}, \hat{M}\right) \tag{11.5}$$

$$\leq H(E) + H\left(\frac{M}{E}, \hat{M}\right)$$
 (11.6)

$$= H_2(P_e) + H\left(\frac{M}{\hat{M}}, E = 0\right) P_E(0) + H\left(\frac{M}{\hat{M}}, E = 1\right) P_E(1)$$
(11.7)

Here, $H\left(\frac{M}{\hat{M}}, E=0\right)=0$ because for E=0, $M=\hat{M}$, and hence entropy =0.

$$= H_2(P_e) + H\left(\frac{M}{\hat{M}}, E = 1\right) P_E(1)$$
(11.8)

$$= H_2(P_e) + H\left(\frac{M}{\hat{M}}, E = 1\right) P_e \tag{11.9}$$

$$\leq H_2(P_e) + H(M)P_e \tag{11.10}$$

$$\leq H_2(P_e) + P_e \log_2 |\mathbb{M}| \text{ Proved.}$$
 (11.11)

11.2 Proof of converse channel coding theorem

Consider any sequence with (ENC_n, DEC_n) such that,

$$\liminf_{n \to \infty} \frac{K_n}{n} \geqslant C + \epsilon$$

then,

$$P_e = \limsup_{n \to \infty} = Pr\left[\hat{M}_i \neq M_i\right] \geqslant \epsilon$$

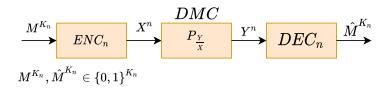


Figure 11.1: Single user over DMC channel

Proof:

we need some well known inequalities to prove this. 1) Fano's inequality and 2) bound on mutual information. Fano's inequality we have discussed just above and now let discuss bound on mutual information. For any P_{X^n} , if Y^n is obtained by passing of X^n through Discrete memoryless channel having transition probability $P_{\frac{Y}{Y}}$ then,

$$I\left(X^{n};Y^{n}\right) \leqslant \sum_{i=1}^{n} I\left(X_{i};Y_{i}\right) \leqslant nC \tag{11.12}$$

Now using expression of mutual information,

$$I\left(X^{n};Y^{n}\right) = H\left(Y^{n}\right) - H\left(\frac{Y^{n}}{X^{n}}\right) \tag{11.13}$$

$$= \sum_{i=1}^{n} \left[H\left(\frac{Y_i}{Y_1, Y_2 \dots Y_{i-1}}\right) - H\left(\frac{Y_i}{Y_1, Y_2 \dots Y_{i-1}, X^n}\right) \right]$$
(11.14)

(Chain rule of entropy)

$$\leq \sum_{i=1}^{n} \left[H(Y_i) - H\left(\frac{Y_i}{Y_1, Y_2 \dots Y_{i-1}, X^n}\right) \right]$$
 (11.15)

(Since conditioning decreases entropy)

Now,

$$H\left(\frac{Y_i}{Y_1, Y_2 \dots Y_{i-1}, X^n}\right) = \sum_{x^n, y_1, \dots, y_n} P\left(\frac{y_i}{y_1, \dots y_{i-1}, x^n}\right) \times \log\left(\frac{1}{P\left(\frac{y_i}{y_1, \dots y_{i-1}, x^n}\right)}\right)$$
(11.16)

Since for dicrete memoryless channel present output depends only on the present input, hence we can write.

$$P\left(\frac{y_i}{y_1, \dots y_{i-1}, x^n}\right) = P\left(\frac{y_i}{x_i, y_1, x_i \dots y_{i-1}, x_{i-1}}\right) = P\left(\frac{y_i}{x_i}\right)$$
(11.17)

Using eq 11.17 in eq 11.16, we can write,

$$H\left(\frac{Y_i}{Y_1, Y_2 \dots Y_{i-1}, X^n}\right) = \sum_{x^n, y_1, \dots, y_n} P\left(\frac{y_i}{x_i}\right) \times \log\left(\frac{1}{P\left(\frac{y_i}{x_i}\right)}\right) = H\left(\frac{Y_i}{X_i}\right)$$
(11.18)

By combining all i.e using eq11.18 in eq11.15 we can write,

$$= \sum_{i=1}^{n} \left[H\left(Y_{i}\right) - H\left(\frac{Y_{i}}{X_{i}}\right) \right] \tag{11.19}$$

And we have,

$$\sum_{i=1}^{n} \left[H(Y_i) - H\left(\frac{Y_i}{X_i}\right) \right] = \sum_{i=1}^{n} I(X_i; Y_i)$$
(11.20)

From eq11.13, eq11.15 and eq11.20 we have,

$$I(X^n; Y^n) \le \sum_{i=1}^n I(X_i; Y_i) \le nC$$
 (11.21)

This conclude the proof of eq 11.12.

