EE7330: Network Information Theory

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Lecture Notes 11: Tools for proving converse results

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11.1 Fano's inequality

Consider M and \hat{M} are jointly distributed and probability of error $P_e = Pr[M \neq \hat{M}]$. And M & $\hat{M} \in \mathbb{M}$. Then

$$H(M/\hat{M}) \leqslant H_2(P_e) + P_e \log|\mathbb{M}| \tag{11.1}$$

Proof:

Consider a indicator function E as :

$$E = \begin{cases} 1 & \text{if } \hat{M} \neq M \\ 0 & \text{if } \hat{M} = M \end{cases}$$
 (11.2)

So, we have $P_E(1) = P_e, P_E(0) = 1 - P_e$, and we can write $H(E) = H_2(P_e)$. Now

$$H\left(M, E|\hat{M}\right) = H\left(M|\hat{M}\right) + H\left(E|M, \hat{M}\right)$$
 (chain rule of entropy) (11.3)

$$H\left(M|\hat{M}\right) = H\left(E|\hat{M}\right) + H\left(E|M,\hat{M}\right) - H\left(E|M\hat{M}\right) \tag{11.4}$$

Since E is independent from M and \hat{M} , we can write,

$$H\left(M|\hat{M}\right) = H\left(E|\hat{M}\right) + H\left(\frac{M}{E}, \hat{M}\right) \tag{11.5}$$

$$\leq H(E) + H\left(M|E,\hat{M}\right)$$
 (11.6)

$$= H_2(P_e) + H(M|\hat{M}, E = 0) P_E(0) + H(M|\hat{M}, E = 1) P_E(1)$$
(11.7)

Here, $H(M|\hat{M}, E=0) = 0$ because for E=0, $M=\hat{M}$, and hence entropy = 0.

$$= H_2(P_e) + H(M|\hat{M}, E = 1) P_E(1)$$
(11.8)

$$= H_2(P_e) + H(M|\hat{M}, E = 1) P_e$$
(11.9)

$$\leqslant H_2\left(P_e\right) + H\left(M\right)P_e\tag{11.10}$$

$$\leq H_2(P_e) + P_e \log_2 |\mathbb{M}| \text{ Proved.}$$
 (11.11)

11.2 Proof of converse channel coding theorem

Theorem : Consider any sequence with (ENC_n, DEC_n) for DMC with transition probability $P_{Y|X}$ such that,

$$\lim_{n \to \infty} \inf \frac{K_n}{n} \geqslant C + \epsilon$$

then,

$$P_e = \limsup_{n \to \infty} = Pr\left[\hat{M}_i \neq M_i\right] \geqslant \frac{\epsilon}{R}$$

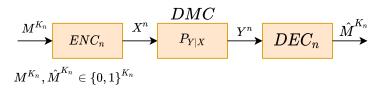


Figure 11.1: Single user over DMC

Proof:

We need some well-known inequalities to prove this. 1) Fano's inequality and 2) bound on mutual information. Fano's inequality we have discussed just above and now let us discuss bound on mutual information. **Lemma**: For any P_{X^n} , if Y^n is obtained by passing X^n through Discrete memoryless channel having transition probability $P_{Y|X}$ then,

$$I\left(X^{n};Y^{n}\right) \leqslant \sum_{i=1}^{n} I\left(X_{i};Y_{i}\right) \leqslant nC \tag{11.12}$$

Proof of lemma:

Consider the expression of mutual information,

$$I(X^{n}; Y^{n}) = H(Y^{n}) - H(Y^{n}|X^{n})$$
(11.13)

$$= \sum_{i=1}^{n} \left[H\left(Y_{i} | Y_{1}, Y_{2} \dots Y_{i-1}\right) - H\left(Y_{i} | Y_{1}, Y_{2} \dots Y_{i-1}, X^{n}\right) \right]$$
(11.14)