Proof of converse:

We have message, which is i.i.d and uniform so we can write,

$$K_n = H\left(M^{K_n}\right) \tag{11.22}$$

$$=H\left(\frac{M^{K_n}}{\hat{M}^{K_n}}\right)+I\left(M^{K_n};\hat{M}^{K_n}\right) \tag{11.23}$$

$$\leq H_2(P_e) + P_e log_2(2^{M^{K_n}}) + I(M^{K_n}: M'^{K_n})$$
 (11.24)

(using Fano's inequality)

$$\leq H_2(P_e) + P_e K_n + I\left(M^{K_n}; \hat{M}^{K_n}\right)$$
 (11.25)

$$\leq H_2(P_e) + P_e K_n + I(X^n; Y^n)$$
 (11.26)

(using data processing inequality)

$$\leq H_2(P_e) + P_e K_n + nC$$
 (From eq 11.21) (11.27)

$$K_n \leqslant H_2(P_e) + P_e K_n + nC \Rightarrow \frac{H_2(P_e)}{n} \geqslant \frac{K_n(1 - P_e)}{n} - C$$

$$\lim_{n \to \infty} \left(\frac{K_n(1 - P_e)}{n} - C \right) \leqslant \lim_{n \to \infty} \frac{H_2(P_e)}{n}$$

Where,
$$P_e = \limsup_{n \to \infty} Pr\left[\hat{M}_i^{K_n} \neq M_i^{K_n}\right]$$

Let $\frac{K_n}{n} = R$,

$$\lim_{n \to \infty} P_e \geqslant \frac{R - C}{R} \tag{11.28}$$

$$\limsup_{n \to \infty} P_e \geqslant \frac{R - C}{R} \tag{11.29}$$

If we operate at the rate more than capacity of the channel say $R = C + \epsilon$

$$\limsup_{n \to \infty} P_e \geqslant \frac{\epsilon}{R} \tag{11.30}$$

That is probability or error is always non-zero. But if we operate the channel below the capacity and for

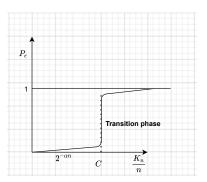


Figure 11.2: Probability of error with capacity constraint

large n, in (11.29) we can observe P_e approaches to 0.

11.3 Mrs Gerber's Leema:

Suppose we have $X = (X_i, ... X_n)$ where $X_i \in \mathbb{M}^n \in \{0,1\}$ be a binary random n-vector. Let $P_X(x) = P(X = x)$, where $x \in \mathbb{M}^n$ define its probability distribution Let say X = Ber(q). Let us consider that the random vector X is the input to a binary symmetric channel with crossover probability P, where $0 < P < \frac{1}{2}$. Let be $Y = (Y_i, ... Y_n)$ where $Y_i \in \mathbb{M}^n \in \{0,1\}$ the corresponding channel output n-vector. The probability

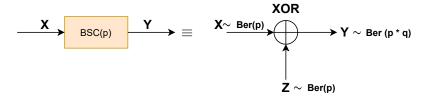


Figure 11.3: BSC channel

distribution of Y is defined using transition probability of channel and X (described in figure 11.3). We will use notation $\mathbf{p} \star \mathbf{q} = \mathbf{p}(\mathbf{1} - \mathbf{q}) + \mathbf{q}(\mathbf{1} - \mathbf{p})$. Then,

$$H(Y) \geqslant H_2 \left(p \star H_2^{-1} (H(X)) \right)$$
 (11.31)

with equality if and only if the $\{X_i\}_1^n$ are independent. and $H\{X_k\}=kp$

Suppose we are generation X which depends on some distribution U, and it passes through BSC(p) (with transition probability p) with output distribution Y. Then,

$$H\left(\frac{Y}{U}\right) \geqslant H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right)$$
 (11.32)

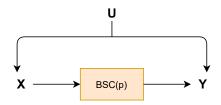


Figure 11.4: BSC channel

In vector form,

$$\frac{H\left(Y^{n}|u\right)}{n} \geqslant H_{2}\left(H_{2}^{-1}\left(H\left(\frac{X^{n}|u}{n}\right)\right) \star p\right) \tag{11.33}$$

proof:

We have claim,

$$H\left(\frac{Y}{U}\right) \geqslant H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right)$$
 (11.34)

Considering R.H.S. of the inequality we have,

$$H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right) \tag{11.35}$$

Suppose we have $f(u) = H_2(H_2^{-1}(u) \star p)$.

Now, eq 11.35 can be written as,

$$f\left(\sum_{u} P_{U}(u).H\left(X|U=u\right)\right) \tag{11.36}$$

Since f is convex function, we can write,

$$f\left(\sum_{u} P_{U}(u).H\left(X|U=u\right)\right) \leqslant \sum_{u} P_{U}(u).f\left(H\left(X|U=u\right)\right) \tag{11.37}$$

Again using the expression of f(u) we can have,

$$f(H(X|U=u)) = H_2(H_2^{-1}(H(X|U=u)) \star p)$$
 (11.38)

For $U = u, X \sim Ber(q_u)$ and $Y \sim (q_u \star p)$.

$$H(X|U=u) = H_2(q - u \star p) = H_2(H_2^{-1}(H(X|U=u)) \star p) = f(H(X|U=u))$$
(11.39)

Combining all above expressions, R.H.S. becomes,

R.H.S
$$\leq \sum_{u} P_{U}(u).H(X|U=u)$$
 (11.40)

Hence,

$$H\left(\frac{Y}{U}\right) \geqslant H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right)$$
 (11.41)

and this conclude the proof.