(Chain rule of entropy)

$$\leq \sum_{i=1}^{n} \left[H(Y_i) - H(Y_i | Y_1, Y_2 \dots Y_{i-1}, X^n) \right]$$
(11.15)

(Since conditioning decreases entropy)

Now,

$$H(Y_{i}|Y_{1}, Y_{2}...Y_{i-1}, X^{n}) = \sum_{x^{n}, y_{1},...,y_{n}} P_{Y|X}(y_{i}|y_{1},...y_{i-1}, x^{n}) \times \log\left(\frac{1}{P_{Y|X}(y_{i}|y_{1},...y_{i-1}, x^{n})}\right)$$
(11.16)

Since for discrete memoryless channel present output depends only on the present input, hence we can write.

$$P_{Y|X}(y_i|y_1, \dots y_{i-1}, x^n) = P_{Y|X}(y_i|x_i, y_1, x_i \dots y_{i-1}, x_{i-1}) = P_{Y|X}(y_i|x_i)$$
(11.17)

Using eq (11.17) in eq (11.16), we can write,

$$H(Y_i|Y_1, Y_2 \dots Y_{i-1}, X^n) = \sum_{x^n, y_1, \dots, y_n} P_{Y|X}(y_i|x_i) \times \log_2\left(\frac{1}{P_{Y|X}(y_i|x_i)}\right) = H(Y_i|X_i)$$
(11.18)

By combining all i.e using eq(11.18) in eq11.15 we can write,

$$= \sum_{i=1}^{n} \left[H(Y_i) - H(Y_i|X_i) \right]$$
 (11.19)

And we have,

$$\sum_{i=1}^{n} \left[H(Y_i) - H(Y_i|X_i) \right] = \sum_{i=1}^{n} I(X_i; Y_i)$$
(11.20)

From eq(11.13), eq(11.15) and eq(11.20) we have,

$$I(X^n; Y^n) \le \sum_{i=1}^n I(X_i; Y_i) \le nC$$
 (11.21)

This conclude the proof of eq (11.12).

Proof of converse:

We have message, which is i.i.d and uniform so we can write,

$$K_n = H\left(M^{K_n}\right) \tag{11.22}$$

$$= H\left(M^{K_n}|\hat{M}^{K_n}\right) + I\left(M^{K_n}; \hat{M}^{K_n}\right)$$
 (11.23)

$$\leq H_2(P_e) + P_e \log_2(2^{M^{K_n}}) + I(M^{K_n} : {M'}^{K_n})$$
 (11.24)

(using Fano's inequality)

$$\leq H_2(P_e) + P_e K_n + I\left(M^{K_n}; \hat{M}^{K_n}\right)$$
 (11.25)

$$\leq H_2(P_e) + P_e K_n + I(X^n; Y^n)$$
 (11.26)

(using data processing inequality)

$$\leq H_2(P_e) + P_e K_n + nC \quad \text{(From eq 11.21)}$$
 (11.27)

$$K_n \leqslant H_2(P_e) + P_e K_n + nC \Rightarrow \frac{H_2(P_e)}{n} \geqslant \frac{K_n(1 - P_e)}{n} - C$$

$$\lim_{n \to \infty} \left(\frac{K_n(1 - P_e)}{n} - C \right) \leqslant \lim_{n \to \infty} \frac{H_2(P_e)}{n}$$

Where,
$$P_e = \limsup_{n \to \infty} Pr\left[\hat{M}_i^{K_n} \neq M_i^{K_n}\right]$$

Let $\frac{K_n}{n} = R$,

$$\lim_{n \to \infty} P_e \geqslant \frac{R - C}{R} \tag{11.28}$$

$$\limsup_{n \to \infty} P_e \geqslant \frac{R - C}{R} \tag{11.29}$$

If we operate at the rate more than capacity of the channel say $R = C + \epsilon$

$$\limsup_{n \to \infty} P_e \geqslant \frac{\epsilon}{R} \tag{11.30}$$

That is probability or error is always non-zero. But if we operate the channel below the capacity and for

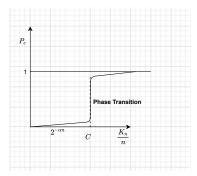


Figure 11.2: Probability of error with capacity constraint

large n, in (11.29) we can observe P_e approaches to 0.

11.3 Mrs Gerber's Lemma:

Suppose we have $X = (X_i, ... X_n)$ where $X_i \in \mathbb{M}^n \in \{0,1\}$ be a binary random n-vector. Let $P_X(x) = P(X = x)$, where $x \in \mathbb{M}^n$ define its probability distribution Let say $X \sim Ber(q)$. Let us consider that the random vector X is the input to a binary symmetric channel with crossover probability p, where $0 . Let be <math>Y = (Y_i, ... Y_n)$ where $Y_i \in \mathbb{M}^n \in \{0,1\}$ the corresponding channel output n-vector. The probability

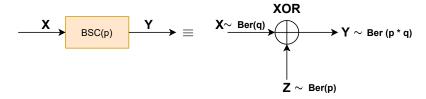


Figure 11.3: BSC channel

distribution of Y is defined using transition probability of channel and X (described in figure 11.3). We will use notation $\mathbf{p} \star \mathbf{q} = \mathbf{p}(\mathbf{1} - \mathbf{q}) + \mathbf{q}(\mathbf{1} - \mathbf{p})$. Then,

$$H(Y) \geqslant H_2 \left(p \star H_2^{-1} (H(X)) \right)$$
 (11.31)

with equality if and only if the $\{X_i\}_1^n$ are independent. and $H\{X^k\}=kp$

Lemma : Suppose we are generating X which depends on some distribution U, and it passes through BSC(p) (with transition probability p) with output distribution Y. Then,

$$H\left(\frac{Y}{U}\right) \geqslant H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right)$$
 (11.32)

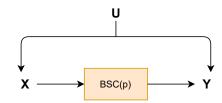


Figure 11.4: Binary symmetric channel

In vector form,

$$\frac{H\left(Y^{n}|U\right)}{n} \geqslant H_{2}\left(H_{2}^{-1}\left(H\left(\frac{X^{n}|U|}{n}\right)\right) \star p\right) \tag{11.33}$$

Proof. We have claim,

$$H\left(\frac{Y}{U}\right) \geqslant H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right)$$
 (11.34)

Considering R.H.S. of the inequality we have,

$$H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right) \tag{11.35}$$

Suppose we have $f(u) = H_2(H_2^{-1}(u) \star p)$.

Now, eq (11.35) can be written as,

$$f\left(\sum_{u} P_{U}(u).H\left(X|U=u\right)\right) \tag{11.36}$$

Since f is convex function, we can write,

$$f\left(\sum_{u} P_{U}(u).H\left(X|U=u\right)\right) \leqslant \sum_{u} P_{U}(u).f\left(H\left(X|U=u\right)\right) \tag{11.37}$$

Again using the expression of f(u) we can have,

$$f(H(X|U=u)) = H_2(H_2^{-1}(H(X|U=u)) \star p)$$
 (11.38)

For $U = u, X \sim Ber(q_u)$ and $Y \sim (q_u \star p)$.

$$H(X|U=u) = H_2(q - u \star p) = H_2(H_2^{-1}(H(X|U=u)) \star p) = f(H(X|U=u))$$
(11.39)

Combining all above expressions, R.H.S. becomes,

R.H.S
$$\leq \sum_{u} P_{U}(u).H(X|U=u)$$
 (11.40)

Hence,

$$H\left(\frac{Y}{U}\right) \geqslant H_2\left(H_2^{-1}\left(H\left(\frac{X}{U}\right)\right) \star p\right)$$
 (11.41)

and this concludes the proof.